Practical guide to sample size calculations: non-inferiority and equivalence trials

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Abstract

A sample size justification is a vital part of any trial design. However, estimating the number of participants required to give a meaningful result is not always straightforward. A number of components are required to facilitate a suitable sample size calculation. In this paper, the steps for conducting sample size calculations for non-inferiority and equivalence trials are summarised. Practical advice and examples are provided that illustrate how to carry out the calculations by hand and using the app SampSize.

1 Introduction

The introduction paper in this series highlighted the key components required to estimate the sample size of a clinical trial [1]. This paper is a practical guide for applying these steps to non-inferiority and equivalence, parallel group clinical trials with a Normally distributed endpoint.

The paper begins by providing a brief explanation of noninferiority and equivalence trials in context with their null and alternative hypotheses. Then the formulae for sample size calculations are presented, highlighting the role of each of the components.

Examples are used to demonstrate how this information is used in practice. The examples are illustrated using the app SampSize [2] and also by hand to emphasise the ease of calculating the sample size from the equations provided. Details of how to obtain the app are provided in Table 1.

Table 1: How to obtain the app SampSize The SampSize app is available on the Apple App Store to download for free and can be used on iPod Touch, iPad and iPhones. The app is also available on the Android Market. It requires Android version 2.3.3 and above. For the calculations in this paper, an iPad is used.

2 Components of a Sample Size

In the introduction to this series, the steps for a sample size calculation are identified. Table 2 gives a summary of these points. An important step before carrying out any sample size calculation is being able to complete this table.

| Step | Summary |
|--------------------------------------|----------------------------|
| Objective | Is the trial aiming to |
| | show superiority, non- |
| | inferiority or equiva- |
| | lence? |
| Endpoint | What endpoint will be |
| | used to show the pri- |
| | mary outcome? Nor- |
| | mal, binary, ordinal or |
| | time to event? |
| Error | Type I error: How much |
| | chance are you willing |
| | to take of rejecting the |
| | null when it is actually |
| | true? |
| | Type II error: How |
| | much chance are you |
| | willing to take of not re- |
| | jecting the null when it |
| | is actually false? |
| Non-inferiority or equivalence limit | What is the minimum |
| ~ _ | clinically acceptable dif- |
| | ference? |
| Population Variance | What is the population |
| | variability? |
| Other | Do you need to account |
| | for dropouts? How |
| | many patients meet the |
| | inclusion criteria? |

Table 2: Summary of steps required for a sample size calculation

3 Non-Inferiority Clinical Trials

In an investigation of a new intervention for a particular disease, a randomised controlled trial can be conducted to investigate if the new treatment is superior to a placebo control. However, once an existing therapy has been established, it may no longer be ethical to undertake placebo controlled trials. Instead, active-controlled trials can be conducted where a new treatment is compared with an established treatment with the objective of demonstrating that the new treatment is non-inferior. For certain trials, the objective therefore is not to demonstrate that a new treatment is superior to placebo or equivalent to an established treatment but rather to demonstrate that a given treatment is clinically not inferior or no worse compared with another. The null (H_0) and alternative (H_1) hypotheses for non-inferiority trials may take the form as follows:

 H_0 : Treatment A is inferior in terms of the mean response $\mu_A - \mu_B \leq -d_{NI}$. H_1 : Treatment A is non-inferior in terms of the mean response

 $\mu_A - \mu_B > -d_{NI}.$

The non-inferiority limit, d_{NI} , is defined to be the difference that is clinically acceptable for us to conclude that there is no difference between treatments. Table 3 summarises how this limit might be selected.

Table 3: Summary of considerations in setting non-inferiority limits.

The setting of a non-inferiority limit is a controversial issue. Regulatory guidelines exist to provide guidance on the topic [3],[4]. The limit is defined as the largest difference that is clinically acceptable, so that a difference bigger than this would matter in practice [7]. This difference also cannot be greater than the smallest effect size that the active (control) drug would be reliably expected to have compared with placebo in the setting of the planned trial [8].

The definition of an acceptable level of non-inferiority can be made by making a retrospective comparison to placebo such that if we can show that a new treatment is non-inferior to standard treatment, we are indirectly demonstrating that the new treatment is superior to the placebo [9]. In such indirect comparisons, the following **ABC** needs to be considered [10][11]:

1. The Assay sensitivity of the active control in both the placebo controlled trials and in the active-controlled non-inferiority trial exists.

3. Bias is minimised through steps such as ensuring that the patient population and the primary efficacy endpoint are essentially the same for the placebo-controlled trial and the active-controlled trial.

2. Constancy assumption of the effect of the common comparator. Such that for two trials in sequence (Trial 1 and Trial 2), the control effect of Treatment B vs placebo in Trial 1 is assumed to be the same as the control effect of Treatment B vs placebo in Trial 2 had placebo been given.

Non-inferiority trials reduce to a simple one-sided hypothesis test. In practice, this is operationally the same as constructing a $(1 - 2\alpha)100\%$ confidence interval (CI) and concluding non-inferiority provided that the lower end of this CI is greater than d_{NI} .

Usually, non-inferiority trials (like equivalence trials discussed later) com-

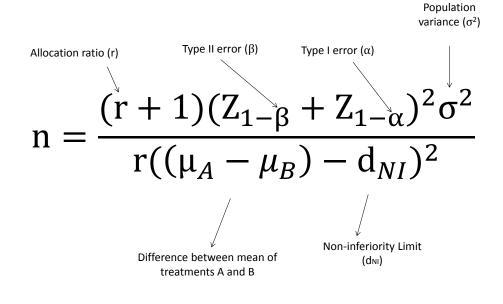


Figure 1: Formula for a non-inferiority parallel group trial.

pare the investigative therapy to an active control. Statistically, they could be considered a special case of equivalence trials. However, they differ in one very important aspect. For a non-inferiority trial, a mean difference a long way from d_{NI} , in a positive sense, is not a negative outcome for the study. While for an equivalence trial, it would make demonstrating equivalence harder.

The distinction between non-inferiority and equivalence is important. Imagine we were designing a role replacement study to assess the work of nurse practitioners against doctors. For an equivalence assessment, we would wish to show that nurse practitioners and doctors are the same with respect to an appropriate clinical outcome, such that simultaneously we would wish to prove that nurse practitioners are as good as doctors and doctors are as good as nurse practitioners. However, with a non-inferiority trial, we would only wish to show that nurse practitioners are as good as doctors. If they are better in other words doctors are worse this would still be a positive outcome.

The sample size required for a non-inferiority clinical trial can be calculated using the formula in Figure 1 [5],[6].

Table 4 gives common Normal deviates for different percentiles. For example, for $\beta = 0.1$, we would have x = 0.1 and $Z_{1-x} = 1.282$, while for $\alpha = 0.05$, we would have x = 0.025 and $Z_{1-x} = 1.96$. These values are useful when calculating a sample size by hand.

As for superiority trials, when the population variance will be assumed unknown in the analysis of the trial, it is best to calculate the power under the assumption of a non-central t distribution [5],[6]. This approach is used in the

Table 4: Normal deviates for common percentiles

| x | Z_{1-x} |
|-------|-----------|
| 0.200 | 0.842 |
| 0.150 | 1.036 |
| 0.100 | 1.282 |
| 0.050 | 1.645 |
| 0.025 | 1.960 |
| 0.010 | 2.326 |
| 0.001 | 3.090 |

app SampSize

$$1 - \beta = Probt\left(t_{1-\alpha}, n_A(r+1) - 2, n_A(r+1) - 2, \tau\right)$$
(1)

where τ is defined as

$$\tau = |\frac{((\mu_A - \mu_B) - d_{NI})\sqrt{rn_A}}{\sqrt{(r+1)\sigma^2}}|$$

Probt(·) here is taken from the function in the package SAS (SAS Institute, Inc., Cary, NC, USA). To calculate the sample size, we can use Table 5, which gives calculated sample sizes for various standardised non-inferiority limits ($\delta_{NI} = d_{NI}/\sigma$).

The percentage mean differences are given for the case where it is anticipated that there may be a non-zero difference between treatments, that is, $\mu_A - \mu_B \neq 0$. For example, if we had $d_{NI} = -10$ (arbitrary scale) but we believed there may be a small difference between treatments, such that $\mu_A - \mu_B = 1$, then the percentage mean difference would be 1/10=10%, that is, the percentage of the non-inferiority limit of the mean difference.

Note though the asymmetric effect on the sample size. This is because we only need to show one bound of a 95% CI that has to exclude d_{NI} . For the same example, if we had $d_{NI} = -10$ (scale arbitrary) but we believed that there may be a small difference between treatments, such that $\mu_A - \mu_B = 1$, then when we are designing the study, we are anticipating the mean difference to be 11 away from the margin. Therefore, a smaller sample size is needed compared with the assumption of $\mu_A - \mu_B = 0$ or $\mu_A - \mu_B = 1$.

For quick calculations (for 90% power and type I error of 2.5%), the following result can be used [5][6]:

$$n_A = \frac{10.5\sigma^2(r+1)}{((\mu_A - \mu_B) - d_{NI})^2)}$$
(2)

In the case of r = 1 (2) resolves to:

$$n_A = \frac{21\sigma^2}{((\mu_A - \mu_B) - d_{NI})^2)}$$
(3)

Table 6 discusses the rationale for setting the type I error in a non-inferiority trial.

Table 5: Sample sizes (n_A) for one arm of a parallel group non-inferiority study with equal allocation for different standardised non-inferiority limits ($\delta_{NI} = d_{NI}/\sigma$) and true mean differences (as a percentage of the non-inferiority limit) for 90% power and type I error rate of 2.5% (from Equation (2)). **Percentage mean difference**

| Percentage mean difference | | | | | | | | | | | |
|----------------------------|------|------|------|------|------|------|------|-------|-------|-------|-------|
| δ_{NI} | 25% | 20% | 15% | 10% | 5% | 0% | 5% | 10% | 15% | 20% | 25% |
| 0.05 | 5381 | 5839 | 6358 | 6949 | 7626 | 8407 | 9316 | 10379 | 11636 | 13136 | 14945 |
| 0.10 | 1346 | 1461 | 1590 | 1738 | 1908 | 2103 | 2330 | 2596 | 2910 | 3285 | 3737 |
| 0.15 | 599 | 650 | 708 | 773 | 849 | 935 | 1036 | 1155 | 1294 | 1461 | 1662 |
| 0.20 | 338 | 366 | 399 | 436 | 478 | 527 | 584 | 650 | 729 | 822 | 935 |
| 0.25 | 217 | 235 | 256 | 279 | 306 | 338 | 374 | 417 | 467 | 527 | 599 |
| 0.30 | 151 | 164 | 178 | 194 | 213 | 235 | 260 | 290 | 325 | 366 | 417 |
| 0.35 | 111 | 121 | 131 | 143 | 157 | 173 | 192 | 213 | 239 | 270 | 306 |
| 0.40 | 86 | 93 | 101 | 110 | 121 | 133 | 147 | 164 | 183 | 207 | 235 |
| 0.45 | 68 | 74 | 80 | 87 | 96 | 105 | 116 | 130 | 145 | 164 | 186 |
| 0.50 | 55 | 60 | 65 | 71 | 78 | 86 | 95 | 105 | 118 | 133 | 151 |
| 0.55 | 46 | 50 | 54 | 59 | 64 | 71 | 78 | 87 | 98 | 110 | 125 |
| 0.60 | 39 | 42 | 46 | 50 | 54 | 60 | 66 | 74 | 82 | 93 | 105 |
| 0.65 | 33 | 36 | 39 | 43 | 47 | 51 | 57 | 63 | 70 | 79 | 90 |
| 0.70 | 29 | 31 | 34 | 37 | 40 | 44 | 49 | 54 | 61 | 68 | 78 |
| 0.75 | 25 | 27 | 30 | 32 | 35 | 39 | 43 | 48 | 53 | 60 | 68 |
| 0.80 | 23 | 24 | 26 | 29 | 31 | 34 | 38 | 42 | 47 | 53 | 60 |
| 0.85 | 20 | 22 | 23 | 26 | 28 | 31 | 34 | 37 | 42 | 47 | 53 |
| 0.90 | 18 | 20 | 21 | 23 | 25 | 27 | 30 | 34 | 37 | 42 | 48 |
| 0.95 | 16 | 18 | 19 | 21 | 23 | 25 | 27 | 30 | 34 | 38 | 43 |
| 1.00 | 15 | 16 | 17 | 19 | 21 | 23 | 25 | 27 | 31 | 34 | 39 |
| | | | | | | | | | | | |

Table 6: Rationale for setting a type I error for a non-inferiority trial.

The convention for non-inferiority trials is to set the type I error rate at half of that used for a two-sided test used in a superiority trial, i.e. $\alpha = 0.025$, a type I error rate of 2.5%. It is conventional for one-sided tests to have half the type I error rate of two-sided tests [12].

It could be argued however that setting the type I error rate for non-inferiority trials at half that for superiority trials is consistent. This is because although in a superiority trial we have a two-sided 5% significance level in practice, for most trials, what we have is a one-sided investigation with a 2.5% level of significance. The reason for this is that we have an investigative therapy and a control therapy and it is only the statistical superiority of the investigative therapy that is of interest [9],[13]. It is logical therefore to formulate superiority trials as a one-sided test - which is inevitable.

Operationally, statistical assessment of non-inferiority is the same as constructing 95% CI and concluding non-inferiority provided that the lower end of this CI is greater than d_{NI} . While for a superiority trial, statistical assessment of superiority is the same as constructing 95% CI and concluding superiority provided that the lower end of this CI is greater than 0.

Through the rest of the sections on equivalence and noninferiority trials, the assumption will be that $\alpha = 0.025$ and that 95% CIs will be used in the final statistical analysis.



Figure 2: Screen shot of SampSize app for non-inferiority worked example with standardised differences.

3.1 Worked Example

A trial is to be undertaken to investigate the effect of a new pain treatment of rheumatoid arthritis. The trial design being considered is a non-inferiority study. The objective is to show that the new treatment is as good as standard therapy. The primary endpoint will be pain measured on a visual analogue scale. The non-inferiority limit is 2.5mm. The standard deviation is anticipated to be 10mm; assuming a one-sided type I error rate of 2.5% and 90% power. The investigator wishes to estimate the sample size per arm. The true mean difference between the treatments is thought to be zero.

After identifying the key points summarised in Table 7, it is possible to plug the values into the formula in Figure 1. This gives:

$$n_A = \frac{(1+1)(Z_{1-0.1} + Z_{1-0.025})^2 (10^2)}{1 \times 2.5^2} \tag{4}$$

Using the percentiles from Table 4, the sample size is calculated to be 336 for each group.

With a non-inferiority limit of 2.5 and a standard deviation of 10 from Table V, the sample size is estimated to be 338 patients per arm.

If we believed that there has to be a small difference of 0.5 between treatments, which equates to (0.5/2.5) 20% of the non-inferiority limit, then the sample size could be reduced to 235 patients per arm. The sample size is reduced as the assumption is that there is a small difference, in favour of the investigative treatment, and so, under this assumption, it should be easier to show non-inferiority.

Repeating the calculation with the SampSize app select Non-Inferiority then Parallel Group, Normal and Calculated Sample Size. Example entries and output using standardised entries (standard deviation of 1 and non-inferiority limit of 2.5/10=0.25) are given in Figure 2 using the assumption of a zero difference.

Example entries and output using the original scale of the non-inferiority limits and standard deviation are shown in Figure 3 using the assumption of a 0.5 difference between treatments. Note that a non-inferiority limit of 2.5 and difference of -0.5 equates to the same calculation as a non-inferiority limit of -2.5 and a difference of 0.5.

The SampSize app also gives the sample size as 338 and 235 patients per arm for a difference of 0 and 0.5, respectively, the same as the results from Table 5.

3.2 Superiority Trials Revisited

In a superiority trial, the objective is to determine whether there is evidence of a statistical difference in the comparator of interest between the regimens with reference to the null hypothesis that the regimens are the same.

Something we glossed over in the earlier papers was that for superiority trials, the alternative hypothesis is such that the treatments are not equal or not equivocal [1],[14], not that one treatment is superior [12]. Formally, a superiority trial should be formulated as follows:



Figure 3: Screen shot of SampSize app for non-inferiority worked example on original scale.

| Step | Summary |
|-----------------------|--|
| Objective | Non-inferiority: H_0 : |
| | $\mu_A - \mu_B \le -d_{NI} \operatorname{vs} H_1:$ |
| | $\mu_A - \mu_B > -d_{NI}$ |
| Endpoint | Improvement in visual |
| | analogue pain |
| Error | Type I error $\alpha = 0.05$ |
| | Type II error $\beta = 0.1$, |
| | power $1 - \beta = 0.9$ |
| Non-Inferiority Limit | $d_{NI} = 2.5mm$ |
| Population Variance | $\sigma = 10mm$ |
| Other | r = 1 |

Table 7: Key components required for sample size for a non-inferiority trial.

 H_0 : There is no difference between treatments in terms of the mean response $\mu_A = \mu_B$.

 H_1 : A given treatment is superior in terms of the mean response $\mu_A > \mu_B$. where in the definition of the null and alternative hypotheses, μ_A and μ_B refer to the mean response on regimens A and B, respectively.

In our discussion, we highlight how, in effect, non-inferiority studies are onesided tests. The same could be said for superiority trials. However, in our earlier paper, we have used the convention of a two-sided test for superiority. It could be argued this is not strictly correct but it is consistent with International Conference on Harmonisation (ICH) E9 guidelines [12].

3.3 As Good As or Better Trials

For some clinical trials, the objective is to demonstrate that either a given treatment is clinically non-inferior or that it is clinically superior when compared with the control, that is, the treatment is as good as or better than the control [5][6][15]. In as good as or better trials, two null and alternative hypotheses are investigated. First, the non-inferiority null and alternative hypotheses:

 H_0 : A given treatment is inferior with respect to the mean response.

 H_1 : The given treatment is non-inferior with respect to the mean response. If this null hypothesis is rejected, then a second, superiority null hypothesis can be investigated:

 H_0 : The two treatments have equal effect with respect to the mean response.

 H_1 : The two treatments are different with respect to the mean response. Operationally, these null hypotheses are investigated through the construction of a 95% CI to investigate where the lower (or upper as appropriate) bound lies. If it excludes the non-inferiority limit, then a conclusion of non-inferiority can be made, while if it excludes zero, a conclusion of superiority can be made. As good as or better trials therefore combine the null hypotheses of superiority and non-inferiority trials into one closed testing procedure while maintaining the overall type I error rate by undertaking a non-inferiority test followed by a test for superiority.

The method used depends on the primary objective of the trial. For example, if the primary objective is to investigate non-inferiority, under the assumption of a small difference between treatments, then the sample size can be calculated for this objective. A calculation can then be made as to what power the study has with the calculated non-inferiority sample size for the superiority objective.

Note that in these calculations consideration would need to be made as to the primary analysis population. For a superiority trial, the primary data set may be that is based on an intention to treat (ITT) data set; for a non-inferiority trial, the primary data set may be both the per protocol and the ITT data set [5].

4 Equivalence Trials

In certain cases, the objective of a study is not to demonstrate superiority of one treatment over another but to show that two treatments have no clinically meaningful difference, that is, they are clinically equivalent. The null (H_0) and alternative (H_1) hypotheses for such equivalence trials take the form as follows:

 H_0 : There is a clinically meaningful difference between treatments in terms of the mean response: $\mu_A - \mu_B \leq -d_E$ and $\mu_A - \mu_B \geq -d_E$.

 H_1 : There is no clinically meaningful difference between treatments in terms of the mean response $-d_E < \mu_A - \mu_B < d_E$.

The equivalence limit d_E , equates to the largest difference clinically acceptable for us to conclude no difference between treatments.

These hypotheses are an example of an intersection-union test, in which the null hypothesis is expressed as a union and the alternative as an intersection [5]. Therefore, to conclude equivalence, we need to reject each component of the null hypothesis.

An approach with equivalence trials is to test each component of the null hypothesis called the two one-sided test (TOST) procedure. In practice, this is operationally the same as constructing a $(1 - 2\alpha)100\%$ CI and concluding equivalence if the CI falls completely within the interval $(-d_E, +d_E)$.

For example, if d_E is set to be 10 (arbitrary scale), and after conducting a trial, a 95% CI for the difference between treatments is found to be (-3, 7). This interval is wholly contained within (-10, 10) so it is possible to conclude that the two treatments are equivalent.

Note that although not covered in this paper, bioequivalence trials are similar, in an operational sense, to the equivalence trials in terms of the null and alternative hypotheses [5]. The main difference statistically is that the type I error is 5% and 90% CIs are used. There are of course bigger differences clinically in their objectives. The rationale for setting the type I error in an equivalence trials is discussed in Table 8.

For equivalence trials, the sample size cannot be derived directly for the general case where the expected true mean difference is non-zero. Instead, the power (and type II error rate) needs to be calculated for a given sample size [5],[6].

$$1 - \beta = \Phi\left(\sqrt{\frac{((\mu_A - \mu_B) - d_E)^2 r n_A}{(r+1)\sigma^2}} - Z_{1-\alpha}\right) + \Phi\left(\sqrt{\frac{((\mu_A - \mu_B) + d_E)^2 r n_A}{(r+1)\sigma^2}} - Z_{1-\alpha} - 1\right)$$
(5)

The sample size is then estimated by iterating until the required power is reached. As with non-inferiority and superiority trials, it is best to use a noncentral t distribution to calculate the type II error and power. From a non-central t distribution, the power can be calculated using the following formula [5],[6]

$$1 - \beta = Probt(-t_{1-\alpha,n_A(r+1)-2}, n_A(r+1) - 2, \tau_2)$$

$$- Probt(t_{1-\alpha,n_A(r+1)-2}, n_A(r+1) - 2, \tau_1)$$
(6)

where τ_1 and τ_2 are non-centrality parameters defined as follows

$$\tau_{1} = \left| \frac{((\mu_{A} - \mu_{B}) + d_{E})\sqrt{rn_{A}}}{\sqrt{(r+1)\sigma^{2}}} \right|$$
$$\tau_{2} = \left| \frac{((\mu_{A} - \mu_{B}) - d_{E})\sqrt{rn_{A}}}{\sqrt{(r+1)\sigma^{2}}} \right|$$

Result (6) is used to estimate the sample sizes in Table 9. For the special case of no treatment difference, a direct estimate of the sample size is given in Figure 4.

While under the assumption of a non-central t distribution, the power can be derived from

$$1 - \beta = 2Probt(-t_{1-\alpha,n_A(r+1)-2}, n_A(r+1) - 2, \tau - 1)$$
(7)

where τ is defined as

$$\tau = \frac{-\sqrt{n_A} r d_E}{\sqrt{(r+1)\sigma^2}}.$$
(8)

For quick calculations of no treatment difference (for 90% power and type I error rate of 2.5%), the following result can be used

$$n_A = \frac{13\sigma^2(r+1)}{d_E^2 r}.$$
 (9)

or for r = 1

$$n_A = \frac{26\sigma^2}{d_E^2}.\tag{10}$$

Table 8: Rationale for setting a type I error for an equivalence trial.

Strictly speaking, when undertaking two simultaneous one-sided tests, setting $\alpha = 0.05$ will maintain an overall type I error rate of 5%. However, convention for equivalence trials like for non-inferiority trials described earlier is to set the type I error rate at half of that which would be employed for a two-sided test used in a superiority trial, i.e. $\alpha = 0.025$, a type I error rate of 2.5%.

This is effectively an investigation of two separate null and alternative hypotheses of the form:

$$H_0: \mu_A - \mu_B \le -d_E$$
$$H_1: \mu_A - \mu_B > -d_E$$

and

$$H_0: \mu_A - \mu_B \ge d_E$$
$$H_1: \mu_A - \mu_B < d_E$$

Each of the null and alternative hypotheses look like a non-inferiority test. If we can reject the null hypothesis for each test so we can declare Treatment A to be non-inferior to Treatment B and Treatment B to be non-inferior to A, we can accept the alternative hypothesis of equivalence. In this context, the setting of the type I error at 2.5% is consistent with non-inferiority trials.

Table 9: Sample sizes (n_A) for one arm of a parallel group equivalence study with equal allocation for different standardised equivalence limits(as a percentage of the equivalence limit) 90% power and type I error rate of 2.5% (from Equation (6)).

| Percentage mean difference | | | | | |
|----------------------------|-------|-------|-------|-------|-------|
| δ_R | 0% | 10% | 15% | 20 | 25% |
| 0.05 | 10397 | 11042 | 11915 | 13218 | 14960 |
| 0.10 | 2600 | 2762 | 2980 | 3306 | 3741 |
| 0.15 | 1157 | 1228 | 1325 | 1470 | 1664 |
| 0.20 | 651 | 691 | 746 | 827 | 936 |
| 0.25 | 417 | 443 | 478 | 530 | 600 |
| 0.30 | 290 | 308 | 332 | 369 | 417 |
| 0.35 | 214 | 227 | 245 | 271 | 307 |
| 0.40 | 164 | 174 | 188 | 208 | 235 |
| 0.45 | 130 | 138 | 149 | 165 | 186 |
| 0.50 | 105 | 112 | 121 | 134 | 151 |
| 0.55 | 87 | 93 | 100 | 111 | 125 |
| 0.60 | 74 | 78 | 84 | 93 | 105 |
| 0.65 | 63 | 67 | 72 | 80 | 90 |
| 0.70 | 55 | 58 | 62 | 69 | 78 |
| 0.75 | 48 | 51 | 54 | 60 | 68 |
| 0.80 | 42 | 45 | 48 | 53 | 60 |
| 0.85 | 37 | 40 | 43 | 47 | 53 |
| 0.90 | 34 | 36 | 38 | 42 | 48 |
| 0.95 | 30 | 32 | 34 | 38 | 43 |
| 1.00 | 27 | 29 | 31 | 35 | 39 |

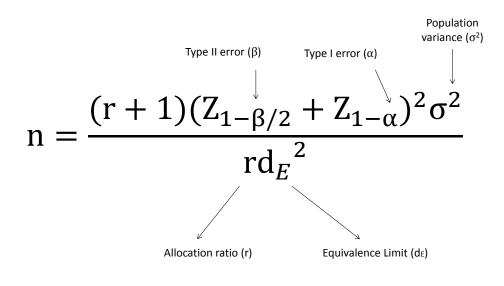


Figure 4: Formula for an equivalence parallel group trial.

4.1 Worked Example

Consider again the trial to investigate the effect of a new pain treatment of rheumatoid arthritis. The objective now is to show that the new treatment is equivalent to a standard therapy. The primary endpoint pain is measured on a visual analogue scale. The largest clinically acceptable effect for which equivalence can be declared is a mean difference of 2.5 mm. The standard deviation is anticipated to be 10 mm; assuming a one-sided type I error of 2.5% and 90% power. The investigator wishes to estimate the sample size per arm. The true mean difference between the treatments is thought to be zero. After identifying these key points summarised in Table 10, it is possible to plug the values into the formula in Figure 4. This gives

$$n_A = \frac{(1+1)(Z_{1-0.1/2} + Z_{1-0.025})^2}{1 \times 2.5^2}.$$
(11)

Using the common percentiles from Table IV, the sample size is calculated to be 416 for each group. With an equivalence limit of 2.5 and a standard deviation of 10 from Table 9, the sample size is estimated to be 417 patients per arm.

If we believed there to be is a small difference of 0.5 between treatments, which equates to 20% of the equivalence limit, then the sample size would need to be increased to 530 patients per arm. The sample size is increased as it is harder to show equivalence with a small difference between treatments than with no difference between treatments.

|--|

| \mathbf{Step} | Summary |
|-------------------------------|---|
| Objective | Superiority: $H_0: \mu_A - \mu_B \leq d_E$ or |
| | $H_0: \mu_A - \mu_B \ge +d_E \text{ vs } H_1: -d_E <$ |
| | $\mu_A - \mu_B > d_E$ |
| Endpoint | Improvement in visual analogue |
| | pain |
| Error | Type I error $\alpha = 0.025$ |
| | Type II error $\beta = 0.1$, power $1 - \beta =$ |
| | 0.9 |
| Equivalence limit | $d_E = 2.5mm$ |
| Population standard deviation | $\sigma = 10mm$ |
| Other | r = 1 |

Repeating the calculation with the SampSize app select Equivalence then Parallel Group, Normal and Sample Size. Example entries and output for the calculation of an equivalence limit of 2.5 and a mean difference of 0.5 are given in Figure 5.

The SampSize app also gives the sample size as 417 and 530 patients per arm for a difference of 0 and 0.5, respectively.

5 Summary

The paper described sample size calculations for trials with a Normally distributed primary endpoint and highlighted how the different non-inferiority and equivalence objectives can impact on calculations. Statistical tables and the app, SampSize, have been used to demonstrate a number of examples of sample size calculations.

While the emphasis of this paper is on trials with a Normally distributed primary endpoint, the principles described can be generally applied to different designs and distributional forms [5],[6].

6 Acknowledgements

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Figure 5: Screen shot of SampSize app for equivalence worked example with standardised differences.

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