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ABSTRACT

Objective. Clinical genetic testing is commercially available for rs61764370, an inherited variant residing in aKRAS 3' UTR microRNA binding site, based on suggested associations with increased ovarian and breast cancer356risk as well as with survival time. However, prior studies, emphasizing particular subgroups, were relatively357small. Therefore, we comprehensively evaluated ovarian and breast cancer risks as well as clinical outcome associated with rs61764370.358

Methods. Centralized genotyping and analysis were performed for 140,012 women enrolled in the Ovarian360Cancer Association Consortium (15,357 ovarian cancer patients; 30,816 controls), the Breast Cancer Association361Consortium (33,530 breast cancer patients; 37,640 controls), and the Consortium of Modifiers of BRCA1 and362BRCA2 (14,765 BRCA1 and 7904 BRCA2 mutation carriers).363

Results. We found no association with risk of ovarian cancer (OR = 0.99, 95% Cl 0.94–1.04, p = 0.74) or breast 364 cancer (OR = 0.98, 95% Cl 0.94–1.01, p = 0.19) and results were consistent among mutation carriers (*BRCA1*, 365 ovarian cancer HR = 1.09, 95\% Cl 0.97–1.23, p = 0.14, breast cancer HR = 1.04, 95\% Cl 0.97–1.12, p = 0.27; 366 *BRCA2*, ovarian cancer HR = 0.89, 95\% Cl 0.71–1.13, p = 0.34, breast cancer HR = 1.06, 95\% Cl 0.94–1.19, p = 367 0.35). Null results were also obtained for associations with overall survival following ovarian cancer (HR = 368 0.94, 95\% Cl 0.83–1.07, p = 0.38), breast cancer (HR = 0.96, 95\% Cl 0.87–1.06, p = 0.38), and all other 369 previously-reported associations.

Conclusions. rs61764370 is not associated with risk of ovarian or breast cancer nor with clinical outcome for 371 patients with these cancers. Therefore, genotyping this variant has no clinical utility related to the prediction or 372 management of these cancers. 373

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393 **1. Introduction**

MicroRNAs (miRNAs) are a class of small non-coding RNA molecules 394that negatively regulate gene expression by binding partially comple-395 mentary sites in the 3' untranslated regions (UTRs) of their target 396 mRNAs. In this way, miRNAs control many cancer-related biological 397 398 pathways involved in cell proliferation, differentiation, and apoptosis [1]. To date, several inherited variants in microRNAs or miRNA target 399 sites have been reported to confer increased cancer risks [2]. One such 400 variant is located in the 3' UTR of the KRAS gene (rs61764370 T > G) 401 for which the rarer G allele has been reported to confer an increased 402 risk of ovarian, breast, and lung cancer [3-7] as well as endometriosis 403 [8], although not consistently [9–11]. 404

For ovarian cancer, the rs61764370 G allele was also reported to be 405associated with increased risk (320 cases, 328 controls). Further in-406 creased risks were observed among 23 BRCA1 mutation carriers and 407 31 women with familial ovarian cancer, but without BRCA1 or BRCA2 408 mutations [3]. In contrast, no association with ovarian cancer risk was 409 seen in another, much larger study, based on 8669 cases, 10,012 con-410 trols, and 2682 BRCA1 mutation carriers [9]. One criticism on the latter 411 412 study was that some of the genotype data were for rs17388148, an 413 imputed proxy for rs61764370; even though rs17388148 is highly correlated with rs61764370 ($r^2 = 0.97$) and was imputed with high accuracy ($r^2 = 0.977$) [12,13]. The minor allele of rs61764370 was also associated with shorter survival time in a study of 279 ovarian cancer patients diagnosed after age 52 years with platinum-resistant disease 417 (28 resistant, 263 not resistant) and with sub-optimal debulking surgery after neoadjuvant chemotherapy (7 sub-optimal, 109 optimal) 419 [14]. However, another study observed no association between 420 rs61764370 and ovarian cancer outcome (329 cases) [15].

For breast cancer, a borderline significant increased frequency of the 422 rs61764370 G allele was observed in 268 *BRCA1* mutation carriers with 423 breast cancer, but not in 127 estrogen receptor (ER)-negative familial 424 non-*BRCA1/BRCA2* breast cancer patients [5]. However, in a subsequent 425 study, the variant was reported to be associated with increased risk of 426 ER/PR negative disease (80 cases, 470 controls), as well as with triple 427 negative breast cancer diagnosed before age 52 (111 cases, 250 controls), 428 regardless of *BRCA1* mutation status [6]. The validity of these findings has 429 been questioned given the very small sample sizes and the number of 430 subgroups tested [16,17]. Another report found no association with spo-431 radic or familial breast cancer risk (695 combined cases, 270 controls), 432 but found that the variant was associated with ERBB2-positive and 433 high grade disease, based on 153 cases who used post-menopausal hor-434 mone replacement therapy [18].

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It has also been reported, based on 232 women with both primary 436 437 ovarian and breast cancer, that the frequency of the G allele at rs61764370 was increased for those who were screened negative for 438 439BRCA1 and BRCA2 (92 cases), particularly among those enrolled within two years of their ovarian cancer diagnosis (to minimize survival bias, 440 30 cases), those diagnosed with post-menopausal ovarian cancer (63 441 cases), those with a family history of ovarian or breast cancer (24 442 cases), and those with a third primary cancer (16 cases) [4]. 443

444 This notable lack of consistency in findings between studies might be 445 expected when appropriate levels of statistical significance are not used to declare positive findings from multiple small subgroup comparisons 446or post-hoc hypotheses [19]. In this respect, the dangers of subgroup 447analyses in the context of clinical trials are well-recognized [20]. These 448are important caveats, particularly since a genetic test for rs61764370 449 is currently marketed in the US for risk prediction testing to women 450who are at increased risk for developing ovarian and/or breast cancer 451or women who have been diagnosed with either ovarian or breast cancer 452453 themselves [21]. In general, much larger studies, with sufficient power to detect positive findings at much more stringent levels of statistical signif-454icance ought to be required to establish the clinical validity of a genetic455test. Therefore, we conducted centralized genotyping of rs61764370456and other variants in the genomic region around the KRAS gene in457140,012 women to examine associations with risk and clinical outcome458of ovarian and breast cancer.459

2. Methods

2.1. Study participants

The following three consortia contributed to these analyses: the 462 Ovarian Cancer Association Consortium (OCAC: 41 studies, Supplemen-463 tary Table S1) [22], the Breast Cancer Association Consortium (BCAC: 37 464 studies, Supplementary Table S2) [23], and the Consortium of Modifiers 465 of *BRCA1* and *BRCA2* (CIMBA: 55 studies, Supplementary Table S3) 466 [24,25]. OCAC and BCAC consisted of case-control studies of unrelated 467 women, and CIMBA consisted of studies of women with germline 468

t1.1 Table 1

t1.2 Associations between KRAS rs61764370 and risk of ovarian and breast cancer.

t1.3 For BRCA1 and BRCA2 mutation carrier analyses, cases are affected BRCA1/BRCA2 mutation carriers and controls are unaffected BRCA1/BRCA2 mutation carriers, and relative risks are estimated by hazard ratios; for other analyses, relative risks are estimated by odds ratios; ovarian cancer analyses used OCAC data adjusted for study, age, and the five European principal components; breast cancer analyses used BCAC data adjusted for study, age, and the seven European principal components; BRCA1 and BRCA2 mutation carrier analyses used CIMBA data

1.6 with age as follow-up time and stratified for country: 95% CI. 95% confidence interval

	Number		Minor allele	frequency	Relative risk (95% CI)	
	Cases	Controls	Cases	Controls		p-Valu
Ovarian cancer						
All invasive	15,357	30,816	0.0914	0.0949	0.99 (0.94-1.04)	0.74
Histology						
High-grade serous	6938	30,816	0.0946	0.0949	1.04 (0.97-1.11)	0.26
Endometrioid	2151	30,816	0.0834	0.0949	0.90 (0.80-1.00)	0.06
Clear cell	1015	30,816	0.0994	0.0949	1.09 (0.94-1.27)	0.27
Mucinous	1000	30,816	0.0902	0.0949	0.99 (0.85-1.16)	0.91
Low-grade serous	485	30,816	0.0705	0.0949	0.76 (0.59-0.97)	0.03
First-degree family history						
Ovarian cancer	483	342	0.0803	0.0849	0.87 (0.60-1.27)	0.47
Breast or ovarian cancer	477	18,442	0.0977	0.0915	1.09 (0.93-1.28)	0.28
BRCA1/2 mutation negative	346	15,492	0.1050	0.0997	1.09 (0.85–1.41)	0.49
BRCA1 mutation carriers	2332	12,433	0.0954	0.0922	1.09 (0.97–1.23)	0.14
BRCA2 mutation carriers	599	7305	0.0952	0.0966	0.89 (0.71–1.13)	0.34
Enrolled within two years of diagnosis						
All invasive	10,121	30,815	0.0942	0.0949	0.99 (0.95-1.04)	0.68
BRCA1 mutation carriers	1095	10,802	0.0950	0.0940	1.05 (0.90–1.23)	0.52
BRCA2 mutation carriers	270	6509	0.0907	0.0979	0.85 (0.60–1.20)	0.36
Menopausal status	270	0000	010007	0.007.0		0.50
Pre- or peri-menopausal	4264	8789	0.0915	0.0927	1.02 (0.92-1.13)	0.68
Post-menopausal	11,058	15,903	0.0916	0.0951	0.99 (0.93–1.06)	0.81
Prior breast cancer	11,000	10,000	010010	010001		0101
Enrolled within two years of diagnosis	426	30,815	0.0943	0.0949	0.91 (0.71-1.17)	0.46
Post-menopausal ovarian cancer	341	15,903	0.0810	0.0951	0.90 (0.68–1.21)	0.49
First degree breast or ovarian cancer family history	202	30,815	0.0916	0.0949	0.99 (0.70–1.40)	0.95
This degree breast of ovarian cancer family history	202	50,015	0.0510	0.05 15	0.55 (0.76 1.16)	0.55
Breast cancer						
All invasive	33,530	37,640	0.0904	0.0930	0.98 (0.94-1.01)	0.19
Receptor status						
ER—/PR—	4009	37,043	0.0940	0.0932	1.04 (0.96-1.13)	0.36
ER –/PR –/ERBB2 –	1673	28,480	0.0885	0.0947	0.97 (0.85-1.10)	0.62
First-degree family history						
Breast cancer	4357	1943	0.0942	0.0954	0.96 (0.84-1.10)	0.59
Ovarian or breast cancer	4593	2265	0.0933	0.0949	0.96 (0.85-1.09)	0.52
Age diagnosis <52						
ER – /PR –	1530	37,043	0.0980	0.0932	1.07 (0.95-1.22)	0.28
ER –/PR –/ERBB2 –	546	27,690	0.0908	0.0948	0.99 (0.81-1.20)	0.90
BRCA1/2 mutation negative	1431	1097	0.0853	0.0925	0.91 (0.75–1.11)	0.35
BRCA1 mutation carriers	7543	7222	0.0935	0.0919	1.04 (0.97–1.12)	0.27
BRCA2 mutation carriers	4138	3766	0.1005	0.0921	1.06 (0.94–1.19)	0.35
Enrolled within two years of diagnosis						
All invasive	20,444	34,349	0.0924	0.0934	0.99 (0.95-1.04)	0.73
BRCA1 mutation carriers	2595	5976	0.0896	0.0924	0.95 (0.85–1.05)	0.30
BRCA2 mutation carriers	1359	3365	0.0960	0.0926	1.05 (0.90–1.23)	0.50
Menopausal status					(,	
Pre- or peri-menopausal	7086	8642	0.0934	0.0933	0.98 (0.91-1.07)	0.70
Post-menopausal	16,346	18,605	0.0904	0.0943	0.98 (0.93–1.03)	0.36

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t2.1

deleterious BRCA1 or BRCA2 mutations primarily identified through 469 470 clinical genetics centers. For the purpose of the current analyses, only 471 participants of European ancestry were included. Following genotyping, 472quality control exclusions (described below), and analysis-specific exclusions, data from the following women were available for analysis: 473 46,173 OCAC participants (15,357 patients with invasive epithelial 474 ovarian cancer and 30,816 controls), 71,170 BCAC participants (33,530 475patients with invasive breast cancer and 37,640 controls), and 22,669 476 477CIMBA participants (for ovarian cancer analyses: 2332 affected and 12,433 unaffected BRCA1 carriers, 599 affected and 7305 unaffected 478 479BRCA2 carriers; for breast cancer analyses: 7543 affected and 7222 unaffected BRCA1 carriers, 4138 affected and 3766 unaffected BRCA2 480481 carriers). For OCAC, overall and progression-free survival data were 482 available for 3096 patients from 13 studies. Overall survival data were available for 28,471 patients from 26 BCAC studies and for 2623 483 mutation carriers with breast cancer from 11 CIMBA studies (excluding 484studies with less than ten deaths) as described previously [26,27]. Each 485study was approved by its relevant governing research ethics commit-486 tee, and all study participants provided written informed consent. 487

488 2.2. Genotyping and imputation

Genotyping for rs61764370 was performed using the custom iCOGS 489 Illumina Infinium iSelect BeadChip, as previously described [22–25]. In 490 total, DNA from 185,443 women of varying ethnic background was ge-491 notyped (47,630 OCAC participants, 114,255 BCAC participants, 23,558 492493 CIMBA participants), along with HapMap2 DNAs for European, African, and Asian populations. Genotype data were also available for three 494OCAC genome-wide association studies (UK GWAS, US GWAS, Mayo 495GWAS) that had been genotyped using either the Illumina 496 497Human610-Quad Beadchip (12,607 participants) [28] or the Illumina HumanOmni2.5-8 Beadchip (883 participants). Raw intensity data 498files underwent centralized genotype calling and quality control 499[22–25]. HapMap2 samples were used to identify women with predict-500ed European intercontinental ancestry; among these women, a set of 501 over 37,000 unlinked markers was used to perform principal compo-502503nent (PC) analysis [29]. The first five and seven European PCs were found to control adequately for residual population stratification in 504OCAC and BCAC data, respectively. Samples with low conversion rate, 505extreme heterozygosity, non-female sex, or one of a first-degree relative 506 507pair (the latter for OCAC and BCAC only) were excluded. Variants were excluded if they were monomorphic or had a call rate <95% (minor al-508lele frequency (MAF) >0.05) or <99% (MAF <0.05), deviation from 509Hardy–Weinberg equilibrium ($p < 10^{-7}$), or >2% duplicate discordance. 510In addition to rs61764370, 54 variants within 100 kb on either side 511512of KRAS on chromosome 12 (25,258,179 to 25,503,854 bp in GRCh37.p12) were genotyped. Moreover, to provide a common set of 513variants across the region for analysis in all the data sets, we also used 514imputation to infer genotypes for another 1056 variants and for variants 515that failed genotyping. We performed imputation separately for OCAC 516517samples, BCAC samples, BRCA1 mutation carriers, BRCA2 mutation car-518riers, and for each of the OCAC GWAS. We imputed variants from the 1000 Genomes Project data using the v3 April 2012 release as the refer-519ence panel [30]. To improve computation efficiency we initially used a 520two-step procedure, which involved pre-phasing using the SHAPEIT 521522software [31] in the first step and imputation of the phased data in the second. We used the IMPUTE version 2 software [32] for the imputation 523for all studies with the exception of the US GWAS for which we used the 524MACH algorithm implemented in the minimac software version 5252012.8.15 and MACH version 1.0.18 [33]. We excluded variants from as-526sociation analyses if their imputation accuracy was $r^2 < 0.30$ or their 527MAF was <0.005, resulting in 974 variants genotyped and imputed for 528OCAC, 989 variants genotyped and imputed for BCAC, and 1001 variants 529genotyped and imputed for CIMBA, including rs61764370 (Supplemen-530531 tary Tables S5, S6, and S7).

2.3. Analysis

Genotypes were coded for genotype dosage as 0, 1, or 2, based on the 533 number of copies of the minor allele. For ovarian cancer case-control 534 analysis (i.e., OCAC studies), logistic regression provided estimated 535 risks of invasive epithelial ovarian cancer with odds ratios (ORs) and 536 95% confidence intervals (CIs) adjusting for study, age, and the five 537 European PCs. Subgroup analyses were conducted by histology, family 538 ovarian and breast cancer history, menopausal status, time between 539 ovarian cancer diagnosis and recruitment, and history of multiple 540 primary cancers. For breast cancer case-control analysis (i.e., BCAC 541 studies), the association between genotype and invasive breast cancer 542 risk was evaluated by logistic regression, adjusting for study, age, and 543 the seven European PCs, providing ORs and 95% CIs. Additional 544 subgroup analyses were based on receptor status, first-degree family 545 ovarian and breast cancer history, BRCA1 and BRCA2 mutation status, 546 enrollment within two years of diagnosis, menopausal status (i.e. last 547 menstruation longer than twelve months ago), age at diagnosis less 548 than 52 years, and history of hormone replacement therapy use (i.e. lon- 549 ger than twelve months use). Risk analysis for BRCA1 and BRCA2 muta- 550 tion carriers (i.e. CIMBA studies) was done using a Cox proportional 551 hazard model to estimate hazard ratios (HRs) per copy of the minor al- 552 lele, with age as follow-up time and stratified by country of residence; 553 US and Canadian strata were further subdivided by self-reported 554 Ashkenazi Jewish ancestry [24,25]. A weighted cohort approach was ap- 555 plied to correct for potential testing bias due to overrepresentation of 556 cases in the study population [34]. We used robust variance estimation 557 to allow for the non-independence of carriers within the same family 558 [35]. To assess associations with ovarian cancer risk, mutation carriers 559 were followed from birth until ovarian cancer diagnosis (event), a 560 risk-reducing salpingo-oophorectomy (RRSO) or the age at enrollment, 561

Table 2

Associations between KRAS rs61764370 and outcome in ovarian and breast cancer. t2.2 Ovarian cancer analyses used OCAC data adjusted for age at diagnosis (overall survival ont2.3 lv), the five European principal components, histology (serous, mucinous, endometrioid, t2.4 clear cell, and other epithelial), grade (low versus high), FIGO stage (I–IV), residual disease t2.5 after debulking surgery (nil versus any), and stratified by study; breast cancer analyses t2.6 used BCAC data adjusted for age at diagnosis, tumor size, nodal status, grade, adjuvant hor- t2.7 monal and/or chemotherapy and was stratified by study: analyses for BRCA1 and BRCA2 t2.8 mutation carriers used CIMBA data adjusted for age at diagnosis, tumor size, nodal status, t2.9 grade, adjuvant hormonal and/or chemotherapy, and preventive bilateral oophorectomy t2.10 and was stratified by study; 95% CI, 95% confidence interval. t2.11

	No. of patients	No. of events	Hazard ratio (95% CI)	p-Value
Ovarian cancer				
Overall survival				
All patients	3096	1421	0.94 (0.83-1.07)	0.38
Patients who were suboptimally debulked after cytoreductive surgery	1114	784	0.94 (0.78–1.13)	0.50
Post-menopausal patients > 52 years	2226	1276	0.97 (0.84–1.12)	0.70
Progression-free survival				
All patients	3096	2144	1.01 (0.90-1.13)	0.84
Patients who were suboptimally	1114	961	1.03 (0.87-1.21)	0.74
debulked after cytoreductive surgery				
Post-menopausal patients >52 years	2226	1603	1.02 (0.90–1.16)	0.76
Breast cancer				
Overall survival				
All patients	28,471	3013	0.96 (0.87-1.06)	0.38
ER-positive patients	20,071	1754	0.96 (0.85–1.10)	0.58
ER-negative patients	4778	771	0.97 (0.81–1.18)	0.78
Breast cancer-specific survival			(
All patients	28,471	1693	0.95 (0.83-1.08)	0.40
Overall survival			. ,	
BRCA1 mutation carriers	1706	241	0.72 (0.48-1.08)	0.11
BRCA2 mutation carriers	917	162	0.98 (0.65-1.46)	0.90

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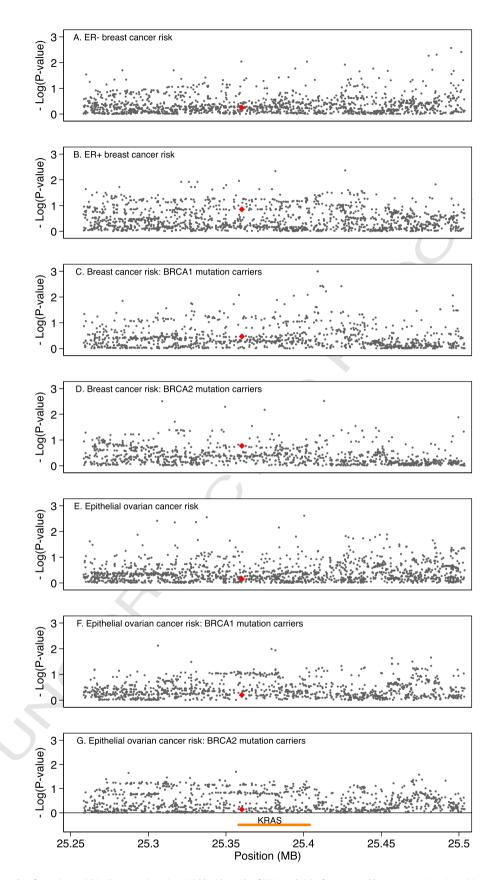


Fig. 1. Regional association plots for variants within the genomic region 100 kb either side of *KRAS* and risk of ovarian and breast cancer. X-axis position is referent to position (bp) on chromosome 12, build GRCh37.p12; yellow line indicates position of *KRAS*; red triangle indicates rs61764370. Y-axis is $-\log_{10}(p-values)$ from association tests for risk of A) ER-negative breast cancer, B) ER-positive breast cancer, C) breast cancer in *BRCA1* mutation carriers, D) breast cancer in *BRCA2* mutation carriers, E) epithelial ovarian cancer, F) epithelial ovarian cancer in *BRCA1* mutation carriers. (For interpretation of the references to color in this figure legend, the reader is referred to the web version of this article.)

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whichever occurred first. We also performed analyses restricted to women diagnosed or censored within two years before their enrollment. To assess associations with breast cancer risk, mutation carriers were followed from birth until a breast cancer diagnosis (i.e. either ductal carcinoma in situ or invasive breast cancer), ovarian cancer diagnosis, a risk-reducing bilateral prophylactic mastectomy or the age at enrollment, whichever occurred first.

569Survival analysis of OCAC patients used Cox proportional hazards 570models estimating HRs and 95% CIs considering overall survival as 571well as progression-free survival following ovarian cancer diagnosis. Overall survival was adjusted for age at diagnosis, the five European 572PCs, histology, grade, FIGO stage, and residual disease after debulking Q7 surgery, and stratified by study, left truncating at the date of study 574575entry and right censoring at five years to minimize events due to other causes. Progression-free survival was analyzed as for overall 576survival, but without adjustment for age and right censoring, and was 577 defined as the time between the date of histologic diagnosis and the 578first confirmed sign of disease recurrence or progression, based on 579GCIG (Gynecological Cancer InterGroup) criteria [36]. We also performed 580subgroup analysis of patients suboptimally debulked after cytoreductive 581surgery (residual disease >1 cm) and of post-menopausal patients (age 582at diagnosis > 52 years). Survival analysis of BCAC patients used Cox pro-583584portional hazard models estimating HRs and 95% CIs considering overall and breast cancer-specific survival following breast cancer diagnosis. 585 Models were adjusted for age at diagnosis, tumor size, nodal status, 586grade, adjuvant hormonal and/or chemotherapy, and stratified by 587study, left-truncating at the date of study entry and right censoring at 588589ten years. In addition, we performed subgroup analysis on ER-positive and ER-negative patients. For CIMBA breast cancer patients associations 590between genotype and overall survival were evaluated using Cox 591592proportional hazard models estimating HRs and 95% CIs. Models were 593adjusted for age at diagnosis, tumor size, nodal status, grade, adjuvant 594hormonal and/or chemotherapy, and preventive bilateral oophorectomy 595and stratified by study, left-truncating at the date of study entry and right censoring at twenty years. Analyses were performed using STATA ver-596sion 12.0 (StataCorp, Texas, USA). 597

598 3. Results

The results of the overall analysis as well as the subgroup analyses investigating the association between the minor allele at rs61764370 and ovarian cancer risk, breast cancer risk, and ovarian and breast cancer risks in *BRCA1* and *BRCA2* mutation carriers are shown in Table 1. Associations with clinical outcomes in and ovarian and breast cancer patients including *BRCA1* and *BRCA2* mutation carriers are shown in Table 2 and Supplementary Table S4.

606 We found no evidence for association between the rs61764370 G allele and ovarian or breast cancer risk. The most statistically significant 607 association was observed for risk of low-grade serous ovarian cancer 608 (n = 485; OR 0.76, 95% CI 0.59-0.97, p = 0.031), but this finding was 609 not significant after Bonferroni correction for multiple testing. We also 610 611 evaluated the association for additional specific subgroups in which an 612 association with rs61764370 had been reported previously [3–6]. Ovarian cancer subgroups considered BRCA1 mutation carriers as well as 613BRCA1 and BRCA2 screened-negative patients with first degree family 614histories of breast or ovarian cancer and patients who had been diag-615 616 nosed with breast cancer before their ovarian cancer diagnoses. For breast cancer these included, among others, BRCA1 mutation carriers, 617 patients diagnosed with ER- and PR-negative tumors, and patients 618 diagnosed with triple negative tumors before age 52 years. Importantly, 619 we observed no evidence for association of rs61764370 with any of 620 these subgroups (detailed in Table 1), with all ORs close to unity and 621 very narrow CIs including unity. 622

523 Similarly, case-only analyses did not reveal any associations 524 between rs61764370 genotype and ovarian and breast cancer clinical 525 features or outcome (Table 2 and Supplementary Table S4). For example, the previously reported association between rs61764370 626 and risk of ERBB2-positive and high grade breast cancer in hormone re- 627 placement therapy users [18] was not replicated (Supplementary 628 Table S4), and in ovarian cancer analyses we found no evidence of re- 629 duced survival among patients diagnosed after age 52 years or patients 630 with suboptimal debulking after cytoreductive surgery (Table 2) [14]. 631 The G allele of rs61764370 was also not associated with survival of 632 breast cancer patients (Table 2). 633

Finally, we evaluated the association between the primary pheno- 634 types of interest and common genetic variation (MAF > 0.02) in the ge- 635 nomic region of *KRAS* (i.e., within 100 kb on either side of the gene), 636 using imputed and genotyped data on 974 variants for OCAC, 989 vari- 637 ants for BCAC, and 1001 variants genotyped and imputed for CIMBA 638 (Supplementary Tables S5, S6, and S7). We found no evidence of associ- 639 ation for any of these variants, including rs61764370 and rs17388148, 640 with these phenotypes that would withstand Bonferroni correction for 641 multiple testing, as detailed in Supplementary Tables S5, S6, and S7 642 and shown in regional association plots (Fig. 1).

4. Discussion

Our analysis of 140,012 women genotyped for inherited variants in 645 the *KRAS* region provides definitive clarification of the role of these var-646 iants in ovarian and breast cancer susceptibility and outcome. We have found no evidence to support an association between rs61764370 and 648 ovarian or breast cancer risk, or clinical outcomes in patients with ovar-649 ian or breast cancer. In the absence of any association and with ORs close to unity we would not typically consider sub-group analyses, particularly sub-groups for which differential associations would not be expected to occur. However, given the previous positive association reported for a myriad of different subgroups, we tested for association among each of these subgroups and found no evidence to support the previously reported associations.

Our study has notable strengths. The vast majority (i.e. >95%) of the 657 samples were genotyped using the same genotyping platform and 658 employing a common approach to genotype calling and quality control; 659 additional samples used denser arrays and nearly identical procedures. 660 The very large sample sizes for all the major phenotypes of interest pro-661 vide substantial statistical power to exclude any clinically relevant associated risks for the major phenotypes of interest (Fig. 2). The null results found here are thus not due to lack of statistical power, and this analysis also had greater than 80% power to detect association for most of the subgroups, although for some subgroups it was not possible to exclude 666

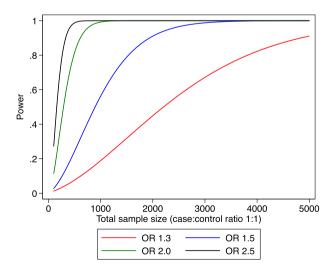


Fig. 2. Power curve for modest risk variants according to the total sample size. X-axis is total sample size for which case–control ratio is 1:1. Y-axis is the statistical power (range $0 \cdot 5 - 1 \cdot 0$) for variants given a range of risks, assuming alpha = $0 \cdot 01$ and minor allele frequency $0 \cdot 09$.

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modest risks. In contrast to the current findings, other genetic associa-667 668 tion analyses using the same genotyping platform and the same studies as included here have identified more than 90 common germline 669 670 variants associated with ovarian or breast cancer risk at $p < 5 \times 10^{-8}$ [22,23,37]. While critiques on a previous null KRAS report have sug-671 gested that inclusion of male controls, use of "prevalent" cases, and reli-672 ance on a surrogate genetic variant may have led to falsely negative 673 conclusions, these are not issues in the present data set. Rather, we 674 675demonstrate the importance of international collaboration to identify 676 true associations as well as to refute false associations, an equally impor-677 tant objective.

The rise of individualized medicine including the use of panels of 678 679 common variants to predict cancer risk more accurately than using fam-680 ily history alone holds great promise [38]. For example, the 31 prostate cancer susceptibility alleles confirmed as of 2011 (at $p < 5 \times 10^{-8}$) can 681 be combined to identify men in the top one percent of the risk distribu-682 tion having a 3.2-fold increased risk [39]. Prediction has since then im-683 proved with now over 70 prostate cancer susceptibility alleles [40] 684 and the utility of these genetic tests is currently under clinical evalua-685 tion. A similar clinical examination in ovarian and breast cancer is not 686 far behind, with now over 18 and 77 confirmed susceptibility alleles, re-687 spectively, for these cancers [22,23]. The genotype at rs61764370, how-688 689 ever, does not predict ovarian or breast cancer risk, even among 690 particular subgroups of women or for particular subtypes of disease, nor is it a marker of differential outcome following diagnosis with 691 these cancers. Therefore, genetic test results for rs61764370 should 692 not be used to counsel women about their ovarian or breast cancer 693 694 risks or outcome. Our results highlight the dangers of developing clinical tests without appropriate data from carefully conducted, large-scale 695 studies to establish clinical validity. 696

697 Conflict of interest statement

698 There are no conflicts of interest to disclose

699 Antoinette Hollestelle and Ellen L. Goode had full access to all of the data in the study and 700 take responsibility for the integrity of the data and the accuracy of the data analysis.

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Appendix A. Supplementary data

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