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A path analysis disentangling determinants of natal dispersal in a cooperatively breeding bird

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Abstract

Delayed offspring dispersal, the prolonged residence in the natal territory after reaching independence and before dispersing to breed elsewhere, is an important aspect of the evolution of cooperative breeding. By applying a path analysis approach to the long-term Seychelles warbler (*Acrocephalus sechellensis*) dataset, we studied whether and how delayed dispersal is affected by territory quality, the presence of helpers and non-helping subordinates, maternal breeding status, age and fecundity, and offspring sex ratio. We found that offspring are more likely to disperse when their genetic mothers are co-breeders, helpers are absent, and territory quality is high, highlighting that a complex interplay of ecological and social factors shapes dispersal decisions. In contrast to earlier findings, our analysis does not support the idea that the offspring sex ratio is affected by territory quality and helper presence. Collectively, this study underscores the importance of considering proximate factors in understanding cooperative breeding dynamics, and it shows that path analyses offer valuable insights into dissecting the intricate relationships influencing dispersal in wild populations.

Keywords: dispersal, cooperative breeding, offspring sex ratio, Seychelles warbler, path analysis, directed acyclic graph, causal network, causal inference

Introduction

Theories on the evolution of group living and cooperative breeding rest on assumptions as to why offspring remain philopatric rather than disperse to seek mating opportunities elsewhere (García-Ruiz et al., 2022; Hatchwell, 2009; Ligon & Stacey, 1989). Numerous hypotheses have been proposed to address this question, primarily framed within cost–benefit analyses of remaining on the natal territory versus dispersal. Potential benefits of philopatry include access to resources within the natal territory, thereby enhancing offspring survival (e.g., Ekman & Griesser, 2002; Kingma et al., 2016; Robles et al., 2022; Suh et al., 2020). Moreover, social dynamics within the natal territory can influence dispersal probabilities. For instance, dispersal may increase when relatedness between parents and offspring is low, such as when parental mortality leads to new step-parents assuming breeding dominance (Suh et al., 2020), as this reduces the inclusive fitness benefits for parents of offspring staying. Conversely, high relatedness between parents and offspring may drive dispersal, as an evolved response to mitigate kin competition (Hamilton, 1967). The presence of conspecifics of similar (dominance) status and sex within the social group may also influence dispersal tendencies by altering territory inheritance probabilities (Pasinelli &

Walters, 2002). Moreover, larger groups could decrease per capita food availability (Dietz et al., 2022).

Factors outside of the natal territory further complicate the cost–benefit analysis of natal dispersal. Territory and mate availability outside of the natal habitat, influenced by factors such as population density and adult sex ratios, can sway dispersal tendencies (Emlen, 1982; Maag et al., 2018; Nelson-Flower et al., 2018; Speelman et al., 2024). Additionally, the costs and benefits of offspring dispersal might differ between parents and offspring, potentially leading to conflict (Port et al., 2020). Therefore, it is important to consider who decides whether offspring disperses (Quiñones et al., 2016) and whether parents and offspring attempt to manipulate each other to influence the dispersal outcome. For instance, male white-fronted bee-eaters (*Merops bullockoides*) actively try to disrupt the independent breeding attempts of their sons (Emlen & Wrege, 1992), to persuade them to return as helpers. Parents could also influence the (social) environment of offspring by changing brood size, and thus sibling competition, which could influence the dispersal tendency of the offspring (Pasinelli & Walters, 2002). Further, parents could alter the offspring sex ratio towards or away from the helping (and thus philopatric) sex. For example, in ants (*Formica exsecta*)

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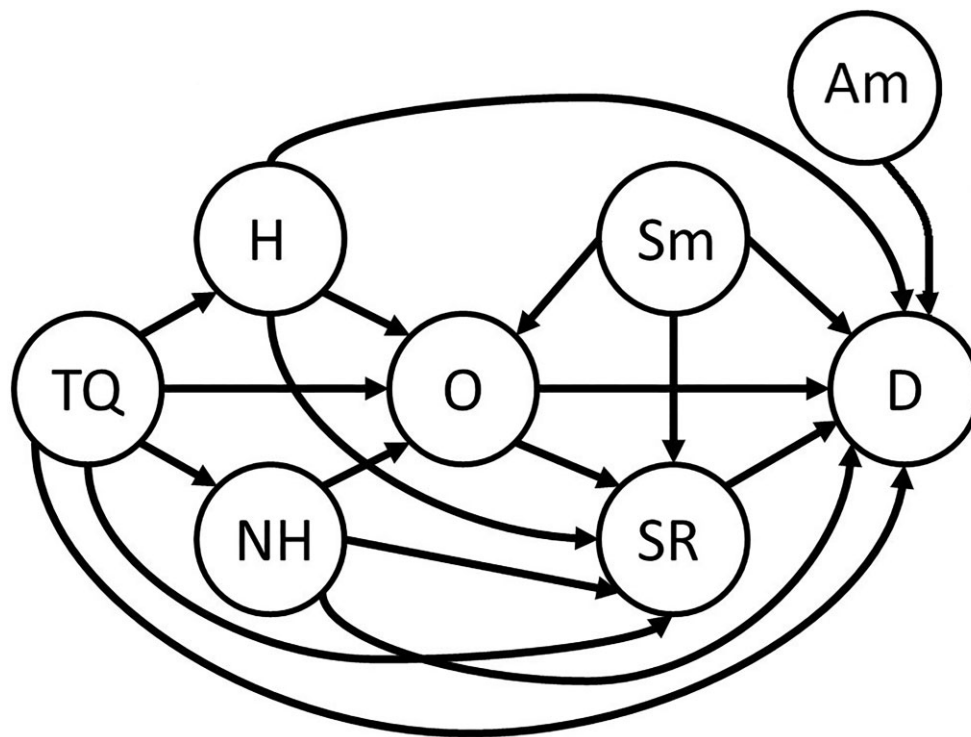


Figure 1. Directed acyclic graph (DAG) representing the hypothetical causal network underlying offspring dispersal. We expected dispersal (D) to be associated with territory quality (TQ), helper presence (H), number of non-helpers (NH), number of offspring (O), status of the mother (Sm), the offspring sex ratio (SR), and the age of the mother (Am). However, these variables were also expected to be associated with each other. Each arrow in the figure indicates such an expected association; see methods for the reasoning behind the arrows. Note that all factors in this DAG are measured on the territory or maternal level.

we expect territory quality to have a positive effect on the number of helpers and non-helpers in a territory, as territories of higher quality contain more resources to support more individuals. Second, the number of offspring might be affected by territory quality, helper presence, the number of subordinate non-helpers, the number of mothers on a territory, and the status of a mother (dominant or subordinate). Territory quality indicates resource availability. Helper presence reflects how many of these resources could be obtained by offspring. Moreover, helpers and non-helpers could deplete the resources available for offspring. The presence of a co-breeder could increase competition among offspring, and the social status of the mother might affect how many offspring the mother can produce in one season, if, for example, co-breeders are limited in their reproductive output by dominant breeders. Third, the offspring sex ratio produced by the focal mother may be affected by territory quality, helper presence, number of non-helpers, the number of offspring of a mother, the number of mothers on a territory, and the status of the mother. Territory quality and helper presence could affect the need for a helper, and therefore the offspring sex ratio. The status of the mother could influence the offspring sex ratio as well, as mothers of different status might obtain different benefits from helpers; e.g., dominant breeders might have more certainty about breeding in the same territory in the future than co-breeders and might therefore benefit more from producing helpers. Fourth, we expect dispersal to be affected by the territory quality, the number of helpers and non-helpers, the number of offspring, and the status and number of mothers within a territory. Territory quality reflects the available resources within a territory, and hence we anticipate that off-

spring opt to remain on high-quality territories and that parents are more likely to allow offspring to stay in these territories. Group size might affect the per capita resource availability and therefore dispersal. Moreover, the status of these other individuals within the territory might be of importance for dispersal tendencies. For instance, when no helper is present on a high-quality territory, offspring might stay to become helper, even when non-helping subordinates are present. Another example would be that offspring disperse early to reduce conflict or are forced earlier to leave when multiple offspring are produced within a season. Furthermore, we explored the potential effects of the status of the mother on offspring dispersal, considering the different costs and benefits associated with staying versus leaving for both categories. For example, offspring of dominant parents might have more incentive to become helpers, or dominant parents might prefer to enhance the survival of their own young versus those of co-breeders by letting their own offspring stay and potentially forcing the offspring of co-breeders disperse. Moreover, we examine the impact of the offspring sex ratio on dispersal, positing that a female-biased sex ratio may promote philopatry due to their higher probability of becoming a helper. Lastly, reviewers pointed out that previous studies found that the age of the mother might have an effect on dispersal of offspring (Ronce et al., 1998). If mothers senesce, their probability to survive to the next year decreases and hence the probability of potential offspring competing with their mothers and future siblings over resources in the next breeding season decreases as well. Therefore, older mothers are expected to have a higher probability to allow their offspring to stay. So far, empirical evidence of a correlation between maternal age and

Table 1. The complete model.

Submodel	Distribution	Response variable levels
Logit(pH) = 1 + TQ + (1 m ID) + (1 y Year)	Bernoulli	0, 1
Logit(pNH) = 1 + TQ + (1 m ID) + (1 y Year)	Ordinal cumulative	0, 1, >1
Logit(pO) = 1 + TQ + H + NH + Sm + (1 m ID) + (1 y Year)	Bernoulli	1, >1
Logit(pS) = 1 + TQ + H + NH + O + Sm + (1 m ID) + (1 y Year)	Binomial	0, 1, 2, 3, 4
Logit(pD) = 1 + TQ + H + NH + O + Sm + SR + Am + (1 m ID) + (1 y Year)	Binomial	0, 1, 2, 3

Note. Each submodel is shown with its respective model distribution, and for each model the possible levels of the response variable are given. Note that for the binomial distributions the levels are given as the number of sons, or the number of dispersing offspring, out of all offspring. TQ stands for territory quality, H for helper presence, NH for the number of non-helpers, O for the number of offspring, Sm for the status of the mother, S for the proportion of sons, D for the proportion of dispersing offspring, SR for the offspring sex ratio, and Am for the age of the mother. ID represents the maternal ID and Year represents the year.

offspring dispersal is limited (although see, e.g., Mayer et al., 2017; Ronce et al., 1998), but given that a decline in survival with age has been previously found in the Seychelles warbler (Hammers et al., 2019), we added maternal age (post hoc) to our analysis. We did not have any hypotheses about correlations between maternal age and any of the other variables, and therefore did not include more causal links in our model.

Data collection

Our analyses are based on data collected in 37 field seasons between 1996 and 2018, including all major (June to September) breeding seasons and 14 minor (February to March) breeding seasons ($n = 719$ data points from 339 mothers, who produced 808 offspring, and this data is complete for all studied traits in our analysis). Birds were caught using mist nets, and ringed with a unique combination of three colour rings and a metal ring issued by the British Trust for Ornithology. Most of the population (>95%) has been ringed since 1997 (Hammers et al., 2019; Richardson et al., 2002). Blood samples (~25 μ l) were obtained via brachial venipuncture. DNA from these samples was used for molecular sexing and parentage analysis (Sparks et al., 2022). Based on observations during each breeding season, territory boundaries, group membership, and social status (dominant breeder, helper, non-helping subordinate, or offspring) were determined. Dominant breeders were identified by contact calls, pair interactions, mate guarding, and intensive breeding effort. Helpers were identified as subordinate individuals that incubate or feed offspring, and non-helping subordinates were all other subordinates on the territory. Co-breeding females were distinguished from helpers once it was found they had produced offspring using genetic parentage analysis (Sparks et al., 2022). Only offspring that were blood sampled were included in our analysis, since DNA samples were necessary for sexing and to determine parentage. While some offspring were sampled as nestlings, many offspring were sampled after fledging. This therefore excludes any offspring that had died before the point of capture, or that were never sampled. Seychelles warblers mostly eat insects from the undersides of leaves (Komdeur, 1996b). Therefore, territory quality is estimated by combining territory size, leaf coverage, and monthly insect abundance counts (sampled on the undersides of leaves) (Spurgin et al., 2018). Territory quality estimates were calculated as $TQ = \alpha \times \sum_{x=1}^n c_x \times i_x$, where α is the territory size in hectares, c_x is the foliage cover for broad-leaved tree species x , i_x is the mean monthly insect count for tree species x per unit leaf area in dm^2 , and n is the number of different tree species in each territory (Komdeur, 1996a). These territory quality estimates were not normally

distributed, but showed a strong right-sided skew, and therefore were standardised by log transformation, mean centring, and dividing by the standard deviation.

We studied the effects of group size on the dispersal probability of offspring per subgroup (helpers, non-helping subordinates, and young offspring) because their presence might influence dispersal probabilities differently. The number of non-helpers was classified into three categories (0, 1, or >1 non-helpers), as more than two non-helpers rarely occurred in the used dataset (0.7% of territories with non-helpers had >2 non-helpers). The number of helpers was classified as helper absence or helper presence (0 or ≥ 1 helpers), again because there was little variance in the number of helpers in our dataset (3.8% of territories had more than one helper).

Statistical analyses

We fitted fully Bayesian path models based on the DAG in Figure 1. With the exception of territory quality (TQ), the status of the mother (Sm), and the age of the mother (Am), all nodes (circles) represent endogenous response variables; i.e., they have at least one directed edge (“arrow”) entering them. For each such response variable, every directly “upstream” node was entered as an additive predictor variable (at the logit scale) in a regression model. Since all response variables are discrete and many breeding females had been observed multiple times, we used generalised linear mixed models, with maternal ID as varying multivariate normal intercepts (“random effects” in frequentist terminology) that were correlated across response variables. Similarly, year was included as a varying multivariate normal intercept to avoid pseudo-replication. Specifically, we implemented the multivariate response model as depicted in Table 1.

For all response variables, logistic models were used for the probability p of observing a particular outcome; hence, logit link functions [$\text{logit}(p) = \log(p/(1 - p))$] were applied in all cases. Each response has a mean intercept indicated by the “1” directly to the right of the equality sign (see Table 1), and each response has a maternal ID-specific and a year-specific Gaussian deviation from the mean intercept, denoted by (1|m|ID) and (1|y|Year) respectively. The shared “m” in this notation indicates that a full covariance matrix was estimated for the maternal deviations, containing the variances of maternal deviations within traits, as well as covariances of maternal deviations between traits. The “y” has the same function, but for year deviations. The assumed conditional probability distributions of the response variables are also given in Table 1.

H is a binary variable that indicates helper presence at the focal female’s territory, with H = 0 encoding an absence of helpers and H = 1 the presence of at least one helper. The

notation $pH = \Pr(H = 1)$ represents the probability of observing $H = 1$. NH is an ordinal variable indicating an absence of non-helpers ($NH = 0$), the presence of a single non-helper ($NH = 1$), or the presence of multiple non-helpers ($NH > 1$). The binary variable O represents the number of offspring of the focal female, indicating a single offspring ($O = 1$) or multiple offspring ($O > 1$). S and D are binomial response variables indicating the number of sons out of all offspring, and the number of dispersing offspring out of all offspring of the focal female produced within a season, respectively. Note that H, NH, and O occur both as response and predictor variables. Additional predictor variables include the ordinal variable Sm the status of the mother, indicating that the focal female is a dominant breeder without co-breeders on the territory ($Sm = 1$), a dominant breeder with co-breeders on the territory ($Sm = 2$), or a co-breeder (in a group with a dominant breeder, $Sm = 3$). Moreover, the discrete variable age of the mother, measured in years, was included as a predictor variable, which ranged from 0 (breeding half a year later than they were born) to 16 years with a mean of 5 years. Finally, SR is a continuous predictor, quantifying the proportion of sons among the focal female's offspring (usually 0 or 1, as often only 1 offspring is produced).

All statistical analyses were conducted using R version 4.3.1 (R Core Team, 2021) with RStudio version 2023.06.0 (RStudio Team, 2020). Bayesian models were fitted using the R package brms version 2.19.0 (Bürkner, 2017), which interfaces the MCMC sampler called by the R package cmdstanr version 2.32.0 (Gabry et al., 2023). We used “weakly informative” Gaussian priors (Lemoine, 2019) for population-level effects (“fixed” effects in frequentist terminology), i.e., normal densities with mean zero and unit standard deviation. For the covariance matrix of the varying intercepts, we used the default priors of brms, i.e., half Student- t priors with 3 degrees of freedom, and LKJ(1) densities for correlation coefficients (Bürkner, 2017).

For each model we ran four chains of 4,000 iterations, including 1,000 “warm-up” iterations. Hence, 12,000 samples from the posterior were stored for analyses. Proper mixing of chains was monitored by visual inspection of trace plots and convergence of chains was verified by inspecting R-hat values, which were all close to 1.000 (two were 1.001, the rest 1.000; Monte Carlo SE were all < 0.005). Goodness-of-fit was inspected using the pp_check function of brms.

To test hypotheses, we used two approaches. First, for each model parameter we calculated the probability of direction (the pd-value), i.e., the posterior probability that the sign of a focal model parameter equals the sign of the marginal posterior distribution's median value for that parameter. A pd-value can range from 0.5 (half of the posterior distribution is on the same side of 0 as the median) to 1.0 (the whole distribution is on the same side as the median). To allow our readers to interpret the results themselves, we have shown all results with a $pd > 0.9$. Yet, for interpretation of the results, we encourage readers to look at both the pd-values and the effect sizes. There were no colliders in the DAG (Cinelli et al., 2022), allowing for estimation of all direct effects from the same model. Colliders are variables that are influenced by the predictor variable as well as by the response variable, and their inclusion in statistical models can in- or deflate the estimate of the effect of the predictor variable on the response variable. Second, for each directed arrow of the DAG, we compared the full model to a model without that arrow using the leave-one-out in-

formation criterion (LOOIC; roughly analogous to AIC/BIC), with the brms function loo_compare (Vehtari et al., 2016), which calculates the difference in LOOIC between the models, with an approximate standard error, and identifies the model, which is theoretically better at out-of-sample prediction. In our results and discussion, we use the terminology “affected by” to indicate all associations with a $pd > 0.9$. However, as this study is correlational, we do not claim to have shown causal relationships between these variables.

Results

We found pd-levels of ≥ 0.9 for the effect of territory quality on the number of non-helpers, for the effect of helper presence on the number of offspring, for the status of the mother and the number of non-helpers on the offspring sex ratio, and for the status of the mother, helper presence and territory quality on the dispersal probability of Seychelles warbler offspring. All other effects had a lower pd (< 0.9) and thus little evidence was found for these associations. As we did not use a clear cut-off for the level of significance, we encourage the reader to make their own judgement about our findings based on the effect sizes, credible intervals, and pd-values combined, especially because some results have a relatively small effect size. Table 2 shows a summary of all results, including the median of the posterior distribution, credible intervals, pd-values, and the output of the leave-one-out comparison. The posterior distributions of all estimated effects are shown in Figure 2, and these results are summarised in a DAG in Figure 3. The effects of all variables with a $pd > 0.9$ are shown in Figure 4, and next we will discuss these effects. For graphs showing the raw data, please see Figure S1A. All the given model estimates (and 95% credible intervals shown in brackets) in the text next are conditional effects; e.g., the values of the offspring sex ratio produced by co-breeders versus dominant females are given conditional on the average values of the other traits.

Helper presence and number of non-helpers

Helper presence was not affected by territory quality, but fewer non-helping subordinates were found with increasing territory quality (Figures 2 and 4A and Table 2).

Number of offspring

The probability to produce more than one offspring was higher when helpers were present than when helpers were absent (Figures 2 and 4B and Table 2), with model estimates of 0.08 (0.02–0.21) versus 0.05 (0.01–0.13), respectively.

Offspring sex ratio

Dominant females with co-breeders and co-breeders both produced more sons than dominant females without co-breeders (Figures 2 and 4C and Table 2), with the estimated probabilities of producing a son being 0.56 (0.44–0.68) for females with co-breeders, 0.55 (0.42–0.68) for co-breeders, and 0.47 (0.39–0.56) for dominant females without co-breeders. Additionally, when more than one non-helper was present on a territory, the offspring sex ratio was more male-biased (Figures 2 and 4G and Table 2). The estimated probability to produce a son changed from 0.51 (0.44–0.58) for zero non-helper and 0.47 (0.39–0.56) for one non-helper to 0.65 (0.50–0.78) for more than one non-helper.

Table 2. The complete model output.

Variable	Posterior median	Credible interval	pd	Complete model wins	Difference	SE of difference
Effects on helper presence						
Territory quality	0.01	-0.18 : 0.20	0.54	No	-0.5	0.3
Effects on number of non-helpers						
Territory quality	-0.17	-0.34 : 0.00	0.97	Yes	-1.0	2.0
Effects on number of offspring						
Territory quality	0.18	-0.13 : 0.50	0.87	Yes	-1.1	1.4
Helper presence	0.50	-0.09 : 1.10	0.95	Yes	0.0	1.7
Number of non-helpers (compared to 0 non-helper)				No	-1.9	0.5
1 non-helper	0.04	-0.61 : 0.67	0.54			
≥2 non-helpers	-0.10	-1.14 : 0.88	0.58			
Status of the mother (compared to dominant breeders without a co-breeder)				No	-2.1	0.5
Dominant breeder with co-breeder	-0.08	-0.79 : 0.60	0.59			
Co-breeder	-0.03	-0.83 : 0.71	0.71			
Effects on the offspring sex ratio						
Territory quality	0.004	-0.15 : 0.16	0.52	No	-1.0	0.3
Helper presence	0.03	-0.33 : 0.38	0.56	No	-1.2	0.4
Number of non-helpers (compared to 0 non-helper)				Yes	-0.6	2.2
1 non-helper	-0.15	-0.50 : 0.20	0.80			
2 non-helpers	0.59	-0.04 : 1.22	0.97			
>1 offspring (compared to 1 offspring)	-0.09	-0.51 : 0.32	0.67	No	-0.8	0.5
Status of the mother (compared to dominant breeders without a co-breeder)				No	-0.3	2.0
Dominant breeder with co-breeder	0.37	-0.05 : 0.79	0.96			
Co-breeder	0.32	-0.14 : 0.79	0.92			
Effects on dispersal probability						
Territory quality	0.12	-0.03 : 0.27	0.93	No	-0.2	1.5
Helper presence	-0.50	-0.86 : -0.15	1.00	Yes	-3.3	2.8
Number of non-helpers (compared to 0 non-helper)				No	-1.3	1.2
1 non-helper	-0.19	-0.57 : 0.16	0.86			
2 non-helpers	0.06	-0.61 : 0.71	0.57			
>1 offspring (compared to 1 offspring)	0.10	-0.33 : 0.53	0.68	No	-0.9	0.6
Status of the mother (compared to dominant breeders without a co-breeder)				Yes	-5.2	3.6
Dominant breeder with co-breeder	0.08	-0.34 : 0.51	0.67			
Co-breeder	0.89	0.39 : 1.42	1.00			
Offspring sex ratio	-0.11	-0.44 : 0.22	0.75	No	-1.0	0.7
Age of the mother	0.003	-0.05 : 0.06	0.54	No	-1.0	0.3

Note. For each submodel and variable, the median of the posterior distribution is given as well as the 95% credible interval and the probability of direction (pd). Additionally, we used the leave-one-out criterion to compare if a model with or without a specific arrow would improve model fit. Here, we indicate whether the complete model wins in this comparison, what the difference is, and what the standard error of that difference is.

Dispersal

Offspring of co-breeders had a higher probability of dispersing (estimate = 0.75, credible interval = 0.62–0.84) compared to those of dominant breeders. The offspring of dominant breeders with co-breeders (estimate = 0.57, credible interval = 0.44–0.69) did not differ substantially from dominant breeders without co-breeders (estimate = 0.55, credible interval = 0.46–0.63) (see [Figures 2](#) and [4F](#) and [Table 2](#)). On territories where helpers were present, offspring had a lower probability of dispersing (estimate = 0.42, credible interval = 0.32–0.54; vs. not present estimate = 0.55, credible interval = 0.46–0.63) (see [Figures 2](#) and [4E](#) and [Table 2](#)). Lastly, offspring dispersal probability increased with increasing territory quality ([Figures 2](#) and [4D](#) and [Table 2](#)).

Discussion

In this study, we found that Seychelles warbler offspring were more likely to disperse when their mothers were co-breeders, territory quality was higher, and helpers were absent. The off-

spring sex ratio of the focal mother was more male-biased when more than one non-helper was present on the territory, and when the mother was a dominant breeding female with a co-breeder. More offspring were born per mother when helpers were present. The number of non-helping subordinates on a territory decreased with increasing territory quality.

The differential dispersal tendencies of offspring from co-breeding and dominant mothers suggest a potential parental influence on dispersal tendencies. This is an intriguing result, as so far no clear evidence exists that birds can recognise genetic kin ([Kempnaers & Sheldon, 1996](#); [Lattore et al., 2019](#)). Possible alternative explanations for the difference in dispersal tendencies include temporal differences in breeding activity between dominant breeders and co-breeders, leading to competitive disparities among offspring, which then could lead to weaker individuals being forced to leave the territory. Co-breeders might lay their egg in the same clutch, but consistently after the dominant breeder, resulting in offspring of dominant breeders having a competitive advantage. Alternatively, co-breeders might lay their egg at a different moment in the breeding season, which could also lead to a compet-

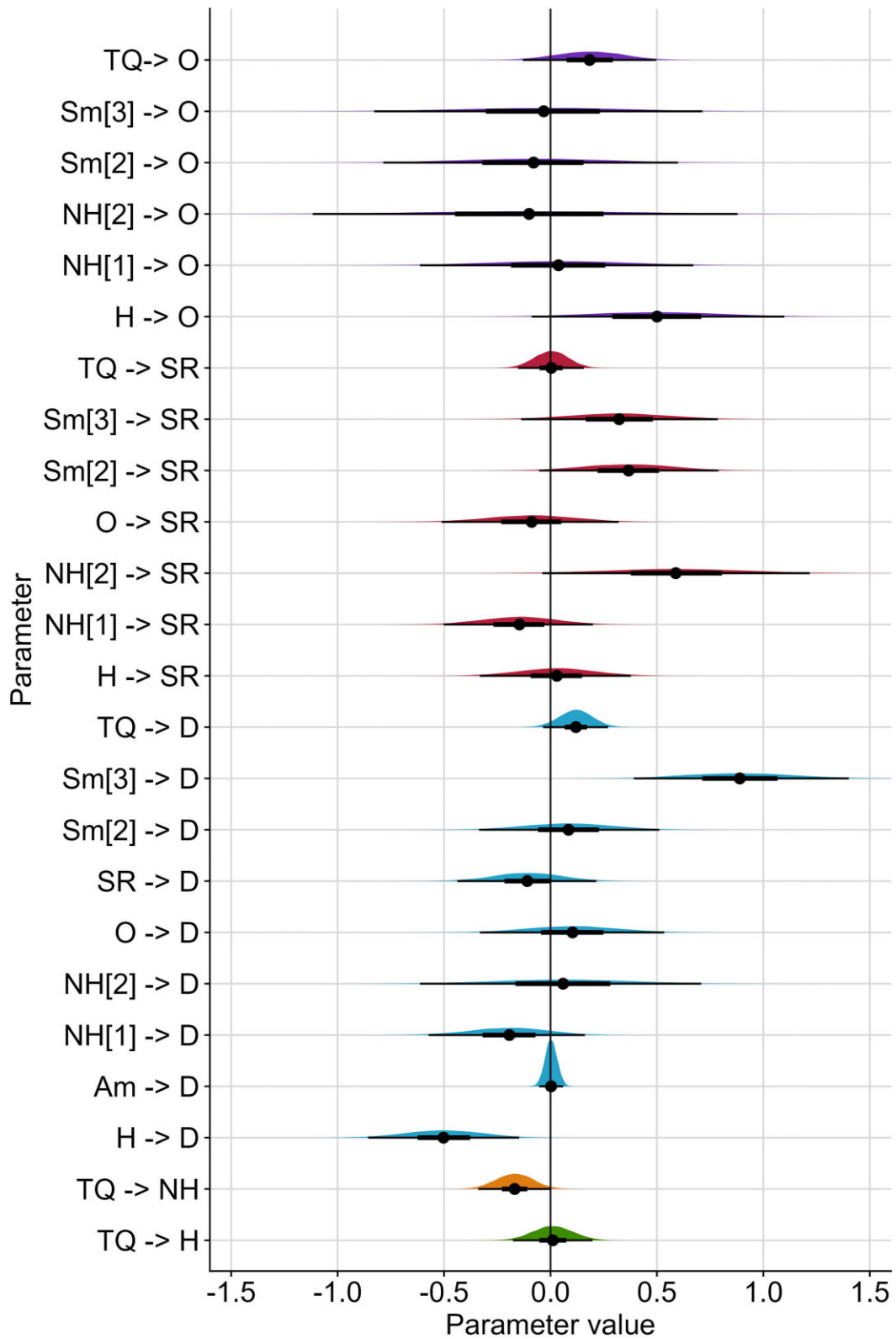


Figure 2. Posterior distributions of all estimated associations. The y-axis shows the associations from which the regression coefficient is estimated. The coloured density plots show the 98% credible interval of this distribution, the thick error bars show the 50% credible intervals (and the thin error bars the 95% credible intervals) of the distribution, and the black points show the median. Number of non-helpers and the status of the mother are variables with three categories, and hence the distributions show the difference between the first category and the other two categories [e.g., Sm[2] compares between categories 1 (breeding female without co-breeder) and 2 (breeding female with co-breeder) of status of the mother]. For abbreviations on the y-axis, see Figure 1.

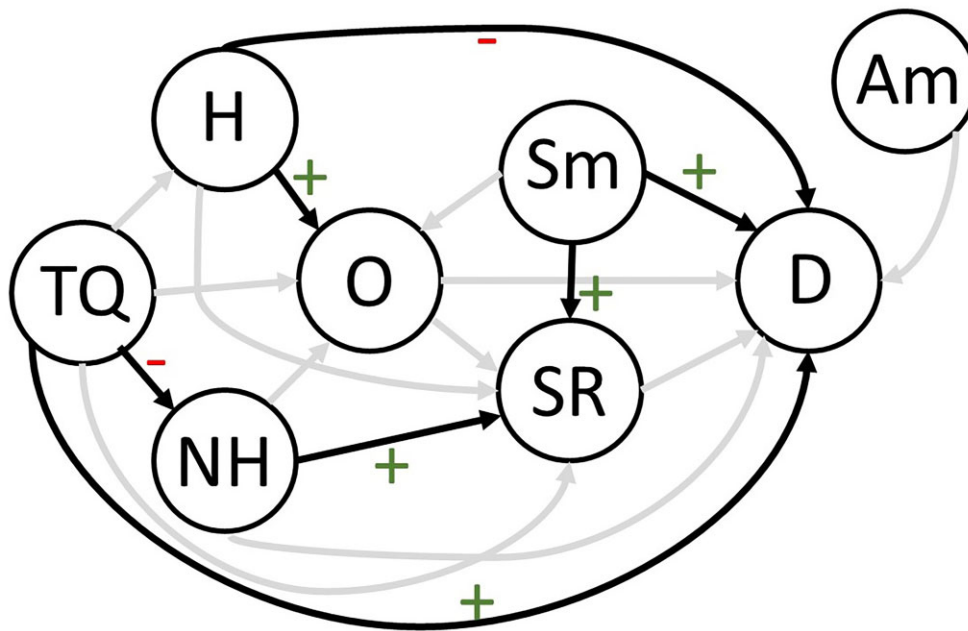


Figure 3. Directed acyclic graph of the hypothetical causal network, showing all results with $pd \geq 0.9$. All direct effects with a $pd < 0.9$ are indicated with light grey lines, while all direct effects with a $pd \geq 0.9$ are indicated with a black line, including a positive (green) or negative (red) sign for the direction of the effect.

itive disparity between dominant and co-breeder offspring. However, the precise mechanisms underlying these differences warrant further investigation. A previous finding in the Seychelles warbler also suggests that dominant breeders have some control on who is residing on the territory. Eikenaar et al. (2007) showed that offspring were more likely to disperse when one or both of the dominant breeders had died or had been replaced. However, in that study the difference in dispersal tendency can also be explained by a difference in familiarity and is not necessarily caused by kin recognition. Alternatively, the difference in dispersal tendency could be explained by a difference in the willingness of offspring to stay when they have a low relatedness to the dominant breeding female. A decrease in willingness to become a helper is expected from a kin-selection perspective, yet it is also possible for individuals to stay and not help. If delaying dispersal is a way to increase condition (as discussed later), then there is no apparent reason for offspring of co-breeders to decrease their willingness to stay. Hence, more research on differences between offspring from dominant and co-breeding females is necessary. Higher dispersal tendencies when there is low relatedness with dominant breeders is also found in Florida scrub-jays (*Aphelocoma coerulescens*; Goldstein et al., 1998; Suh et al., 2020), Siberian jays (*Perisoreus infaustus*; Ekman & Griesser, 2002), and southern pied babblers (*Turdoides bicolor*; Nelson-Flower & Ridley, 2016).

Contrary to our expectations, offspring exhibited higher dispersal probabilities when hatched on high-quality territories. We hypothesise that delaying dispersal could serve as a strategy to enhance individual condition, if dominants tolerate offspring on their territory (Ekman & Griesser, 2002). This might be less necessary for offspring born on high-quality territories, as they might reach a good condition faster, and hence can successfully compete for a position elsewhere. Similar results were found in red kites (*Milvus milvus*), which dispersed earlier in life when food availability was artificially increased

(Scherler et al., 2023). Yet, opposite patterns have also been found in the Florida scrub-jay (*A. coerulescens*; Suh et al., 2020). However, this correlation might differ on the within-territory level compared to the between-territory level. Within territories, offspring of lowest quality might be outcompeted by siblings and thus forced to leave (Dietz et al., 2022), while at a between-territory level, higher-quality offspring might disperse more as they can outcompete others for a position elsewhere. If delaying dispersal is a way of improving condition, it could also explain why fewer non-helpers were found with increasing territory quality. We expect that the condition of offspring may be important for their dispersal behaviour and especially their probability to successfully disperse. However, unfortunately we were not able to include a proxy of offspring condition in our model, as we did not have reliable data for most of our sampled offspring, and hence more research is necessary to understand the effect of condition and its links with territory quality and dispersal. Lastly, offspring were less likely to disperse when helpers were present, which is puzzling, especially as it cannot be caused by an indirect territory quality effect.

Interestingly, we were unable to reproduce the results of Komdeur et al. (1997), who found that the offspring sex ratio correlated with territory quality (more daughters were produced in high-quality territories and more sons in low-quality territories) and with the number of helpers (when two or more helpers were present, sons were produced, independent of territory quality) in the same study population we used, for the years 1993–1995. However, in our study (using data from 1996–2018), we did not find any effect of territory quality and helper presence on the offspring sex ratio (see Table 2, and Figure S2A for a graphical representation of these effects). Yet, we did not test whether our and Komdeur et al.'s results significantly differed from each other, because not all data necessary for our model are available for the years used in Komdeur et al.'s study (1993–1995; e.g., genetic verification of

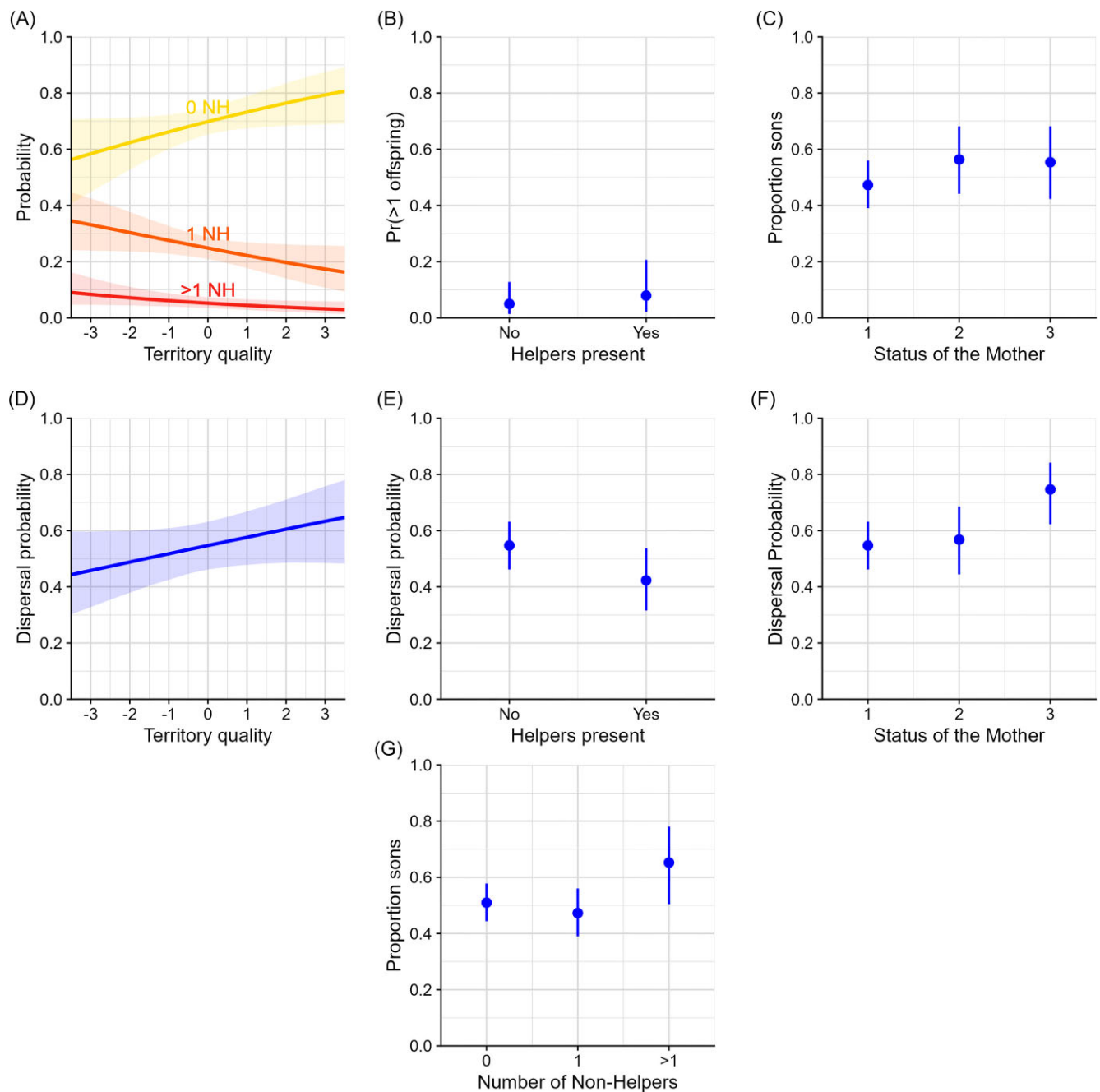


Figure 4. All effects of the causal network with a $pd > 0.9$. (A) The probability of zero non-helper (0 NH), one non-helper (1 NH), or more than one non-helper (>1 NH) is affected by territory quality. (B) The probability of a mother to produce more than one offspring increases when helpers are present compared to when helpers are absent. (C) The proportion of sons a mother produces is affected by the status of the mother. (D) The probability that offspring disperses increases with increasing territory quality. (E) The dispersal probability of offspring decreases when helpers are present compared to when these are absent. (F) The probability that offspring disperse is affected by the status of the mother. (G) The proportion of sons a mother produces is affected by the number of non-helping subordinates. The shadows in panels (A) and (D) indicate the 95% credible intervals. In the other graphs, the median values of the distributions are indicated with the dot, and the error bars indicate the 95% credible intervals. In panels (C) and (F), the status of the mother is indicated by numbers, 1 indicates a dominant breeding female without co-breeders, 2 indicates a dominant breeding female with co-breeders, and 3 indicates co-breeders.

number of offspring and co-breeder status). These different results could be caused by differences in territory quality between both study periods. Since the island was restored in 1986, the natural vegetation has been improving in quality and volume. Hence, all territories could have improved in quality and variance in territory quality may have reduced over the years, and consequently it is possible that a poten-

tial adaptive sex ratio strategy has changed over time. Since many years were included in our analysis, the effect of territory quality on offspring sex ratio could have changed, or weakened, resulting in no “significant” effect of territory quality on offspring sex ratios. In our dataset, the territory quality measure did not seem to improve over years (see [Figure S3A](#) and [Table S3A](#)), but in [Komdeur et al. \(2016\)](#) an increase in

territory quality was found over time, where years were divided into blocks, while we did not do so. We also ran our model on subsets of the data to check for a temporal disappearance or weakening of this trend (see [Supplement 4](#) for the results; 1996–2000 in [Table S4A](#), 2001–2005 in [Table S4B](#), 2006–2010 in [Table S4C](#), and 2011–2015 in [Table S4D](#)) and found no effect of territory quality or helper presence on the sex ratio in any of these subsets. Thus, a disappearance or weakening of this effect cannot be concluded. Alternatively, the difference in timing of measurement of the offspring sex ratio between [Komdeur et al. \(1997\)](#) and our study could cause the different results, as we only included chicks that are part of the pedigree and therefore generally at least have survived until fledging. In contrast, [Komdeur et al. \(1997\)](#) looked at nestlings. Hence, any effect could have disappeared due to sex-specific mortality in early development, though this is impossible to study in new data, as most nests are not reachable (currently most nests are over 8 m high in extremely soft wood trees). Moreover, [Komdeur et al. \(1997\)](#) did not distinguish between helpers and co-breeders, as parentage analysis was started later, and there was less information about the helping behaviour of individuals, which could have made a distinction between helpers and non-helping subordinates more difficult. In conclusion, more research is necessary to fully understand offspring sex ratio trends in the Seychelles warbler.

We also found that mothers produced more offspring when helpers were present, which is in accordance with previous findings ([Komdeur, 1994](#)). Such a correlation has also been found in cichlids (*Neolamprologus obscurus*; [Tanaka et al., 2018](#)) and meerkats (*Suricata suricatta*; [Russell et al., 2003](#)).

Identifying which factors are associated with dispersal decisions when only observational data can be collected should be done using path analyses or structural equation models when it is expected that these factors could also be causally linked among each other ([Busana, 2021](#); [Streiner, 2005](#); [Wright, 1934](#)). Alternatively, a causal inference justification should be provided about which factors should be included (e.g., confounders) or excluded (e.g., colliders) in the model ([Cinelli et al., 2022](#); [McElreath, 2020](#)). These methods can disentangle direct and indirect effects of each factor, and can separate correlated factors. Thus, they provide additional insights compared to modelling responses separately within linear models, and interesting insights can be obtained when total effects show different patterns than direct effects. Moreover, hypotheses that explicitly mention causal pathways and whether direct or total effects are studied are easier to test (as their underlying assumptions are communicated more clearly) and hence can improve knowledge on dispersal decisions and the evolution of cooperation. Additionally, DAGs can improve science communication about the (hypothesised) ecology of a species, and therefore can highlight the differences and similarities between different cooperative breeding species, and thus can enhance our general understanding of the evolution of cooperative breeding and sociality.

In conclusion, using path analyses, we found that individuals in fact disperse more when territory quality is higher, that the status of the mother could be important for dispersal decisions, and that we could not reproduce any sex ratio trends in the Seychelles warbler. Studying proximate factors influencing dispersal (with appropriate statistical tools) is thus important to disentangle the actual causes of delayed dispersal, and eventually the underlying mechanisms causing the evolution of cooperative breeding.

Supplementary material

Supplementary material is available at *Journal of Evolutionary Biology* online.

Data availability

The data are stored on the Dataverse of the University of Groningen, and are available using the link: <https://doi.org/10.34894/LLSZWJ>.

Author contributions

Mirjam J. Borger (Conceptualization [equal], Formal analysis [equal], Investigation [equal], Methodology [equal], Software [equal], Supervision [equal], Validation [equal], Visualization [equal], Writing—original draft [equal], Writing—review & editing [equal]), Kiran G.L. Lee (Conceptualization [equal], Formal analysis [equal], Writing—original draft [equal]), Franz J. Weissing (Methodology [equal], Resources [equal], Visualization [equal], Writing—review & editing [equal]), David S. Richardson (Data curation [equal], Funding acquisition [equal], Resources [equal], Writing—review & editing [equal]), Hannah L. Dugdale (Data curation [equal], Funding acquisition [equal], Writing—review & editing [equal]), Terry Burke (Data curation [equal], Funding acquisition [equal], Project administration [equal], Resources [equal], Writing—review & editing [equal]), Ido Pen (Conceptualization [equal], Formal analysis [equal], Investigation [equal], Methodology [equal], Software [equal], Supervision [equal], Validation [equal], Visualization [equal], Writing—review & editing [equal]), and Jan Komdeur (Data curation [equal], Funding acquisition [equal], Project administration [equal], Resources [equal], Supervision [equal], Writing—review & editing [equal])

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Conflicts of interest

None declared.

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