

This is a repository copy of Monitoring for 5-aminosalicylate nephrotoxicity in adults with inflammatory bowel disease: prognostic model development and validation using data from the Clinical Practice Research Datalink.

White Rose Research Online URL for this paper: <u>https://eprints.whiterose.ac.uk/222483/</u>

Version: Published Version

Article:

Abhishek, A., Nakafero, G. orcid.org/0000-0002-3859-7354, Card, T. orcid.org/0000-0003-2555-2250 et al. (6 more authors) (2025) Monitoring for 5-aminosalicylate nephrotoxicity in adults with inflammatory bowel disease: prognostic model development and validation using data from the Clinical Practice Research Datalink. BMJ Open Gastroenterology, 12. e001627. ISSN 2054-4774

https://doi.org/10.1136/bmjgast-2024-001627

Reuse

This article is distributed under the terms of the Creative Commons Attribution (CC BY) licence. This licence allows you to distribute, remix, tweak, and build upon the work, even commercially, as long as you credit the authors for the original work. More information and the full terms of the licence here: https://creativecommons.org/licenses/

Takedown

If you consider content in White Rose Research Online to be in breach of UK law, please notify us by emailing eprints@whiterose.ac.uk including the URL of the record and the reason for the withdrawal request.



eprints@whiterose.ac.uk https://eprints.whiterose.ac.uk/ BMJ Open Gastroenterology **Monitoring for 5-aminosalicylate** nephrotoxicity in adults with inflammatory bowel disease: prognostic model development and validation using data from the Clinical Practice **Research Datalink**

Abhishek Abhishek,^{1,2} Georgina Nakafero ⁽¹⁾,^{1,2} Tim Card ⁽¹⁾, ³ Maarten W Taal,⁴ Matthew J Grainge,³ Guruprasad P Aithal,^{2,5} Christian D Mallen,⁶ Matthew D Stevenson,⁷ Richard D Riley^{8,9}

To cite: Abhishek A, Nakafero G, Card T. et al. Monitoring for 5-aminosalicylate nephrotoxicity in adults with inflammatory bowel disease: prognostic model development and validation using data from the Clinical Practice Research Datalink. BMJ Open Gastroenterol 2025;12:e001627. doi:10.1136/ bmjgast-2024-001627

 Additional supplemental material is published online only. To view, please visit the journal online (https://doi. org/10.1136/bmjgast-2024-001627).

AA and GN are joint first authors.

Received 8 October 2024 Accepted 7 January 2025

Check for updates

C Author(s) (or their employer(s)) 2025. Re-use permitted under CC BY. Published by BMJ Group.

For numbered affiliations see end of article.

Correspondence to

Dr Georgina Nakafero; georgina.nakafero@ nottingham.ac.uk

ABSTRACT

Objective To develop and validate a prognostic model for risk-stratified monitoring of 5-aminosalicylate nephrotoxicity.

Methods This UK retrospective cohort study used data from the Clinical Practice Research Datalink Aurum and Gold for model development and validation respectively. It included adults newly diagnosed with inflammatory bowel disease and established on 5-aminosalicylic acid (5-ASA) treatment between 1 January 2007 and 31 December 2019. Drug discontinuation associated with 5-ASA nephrotoxicity defined as a prescription gap of \geq 90 days with decline in kidney function was the outcome. Patients prescribed 5-ASAs for \geq 6 months were followed-up for up to 5 years. Penalised Cox regression was used to develop the risk equation with bootstrapping for internal validation and optimism adjustment. Model performance was assessed in terms of calibration and discrimination.

Results 13728 and 7318 participants who contributed 40 378 and 20 679 person-years followup formed the development and validation cohorts with 170 (1.2%) and 98 (1.3%) outcome events respectively. Nine predictors were included in the final model, including chronic kidney disease stage 3 and hazardous alcohol use as strong predictors. Age and Body Mass Index were weak predictors. The optimismadjusted calibration slope, C and D statistics in the development and validation data were 0.90, 0.64 and 0.98, and 1.01, 0.66 and 0.94 respectively. Conclusion This prognostic model used information from routine clinical care and performed well in an independent validation cohort. It can be used to riskstratify blood test monitoring during established 5-ASA treatment. A key limitation is that the decline in kidney function could have been due to factors other than 5-ASA nephrotoxicity.

WHAT IS ALREADY KNOWN ON THIS TOPIC

 \Rightarrow Nephrotoxicity is an uncommon but well-recognised side effect during long-term 5-aminosalicylate treatment for which there are no mechanisms to risk-stratify monitoring.

WHAT THIS STUDY ADDS

- \Rightarrow This study developed and externally validated a prognostic model to estimate the risk of nephrotoxicity due to 5-aminosalicylates during established treatment using a large dataset from the UK that originated during routine clinical care.
- \Rightarrow Most patients were at low risk of 5-aminosalicylate nephrotoxicity and could continue with annual monitoring blood tests while others at high risk may reauire frequent monitorina.
- \Rightarrow This prognostic model can be used to help inform decisions on the interval between monitoring blood tests, thereby, improving patient safety.

HOW THIS STUDY MIGHT AFFECT RESEARCH, **PRACTICE OR POLICY**

 \Rightarrow These findings should be considered by guideline writing groups to provide risk-based recommendations for the frequency of monitoring of kidney function in people with inflammatory bowel disease receiving treatment with 5-aminosalicylic acid.

INTRODUCTION

5-aminosalicylates (5-ASA) are the mainstay of treatment for mild to moderate ulcerative colitis and are often combined with biologics in the treatment of severe disease.¹⁻³ Although effective, they can cause potentially serious adverse effects such as acute interstitial nephritis.⁴⁻⁹ Clinically significant interstitial nephritis due to 5-ASA drugs is estimated to

occur in 1 in 500 patients and may reverse on drug discontinuation.^{7 10} This makes screening for renal toxicity important. Although 5-ASA-induced interstitial nephritis is more common in the first year of treatment,^{8 9} it often occurs later in the treatment course with a median time to onset of between 2.3 and 3 years reported.⁶⁷

Patients established on 5-ASA treatment are recommended to undergo periodic monitoring of kidney function. There is currently no consensus on the frequency of testing due to lack of high-quality data. The British Society of Gastroenterology (BSG) guideline and the British National Formulary (BNF) recommend monitoring of kidney function at annual intervals in those established on treatment, with consideration of increased frequency in those with impaired renal function.^{1 11} Others advise either 6 monthly, or annual or 'periodic monitoring at the clinician's discretion'.^{1-3 12} The recommendation to monitor is often not followed, potentially, due to the rarity of 5-ASA nephrotoxicity. For example, only a fifth of patients underwent monitoring blood tests in a large UK study.¹³

Current recommendations generally adopt a one-sizefits-all approach to monitoring for 5-ASA nephrotoxicity. This is potentially a wasteful use of resources while disadvantaging others at high risk. It would be beneficial to predict nephrotoxicity during established 5-ASA treatment to allow risk-stratified monitoring. In this study, we developed and validated a prognostic model for 5-ASA nephrotoxicity during established treatment to aid riskstratified long-term monitoring strategies.

METHODS Data source

Clinical Practice Research Datalink (CPRD).^{14 15} CPRD is an anonymised longitudinal database of electronic health records collected during routine clinical care in the National Health Service. With almost universal coverage of UK residents, participants that contributed data to the CPRD are representative of the UK population.¹⁴ CPRD Aurum and CPRD Gold are two separate CPRD data sets which use different practice management software and complement each other in terms of coverage of general practices. CPRD Aurum only includes practices from England while GOLD contains practices from other parts of the UK. For the purposes of this study, CPRD Aurum and Gold were used for model development and validation, respectively. Some general practices which have contributed data to both databases at different time were only included in the model development cohort.

Study design

Retrospective cohort study.

Study period

1 January 2007 to 31 December 2019.

Study population

Participants aged \geq 18 years with a new diagnosis of inflammatory bowel disease (IBD) and newly prescribed 5-ASA by their general practitioner (GP) for at least 180 days were eligible. Patients with chronic kidney disease (CKD) stage \geq 4 prior to first 5-ASA prescription were excluded (online supplemental figures S1 and S2).

Follow-up

Patients were followed-up from 180 days after their first GP prescription until the earliest of outcome, death, transfer out of practice, 90-day prescription gap, last data collection from practice, 31 December 2019 or 5 years from the start of follow-up.

Follow-up began 180 days after first GP prescription as our purpose was to predict the risk of 5-ASA nephrotoxicity during established treatment, a period in which there is greatest need to risk-stratify monitoring as the rate of nephrotoxicity is lower than early in the treatment course.

In the UK, 5-ASA drugs are commenced in hospital clinics. Hospital specialists prescribe and monitor these drugs till a stable well-tolerated dose is reached, and the monitoring blood tests are satisfactory. This takes approximately 6 months. After this, the responsibility for prescribing and monitoring is handed over to GPs. Consequently, we started follow-up from 180 days after the first GP prescription to only include patients treated with 5-ASA drugs for approximately 1 year.

Outcome

5-ASA nephrotoxicity-associated drug discontinuation was the outcome of interest. This was defined as a prescription gap of \geq 90 days with decline in kidney function, defined as either progression of CKD based on medical codes recorded by the GP, or >26 µmol/L increase in creatinine concentration, the threshold for consideration of acute kidney injury (AKI)¹⁶ within ±60 days of the last prescription.

To assess the positive predictive value of the outcome, a random sample of 5-ASA discontinuations with abnormal blood test results was drawn. Data for all diagnostic codes entered during primary-care consultations within ± 60 days of the abnormal blood test result were extracted. A.A. screened the list to identify outcomes that could potentially be explained by an alternative condition or its treatment.

Candidate predictors

Candidate predictors were selected based on clinical expertise and knowledge of the published literature (table 1). Age, sex Body Mass Index (BMI), alcohol intake and diabetes were included as they are associated with worsening renal function.¹⁷ Additionally, diabetes has been associated with 5-ASA nephrotoxicity.¹⁸ CKD stage 3 was included as it reduces 5-ASA clearance and has been associated with 5-ASA nephrotoxicity.^{11 19} Prescriptions of ACE inhibitors, aspirin and non-steroidal

| Table 1 | Distribution of candidate predictors in | | | |
|------------------------------------|---|--|--|--|
| development and validation cohorts | | | | |

| Predictor* | Development cohort (CPRD Aurum) n=13728 | Validation cohort (CPRD Gold) n=7318 | | | | |
|-----------------------------------|---|--|--|--|--|--|
| Age, mean (SD) year | 46 (18) | 47 (18) | | | | |
| Male sex | 7075 (51.5) | 3772 (51.5) | | | | |
| Body Mass Index | | | | | | |
| <18.5 kg/m ² | 430 (3.1) | 229 (3.1) | | | | |
| 18.5–24.9 kg/m ² | 4914 (35.8) | 2480 (33.9) | | | | |
| 25.0-29.9 kg/m ² | 3934 (28.7) | 2172 (29.7) | | | | |
| \geq 30 kg/m ² | 2563 (18.7) | 1436 (19.6) | | | | |
| Missing | 1887 (13.8) | 1001 (13.7) | | | | |
| Alcohol use | | | | | | |
| Non-user | 298 (16.7) | 910 (12.4) | | | | |
| Low (1–14 units/week) | 5688 (41.4) | 3707 (50.7) | | | | |
| Moderate (15-21 units/week) | 819 (6.0) | 421 (5.8) | | | | |
| Hazardous (>21 units/week) | 986 (7.2) | 400 (5.5) | | | | |
| Ex-user | 1041 (7.6) | 384 (5.3) | | | | |
| Missing | 2896 (21.1) | 1496 (20.4) | | | | |
| Comorbidities | | | | | | |
| Diabetes | 1043 (7.6) | 554 (7.6) | | | | |
| Chronic kidney disease stage 3 | 797 (5.8) | 399 (5.4) | | | | |
| Other drugs | | | | | | |
| ACE inhibitors | 1279 (9.3) | 732 (10.0) | | | | |
| Aspirin | 991 (7.2) | 621 (8.5) | | | | |
| NSAIDs including COX-2 inhibitors | 635 (4.6) | 308 (4.2) | | | | |

*Values are numbers (percentage) unless stated otherwise

CPRD, Clinical Practice Research Datalink; NSAIDs, non-steroidal anti-inflammatory drugs.

anti-inflammatory drugs (including COX-2 inhibitors) were considered as their use is associated with 5-ASA nephrotoxicity as per the British National Formulary¹¹ and a previous study.¹⁸ 5-ASA dose was not included in the model as 5-ASA-induced nephrotoxicity is reported to be idiosyncratic and not dose-related.⁹

Sample size

The reported incidence of interstitial nephritis due to 5-ASAs ranged between 1.2 and 1.7 per 1000 personyears.^{4 20} Using the formulae of Riley *et al*,²¹ the minimum sample size required for new model development was 1049 participants (43 events) based on a maximum of 15 parameters, and assumed Cox-Snell R² value of 0.12, estimated event rate of 0.005/person-year and a mean follow-up period of 3 years to minimise model overfitting and ensure precise estimation of overall risk at 5 years. Our sample size within CPRD far exceeded this (online supplemental figures S1 and S2).

Statistical analysis

Multiple imputations handled missing data on BMI and alcohol intake using chained equations.²² We carried out 10 imputations in the development dataset and 5 imputations in the validation dataset—a pragmatic approach

considering the large size of CPRD. The imputation model included all candidate predictors, baseline Nelson-Aalen cumulative hazard function and outcome variable.

Model development: All candidate predictors and parameters were included in the Cox model (ie, no variable selection was used) and coefficients of each parameter were estimated and combined using Rubin's rule across the imputed datasets. We compared nested models containing each continuous predictor in turn with that containing the best-fitting first-order polynomial. It was found that for both age and BMI, fractional terms did not improve model fit (p value=0.217 and 0.359 respectively); hence, these variables were not transformed. Patients who died were censored at their death; thus, the model gives estimated 5-year risks conditional on being alive at 5 years which is highly likely for most patients, see Discussion.

The risk equation for predicting an individual's risk of 5-ASA discontinuation with nephrotoxicity-associated drug discontinuation by 5 years follow-up was formulated in the development data. The baseline survival function at t=5 years, a non-parametric estimate of survival function when all predictor values are set to zero, which is equivalent to the Kaplan-Meier product-limit estimate, was estimated along with the estimated Cox regression coefficients (β). This led to the equation for the predicted absolute risk over time.²³

Predicted risk=1 – $S_0(t_{=5})^{exp(X\beta)}$, where $S_0(t_{=5})$ is the baseline survival function at 5 years of follow-up and βX is the linear predictor, $\beta_1 x_1 + \beta_2 x_2 + \ldots + \beta_p x_p$, with the individual's predictor values denoted by X.

Model internal validation and shrinkage

Apparent performance (in the development dataset) of the model was assessed in terms of calibration and discrimination. Calibration examines the agreement between predicted and observed risks and can be summarised by a calibration plot (with smooth calibration curve) and calibration slope (ideal value of 1). Discrimination was measured using the Harrell's C-statistic, a measure of the model's predictive accuracy, and Royston D statistic, interpreted as a log HR comparing the outcome rate of two groups defined by the median of the model's linear predictor.^{24 25} Overall fit was quantified by $R^2_{\ D}$, a measure of variation explained by the model.

Internal validation was performed using bootstrapping with replacement, with 500 bootstrap samples used to estimate and adjust performance estimates for optimism due to potential overfitting. The full model was fitted in each bootstrap sample and then its performance quantified in the bootstrap sample (apparent performance) and the original sample (test model performance), and the optimism calculated (difference in test performance and apparent performance) for each measure of calibration and discrimination performance. A uniform shrinkage factor was estimated as the average of optimism-adjusted calibration slopes from the bootstrap sample. This process was repeated for all 10 imputed datasets, and the final uniform shrinkage calculated by averaging across the estimated shrinkage estimates from each imputation. Optimism-adjusted estimates of performance for the original model were then calculated, as the original apparent performance minus the average optimism.

To account for overfitting, the original β coefficients were multiplied by the final uniform shrinkage factor and the baseline hazard re-estimated with the shrunken β coefficients as an offset; this ensured that overall calibration was maintained and produced the final model.

Model external validation

The final developed model equation was applied to individuals in the validation dataset, and calibration and discrimination were examined.^{24 25} Calibration of 5-year risks was examined by plotting agreement between estimated risk from the model and 'observed' outcome risks. Pseudo-observations were used to construct smooth calibration curves across all individuals via a running non-parametric smoother. We also examined predicted and observed risks within 10 equally sized groups defined by tenths of predicted risk. Separate calibration plots were considered for each imputation. Stata-MP version 17 was used for statistical analyses. This study was reported in line with the TRIPOD guidelines (see TRIPOD checklist [online supplemental]).²⁶

RESULTS

Participants

Data for 13728 and 7318 participants that contributed 40378 and 20679 person-years follow-up were included

in the development and validation cohorts, respectively (online supplemental figures S1 and S2). In the development cohort, 9096 (66%) had ulcerative colitis, 2769 (20%) had Crohn's disease, and 1863 (14%) had IBD without any specific coding for subtype. In the validation cohort, 4449 (61%) had ulcerative colitis, 1539 (21%) had Crohn's disease, and 1330 (18%) had IBD without any specific coding for subtype. Over 50% were recorded as male and had similar lifestyle factors, comorbidities and drug treatments (table 1). Nine candidate predictors (12 predictor parameters) were included in the model (table 2). We were unable to include lithium, ciclosporin, voclosporin and methotrexate, drugs associated with 5-ASA nephrotoxicity,^{11 18} as potential predictors in the prognostic model because there were a small number of patients prescribed these drugs in the development cohort and none of them had the outcome of interest.

Model development

In the development dataset, there were 25 (0.18%) deaths. 170 outcomes occurred in n=13728 patients (1.2%) during follow-up at a rate (95% CI) of 4.21 (3.62–4.89) per 1000 person-years. These outcomes occurred at a similar rate throughout the follow-up period (online supplemental figure S3). The median (IQR) time to occurrence of outcome was 1.66 (0.80–3.04) years. Outcome validation exercise revealed that 13.7% outcomes (n=10/73) could potentially be explained by another contemporaneous illness (vomiting, diarrhoea, membranoproliferative nephritis unspecified, gastroenteritis, acute myocardial infarction, congestive cardiac failure, urology

| Table 2 Cox model HRs and β-coefficients before penalisation (shrinkage) | | | | | | | |
|--|-------------------------|--------------|------------------------|--|--|--|--|
| Predictors | Adjusted HR (95% CI) | Coefficients | Shrunken coefficients* | | | | |
| | | Coemolents | official coefficients | | | | |
| Age, year | 1.01 (1.00, 1.02) | 0.0096984 | 0.00872856 | | | | |
| Female sex | 0.73 (0.52, 1.01) | -0.3190717 | -0.28716453 | | | | |
| Body Mass Index (kg/m ²) | 1.03 (1.00, 1.05) | 0.0259328 | 0.02333952 | | | | |
| Alcohol use | | | | | | | |
| Non-user | Reference | - | | | | | |
| Low (1–14 units/week) | 1.38 (0.84, 2.56) | 0.320509 | 0.2884581 | | | | |
| Moderate (15-21 units/week) | 1.10 (0.46, 2.61) | 0.0962384 | 0.08661456 | | | | |
| Hazardous (>21 units/week) | 2.26 (1.17, 4.36) | 0.8169186 | 0.73522674 | | | | |
| Ex-user | 2.01 (1.08, 3.77) | 0.7002583 | 0.63023247 | | | | |
| Comorbidities | | | | | | | |
| Diabetes | 1.26 (0.79, 1.99) | 0.2276251 | 0.20486259 | | | | |
| Chronic kidney disease stage 3 | 2.98 (1.96, 4.53) | 1.09148 | 0.982332 | | | | |
| Other drugs | | | | | | | |
| ACE inhibitors | 1.34 (0.88, 2.03) | 0.2938861 | 0.26449749 | | | | |
| Aspirin | 1.30 (0.84, 2.03) | 0.2653679 | 0.23883111 | | | | |
| NSAIDs | 1.14 (0.63, 2.06) | 0.1298207 | 0.11683863 | | | | |
| | | | | | | | |

*Multiplied by the shrinkage factor (0.9).

NSAIDs, non-steroidal anti-inflammatory drugs.

Box 1 Equation to predict the risk of 5-ASA discontinuation after 6 months of primary care prescription and within the next 5 years

Risk score = 1 – 0.994 ^{exp(0.90βX)}, where β X= 0.0096984 * age in years at first primary-care prescription – 0.3190717 * female sex + 0.0259328 * Body Mass Index + 0.320509 * Iow alcohol intake + 0.0962384 * moderate alcohol intake + 0.8169186 * hazardous alcohol intake + 0.7002583 * ex-alcohol intake + 0.2276251 * diabetes + 1.09148 * chronic kidney disease stage 3 + 0.2938861 * ACE inhibitors + 0.2653679 * aspirin + 0.1298207 * NSAIDs. All variables are code 0, and 1 if absent or present respectively, except for BMI and age that are continuous variables. 0.994 is the baseline survival function at 5 years, 0.90 is the shrinkage factor and the other numbers are the estimated regression coefficients for the predictors, which indicate their mutually adjusted relative contribution to the outcome risk.

referral and admission to medical emergency ward (with unspecified reason) with a positive predictive value of 86%. Pre-existing CKD stage 3, hazardous alcohol use and ex-alcohol use were strong predictors of drug discontinuation with adjusted HR (95% CI) before shrinkage of 2.98 (1.96, 4.53), 2.26 (1.17, 4.36) and 2.01 (1.08, 3.77) respectively (table 2). Before shrinkage, the calibration slope in the development data was 1.00 (95% CI 0.81 to 1.19). From the bootstrap process, the estimated uniform shrinkage factor (optimism-adjusted calibration slope) was 0.90 and this was used to shrink predictor coefficients for optimism, to form the final model shown in box 1, including, after re-estimation, the cumulative baseline survival function (S₀) of 0.994 at 5 years of follow-up.

Model performance in the development cohort

Calibration plot of the final (ie, after shrinkage) model at 5 years is shown in online supplemental figure S4. As most patients had low risk of outcome, the groups defined by

tenths of predicted risk are mostly clustered at the bottom left of the calibration plot (online supplemental figure S5). The smoothed calibration curve at 5 years showed alignment of observed risk to the predicted risk with wide CIs at >0.1 risk probabilities (online supplemental figure S4). The Royston D statistic (95% CI) was 1.13 (0.88, 1.38) corresponding to an HR of 3.10 (2.44–3.97) when comparing the risk groups above and below the median of linear predictor. The optimism-adjusted Royston D statistic was 0.98 corresponding to an HR of 2.66. The optimism-adjusted C-statistic and $R^2_{\ D}$ were 0.64 and 0.18 respectively (table 3).

Model performance in the validation cohort

There were 10 (0.14% deaths). 98 outcomes occurred in n=7318 (1.3%) patients at a rate (95% CI) of 4.74 (3.89-5.77) per 1000 person-years in the validation cohort. The calibration slope (95% CI) at the 5-year follow-up period was 1.01 (0.71-1.30). The calibration plot showed reasonable correspondence between observed and predicted risk at 5 years across the tenths of risk with CIs crossing the ideal line of agreement (online supplemental figure S6). As above, most groups clustered at the bottom left of the calibration plot owing to low risk of outcome for most patients (online supplemental figureS7). The smoothed calibration curve showed miscalibration in those at high risk, with under-prediction of risk, although this could be limited by a small number of people at high risk reflected by a wide 95% CI (figure 1). Model performance was also tested at years 1, 2, 3 and 4 and showed a similar pattern (online supplemental figures S11 and S12). Model discrimination in the validation data was broadly like the development data (table 3). The Royston D statistic in the validation data was 0.94 (0.62, 1.26), corresponding to HR (95% CI) 2.56 (1.86–3.53). The R² score was 0.17 (95% CI 0.08, 0.27). Harrell's C-statistic was 0.66 (95% CI 0.60, 0.72).

| Table 3 Model diagnostics | | | | | | | | |
|-------------------------------|--------------------------|----------------------|----------------------|---------------------------------------|---------------------------------------|--|--|--|
| Measure | Apparent performance* | Test performance† | Average optimism‡ | Optimism corrected performance§ | External validation (CPRD Gold) | | | |
| Overall calibration slope | 1.00 (0.81, 1.19) | 0.90 | 0.10 | 0.90 | 1.01 (0.71, 1.30) | | | |
| Harrel's C-statistic | 0.67 (0.63, 0.72) | 0.66 | 0.02 | 0.64 | 0.66 (0.60, 0.72) | | | |
| R ² _D | 0.23 (0.16, 0.31) | 0.21 | 0.05 | 0.18 | 0.17 (0.08, 0.27) | | | |
| Royston D statistic | 1.13 (0.88, 1.38) | 1.05 | 0.15 | 0.98 | 0.94 (0.62, 1.26) | | | |

*Estimated directly from data that was used to develop the model.

†Determined by executing full model in each bootstrap sample of the development dataset (500 samples with replacement), calculating bootstrap performance and applying same model in original sample.

‡Average difference between model performance in bootstrap sample of the development dataset and performance in the development dataset.

§Obtained by subtracting average optimism from apparent performance.

CPRD, Clinical Practice Research Datalink;



Figure 1 Calibration of a prognostic model for 5-aminosalicylic acid (5-ASA) discontinuation with abnormal monitoring blood test results at 5 years in the validation cohort. Data from a single imputed dataset were used for illustration. Baseline survival function(So) was 0.994 at 5 years. Black line reflects perfect prediction. Grey line shows model prediction with 95% CI in grey shade.

Worked examples

10 anonymised patient profiles, one from the middle of each tenth of predicted risk, were selected from the development cohort (online supplemental table S1). The cumulative probability of outcome over 5 years ranged from 0.9% in the middle of the first group to 1.9% in the middle of the seventh group, and 13.2% in the middle of the 10th group.

DISCUSSION

Summary

This study developed and externally validated a prognostic model for 5-ASA discontinuation due to nephrotoxicity defined using a gap in prescription and renal function decline in people with IBD established on 5-ASA treatment that is, after the first 6 months of primary-care prescription. We were able to predict 5-ASA discontinuation due to nephrotoxicity using readily available data. Currently, everyone prescribed 5-ASA is recommended to undergo annual monitoring blood tests for the early detection of 5-ASA nephrotoxicity. As early discontinuation of the drug and treatment with steroids may result in rapid improvement in kidney function, there is a compelling case for monitoring those at a high risk of nephrotoxicity more frequently.⁶ This study also provides data on the incidence of nephrotoxicity due to 5-ASA in a large population established on 5-ASA in primary care. Due to the use of a pragmatic definition, these rates are likely to be an overestimate.

Strengths and limitations

Strengths of this study include model development and validation using a large real-world and nationally representative dataset that originated during routine clinical care in the UK and therefore, has high generalisability to the UK population. Additionally, the model had good performance characteristics in an independent dataset and predicted clinically relevant nephrotoxicity that is, one that required 5-ASA cessation, providing face validity. The prognostic factors included those that are readily available during routine consultations, making the model easy to use, especially, once incorporated into clinical decision support software.

However, the findings of this study should be interpreted considering several limitations. First, there is some information which we would ideally want but was not available. The date when 5-ASA was first prescribed from hospital was unavailable. We think this is not a serious issue as the model can be applied once the patient has completed 6 months of 5-ASA treatment from primary care after treatment initiation and initial stabilisation in a hospital clinic. Second, we did not have information on the use of biologics as these are hospital prescribed. There is however no reason to believe that concurrent biologic prescription would increase the risk of 5-ASA nephrotoxicity. Third, we also did not have data on the severity of IBD as this is not recorded in CPRD. However, there is no evidence that suggests IBD disease activity is associated with drug-induced side effects.^{27 28} Fourth, we must consider whether the ascertainment of discontinuation being due to nephrotoxicity is reliable. The decline in kidney function could have been due to factors other than 5-ASA-induced nephrotoxicity. Though, a review of potential causes of outcomes in the development cohort revealed that only a minority could potentially be explained by another contemporaneous illness. Our definition of CKD progression relied on coding information and may therefore have been applied variably by different practitioners, and the minimum serum creatinine rise we required to consider a diagnosis of AKI may have included some cases of CKD progression. No kidney biopsy confirmation of 5-ASA nephrotoxicity was available, but this is in keeping with the real-world situation where a decision to stop treatment is usually made before kidney biopsy is considered. Fifth, we did not include genetic factors such as human leucocyte antigen polymorphism that have been associated with 5-ASA nephrotoxicity⁷ as a prognostic factor because such data are not available in the CPRD. Future research should include genetic factors to find out if these would improve the prediction of 5-ASA-associated nephrotoxicity. Sixth, although the external validation dataset was distinct from the model development dataset, it also originated from the UK meaning that generalisability beyond the UK cannot be guaranteed. Therefore, before considering the model in other countries, further validation is required in non-UK data. Finally, we censored individuals who died at their death time rather than modelling death

as a competing risk. This means risk predictions at 5 years assume the individual will be alive at that time, which we considered to be clinically more informative for this situation where monitoring decisions are being made.

Our approach as outlined above was inclusive capturing all drug discontinuations possibly due to nephrotoxicity, and hence, our estimates represent a 'worst-case' scenario and provide a firm basis for guiding safe practice. Our data show that just over 4 in 1000 IBD patients on 5-ASA will have their medication discontinued possibly due to nephrotoxicity in any year. That this level is rather higher than that of the best previous estimates of the risk of 5-ASA nephrotoxicity may partly be explained by our estimates reflecting the real world and not randomised controlled trials, but equally doubtless implies that in some cases, the drug is discontinued when it is not to blame.⁴⁵

Comparison with existing literature

We observed the rate of drug discontinuation due to apparent nephrotoxicity to be 4.21 and 4.74/1000 person-years in the development and validation cohorts, confirming a low risk of nephrotoxicity associated with 5-ASA treatment. Nevertheless, acute interstitial nephritis is a serious complication which results in end-stage kidney disease in 14.6% of those affected.⁶ Our observed incidence of 5-ASA nephrotoxicity is about threefold lower than the rate reported by Jairath *et al* using data from the THIN database.¹⁸ However, in their analysis, nephrotoxicity was defined using any diagnostic code for 'chronic, acute and unspecified renal impairment' and did not consider drug discontinuation so their definition likely would have included a variety of unrelated kidney diseases.¹⁸

Implications for research and/or practice

Our analysis also confirms that, although the risk of nephrotoxicity with 5-ASA treatment is variable across individuals, only a small number of readily available demographic and clinical variables are needed to tailor risk predictions to individuals. This allows the prediction model to be easily used in clinical practice that is, either as an online calculator or embedded in primary care computer clinical decision software solutions. Based on the level of risk we report, current guidelines on monitoring may seem reasonable at a population level. However, we have shown there is a small predictable subset with an appreciably higher risk in whom more frequent monitoring may be appropriate. It is beyond the scope of this study to propose specific risk thresholds at which current clinical practice may be changed. Such recommendations are best made by national and international guideline writing groups considering the views of clinical experts, patients and commissioners. In the UK, the frequency of monitoring blood tests is typically decided according to the recommendations of specialist societies such as the BSG. These findings ought to be considered by them when formulating their guidance.

We recommend that the prognostic model be validated using data from other countries. Further research is needed to examine the impact on patient outcomes of using the model to direct monitoring decisions, for example, using decision analysis methods, costeffectiveness models and cluster-randomised trials comparing practices that do or do not implement the model.

CONCLUSION

We have developed and externally validated a prognostic model for 5-ASA discontinuation due to nephrotoxicity that may be used to individualise blood test monitoring using principles of shared decision making between the patient and the physician. These findings need to be considered by guideline writing groups to provide riskbased recommendations for the frequency of monitoring of kidney function in people with IBD receiving treatment with 5-ASA.

Author affiliations

¹Academic Rheumatology, University of Nottingham, Nottingham, East Midlands, UK ²Nottingham NIHR BRC, Nottingham, UK

³Lifespan and Population Health, School of Medicine, University of Nottingham, Nottingham, UK

⁴Centre for Kidney Research and Innovation, Translational Medical Sciences, University of Nottingham, Nottingham, UK

⁵Nottingham Digestive Diseases Centre, Translational Medical Sciences, School of Medicine, University of Nottingham, Nottingham, UK

⁶Primary Care Centre Versus Arthritis, Keele University, Keele, UK

⁷School of Health and Related Research, University of Sheffield, Sheffield, UK ⁸Institute of Applied Health Research, College of Medical and Dental Sciences, University of Birmingham, Birmingham, UK

Contributors GN, MJG, TC, MWT, GPA, CDM, MDS, RDR and AA designed the study. GN analysed the data supervised by MJG, RDR and AA. GN, MJG, TC, MWT, GPA, CDM, MDS, RDR and AA interpreted the data. AA and GN drafted the manuscript. AA and GN revised the manuscript. AII authors critically evaluated and revised the manuscript. AA is the guarantor. The corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted.

Funding This research was funded by National Institute for Health and Care Research (NIHR) grant NIHR130580. The funders had no role in conducting and/or reporting this study.

Competing interests AA has received Institutional research grants from AstraZeneca and Oxford Immunotech; and personal fees from UpToDate (royalty), Springer (royalty), Cadilla Pharmaceuticals (lecture fees), NGM Bio (consulting), Limbic (consulting) and personal fees from Inflazome (consulting) unrelated to the work. GPA has received consulting fees from Abbott, Albereo, Amryth, AstraZeneca, BenevolentAI, DNDI, GlaxoSmithKline, NuCANA, Pfizer, Roche Diagnostics, Servier Pharmaceuticals, W.L Gore & Associates paid to the University of Nottingham unrelated to the work. CPF has received Consultancy/Advisory board fees from Abbvie, GenMab, Incyte, Morphosys, Roche, Takeda, Ono, Kite/Gilead, BMS/Celgene, BTG/Veriton and departmental research funding from BeiGene unrelated to the work. RR is a BMJ Statistical Editor for which he receives consulting fees, and he receives royalties for two textbooks. The other authors have no conflict of interest to declare.

Patient consent for publication Not applicable.

Ethics approval This study was approved by the independent scientific advisory committee of Clinical Practice Research Datalink (CPRD, reference: 20_000236R). CPRD has overarching research ethics committee approval, and individual studies do not require separate ethical approval.

Provenance and peer review Not commissioned; externally peer reviewed.

⁹NIHR, Birmingham Biomedical Research Centre, Birmingham, UK

Data availability statement Data may be obtained from a third party and are not publicly available. Data used in the study are from the Clinical Practice Research Datalink and cannot be shared due to licensing restrictions. Study protocol is available from www.cprd.com.

Supplemental material This content has been supplied by the author(s). It has not been vetted by BMJ Publishing Group Limited (BMJ) and may not have been peer-reviewed. Any opinions or recommendations discussed are solely those of the author(s) and are not endorsed by BMJ. BMJ disclaims all liability and responsibility arising from any reliance placed on the content. Where the content includes any translated material, BMJ does not warrant the accuracy and reliability of the translations (including but not limited to local regulations, clinical guidelines, terminology, drug names and drug dosages), and is not responsible for any error and/or omissions arising from translation and adaptation or otherwise.

Open access This is an open access article distributed in accordance with the Creative Commons Attribution 4.0 Unported (CC BY 4.0) license, which permits others to copy, redistribute, remix, transform and build upon this work for any purpose, provided the original work is properly cited, a link to the licence is given, and indication of whether changes were made. See: https://creativecommons.org/licenses/by/4.0/.

ORCID iDs

Georgina Nakafero http://orcid.org/0000-0002-3859-7354 Tim Card http://orcid.org/0000-0003-2555-2250

REFERENCES

- Lamb CA, Kennedy NA, Raine T, et al. British Society of Gastroenterology consensus guidelines on the management of inflammatory bowel disease in adults. Gut 2019;68:s1–106.
- Ko CW, Singh S, Feuerstein JD, et al. AGA Clinical Practice Guidelines on the Management of Mild-to-Moderate Ulcerative Colitis. *Gastroenterology* 2019;156:748–64.
 Dignass A, Van Assche G, Lindsay JO, et al. The second European
- 3 Dignass A, Van Assche G, Lindsay JO, *et al*. The second European evidence-based Consensus on the diagnosis and management of Crohn's disease: Current management. *Journal of Crohn's and Colitis* 2010;4:28–62.
- 4 Van Staa TP, Travis S, Leufkens HGM, et al. 5-aminosalicylic acids and the risk of renal disease: A large British epidemiologic study☆. Gastroenterology 2004;126:1733–9.
- 5 Muller AF, Stevens PE, McIntyre AS, et al. Experience of 5aminosalicylate nephrotoxicity in the United Kingdom. Aliment Pharmacol Ther 2005;21:1217–24.
- 6 Moss JG, Parry CM, Holt RCL, et al. 5-ASA induced interstitial nephritis in patients with inflammatory bowel disease: a systematic review. Eur J Med Res 2022;27:61.
- 7 Heap GA, So K, Weedon M, *et al.* Clinical Features and HLA Association of 5-Aminosalicylate (5-ASA)-induced Nephrotoxicity in Inflammatory Bowel Disease. *J Crohns Colitis* 2016;10:149–58.
- 8 Corrigan G, Stevens PE. Review article: interstitial nephritis associated with the use of mesalazine in inflammatory bowel disease. *Aliment Pharmacol Ther* 2000;14:1–6.

- 9 Gisbert JP, González-Lama Y, Maté J. 5-Aminosalicylates and renal function in inflammatory bowel disease: a systematic review. *Inflamm Bowel Dis* 2007;13:629–38.
- 10 World MJ, Stevens PE, Ashton MA, et al. Mesalazine-associated interstitial nephritis. *Nephrol Dial Transplant* 1996;11:614–21.
- 11 Committee JF. *British national formulary (85)*. London: BMJ Group and Pharmaceutical Press, 2023.
- 12 Guillo L, Delanaye P, Flamant M, et al. Kidney function monitoring in inflammatory bowel disease: The MONITORED consensus. *Dig Liver Dis* 2022;54:309–15.
- 13 Siddique N, Farmer C, Muller AF. Do gastroenterologists monitor their patients taking 5-amino-salicylates following initiation of treatment. *Frontline Gastroenterol* 2015;6:27–31.
- 14 Herrett E, Gallagher AM, Bhaskaran K, et al. Data Resource Profile: Clinical Practice Research Datalink (CPRD). Int J Epidemiol 2015;44:827–36.
- 15 Wolf A, Dedman D, Campbell J, et al. Data resource profile: Clinical Practice Research Datalink (CPRD) Aurum. Int J Epidemiol 2019;48:1740–1740g.
- 16 Khwaja A. KDIGO clinical practice guidelines for acute kidney injury. Nephron Clin Pract 2012;120:c179–84.
- 17 Kazancioğlu R. Risk factors for chronic kidney disease: an update. Kidney Int Suppl (2011) 2013;3:368–71.
- 18 Jairath V, Hokkanen SRK, Guizzetti L, et al. No increased risk of nephrotoxicity associated with 5-aminosalicylic acid in IBD: a population-based cohort and nested case-control study. *Aliment Pharmacol Ther* 2019;50:416–24.
- 19 Patel H, Barr A, Jeejeebhoy KN. Renal effects of long-term treatment with 5-aminosalicylic acid. Can J Gastroenterol 2009;23:170–6.
- 20 Logan RF, Staa TP. Sulphasalazine and mesalazine: serious adverse reactions re-evaluated on the basis of suspected adverse reaction reports to the Committee on Safety of Medicines * Author's reply. *Gut* 2003;52:1530–1.
- 21 Riley RD, Ensor J, Snell KIE, et al. Calculating the sample size required for developing a clinical prediction model. BMJ 2020;368:m441.
- 22 Schafer JL. Multiple imputation: a primer. *Stat Methods Med Res* 1999;8:3–15.
- 23 Steyerberg EW. Clinical prediction models: a practical approach to development, validation, and updating. Cham, Switzerland: Springer International Publishing AG, 2019.
- 24 Royston P, Altman DG. External validation of a Cox prognostic model: principles and methods. *BMC Med Res Methodol* 2013;13:33.
- 25 Cox DR. Note on Grouping. J Am Stat Assoc 1957;52:543-7.
- 26 Moons KGM, Altman DG, Reitsma JB, et al. Transparent Reporting of a multivariable prediction model for Individual Prognosis or Diagnosis (TRIPOD): explanation and elaboration. *Ann Intern Med* 2015;162:W1–73.
- 27 Broekman MMTJ, Coenen MJH, Wanten GJ, et al. Risk factors for thiopurine-induced myelosuppression and infections in inflammatory bowel disease patients with a normal TPMT genotype. Aliment Pharmacol Ther 2017;46:953–63.
- 28 Labidi A, Hafi M, Ben Mustapha N, et al. Toxicity profile of thiopurines in inflammatory bowel disease: a retrospective cohort analysis. *Tunis Med* 2020;98:404–12.