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1 Experimental evolution of local adaptation under
2 unidimensional and multidimensional selection

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6 Summary

7 Local adaptation is a fundamental evolutionary process generating biological diversity and potentially
8 enabling ecological speciation. Divergent selection underlies the evolution of local adaptation in
9 spatially structured populations by driving their adaptation towards local optima. Environments rarely
10 differ along just one environmental axis, and so divergent selection may often be multidimensional.
11 How the dimensionality of divergent selection affects local adaptation is unclear: Evolutionary theory
12 predicts that increasing dimensionality will increase local adaptation when associated with stronger
13 overall selection but may have less predictable effects if selection strengths are equal. Experiments
14 are required that allow the effect of the dimensionality of selection on local adaptation to be tested
15 independently of the total strength of selection. We experimentally evolved 32 pairs of monogonont
16 rotifer populations under either unidimensional divergent selection (a single pair of stressors) or
17 multidimensional divergent selection (three pairs of stressors), keeping the total strength of selection
18 equal between treatments. At regular intervals, we assayed fitness in home and away environments
19 to assess local adaptation. We observed an initial increase and subsequent decline of local adaptation
20 in populations exposed to multidimensional selection, compared to a slower but eventually stronger
21 increase in local adaptation in populations exposed to unidimensional selection. Our results contrast
22 with existing predictions such as the 'weak multifarious' and 'stronger selection' hypotheses. Instead,
23 we hypothesise that adaptation to multidimensional divergent selection may favour generalist
24 genotypes and only produce transient local adaptation.

25

26 Keywords

27 Divergent selection, Dimensionality, Fitness, Generalists, Specialists, Divergence

28 Introduction

29 When faced with selection pressures that vary spatially, populations may adapt to local conditions¹.
30 Such local adaptation produces a pattern in which populations have higher fitness in their home
31 environment than they would if transplanted to a different environment². Local adaptation is an
32 important component of within-species diversity and an increase in local adaptation is also a vital step
33 towards ecological speciation, providing an extrinsic form of reproductive isolation^{3,4}. The strength of
34 local adaptation, and the speed at which it increases, depends on environmental factors such as the
35 strength of selection and gene flow^{5,6}, and genomic conditions such as the amount and distribution of
36 standing genetic variation^{7,8} and the genomic architecture^{9,10}.

37 One variable that has received far less attention, however, is the number of simultaneous divergent
38 selection pressures experienced by a population, i.e. the dimensionality of divergent selection¹¹⁻¹³.
39 Local adaptation can be driven by between-population heterogeneity in a single selection pressure
40 (unidimensional divergent selection) or by multiple selection pressures (multidimensional divergent
41 selection). Multidimensional divergent selection is generally predicted to drive the evolution of
42 reproductive barriers, including local adaptation, more effectively than unidimensional selection and
43 so to promote ecological speciation¹¹⁻¹⁴. However, to date, no study has tested how local adaptation
44 builds under different dimensionalities of selection¹⁵.

45 It is important to distinguish the dimensionality of divergent selection from the overall dimensionality
46 of selection (sometimes referred to as 'niche dimensionality')¹⁶⁻¹⁹. Overall dimensionality reflects the
47 total number of selection pressures impacting a population, many of which will act via stabilising
48 selection¹⁸⁻²¹. In any environment selection is likely to be imposed by multiple environmental variables
49 and to act on multiple traits that are determined by many genes. The dimensionality of selection then
50 depends on the extent to which environmental variables, and organismal traits are independent. For
51 example, extensive pleiotropy reduces the effective dimensionality because generalist genotypes can
52 increase fitness in response to multiple selection pressures simultaneously^{22,23}.

53 However, the dimensionality of divergent selection considers only selection pressures which vary
54 between environments. Using a framework such as Fisher's Geometric Model²⁴, any two
55 environments are separated by a single vector in multidimensional space. We can, therefore, define
56 the dimensionality of divergent selection as the number of orthogonal environmental variables that
57 are correlated with this vector. The dimensionality of divergent selection is also likely to be related to
58 the number of traits and number of loci that must evolve to achieve local adaptation, although these
59 ways of defining the available space do not map directly, for example because genetic correlations (G-

60 matrix covariances) mean that fewer genetic axes may underlie responses to multiple environmental
61 variables or traits²⁵⁻²⁷.

62 Predictions of how dimensionality might impact local adaptation and speciation, such as Nosil *et al.*'s¹¹
63 'stronger selection' vs 'weak multifarious' hypotheses, describe how multidimensionality might
64 distribute divergent selection over more traits. In reality, there are several different ways in which the
65 dimensionality of divergent selection might impact these processes, though these ideas currently lack
66 empirical support. Firstly, increasing the dimensionality of divergent selection could drive the
67 evolution of local adaptation in a variety of ways²⁵. For instance, multidimensionality might be
68 associated with an increase in the overall strength of divergent selection, driving local adaptation via
69 reduced fitness of migrants and hybrids, thus decreasing effective gene flow^{12,27}. Moreover, more
70 selection pressures might produce divergence in more orthogonal traits^{26,28}. This may speed local
71 adaptation as a greater number of orthogonal traits implies more genes, and hence more standing
72 additive genetic variation, enabling more rapid increase of local adaptation^{7,29}.

73 Alternatively, increasing the dimensionality of divergent selection might plausibly slow, or even
74 prevent, local adaptation if the total strength of divergent selection does not increase with
75 dimensionality²⁵. In this scenario, if increased dimensionality leads to more loci contributing to genetic
76 variation, but the overall strength of selection does not increase, selection may become diluted across
77 many loci, leading to weaker per-locus selection coefficients. Here, outcomes are harder to predict³⁰.
78 Greater total genetic variance, due to increased mutational input, may allow faster and stronger local
79 adaptation^{30,31}. Alternatively, per-locus selection may be too low to overcome the homogenising force
80 of gene flow, preventing local adaptation, whereas unidimensional selection concentrated onto few
81 loci may provide a stronger response¹¹. Furthermore, local adaptation may take longer, as a greater
82 number of new, locally-adaptive alleles must be brought together by recombination arising during
83 sexual reproduction (overcoming Hill-Robertson interference³²) to produce locally-adapted
84 genotypes.

85 Attempting to quantify the dimensionality of divergent selection is difficult in natural populations.
86 Studies that determine the relationship between local adaptation and environmental heterogeneity
87 tend to focus on single environmental differences between populations³³⁻³⁶. This may be because it is
88 considerably easier to identify adaptation along a single consistent environmental gradient, although
89 there are excellent examples of studies identifying multidimensional divergent selection³⁷⁻³⁹.
90 Comparison of these studies indicates that multidimensionality is associated with stronger local
91 adaptation. A meta-analysis of 35 reciprocal transplant studies²⁷ identified that the dimensionality of
92 environmental divergence (maximum likelihood estimate from available data, as defined in

93 Hohenlohe and Arnold⁴⁰) accounted for a larger proportion of variance in local adaptation (20%) than
94 environmental variables alone (4%)⁴¹.

95 Laboratory experiments have been proposed as one of the most effective methods available to assess
96 the impact of dimensionality on local adaptation¹. Using experimental evolution, environmental
97 selection pressures and levels of ongoing gene flow can be manipulated and controlled⁴². However, a
98 review identified no single experimental evolution study that has varied the dimensionality of
99 selection¹⁵. Although comparisons between experiments have been made^{12,13}, only 5 studies have
100 imposed divergent selection along more than one axis^{43–47}. Although three of these produced strong
101 reproductive isolation, it has been argued that this was due to selection on multiple-effect traits (traits
102 that underlie more than one component of reproductive isolation⁴⁸) rather than multidimensional
103 divergent selection^{25,42}.

104 In this study, we used experimental evolution to test how the dimensionality of divergent
105 environmental selection affected the speed and magnitude of evolution of local adaptation. We
106 exposed populations of the monogonont rotifer, *Brachionus plicatilis*, to unidimensional and
107 multidimensional divergent selection pressures (Table 1) whilst keeping total selection equal (Figure
108 1). Populations were paired and connected via gene flow to form metapopulations with ‘home’ and
109 ‘away’ environments for each subpopulation. We tested the strength of local adaptation in each
110 metapopulation at regular intervals over the course of the experiment, defined by differences in
111 fitness when exposed to home vs. away environments^{1,49}. Our planned analyses sought to answer
112 three questions: 1) Do the speed and magnitude of local adaptation vary by dimensionality, 2) if so,
113 are these patterns specific to dimensionality, or could they be explained by the individual stressors
114 used, and 3) what underlying patterns of fitness in the home and away environments are responsible
115 for changes in local adaptation?

116 Because we kept the total strength of selection equal between dimensionality treatments, the
117 patterns of local adaptation were expected to depend primarily on the genetic basis of local
118 adaptation. We predicted that unidimensional divergent selection would result in stronger local
119 adaptation, at least initially, given the expectation that selection would be concentrated upon fewer
120 loci and so might overcome the opposing effect of gene flow more easily. We further predicted that,
121 under multidimensional divergent selection, the selection per locus may be too dilute to overcome
122 gene flow, with conflicting fitness effects among loci, and that this effect would dominate over the
123 greater availability of genetic variation, producing only weak or slow local adaptation. Our results
124 contradicted these predictions: local adaptation evolved more slowly under unidimensional divergent

125 selection but ultimately led to stronger local adaptation, whereas multidimensional selection resulted
126 in initially stronger local adaptation that was transient and eventually declined to low levels.

127 Results

128 **Laboratory adaptation**

129 To assess the baseline of laboratory adaptation, we analysed population density of all demes
130 throughout the experiment. We counted population density pre-shock at four-cycle intervals
131 throughout the experiment whenever local adaptation was being assayed. Additionally, we counted
132 population density post-shock each week for passaging purposes. We observed a significant increase
133 in day 7 (post-shock) population density in all metapopulations at the start of the experiment which
134 levelled off after approximately 10 cycles (Figure S2). During this laboratory adaptation period, shock
135 survival was estimated at cycles 5 and 9 but did not significantly differ from the calibrated value in
136 pilot experiments (1-sample t-test with expectation of 0.45; cycle 5 mean = 0.441, $t = -0.343$, $df = 63$,
137 $p = 0.737$; cycle 9 mean = 0.406, $t = -1.916$, $df = 63$, $p = 0.060$). This indicated that the rise in density
138 was due to increased asexual growth rather than increased shock survival.

139 **Local adaptation**

140 To determine the time-course and strength of local adaptation under each of the two dimensionality
141 treatments, we used repeated reciprocal transplant assays. Using the SA contrast as a metric for local
142 adaptation (see Methods), over 45 cycles of experimental evolution, we found that both
143 unidimensional and multidimensional treatments led to local adaptation, but with contrasting
144 temporal dynamics (Figure 2).

145 Firstly, our analysis using dimensionality treatment as a fixed effect, revealed a significant effect of
146 dimensionality on the way local adaptation (SA) evolved (interaction between treatment and second-
147 order polynomial effect of cycle: $F = 7.17$, $df = 2$, 28.162, $p = 3.04 \times 10^{-3}$; full ANOVA table presented in
148 Table 2, model coefficients provided in Table S2). In populations exposed to multidimensional
149 selection, local adaptation increased rapidly during the early stages of the experiment but
150 subsequently declined to low levels. In contrast, local adaptation increased more slowly in populations
151 exposed to unidimensional selection but ultimately reached higher levels by the end of the experiment
152 (Figure 2).

153 Secondly, we tested whether the effect of the identity of the stressor pair on SA varied by cycle. This
154 model identified no significant interaction between stressor pair and the second order polynomial of
155 cycle ($F = 1.40$, $df = 14$, 29.01, $p = 0.217$), on local adaptation.

156 However, to test the *a priori* hypothesis that treatment (unidimensional vs. multidimensional
157 divergent selection), would influence the evolution of local adaptation, regardless of the specific
158 stressors involved, we performed planned contrasts of marginal means from this model. As detailed

159 above, we predicted differences between the two dimensionality treatments, with unidimensional
160 divergent selection producing faster and stronger local adaptation. Although we did observe clear
161 differences between treatments, they ran counter to this prediction. At the start of the experiment
162 (cycle 5), there was no difference between dimensionalities ($t = -0.155$, $df = 23.97$, $p = 0.878$).
163 However, by the midpoint (cycle 21), multidimensional stressor pairs had produced significantly
164 greater local adaptation than unidimensional stressor pairs ($t = 3.24$, $df = 24.20$, $p = 3.40 \times 10^{-3}$), and by
165 the end (cycle 41) this pattern had reversed so that unidimensional metapopulations now had
166 produced significantly greater local adaptation than multidimensional metapopulations ($t = -2.65$, df
167 $= 23.01$ $p = 0.0142$; Figure S3; Table S3).

168 Repeating this analysis using the 'local-foreign' criterion for local adaptation instead of SA confirmed
169 these results and replicated the pattern of trajectories (Figure S3; Figure S4; Table S4; Table S5. We
170 also observed significantly higher local adaptation at the midpoint ($t = 3.14$, $df = 24.06$, $p = 4.37 \times 10^{-3}$)
171 and lower local adaptation at the end ($t = -3.19$, $df = 23.93$, $p = 3.96 \times 10^{-3}$) in multidimensional lines
172 compared to unidimensional lines when contrasting estimated means from a model using stressor
173 pair, quadratic effect of time and their interaction (Figure S3; Table S5). Finally, although there was a
174 significant quadratic effect of cycle on LF asymmetry ($F = 61.9$, $df = 2, 29.06$, $p = 5.76 \times 10^{-3}$), with local
175 adaptation becoming initially less, then gradually more asymmetric over time, there was no effect of,
176 nor interaction with, stressor pair or dimensionality (Figure S5).

177 Thirdly, we tested whether home fitness, away fitness, or some combination of both was responsible
178 for the local adaptation. We found that home fitness increased linearly and at similar rates in response
179 to all stressor pairs, and hence is unlikely to be responsible for the observed differences in local
180 adaptation. Meanwhile, away adaptation displayed more idiosyncratic behaviour when comparing
181 between stressor pairs and between dimensionality treatments and is, therefore, likely to be
182 responsible for differences in local adaptation.

183 Fitness in home environments was found to increase linearly with time, without any rate variation
184 with respect to stressor pair; there was a significant effect of cycle ($F = 14.94$, $df = 1, 30.53$, $p = 5.40 \times 10^{-4}$)
185 and stressor pair identity ($F = 5.07$, $df = 7, 23.82$, $p = 1.25 \times 10^{-3}$), but with no interaction or quadratic
186 effects (Figure 3). Repeating this analysis with dimensionality treatment rather than stressor pair
187 yielded the same pattern of results; a significant effect of cycle ($F = 14.05$, $df = 1, 30.54$, $p = 7.44 \times 10^{-4}$)
188 and dimensionality ($F = 9.11$, $df = 1, 29.92$, $p = 5.15 \times 10^{-3}$) but no interaction or quadratic effects. Fitness
189 in the home environment was higher for unidimensional than for multidimensional treatments,
190 throughout the experiment ($\Delta w = 0.066 \pm 0.022$).

191 The differences in evolution of local adaptation between treatments were driven mainly by fitness in
192 the away environments. As for home fitness, there was a significant effect of stressor pair on fitness
193 ($F = 5.01$, $df = 7$, 23.72 , $p = 1.35 \times 10^{-3}$). Here, the interaction between stressor pair and a quadratic
194 effect of cycle approached significance ($F = 2.03$, $df = 14$, 29.02 , $p = 0.053$), suggesting variation in
195 patterns of evolution of away fitness according to stressor pair. The four multidimensional stressor
196 pairs showed consistent quadratic effects, an initial fall in away fitness followed by a rise. Together
197 with a linear increase in home fitness, this explains the observed pattern of local adaptation: away
198 fitness was lower than home fitness in the middle of the experiment, but not by the end (Figure 3).
199 Unidimensional stressor pairs displayed a less consistent pattern, with two stressor pairs displaying
200 negative quadratic effects and two stressor pairs displaying positive quadratic effects. The increase in
201 local adaptation of unidimensional metapopulations towards the end of the experiment is therefore
202 due to the linear increase in home fitness with little overall change in away fitness (Figure 3). Using
203 dimensionality treatment rather than stressor pair yielded a significant effect of dimensionality ($F =$
204 14.11 , $df = 1$, 29.93 , $p = 7.44 \times 10^{-4}$), with unidimensional away fitness higher than multidimensional
205 away fitness ($\Delta w = 0.093 \pm 0.024$) but no effect of, nor interaction with, cycle.

206

207 Discussion

208 In natural environments, divergent selection is frequently multidimensional, but how the
209 dimensionality of selection affects the evolution of local adaptation remains unknown. Our
210 experiment directly tests how the dimensionality of divergent selection affects the evolution of local
211 adaptation whilst controlling for the overall strength of selection, identity of stressors, gene flow and
212 recombination^{15,25,42}. Our first prediction of faster evolution of local adaptation under unidimensional
213 divergent selection, was not realised. Our second prediction of weak response to multidimensional
214 selection was also inconsistent with the results. Metapopulations evolving under unidimensional
215 divergent selection did not show local adaptation until late in the experiment. Meanwhile,
216 metapopulations evolving under multidimensional divergent selection produced rapid, but ultimately
217 transient local adaptation.

218 Fitness was generally higher in unidimensional lines than multidimensional lines from the outset of
219 the experiment (Figure 3). This could reflect rapid adaptation in unidimensional lines over the first
220 four cycles (between the start of experimental evolution and the first assay), but to both
221 environments, since local adaptation did not increase. Alternatively, it could be that, despite
222 calibration, unidimensional selection was generally slightly weaker than multidimensional selection
223 due to some experimental artefact such as the filtering process.

224 The patterns of home and away fitness in our data indicate that the factor which determined local
225 adaptation was fitness of populations in the away environment since home fitness increased linearly,
226 at similar rates, under both dimensionality treatments. The challenge is to understand the different
227 patterns of fitness in away environments under unidimensional and multidimensional divergent
228 selection. These patterns are inconsistent with the expectation that unidimensional divergent
229 selection might concentrate selection strongly onto a few loci, leading to rapid local adaptation, and
230 the opposite for multidimensional selection. What property of dimensionality limits fitness in away
231 environments at an early stage for multidimensional divergent selection, but at a later stage for
232 unidimensional divergent selection? Here, we propose an explanation based on the idea that locally-
233 adaptive genotypes may be expected to have different fitness effects in away environments
234 depending on the dimensionality of the environmental difference. This hypothesis is consistent with
235 our results but will need to be verified by further experiments.

236 **Specialists vs Generalists**

237 Genotypes increased in frequency because they were advantageous under their home conditions, and
238 indeed in both treatments we saw a gradual increase in home fitness across the experiment. Given

239 that the migration rate was low and hence gene flow from the other deme was not expected to be a
240 strong opposing force⁶, we suggest that the main determinant of local adaptation was the fitness of
241 these genotypes in the away environment. With respect to their performance under away conditions,
242 three possible genotype classes might compete. A genotype class with positive fitness effects in an
243 away environment suggests a generalist genotype, whereas a genotype class with negative away
244 fitness effects suggests a specialist in the home environment with antagonistic pleiotropy. The third
245 genotype class represents conditional neutrality, where the genotype is a specialist in the home
246 environment but there are no observed costs in the away environment⁵⁰. Interpreting patterns of local
247 adaptation in our experiment requires consideration of the away fitness of genotypes that were
248 available as standing variation at the start of the experiment and those that were generated by
249 mutation and recombination as the experiment progressed.

250 Our data suggest that the relative contribution of these genotypic classes to adaptation may have
251 varied by dimensionality and over the course of experimental evolution. Overall patterns in home vs
252 away fitness of metapopulations under unidimensional divergent selection suggest conditional
253 neutrality because home fitness increased with little change in away fitness. However, this varied
254 among stressor pairs, with some evidence for generalist effects early and antagonistic effects late in
255 the experiment for two stressor pairs (Sal and Misc). Meanwhile, patterns in metapopulations under
256 all four stressor pairs from the multidimensional divergent selection treatment were consistent with
257 early local adaptation driven by specialist genotypes with small (conditional neutrality) or negative
258 (antagonistic) effects on fitness in the away environment. Generalist genotypes that spread later could
259 explain reduced local adaptation without loss in home fitness. These genotypes could have been
260 created by recombination between conditionally-neutral genotypes that were exchanged between
261 demes.

262 The initial response to divergent selection is likely to depend on standing genetic variation, in this case
263 the diverse set of 500 clones used to initiate the experimental metapopulations²⁶. Our results imply
264 that this included clones that increased fitness under all experimental treatments. They further imply
265 that clones with high fitness in response to one stressor also tended to have increased fitness in
266 response to at least one other stressor or to be neutral in the away environment and, hence, they
267 behaved as generalists in the context of the unidimensional treatments²². However, this generalist
268 behaviour appears to be limited: clones with positive fitness effects on average across the three
269 stressors used in a multidimensional treatment were neutral or antagonistic with respect to average
270 fitness across the three different stressors used in the away environment, i.e. they exhibited
271 conditional neutrality^{50,51} or were specialist. This would explain the initial increase in home fitness

272 under both treatments, accompanied by local adaptation under multidimensional divergent selection
273 but not under unidimensional selection.

274 Later in experimental evolution, responses are likely to have changed due to increased frequency of
275 initially rare clones, or generation of new clones by recombination or mutation^{26,28}. For unidimensional
276 metapopulations this resulted, on average, in a continued increase in fitness in the home environment
277 but either little change or a decrease in fitness in the away environment. This implies that further
278 evolution involved greater specialisation, i.e. that the clones with the highest fitness in one
279 environment now tended to have antagonistic fitness effects in the other environment, unlike those
280 clones initially available. Not surprisingly, this pattern was somewhat dependent on the specific
281 stressors used (Figure 3). By contrast, under multidimensional divergent selection, conditionally
282 neutral specialist genotypes might have flowed freely between demes, owing to the lack of negative
283 effects in the away environment, resulting in a progressive loss of local adaptation without a reduction
284 in home fitness^{52,53}. It is conceivable that generalist clones might have been generated by
285 recombination between specialist conditionally neutral genotypes which were adaptive in different
286 environments. This pattern was consistent across stressor combinations (Figure 3).

287 Another way to picture this behaviour is to consider adaptive trajectories on a multidimensional
288 adaptive landscape, using Fisher's Geometric Model²⁴. As a result of calibration in pilot experiments,
289 the starting populations were equally maladapted to each environment: the fitness reduction in
290 response to each stressor or stressor combination was initially equal. Therefore the two demes had
291 to adapt along trajectories of equal length, but the angle between these trajectories differed^{54,55}. The
292 divergence between these adaptive trajectories was shaped by the co-variance of environmental
293 variables and the genetic covariance of adaptive traits at the point of divergence (as could be
294 respectively described by an E-matrix or G-matrix^{28,56,57}). The angle of divergence between adaptive
295 trajectories is extremely challenging to predict *a priori* and may not be as wide for unidimensional
296 divergent selection as one might assume⁵⁸. For instance, high and low salinity shocks, as used in this
297 study, might intuitively imply adaptive trajectories with a 180° angle of divergence (i.e. antiparallel
298 trajectories *sensu* Bolnick *et al.*⁵⁵). However, generalist 'osmotic shock' alleles could theoretically
299 provide adaptation in both environments, producing adaptive trajectories with little divergence, and
300 limited opportunity for local adaptation. For multidimensional divergent selection, the angle between
301 trajectories is even more difficult to predict, although we argue that it is likely to be wider given a
302 lower probability of 'one-size-fits-all' generalist genotypes, at least in the short-term. In the longer
303 term, this would also depend on the nature of mutational input; do new alleles tend to have
304 pleiotropic effects, and if so, in which direction⁵⁹?

305 Specific features of our experiments might restrict the generality of the results. Firstly, *B. plicatilis* is a
306 facultatively sexual organism. This is not an uncommon reproductive strategy⁶⁰ but it may influence
307 the dynamics of local adaptation. Sex and recombination both create and decompose locally adapted
308 genotypes^{23,48,61}. However, a small amount of sex can be sufficient to gain most of its advantage and
309 so we would not expect major differences in behaviour from obligate-sexual species⁶².

310 Secondly, our means of delivering multidimensional selection mimics spatially and temporally
311 heterogeneous selection within a single population, such that stressors are experienced separately
312 and do not interact²³. This was necessary for experimental tractability but reality is often likely to be
313 more complex. Selection pressures often do not influence all individuals in all generations, but they
314 are rarely independent and might often have interactive effects on fitness, whether experienced
315 simultaneously or successively^{63,64}. How this might influence the outcome of local adaptation is
316 unknown but it is unlikely to change the broad distinction between low and high dimensionality of
317 divergent selection that we sought to address.

318 Previously, unidimensional vs multidimensional divergent selection comparisons have focused on the
319 effects of diluting divergent selection over few vs many traits or loci. For instance, Nosil *et al.*¹¹ present
320 the 'strong selection' and 'weak multifarious' hypotheses in terms of distributing a fixed quantity of
321 divergent selection over few vs many traits, and the consequences this may have for overcoming the
322 homogenising force of gene flow.

323 The interpretation of our study modifies this argument. In addition to the way selection is
324 concentrated or diluted across the genome, it may be important to consider the performance of alleles
325 when exposed to a range of environments. Models and perspectives have tended to focus on alleles
326 with antagonistic effects¹². However, outcomes are likely to be different if conditionally neutral and
327 generalist alleles are available. The strength of the divergent component selection depends on the
328 complex interaction between the environment and the available genetic variation, in ways that change
329 as evolution proceeds. Unidimensional divergent selection may, in the short term, be less divergent
330 than one might think due to the availability of generalist alleles. However, we suggest that on long
331 evolutionary time-scales, this is likely to act only as a short-term barrier to local adaptation. Once
332 generalist alleles fix, the G-matrix will have in effect been rotated to align with the E-matrix, allowing
333 specialist alleles to drive local adaptation. We suggest that, where gene flow is ongoing,
334 multidimensional divergent selection may be more likely to create an ecological generalist in the long
335 term than locally adapted genotypes. However, in allopatry, multidimensional divergent selection
336 remains likely to form locally adapted genotypes. Furthermore, by potentially impacting more loci,
337 multidimensional divergent selection may access more standing or mutational genetic variation and

338 so drive long-term divergence and ecological speciation, through coupling with other barriers such as
339 assortative mating⁶⁵. Further studies that connect the dimensionality of novel selection pressures, the
340 genetic basis of adaptation, and fitness in a range of environments are needed to test these ideas.

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349 Author contributions

350 RKB conceived the project. NJW and IE designed the experiment with contributions from all authors.
351 IE established the starting populations. NJW performed the experiments. NJW and APB analysed the
352 data. NJW wrote the manuscript with contributions from all authors.

353 Declaration of interests

354 The authors declare no conflicts of interest

355 Figure legends

356 **Figure 1: Experimental design**

357 Panel A depicts the treatment-replicate-metapopulation structure. Two demes formed a
358 metapopulation, with divergent selection applied across the metapopulation. There were four
359 unidimensional forms of divergent selection (four stressor pairs) and four multidimensional forms of
360 divergent selection (four combinations of three stressor pairs). Each of these was replicated four
361 times, yielding 16 metapopulations per dimensionality treatment. Panel B details the weekly cycle of
362 experimental evolution. Panel C depicts a sympatric-allopatric (SA) local adaptation test in which the
363 fitness of each deme of the metapopulation was assayed in each environment.

364 **Figure 2: Local adaptation under experimental divergent selection with gene flow.**

365 Points represent SA estimates for each metapopulation. Curved thick lines display model fit using
366 the first and second order polynomial of cycle, dimensionality, and their interaction. Random effects
367 not presented in this Figure. See also supplementary Figures S3-S5 and supplementary Tables S2-S5.

368 **Figure 3: Home vs away fitness of metapopulations.**

369 A: Loess fits to data grouped by dimensionality. Red/blue lines represent loess fits (span = 3) for each
370 environment (without allowance for replicate effects). Data shown are the average fitnesses of the
371 two demes per metapopulation in their home/away environments. Grey lines connect fitness
372 through time for each metapopulation (see also supplementary Figure S2).

373 B: Data separated by stressor pair: unidimensional treatment on the top row, multidimensional
374 treatment on the bottom row. Lines represent the best model fits for home and away fitnesses,
375 separately.

376

Treatment	Stressor pair	Environment 'A'	Environment 'B'
Unidimensional	pH	Alkali (Alk) 0.0288 M NaOH solution	Acid 0.025 M HCl solution
	Salinity (Sal)	High Salinity (HS) 60 g/L salt solution	Low Salinity (LS) 0 g/L salt solution (pure water)
	Insecticide (Ins)	Neurotoxin (Neuro) 31.25 µg/L permethrin solution (Lignum)	Digestive Inhibition (DI) 6.6 g/L <i>Bacillus thuringiensis</i> subsp. <i>kurstaki</i> 54% w/w granule (DiPel DF)
	Miscellaneous (Misc)	Hot 40°C	Ethanol (Eth) 6% ethanol solution
Multidimensional	Sal-Ins-Misc	HS-Neuro-Hot	LS-DI-Eth
	pH-Ins-Misc	Alk-Neuro-Hot	Acid-DI-Eth
	pH-Sal-Misc	Alk-HS-Hot	Acid-LS-Eth
	pH-Sal-Ins	Alk-HS-Neuro	Acid-LS-DI

Table 1: Stressors used in this experiment.

Individual stressor pairs (unidimensional), and combinations of stressor pairs (multidimensional) given in bold. Each stressor, other than those on the salinity axis, is given as relating to a standard solution of 12 g/L saltwater. Any addition of, for instance, HCl to form an acidic shock media was sufficiently small as to not impact salinity to the degree that fitness would be affected.

Variable	F	DF	Residual DF	p-value
Poly(Cycle, 2)	3.54	2	27.86	0.043
Dimensionality	2.20	1	29.96	0.147
Poly(Cycle, 2) : Dimensionality	7.17	2	28.16	0.003

Table 2: Model statistics. Analysis of deviance table for a linear mixed effects model explaining variation in SA using the first and second order polynomial of cycle, dimensionality, and their interaction. P-values obtained through type II Wald F tests with Kenward-Roger degrees of freedom.

381 STAR Methods

382 Resource Availability

383 Lead contact

384 Further information and requests for resources and reagents should be directed to and will be
385 fulfilled by the lead contact, Nathan White (nwhite3@sheffield.ac.uk)

386 Material availability

387 This study did not generate any new unique reagents

388 Data and code availability

- 389 • Local adaptation data has been deposited at DataDyrad and are publicly available as of the
390 date of publication. DOIs will be listed in the key resources table
- 391 • All original code has been deposited at DataDyrad and is publicly available as of the date of
392 publication. DOIs will be listed in the key resources table.
- 393 • Any additional information required to reanalyse the data reported in this paper is available
394 from the lead contact upon request.

395 Experimental Model and Subject Details

396 ***Brachionus plicatilis***

397 *Brachionus plicatilis* is a facultatively sexual aquatic metazoan (Rotifera, Monogononta). It reproduces
398 asexually at low population density, then switches to sexual reproduction at high population density⁶⁶.
399 Sexual reproduction in monogonont rotifers produces diapausing embryos (also known as ‘resting
400 eggs’) that do not hatch under normal culture conditions, requiring a short period of dormancy and
401 specific conditions to trigger hatching⁶⁷. *B. plicatilis* rotifers used in this study were derived from
402 diapausing embryos in sediment samples from two brackish ponds in the Juncar-Segura basin,
403 Albacete province, Spain: Laguna del Salobrejo (38°54.765’N, 1°28.275’W, 0.36 km² surface area) and
404 Hoy Yerba (38°46.7667’N, 1°26.1167’W, 0.03 km² surface area). Diapausing embryos were hatched in
405 isolation under laboratory conditions and allowed to form clonal cultures. Due to the *B. plicatilis*
406 species complex comprising 15 species⁶⁸, sequencing of mitochondrial cytochrome c oxidase subunit
407 1 (CO1) was performed to confirm species identity as *B. plicatilis sensu stricto*⁶⁹. In total, 54 clonal
408 cultures were selected for the experiment, 27 originating from each pond, on the basis that they grew
409 well under laboratory conditions so limiting later laboratory adaptation. Further details of clones and
410 collection methods are in García-Roger and Ortells⁷⁰. These 54 cultures were combined in equal
411 proportions and grown to high density to produce diapausing embryos via sexual reproduction. These
412 embryos were hatched and cultured to form a genetically diverse set of 500 clonal cultures. These

413 cultures were pooled, and samples were taken from this pool over four successive weeks to form
414 replicate experimental populations.

415 All cultures were maintained in 12 g/L artificial seawater (TetraMarine) at 25°C on a 12:12 light-dark
416 cycle. Cultures were grown on a seven-day cycle. For each culture, after 7 days of population growth,
417 a new culture was established through passaging individuals from the old high-density culture to
418 establish a new culture at population density of 6 rotifers ml⁻¹. Total culture volume was made up to
419 400ml with fresh media (12g/L artificial seawater) containing 250µl/L *Nannochloropsis* paste as food
420 (Seahorsebreeder). 50ml of fresh media containing concentrated *Nannochloropsis* paste (4ml/L) was
421 added to each culture on day 3, increasing culture volume to 450ml.

422 Method Details

423 *Pilot experiments to calibrate shock duration for each stressor*

424 This study required the fitness effects of multiple forms of selection to be calibrated in order to
425 deliver an equally strong shock, measured as 24-hour survival rate (see Main Text), and thus
426 generate divergent selection of comparable strength across metapopulations and treatments. To
427 calibrate these shocks, we performed pilot experiments using the same starting population as for
428 experimental evolution. We aimed for a 45% survival rate. This rate was selected as preliminary
429 experiments showed an approximately 90% survival rate when filtered and ‘shocked’ with standard
430 growth media, hence 45% survival would represent a 50% mortality rate due to the shock media. To
431 standardise each shock so that a 45% survival rate was achieved, we varied the incubation duration
432 in the shock medium. The eight forms of selection are listed in Table 1.

433 Populations were grown on a weekly cycle over four weeks as described in the Main Text. At
434 passage, rotifers from all populations were pooled together and passaged to establish a mixed
435 population from which new populations were established. Shocks were delivered as described in the
436 Main Text, but with variable incubation durations from 30-70 minutes. We calculated survival rate as
437 described in the Main Text and fit a linear model whereby incubation duration, stressor and their
438 interaction were used to explain the variance in survival rate. We interpolated from these fits to
439 identify the predicted duration that would produce 45% survival.

440 All stressors displayed a negative linear relationship between incubation duration and survival
441 (Figure S1). The durations predicted to produce 45% survival are presented in Table S1. These
442 durations were used as the incubation periods for each shock in experimental evolution.

443 *Unidimensional vs multidimensional divergent selection*

444 To test the hypothesis that local adaptation would vary with the dimensionality of divergent selection,
445 we experimentally evolved populations of *B. plicatilis* under two selection treatments:
446 ‘unidimensional’ divergent selection and ‘multidimensional’ divergent selection. Unidimensional
447 divergent selection, as defined here, imposed selection using pairs of environmental stressors, each
448 differing from the ancestral environments in one way. Multidimensional divergent selection, as
449 defined here, imposed selection using paired combinations of three environmental stressors each of
450 which differed from the ancestral environment. The two alternative environmental conditions of a
451 stressor are hereafter referred to as ‘stressor pairs’ (e.g. unidimensional high and low salinity are a
452 stressor pair; Table 1). Stressors in a pair may not generate strictly antiparallel selection and a single
453 stressor might require adaptive responses in more than one trait. Therefore, our unidimensional
454 treatments may not impose divergent selection on a single environmental or phenotypic axis.
455 Nevertheless, our design contrasts low dimensionality of divergent selection in the unidimensional
456 treatment with higher dimensionality in the multidimensional treatment.

457 In order to test the effects of dimensionality *per se*, rather than the effects of specific environmental
458 variables, we nested four different stressor pairs (or stressor pair combinations) within each
459 treatment. Our unidimensional treatment was replicated four times using the stressor pairs: ‘pH’ (acid
460 vs alkali), ‘Salinity’ (“Sal” = high vs low salinity), ‘Insecticide’ (“Ins” = neurotoxin vs digestive inhibition)
461 and ‘Miscellaneous’ (“Misc” = hot vs ethanol). Our multidimensional treatments were also replicated
462 four times using the four unique combinations of three of these four stressor pairs: ‘Sal-Ins-Misc’, ‘pH-
463 Ins-Misc’, ‘pH-Sal-Misc’ and ‘pH-Sal-Ins’ (Table 1).

464 Each pair of stressors was applied over a pair of cultures such that each culture was exposed to a
465 unique stress or stress combination and the two cultures in the pair were linked by migration (see
466 below). We will refer to the individual cultures as ‘demes’ and the pairs as ‘metapopulations’. Each
467 stressor pair was replicated four times: experimental demes were populated by sampling from pools
468 of the 500 clones described above, over four successive weeks. Therefore, in total there were 16
469 unidimensional and 16 multidimensional metapopulations (four pairs of stressors per dimensionality
470 x four culture replicates; Figure 1a).

471 ***Experimental cycle***

472 Metapopulations experienced a weekly cycle of growth, selection, passaging and migration (Figure
473 1b). Demes were established at 6 rotifers ml⁻¹ (asexual females; 400ml cultures, hence an expected
474 population of 2400 individuals). After 6 days of asexual growth each deme was exposed to selection
475 via a shock stressor designed to produce a large reduction in population density from pre-shock to 24
476 hours post-shock. For brevity, we refer to this reduction in population density metric as ‘survival’,

477 although it may include some asexual reproduction. All shock stressors were calibrated in pilot studies
478 to produce 45% survival relative to the control at the outset of experimental evolution, i.e. before any
479 evolutionary response, using an identical source population (Supplementary Information; Figure S1).
480 In this way, the strength of selection was standardised across all metapopulations, including between
481 unidimensional and multidimensional treatments.

482 To deliver each shock, rotifers were removed from culture by filtration through a 50 μ m mesh bag that
483 was rapidly transferred to a shock medium for the specified duration (Table S1). Rotifers were then
484 washed back into their original culture medium. For multidimensional selection, prior to filtration,
485 each deme was split into three equally sized cultures, each of which was filtered independently and
486 exposed to a different shock. Cultures were merged after the shock so that, in the long term, the whole
487 population experienced all three shocks. This strategy made it possible to impose shocks that did not
488 interact in their effects on survival and, therefore, to ensure that each stressor contributed equally to
489 selection on the deme, with overall strength equal to the unidimensional treatments. This would not
490 have been possible with simultaneous shocks. In nature, it is often the case that selection pressures
491 do not impact all individuals in all generations and yet adaptation to multiple challenges is necessary
492 for long-term success. Therefore, our experimental strategy is not actually unlike the selection
493 experienced by natural populations.

494 At day 7 (24 hours post-shock), population density was measured, and demes were passaged to form
495 new cultures. Following passage, exactly 24 of the approximately 2,400 rotifers were reciprocally
496 transferred between demes (1% migration). Migration did not use filtration; hence the microbiota was
497 also shared between demes. A 1% migration rate was selected based on experimental evolution
498 guidance^{71,72} to achieve homogenisation of neutral genomic regions ($Nm > 1$) without impeding
499 response to moderate selection pressures per locus.

500 With this passaging routine, there is no contribution from sexual reproduction. To include the effects
501 of recombination arising from sexual reproduction, the remaining non-passaged culture was retained
502 for three more days to allow for additional sexual reproduction. On day 10, diapausing embryos were
503 collected from culture sediment and incubated in 1.2ml 50 g/L artificial seawater at 4°C in the dark for
504 a 2-week dormancy period, after which they were hatched via incubation in 6 g/L artificial seawater
505 at 25°C under constant illumination. 24 hatchlings per deme were transferred to tubes containing
506 30ml 12g/L seawater for a further 4 days, then added to the experimental demes coinciding with
507 weekly passage, to form a 1% sexually produced contribution of new clonal genotypes.

508 ***Population density estimates***

509 Population density was determined by counting the number of rotifers in four 1ml samples under a
510 stereo microscope. This was done at regular intervals throughout the experiment for use in fitness
511 estimation, passaging at a consistent population size and monitoring adaptation to the laboratory and
512 experimental regime. Densities at day 6 were also determined at 4-week intervals in line with local
513 adaptation assays, before shocks were delivered. These enabled comparisons of growth rates without
514 the additional effect of the stressors. Densities at day 7 were determined each week, although these
515 are influenced by both growth rate and shock survival.

516 ***Local adaptation assays***

517 To measure the evolution of local adaptation over time, we performed classic ‘home vs away’
518 reciprocal transplant assays at regular intervals every four passages throughout the experiment
519 (Figure 1c). During passaging, a reciprocal ‘away’ deme was established in addition to the
520 experimental ‘home’ deme. Whilst the experimental home deme received the same stressor (or
521 stressor combination) treatment as usual, the away deme received the alternate
522 stressor/combination within the stressor pair. Thus, both demes in the metapopulation were assayed
523 under each stressor. Stressors were delivered as described above with unidimensional populations
524 receiving a single shock and multidimensional populations being split into thirds to receive separate
525 shocks before being re-combined. Population density was counted pre-shock on day 6, and 24-hours
526 post-shock on day 7 (after re-combining the samples exposed to different shocks for multidimensional
527 populations) to calculate a 24-hour survival rate for each combination. Because selection was
528 delivered via short-term exposure to stressors, and all cultures were maintained in the same
529 conditions, no separate acclimation period was needed before these assays.

530 [Quantification and Statistical Analysis](#)

531 ***Quantifying local adaptation***

532 Survival at 24 hrs was used as a measure of fitness in response to a given shock. A quantitative measure
533 of local adaptation describes the fitness interaction between population and environment, ideally
534 capturing both population-level and environment-level sources of variation⁴⁹. The measure which
535 achieves this is known as the ‘sympatric-allopatric’ (SA) contrast. At 4-cycle intervals and where all
536 local adaptation assays were successfully performed within a metapopulation, we calculated an SA
537 contrast per metapopulation as;

$$538 \quad SA_{AB} = \frac{w(A_a) + w(B_b)}{2} - \frac{w(A_b) + w(B_a)}{2}$$

539 where w is fitness $\left(\frac{\text{density day 7}}{\text{density day 6}}\right)$, A/B are demes of the AB metapopulation and a/b are the local
540 environments for A and B respectively.

541 **Statistical methods**

542 Statistical analyses were performed in R v1.4.1106⁷³. We modelled local adaptation variables using
543 linear mixed-effects models with random slopes and intercepts per metapopulation to account for
544 longitudinal non-independence and nesting of replicates within stressor pairs.

545 Do the speed and magnitude of local adaptation vary by dimensionality?

546 To test whether the speed and magnitude of local adaptation (SA) varied by the treatment
547 (dimensionality) over time, we fitted a fixed effects structure in which SA could vary by a second order
548 polynomial of cycle (i.e. duration of experimental evolution) that varied among stressor pairs;

549 $SA \sim poly(cycle, 2) * dimensionality + random(poly(cycle, 2) | metapopulation)$

550 We tested the fixed effects using Type II sums of squares F tests with Kenward-Roger adjustment for
551 degrees of freedom defined by the *Anova()* function in the *car* package for R⁷⁴. Model comparisons via
552 significance testing of fixed effects and data visualisation indicated that the polynomial was justified.
553 This polynomial also was used in the random effect structure.

554 Could identified effects be explained by differences among stressor pairs alone, or by other factors that
555 bias the SA contrasts?

556 We repeated the above sequence of model testing using stressor pair as a predictor variable instead
557 of dimensionality;

558 $SA \sim poly(cycle, 2) * stressor\ pair + random(poly(cycle, 2) | metapopulation)$

559 To confirm the non-linear evolution of local adaptation (SA) and its dependence on dimensionality,
560 we defined *a priori* contrasts of estimated marginal means from this model between unidimensional
561 stressor pairs and multidimensional stressor pairs at three timepoints (passages 5, 21 & 41) using the
562 *emmeans* package for R⁷⁵. These three contrasts correspond to the prediction that, if the trajectories
563 of SA over time varied by dimensionality, we would expect to observe differences between stressor
564 pairs grouped by dimensionality at the start, midpoint or end of the experiment.

565 For validation, we repeated this statistical analysis using a 'local-foreign' (LF) measure of local
566 adaptation. The 'local-foreign' (LF) contrast^{1,49} quantifies differences between the fitnesses of two
567 populations when in the same environment. It is therefore a measure of local adaptation for a specific
568 environment, rather than across a metapopulation. These were calculated as;

569 $LF_a = w(A_a) - w(B_a), \quad LF_b = w(B_b) - w(A_b)$

570 We repeated our above statistical analyses using LF rather than SA as response variable.

571 Finally, SA measures of local adaptation can be positive because of large asymmetry between the two
572 estimates of LF within a metapopulation, in which case there would be increased gene flow from one
573 deme to another^{1,49}. To test whether LF asymmetry might be driving patterns of SA, we also fitted
574 models for the absolute difference between LF statistics within each metapopulation (LF Asymmetry
575 = $|LFa - LFb|$).

576 *What components of home vs away fitness are responsible for any observed differences?*

577 SA is a contrast between fitness of 'home' and 'away' combinations of demes and environments.
578 Patterns of local adaptation (measured as SA) over time can be due to variation in home fitness, away
579 fitness, or some combination. Therefore, to identify the cause(s) of the patterns in SA, we modelled
580 home and away fitness estimates separately. This used the same fixed and random effects structure
581 as above, but using fitness in either home or away environment as the response variable instead of
582 SA. Mean fitness over both demes within each metapopulation, rather than the fitness of each deme
583 individually, under home or away conditions was used in this analysis. This was to control for non-
584 independence of the two demes within a metapopulation due to reciprocal migration

585

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