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BMC Medicine

The importance of transdiagnostic symptom level assessment to understanding prognosis for depressed adults: analysis of data from six randomized control trials. --Manuscript Draft--

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| Abstract: | <p>Background: Depression is commonly perceived as a single underlying disease with a number of potential treatment options. However, patients with major depression differ dramatically in their symptom presentation and comorbidities, e.g. with anxiety disorders. There are also large variations in treatment outcomes and associations of some anxiety comorbidities with poorer prognoses, but limited understanding as to why, and little information to inform the clinical management of depression. There is a need to improve our understanding of depression, incorporating anxiety comorbidity, and consider the association of a wide range of symptoms with treatment outcomes.</p> <p>Method: Individual patient data from six RCTs of depressed patients (total n=2858) were used to estimate the differential impact symptoms have on outcomes at three post intervention timepoints using individual items and sum scores. Symptom networks (Graphical Gaussian Model) were estimated to explore the functional relations among symptoms of depression and anxiety and compare networks for treatment remitters and those with persistent symptoms to identify potential prognostic indicators.</p> <p>Results: Item-level prediction performed similarly to sum scores when predicting outcomes at 3 to 4 months and 6 to 8 months, but outperformed sum scores for 9 to 12 months. Pessimism emerged as the most important predictive symptom (relative to all other symptoms), across these time points. In the network structure at study entry, symptoms clustered into physical symptoms, cognitive symptoms, and anxiety symptoms. Sadness, pessimism, and indecision acted as bridges between communities, with sadness and failure/worthlessness being the most central (i.e. interconnected) symptoms. Connectivity of networks at study entry did not differ for future remitters vs. those with persistent symptoms.</p> <p>Conclusion: The relative importance of specific symptoms in association with outcomes and the interactions within the network highlight the value of transdiagnostic assessment and formulation of symptoms to both treatment and prognosis. We discuss the potential for complementary statistical approaches to improve our understanding of psychopathology.</p> | |

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| Response to Reviewers: | <p>Thank you for the comments. We have addressed the revisions and hope the manuscript is now suitable for acceptance.</p> <p>Reviewer reports:</p> <p>Reviewer #1: The authors have adequately addressed my comments.</p> <p>Reviewer #4: The authors have sufficiently addressed my concerns. The following are optional suggestions. I found this phrasing much clearer than the phrasing used in the manuscript; using this direct phrasing in the manuscript could clear up a lot of potential confusion about the sample.</p> <p>"There is an ongoing debate in the field whether the most central items derived from network models offer predictive utility"</p> <p>I think I made my point poorly in my previous review. I fully agree with the authors that the usefulness of network centrality is under contention. My concern was just a minor phrasing issue - all relevant items offer some predictive utility, hence central items (and noncentral items) offer predictive utility in the abstract. The debate is over whether central items are different than noncentral items in any meaningful way. I thought the authors point would be clearer phrased in a different way, e.g.:</p> |

"There is an ongoing debate in the field whether central items derived from network models have any special predictive utility" [any predictive utility beyond other items / etc]

We appreciate the clarification and have altered the sentence.

Line 389: "There is an ongoing debate in the field whether central items derived from network models offer predictive utility beyond other items (71–73)."

"The colours correspond to symptoms. We have tried a side by side bar graph but with so many variables we found that to be more difficult to read."

I couldn't find Figure 1 in the revised version, but if the colours correspond to symptoms, a legend should be added.

While we appreciate the suggestion, the symptoms are clearly labelled in each plot in Figure 1. Adding a legend would appear to be superfluous, crowd the figure, and conversely reduce the readability.

Editorial requests:

- Figure 1 is missing - unless 'Figure 2' is a combination of both figures - needs clarification

Apologies for the omission, it appeared on the pdf proof we had from the last submission, immediately before Figure 2, but we have uploaded afresh to ensure it is in the system.

- The following authors have not yet confirmed authorship:

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Can you please ask them to respond to our automated email, or to contact me directly, to confirm whether they agree to be listed as co-authors?
(alessandro.recchioni@biomedcentral.com)

The authors have notified us that they have confirmed their authorship by contacting you directly.

- Please make the following change to your section headings: change 'Introduction' to 'Background'.

We have made this change.

1 1 The importance of transdiagnostic symptom level assessment
2 2 to understanding prognosis for depressed adults: analysis of
3 3 data from six randomized control trials.
4 4
5 5

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35 Abstract

1 36 Background: Depression is commonly perceived as a single underlying disease with a number
2 37 of potential treatment options. However, patients with major depression differ dramatically in
3 38 their symptom presentation and comorbidities, e.g. with anxiety disorders. There are also large
4 39 variations in treatment outcomes and associations of some anxiety comorbidities with poorer
5 40 prognoses, but limited understanding as to why, and little information to inform the clinical
6 41 management of depression. There is a need to improve our understanding of depression,
7 42 incorporating anxiety comorbidity, and consider the association of a wide range of symptoms
8 43 with treatment outcomes.

11 44 Method: Individual patient data from six RCTs of depressed patients (total n=2858) were used
12 45 to estimate the differential impact symptoms have on outcomes at three post intervention
13 46 timepoints using individual items and sum scores. Symptom networks (Graphical Gaussian
14 47 Model) were estimated to explore the functional relations among symptoms of depression and
15 48 anxiety and compare networks for treatment remitters and those with persistent symptoms to
16 49 identify potential prognostic indicators.

19 50 Results: Item-level prediction performed similarly to sum scores when predicting outcomes at
20 51 3 to 4 months and 6 to 8 months, but outperformed sum scores for 9 to 12 months. Pessimism
21 52 emerged as the most important predictive symptom (relative to all other symptoms), across
22 53 these time points. In the network structure at study entry, symptoms clustered into physical
23 54 symptoms, cognitive symptoms, and anxiety symptoms. Sadness, pessimism, and indecision
24 55 acted as bridges between communities, with sadness and failure/worthlessness being the
25 56 most central (i.e. interconnected) symptoms. Connectivity of networks at study entry did not
26 57 differ for future remitters vs. those with persistent symptoms.

30 58 Conclusion: The relative importance of specific symptoms in association with outcomes and
31 59 the interactions within the network highlight the value of transdiagnostic assessment and
32 60 formulation of symptoms to both treatment and prognosis. We discuss the potential for
33 61 complementary statistical approaches to improve our understanding of psychopathology.

35 62

37 63 Keywords: item level analysis, network modelling, comorbidity, depression, anxiety
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Background

Psychological therapies and medication are effective treatments for depression (e.g., 1,2). However, effect sizes have been modest and gains in treatment outcomes have plateaued (3). Interventions for depression target a broad range of symptoms, and knowledge of ‘what’ is being intervened upon is not necessary to the delivery of most treatments, and poses problems for causal inference (4). To improve interventions, we may need to improve our knowledge of the structure of depression (5).

Depression is heterogeneous in terms of aetiology and symptom profile (6–8). Mood disorders are highly comorbid with anxiety disorders, and may share psychological and biological vulnerabilities (9,10). The risk of one disorder can increase the risk of another (11) and the same end state may be achieved via many different paths (equifinality) (12,13). These disorders are not discrete entities and as such, neglecting the symptomatic heterogeneity discards potential insights (14).

There is strong evidence that different symptoms are not equivalent or interchangeable (15) and studies of individual symptoms in the last decade have brought important understanding. For example, individual symptoms may differ in response to treatment (16,17), and have been shown to have a differential impact on functioning (18–20). Depression is a recurrent disorder with the probability of relapse strongly associated with the presence of residual depressive symptoms at the end of treatment (21,22). Comorbid anxiety disorders are related both to worse treatment outcomes (23) and to an increased risk of relapse (21). An assumed unidimensional view of depression, characterized by sum score (sum of symptom severity scores) measurement and prediction models conceals the variability within depression (24). Understanding the relative importance of comorbid symptoms may offer more information than severity of disorder alone and provide additional treatment and prognostic information (25). Large-scale, multisite clinical trial data, coupled with innovative statistical methods can provide categorisation and treatment optimisation to provide immediate benefits by informing clinical decisions (26–28).

There is also value in studying the relations among these symptoms. Network theory posits that the relationships between common affective, cognitive, and somatic symptoms of these disorders, may reflect potential causal pathways and elucidate maintenance mechanisms (29). Depression and anxiety have been modelled as symptom networks using cross-sectional and longitudinal data, demonstrating the interrelation between the symptoms of each disorder, where comorbidity results from mutually reinforcing interrelation between symptoms of each disorder (30,31). Anhedonia, anxiety, worry, fatigue and sadness are predominantly influential symptoms in these networks (5,32,33). The relationship between symptoms / mechanisms can help to predict outcome and potentially inform treatment targets and the development of treatments targeting specific mechanisms (34).

There are inconsistencies in the network literature exploring depression and anxiety, due to design, sampling, and variability arising from differing measurement (15,35). When attempting to discriminate between groups for the purposes of identifying whom may benefit from treatment (prognosis at group level), there are varying results from network comparison studies, where it has been suggested that densely connected networks may be less likely to recover (36). However, these differences are not always observed (37) and require large sample sizes to detect any effect. It is also unclear how these networks generalize to idiographic networks at the present stage. Past research has been conducted on small samples with low quality assessment of patients (or non-clinical samples) and lack of adequate consideration of comorbidity (missing out on the wider spectrum of anxiety disorders).

In this study we aim to:

- 116
 1 117 1) Identify important symptoms for outcome by examining the (differential) impact of
 2 118 individual symptoms on prognosis for adults with depression that took part in
 3 119 randomized controlled trials after seeking treatment in primary care; and assess
 4 120 whether individual symptoms offer predictive value above sum scores.
 5 121 2) Discern the functional relations among symptoms and clarify the interplay between
 6 122 highly comorbid symptoms of depression and anxiety disorders.
 7 123 3) Consider whether there are differences in the baseline symptom networks of patients
 8 124 that remitted vs those whose depression persisted, after treatment.

Method

Datasets

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 128 Data were drawn from a subset of the Dep-GP individual patient data (IPD) database (36).
 129 The formation of the Dep-GP IPD dataset has been described elsewhere (36). Bibliographic
 130 databases were searched up to 29th April 2020 for RCTs of unipolar depressed adults seeking
 131 treatment for depression or with depressive symptoms significant enough for them to seek
 132 treatment, recruited from primary care; had at least one active treatment arm; and used the
 133 CIS-R at baseline.

134 Studies were excluded if they were studies of: patients with depression secondary to a
 135 diagnosis of personality disorder, psychotic conditions, or neurological conditions; bi-polar or
 136 psychotic depressions; children or adolescents; feasibility studies; or were studies of adults
 137 with either depression or an anxiety disorder, rather than a primary depression with or without
 138 comorbid anxiety. Additional inclusion criteria for the present study were the use of the Beck
 139 Depression Inventory (2nd Edition) (BDI-II) (37) at study entry. The inclusion criteria ensured
 140 uniformity in the measurement of depressive and anxiety symptoms, chronicity of problems
 141 and determination of diagnoses including anxiety comorbidities.

142
 143 Data on all individual patients from all six eligible RCTs were included in the current study,
 144 these were: COBALT (38), GENPOD (39), PANDA (40), TREAD (41), MIR (42) and IPCRESS
 145 (43).

Measures

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 147
 148 Individual items from the BDI-II (37); and individual symptom subscales of the CIS-R (44),
 149 including duration of depression and anxiety which have been shown to be independently
 150 associated with prognosis for depressed adults (45).

Outcomes

151
 152 The primary outcome was endpoint depressive symptoms at three to four months post-study
 153 entry. Five of the studies used the BDI-II at three to four months, and one used the PHQ-9. A
 154 continuous 'depression severity' score was developed by converting the responses on each
 155 measure to a latent trait depressive symptom severity score (PROMIS T-Score) (46), using
 156 the expected a posteriori parameter from a multidimensional item-response theory based
 157 score conversion tool (47). Depressive symptoms (PROMIS T-Score) at six to eight months
 158 post-study entry, and nine to twelve months were secondary outcomes.

159 As a sensitivity analysis, the BDI-II scores were used as outcomes for the three timepoints;
 160 (five studies at three to four months; four studies at six to eight months, and three studies at
 161 nine to twelve months).

162 Data analysis

1 163 All analyses were performed in R 3.6 (48) and Stata 16.0 (49). Analysis code is available from
 2 164 <https://osf.io/wck6b/>. The data that support the findings of this study are available from the
 3 165 lead author of the Dep-GP (JB) subject to agreement from the chief investigators or data
 4 166 controllers of the individual RCTs. Restrictions apply to the availability of these data, which
 5 167 were used under license for this study.

8 169 Pre-processing

10 170 Datasets were combined and pre-processed together. There was no missing data at study
 11 171 entry. All items were investigated to ensure they met assumptions for inclusion in the network
 12 172 models, including assessing for: near zero variance; roughly equal variance of items;
 13 173 asymmetrical distributions; and topological overlap (50). Items were removed if they violated
 14 174 assumptions across all studies. We aimed to address topological overlap using the
 15 175 'goldbricker' function in R (51) with a threshold of 25% (correlations between items should
 16 176 have significantly different correlations with 25% of the other items), accepting minimal
 17 177 correlation of 0.5. The respective pair of items were combined into a single variable using
 18 178 Principal Component Analysis (PCA) if reasonable to combine from a clinical perspective.
 19 179 Items were afterwards rescaled to their original Likert scale values to make variances
 20 180 comparable across items (52).
 21 181

24 182 Association with outcomes

26 183 We aimed to examine the differential impact of individual symptoms on outcomes; and assess
 27 184 whether individual symptoms offer predictive value above sum scores. Sum score totals were
 28 185 entered into a linear regression model, while the item severity scores were entered into an
 29 186 elastic net generalized linear model (ENR) (53). ENR, a statistical method combining lasso
 30 187 and ridge regression approaches, minimizes overfitting and the use of ten separate, ten-fold
 31 188 repeated cross validation aids in assessing the effectiveness of the model. The item-level and
 32 189 sum-score models were compared using root mean squared error, mean absolute error and
 33 190 R^2 .
 34 191

36 192 As the item-level predictors were assumed to be correlated and that we wished to assess the
 37 193 explanatory power of individual predictors, we estimated the contribution of each item to the
 38 194 outcome prediction using Shapley Additive exPlanations (54), following ENR model
 39 195 estimation. Five hundred Monte Carlo repetitions were used to estimate each Shapley value.
 40 196 This metric is more accurate than other variable importance metrics when predictors are
 41 197 dependent (55). Items with large Shapley values are 'important', indicating the relative
 42 198 contribution of an item to the model while accounting for correlated features in the data.
 43 199

47 200 Network modelling

49 201 A Graphical Gaussian Model (GGM) aims to capture the direct effects (edges) between items
 50 202 while controlling for all other items in the network. A network was estimated by combining data
 51 203 from the six RCTs. The sample was then split into two networks (those with persistent
 52 204 symptoms vs. remitters: BDI-II score <10 at 3-4 months), the networks were re-estimated and
 53 205 compared using the network comparison test with 1000 iterations (56).
 54 206

56 207 We performed a number of analyses to test the robustness of the networks we estimated.
 57 208 While lasso (57), regularized GGMs (58) are most frequently reported in the network literature,
 58 209 lasso specificity has recently been shown to be lower than expected in dense networks with
 59 210 many small edges, leading to an increase in false positives (59). We also estimated an
 60 211 unregularized GGM using an iterative modelling procedure: the Extended Bayesian

212 Information Criterion (EBIC). Selecting unregularized GGMs according to EBIC has been
1 213 shown to converge to the true model (60). The algorithm runs 100 glasso models, re-fits all
2 214 models without regularization, and subsequently adds and removes edges until EBIC can no
3 215 longer be improved. The best performing model (EBIC parameter) was selected to provide a
4 216 conservative GGM estimation (high specificity).

5 217
6 218 Chronicity of disorders has been shown to interact with symptom severity (45,61). We
7 219 corrected for the potential confounding effects of duration of depression and anxiety within the
8 220 network models.

9 221
10 222 Combining data obtained from different studies holds the potential for between-study
11 223 differences to influence estimation. A network estimation procedure (fused graphical lasso:
12 224 FGL) (62) has been designed to manage this issue, however, this involves estimating
13 225 networks individually and penalizing between study differences. Where study size affects the
14 226 estimation of edges, this can lead to penalization based on sample size rather than on true
15 227 differences between the network structures (63). As such, it was decided to estimate based
16 228 on the combined sample and to compare this to the FGL network (joint estimation using a
17 229 fused penalty, and 10-fold cross validation), to assess the potential influence of group level
18 230 differences.

19 231
20 232 Finally, the network model was tested for the stability of expected influence centrality and the
21 233 accuracy of interrelations using a nonparametric bootstrapping procedure (1000 iterations)
22 234 (64). For details of these see the Supplementary material.

23 235 We obtained two types of information from the resulting network structures. First, symptoms
24 236 can form clusters or communities with other symptoms to which they are connected reflecting
25 237 commonalities between them. We estimated the community structure by using a bootstrapped
26 238 walktrap algorithm (65), investigated for item stability before selecting communities. Second,
27 239 the overall connectivity of a symptom, i.e. its connection to other symptoms, can be quantified
28 240 in a number of ways and is referred to as centrality. Some scholars have argued that activation
29 241 of a central symptom has the potential to activate associated symptoms in the network (66),
30 242 where symptom centrality is then interpreted as symptom importance, given that identifying
31 243 such symptoms may have the potential to elucidate the processes underlying comorbidity and
32 244 implications for treatment. Within the context of communities specifically, symptoms which
33 245 connected to more than one community are referred to as bridge symptoms. Within cross-
34 246 sectional networks (as explored here), we refer to centrality as a statistical parameter, i.e. the
35 247 strength of predictive associations between symptoms. Centrality does not automatically
36 248 translate into clinical relevance (67) and cautious interpretation is warranted (63). It requires
37 249 consideration of: how the symptoms activate within the network (flow or transfer); the
38 250 conceptual similarity between symptoms; and whether there is missing information on the
39 251 shared variance (68). Symptom centrality was calculated using: Expected Influence (EI:
40 252 strength of the relationships a given node has with other node); and the geometric mean of
41 253 the Participation Ratio (PR) and Participation Coefficient (PC); and normalized bridge
42 254 expected influence centrality (69). The PR quantifies the number and strength of edges, while
43 255 the PC takes the community structure into account (70).

256 257 258 259 260 261 262 263 264 265 266 267 268 269 270 271 272 273 274 275 276 277 278 279 280 281 282 283 284 285 286 287 288 289 290 291 292 293 294 295 296 297 298 299 300 301 302 303 304 305 306 307 308 309 310 311 312 313 314 315 316 317 318 319 320 321 322 323 324 325 326 327 328 329 330 331 332 333 334 335 336 337 338 339 340 341 342 343 344 345 346 347 348 349 350 351 352 353 354 355 356 357 358 359 360 361 362 363 364 365

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2 264 *Table 1: Descriptive table of studies included in the dataset. Summary of included variables*
 3 265 *provided in supplementary materials. * International Baccalaureate equivalent ** High school*
 4 266 *diploma equivalent.*

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6 267

7

8 268 Demographic details for the studies are presented in Table 1. Overall samples were
 9 269 comparable. The severity of depressive symptoms captured by BDI-II scores at baseline in
 10 270 the PANDA sample was lower than the other trials. Descriptive results are reported in the
 11 271 supplementary materials.

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*Table 1: Descriptive table of studies included in the dataset. Summary of included variables provided in supplementary materials. * International Baccalaureate equivalent ** High school diploma equivalent.*

Demographic details for the studies are presented in Table 1. Overall samples were comparable. The severity of depressive symptoms captured by BDI-II scores at baseline in the PANDA sample was lower than the other trials. Descriptive results are reported in the supplementary materials.

Association with outcomes

In order to assess the utility of item level models, we compared them to sum score models. For all item level models (Table 2), the optimal shrinkage parameters for the elastic net regression model were selected via minimum cross-validated error criterion ($\alpha = 0.1$ and $\lambda = 0.05$). While models performed similarly at three to four months and six to eight months, the item level elastic net regression model outperformed linear regression with BDI-II and CIS-R (sum of anxiety items) totals at the nine to twelve month time point. The sensitivity analysis performed similarly. Due to the absence of two studies (IPCRESS and PANDA) at the nine to twelve month endpoint, we reran the analyses for the earlier time points without these studies. This sensitivity analysis did not reveal any difference in the pattern of model performance.

Pessimism (Figure 1) was consistently the most important item; health anxiety was in the upper quartile at each time point; and concentration, failure/worthlessness, also in the upper quartile at three to four months; guilt and sleep at six to eight months; and somatic symptoms at nine to twelve months.

Table 2: Performance of the regression models. Sum scores: BDI-II and CIS-R; RMSE root mean squared error; MAE mean absolute error; R^2 proportion of the variance explained.

Figure 1: Shapley values for variable importance are plotted: (showing the difference contribution of items to predictions).

Network Modelling

For the individual items in the network model, near zero variance (e.g. due to floor and ceiling effects) was not observed. However, we saw asymmetric distributions (skew) on a number of items. As such, a Spearman covariance matrix was estimated and used to estimate the network model. Multi-collinearity was identified for two pairs of items (loss of pleasure with loss of interest, failure with worthlessness). New items were constructed using PCA for each

303 pair. The optimal model for the network analysis was an unregularized Graphical Gaussian
 304 Model using the EBIC.

305 A walktrap algorithm identified three, stable, symptom communities (median = 3, SD = 0.15,
 306 95% CI [2.71,3.29]). The three communities split into anxiety items, depressive cognitions and
 307 depressive physical symptoms. Bridging EI elucidated three bridging symptoms between the
 308 communities: sadness and indecisiveness (from the physical symptoms community); and
 309 pessimism (cognitive symptoms community);

310
 311 *Figure 2: Network plot (top) with communities. Bridge symptoms are categorized separately,*
 312 *however sadness and indecisiveness fall into community 1, and pessimism into community 3.*
 313 *The thickness of the edges indicates to what degree items are related, and the colour of the*
 314 *edges indicates the relationship sign (i.e. positive = blue, negative = red). Centrality estimates:*
 315 *PC/PR and EI (bottom).*

316 Centrality estimates (i.e. measures of the strength of connection to other items) are reported
 317 in figure 2. The EI correlation stability coefficient was high (0.75), suggesting that the ordering
 318 of items based on centrality remained the same after re-estimating the network with fewer
 319 cases (the probability the correlation between original centrality indices and centrality of
 320 networks based on subsets was 0.7 or higher) and can be reliably interpreted.

321 The estimates from the different metrics (EI and PC/PR) were correlated ($r = 0.58$). The most
 322 central symptoms were Sadness (PC/PR) and Failure/Worthlessness (EI).
 323 Failure/Worthlessness had a significantly higher EI centrality than twenty-one other symptoms
 324 (see supplementary material). The next most central nodes (EI) were sadness, self-criticism,
 325 and loss of energy (all z-score > 1), followed by concentration, loss of pleasure/interest, and
 326 fatigue (z-score > 0.96). While the next most central nodes when using PC/PR were
 327 pessimism, failure/worthlessness, and punishment (all z-score > 1), then guilt, indecisiveness,
 328 and suicidal thoughts (all z-score > 0.80). Notably, while suicidal thoughts were highly central
 329 according the PC/PR metric (z-score = 0.80) it was much less central using EI (z-score = -
 330 0.67). Loss of energy displayed the opposite relationship, more central for EI (z-score = 1.01)
 331 than PC/PR (z-score = -2.03). Loss of energy and obsessions were jointly the least central
 332 nodes using PC/PR, and obsessions was also the least central when using EI.

333 Robustness checks suggest the resulting Graphical Gaussian Model was stable and accurate.
 334 Stability and accuracy plots, individual networks (with the fused penalty) and the fused network
 335 model are supplied in the supplementary materials. Mean severity was not significantly
 336 correlated ($p < 0.05$) with EI ($r = 0.21$) or PC/PR ($r = -0.05$), while the standard deviation was
 337 significantly correlated for both EI ($r = -0.56$) and PC/PR ($r = -0.41$). Symptom severity was not
 338 associated with nodes being interconnected. Lower variability was associated with variability,
 339 which is the reverse of a more typical concern: differential variability driving the centrality of
 340 nodes (52).

341 The interrelation of the network and the FGL network were compared ($r = 0.72$), suggesting
 342 that between study differences had a small effect on network estimation. The network was
 343 corrected for the influence of duration of depression and anxiety, however the overall influence
 344 on edge estimation was negligible (interrelation between the corrected network and a network
 345 estimated without duration variables: $r = 0.997$). Overall, the resulting network model can be
 346 considered robust.

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348 **Network Comparison Test**

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2 349 Networks (unregularized) were compared (1000 iterations) for those who were classified as in
3 350 remission (n= 956) and those who were not (n = 1466). Mean severity differences at baseline
4 351 were significant for all items ($p < 0.001$). The correlation between networks was high ($r = 0.67$).
5 352 While there were significant difference between edges, the overall networks (see
6 353 supplementary material) did not differ in connectivity (global strength invariance: $p < 0.08$) and
7 354 post hoc tests were not warranted. There was only evidence of one difference in centrality
8 355 between the networks: somatic symptoms were more connected in the remitter network than
9 356 the persister network ($p < 0.001$).

14 358 **Discussion**

15 359 Individuals with depression also present with comorbidity and this could present an issue for
16 360 depression treatment. Understanding how symptoms influence one another across traditional
17 361 diagnostic boundaries, and how they influence important outcomes, may provide insights
18 362 relevant to the assessment and treatment of mood disorders. In this study we initially
19 362 examined the differential impact of individual symptoms on prognosis and assessed whether
20 363 individual symptoms offer predictive value above sum scores. The item level models of
21 364 outcomes post-treatment and the sum score models were similarly associated with outcomes
22 365 at three to four and six to eight months but explained considerably more variance at nine to
23 366 twelve months. Pessimism was consistently the most important predictor of future outcomes
24 367 (independent of its mean), indicating that experiencing pessimism rather than severity of the
25 368 symptom is responsible for this association. Secondly, we explored the functional relations
26 369 among comorbid symptoms of depression and anxiety disorders using network analysis. The
27 370 symptom network comprised of three communities clearly clustering into: anxiety items;
28 371 depressive cognitions; and depressive physical symptoms. The primary bridge symptoms
29 372 between communities were sadness; pessimism; and indecision. The most central symptoms
30 373 across both centrality metrics were sadness and failure/worthless. Finally, we analysed
31 374 differences in the symptom networks at study entry for patients that remitted vs. those whose
32 375 depression persisted, after treatment. Network comparison revealed no overall differences in
33 376 connectivity. Together, the present findings suggest the utility of item-level analysis in
34 377 informing the content of assessments and consideration of individual items over and above
35 378 scale scores when predicting prognosis.
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41 380 **Findings in context**

42
43 381 Exploring the associations with treatment outcomes revealed that item-level prediction
44 382 methods performed similarly to sum scores, and outperformed sum score models at the nine
45 383 to twelve month endpoint. It's not clear why there is a difference at this timepoint, while it was
46 384 not due to attrition between endpoints, it could be due to random variation. It may also reflect
47 385 the course of depression following intervention, or the cyclical nature of depression such that
48 386 individual items are better at predicting the relapse or maintenance of symptoms after benefits
49 387 of treatment have faded, or where an amelioration of symptoms occurred due to further
50 388 treatment post randomisation. There is an ongoing debate in the field whether central items
51 389 derived from network models offer predictive utility beyond other items (71–73). Pessimism
52 390 was not only the best predictor across outcomes, it was a central item (ranked 2nd on PC/PR
53 391 and 6th on EI centrality) that acted as a bridge between communities and showed strong
54 392 associations with sadness and failure/worthlessness. Sadness, comparatively, did not predict
55 393 well across time points. It is worth noting, that sadness falls within the physical symptom
56 394 community and pessimism within the cognitive community. The amenability to act on an
57 395 emotion (sadness), is understandably less than that of a cognition (pessimism), a target of
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396 cognitive therapy. While pessimism in association with a sense of failure / worthlessness may
1 397 negatively impact treatment engagement (i.e. the motivation to sustain goal pursuit in the face
2 398 of obstacles) (74). Given the central role and prognostic value of pessimism we might
3 399 speculate that it is associated with treatment factors, where pessimism hinders some people
4 400 making progress and may not be directly addressed by some psychological treatments.

6 401 Symptoms of anxiety and depression clustered into separate communities with certain
7 402 symptoms acting as bridges between communities. The bridge symptoms are statistically
8 403 relevant and theoretically linked: indecision is a symptom in the classifications of both
9 404 depression and generalized anxiety disorder; pessimism overlaps with worry (75); and the
10 405 strong cross-community edge of sadness to worry, was similar to findings in other studies
11 406 (32,76). The results therefore provide evidence that these bridging symptoms may be
12 407 important in the emergence of comorbidity between anxiety disorders and depression.

15 408 Planned comparisons of networks at study entry for those whose depression would persist
16 409 versus those who would be in remission, revealed no overall difference in connectivity, in
17 410 contrast to Van Borkulo et al. (77), but similar to Schweren et al. (78).

19 411 Overall, we found no correlation between centrality metrics and Shapley values. This extends
20 412 prior work on the association between centrality and the prognostic utility of items (71). Failure
21 413 / worthlessness was predictively important at three to four months, displayed high centrality
22 414 and is suggested to be a key symptom in depression and anxiety (30). The predictive utility of
23 415 health anxiety and somatic concerns may be considered alongside the observation from the
24 416 network comparison where there was a difference in centrality with somatic concerns more
25 417 connected in the remitter network. Health anxiety was in the upper quintile of variable
26 418 importance across timepoints, but relatively unimportant in terms of centrality. Not surprisingly,
27 419 given the conceptual overlap, with health anxiety, the strongest edge was with somatic
28 420 concerns. As such, the degree of concern for one's health, or attention to somatic symptoms,
29 421 whilst not playing a significant role within the maintenance of depression, may act as a
30 422 motivational spur to engage with treatment (in this way enabling rather than disabling the
31 423 individual). The absence of this anxiety may reflect an apathy about one's health which is not
32 424 captured by the motivational item in the BDI. While the predictive modelling did consider the
33 425 influence of each item independent of the other items, modelling the predictive value of
34 426 individual items may be improved by examining the association between the changes at
35 427 symptom level and the overall network (79,80).

38 428 The network derived in this study provides empirical phenomena that can be explained by
39 429 principles of network theory. This requires interpreting the network as a causal system, even
40 430 though we cannot infer temporal relationship between symptoms and there is an absence of
41 431 causal mechanisms within the external field (e.g. environmental factors) (29). These
42 432 limitations apply to most of the findings in the network literature, although overinterpretation is
43 433 common (81). Holding this in mind, we can consider possible pathways and mediating role of
44 434 symptoms through the network. For example, taking suicidal ideation as a clinically severe
45 435 symptom, we can identify the shortest path from worry (82) passing through sadness (bridge),
46 436 and from loss of pleasure/interest (83) to suicidal thoughts, passed through pessimism
47 437 (bridge). It is possible that any causal effect between these connections may be part of a
48 438 longer pathway within the network highlighting a need for greater attention to be given to
49 439 symptom interactions.

52 440 The statistical model investigates a symptom level, transdiagnostic conceptualization of the
53 441 symptom interactions for individuals diagnosed with depression participating in RCTs. These
54 442 interventions are based on biological or psychological theories, most notably Beck's cognitive
55 443 of theory of depression (84). Clinically, pragmatism trumps theoretical completeness; simple

444 interventions which achieve rapid change do not require a detailed appreciation of the potential
1 445 underlying mechanisms. However, oversimplified theories may restrict the ability to identify
2 446 causal patterns; and gaps emerge in practice where the model is suggested to not fit the
3 447 patient (85). More process-driven interventions targeting shared features of disorders have
4 448 been developed (86,87), yet there is no unifying theory. The findings presented may help
5 449 bridge the gap between disorder-specific theories and more transdiagnostic theories.
6 450 Considering how symptoms may interact can help clinicians and researchers to understand
7 451 underlying processes, and in turn to conceptualise their patients' difficulties in a way that
8 452 supplements existing knowledge. A functional analysis which integrates the association
9 453 between sadness and worry does not need to conceptualise the individual as having two
10 454 disorders, but can consider how, for the individual, this interaction is being fueled and may be
11 455 contributing to their distress.

12 456 The journey to develop models that provide both explanatory and predictive utility, will lead to
13 457 greater understanding of psychopathology (88). While the analysis presented is primarily
14 458 exploratory, it sets up clear testable hypotheses. These can be derived by examining the
15 459 central structures within the network, formulating hypotheses and testing on an independent
16 460 sample (89). For instance, whether the bridge edges belonging to pessimism, sadness and
17 461 indecisiveness re-emerge in an independent sample, or whether a discrete intervention
18 462 targeting pessimism would alter the network structure and lead to improved outcome. These
19 463 statistical methods may help inform how identifying pathways and targets may lead to
20 464 improved treatments all dependent on better assessment of symptoms.

21 465 **Strengths and Limitations**

22 466 This study has clear strengths, making use of a large sample of individuals participating in
23 467 RCTs for depression in primary care. The use of same assessment measures at study entry
24 468 removed the need to harmonise data across different measures for the network. While this is
25 469 less true of outcomes where issues of measurement errors arise from the use of PROMIS T-
26 470 Score, the sensitivity analyses provided confidence in the results.

27 471 The demographic balance across samples may affect generalisability however five of the six
28 472 trials were pragmatic trials more closely representative of patient populations. Most cases of
29 473 depression are treated in primary care, and the studies being set in primary care, improve the
30 474 potential generalisability to patients seen in this setting (90).

31 475 This study was limited to the use of aggregate/group level findings to inform within person
32 476 processes. However, the presence of an RCT outcome variable affords us the ability to detect
33 477 changes from one state (e.g. depressed) to another (e.g. remitted), which is typically not the
34 478 case with idiographic research studies that collect cross-sectional data. Exploring the
35 479 prognostics value of networks on deterioration of symptoms would extend the utility of network
36 480 analysis. This would however require generating idiographic networks, where reliable
37 481 estimation necessitates many timepoints (low sensitivity at 100 timepoints; (91)).

38 482 The accuracy of the network is limited by the items included and those omitted. The network
39 483 does not cover the breadth of comorbidity of symptoms across psychopathology and is
40 484 missing other environmental variables. Social adversity is associated with worse treatment
41 485 outcomes for some patients with depression, it can be important to assess for and address
42 486 these issues in clinic, where possible, to mitigate the risks of poor prognoses (92). There is
43 487 also the possibility that the centrality of sadness particularly, represents a strong association
44 488 with a latent variable rather than a specific role within the network (93).

45 489 The network models adjusted for duration of depression and anxiety, and a sensitivity analysis
46 490 assessed for the influence of between study variability, adding robustness to the findings.
47 491 While RCTs are used in the analysis, treatment arms were not factored in, and treated as
48 492 equivalent when estimating outcome. This may make the findings generalizable where
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findings are applicable regardless of treatment offered especially as the treatments included reflect those commonly available in primary care. Controlling for treatment group within the outcome modelling and controlling for relevant covariates (e.g. age, gender and social economic status) would also have improved the robustness of the findings. Such adjustments would have been fitting where the emphasis was on developing the best predictive model, instead of comparing the predictive ability of symptoms vs. total scores. More comprehensive prediction modelling using the Dep-GP dataset has been conducted (94). Additionally, our modelling did not include train/test split, as the whole sample was used in estimation of the network models. While a true out-of-sample 'holdout' dataset would have provided an unbiased evaluation of model fit, and is the preferred method for evaluating such models (95), the internal cross-fold validation employed in the symptom level model offers a layer of robustness supporting the final model estimates (where overfitting presents an issue). This study focussed on item-level analysis in comparison to sum-scores, future comparisons with models which may measure latent constructs in other ways, could be informative.

Single item symptom measurement will have unknown reliability and construct validity. Equally, the restricted range (e.g. a four-point scale) may not adequately capture the range of symptom variance occurring in the sample. Symptom measurement on a broader scale may improve the prediction of changes over time.

Conclusions

Our study used samples from high-quality randomised controlled trials, and the findings can be generalised to adults with depression being treated in primary care. This study has reiterated the importance of assessing for both depressive and anxious symptoms among adults seeking treatment for depression, and that valuable information about prognosis can be gained by understanding the interrelations between individual symptoms; information which is not available when considering sum scores or baseline symptom severity alone. This may be particularly important to longer term outcomes from treatment. Treatment selection and application is often hampered by comorbid symptoms and considered to introduce 'complexity' (96). Considering the bidirectional relationship between symptoms, and associations which may be mediated by another symptom (e.g. a bridge symptom) may help to consider comorbidity as normative.

While specific symptoms and associations have been highlighted, the aim is not to offer simple heuristics to inform clinical judgement and decision making. The relative importance of the highlighted associations should not be overweighed. The aim is not to identify individual items, but to consider the network of interactions. The critical role of individual symptoms and their interactions give rise to the activation of the network through pathways and anxiety and depressive cognitive and physical symptoms may activate one another via these pathways. This network highlights how symptoms of depression and anxiety disorders influence one another. Clinically, there is a need for treatments to adequately assess and address comorbidity.

List of abbreviations

- BDI-II – Beck depression invention (2nd edition)
- CIS-R - revised clinical interview schedule
- EBIC – extended bayesian information criterion
- EI – expected influence

541 ENR – Elastic net regression
 1
 2 542 FGL – fused graphical lasso
 3
 4 543 GGM – gaussian graphical model
 5
 6 544 IPD – individual patient data
 7
 8 545 MAE – mean absolute error
 9
 10 546 PC - participation coefficient
 11
 12 547 PCA – principal component analysis
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 14 548 PHQ-9 – patient health questionnaire 9
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 16 549 PR - participation ratio
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 18 550 RMSE – root mean squared error
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21 552 **Declarations**

22 553 **Ethics approval and consent to participate.**

23 554 Not applicable

24 555 **Consent for publication**

25 556 Not applicable

26 557 **Availability of data and materials**

27 558 The data that support the findings of this study are available from the authors of the
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 29 560 under license for the current study, and so are not publicly available.

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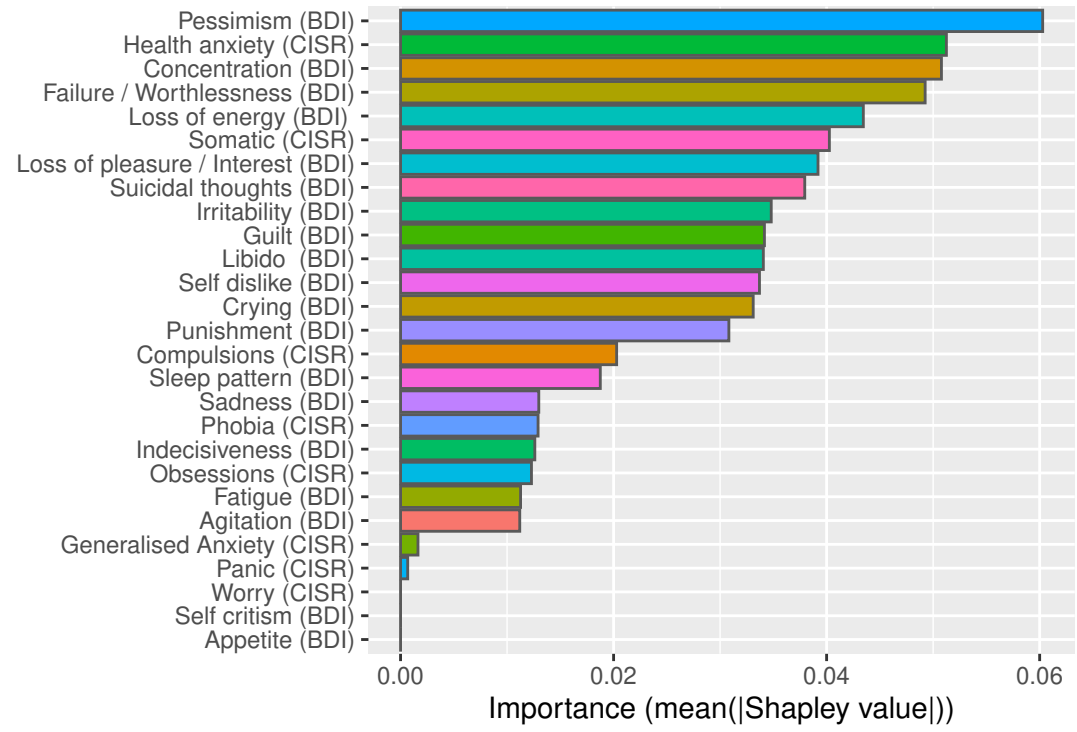
| | COBALT (N=469) | GENPOD (N=601) | IPCRESS (N=295) |
|------------------------------|--------------------------|--------------------------|---------------------------|
| Baseline BDI-II total | | | |
| Mean (SD) | 31.8 (10.7) | 33.7 (9.67) | 33.2 (8.80) |
| Median [Min, Max] | 30.0 [14.0, 60.0] | 33.0 [15.0, 60.0] | 33.0 [15.0, 58.0] |
| Gender | | | |
| Female | 339 (72.3%) | 408 (67.9%) | 200 (67.8%) |
| Male | 130 (27.7%) | 193 (32.1%) | 95 (32.2%) |
| Age | | | |
| Mean (SD) | 49.6 (11.7) | 38.8 (12.4) | 34.9 (11.6) |
| Median [Min, Max] | 50.0 [18.0, 74.0] | 38.0 [18.0, 74.0] | 34.0 [18.8, 74.6] |
| Employment Status | | | |
| Employed | 206 (43.9%) | 357 (59.4%) | 178 (60.3%) |
| Seeking employment | 151 (32.2%) | 123 (20.5%) | 35 (11.9%) |
| Not seeking employment | 112 (23.9%) | 121 (20.1%) | 82 (27.8%) |
| Education | | | |
| Degree or higher | 95 (20.3%) | 0 (0%) | 102 (34.6%) |
| A-level or Diplomas* | 123 (26.2%) | 0 (0%) | 88 (29.8%) |
| GCSE** | 131 (27.9%) | 0 (0%) | 62 (21.0%) |
| None or Other | 120 (25.6%) | 0 (0%) | 43 (14.6%) |
| Missing | 0 (0%) | 601 (100%) | 0 (0%) |
| Ethnicity | | | |
| White | 459 (97.9%) | 575 (95.7%) | 281 (95.3%) |
| Non-White | 10 (2.1%) | 26 (4.3%) | 14 (4.7%) |
| Diagnoses | | | |
| Number of Comorbid Diagnoses | 2.40 (1.09) | 2.39 (0.92) | 2.32 (0.99) |
| Generalized Anxiety Disorder | 312 (66.52%) | 410 (68.22%) | 186 (63.05%) |
| OCD | 79 (16.84%) | 114 (18.97%) | 62 (21.02%) |
| Panic Disorder | 67 (14.29%) | 51 (8.49%) | 16 (5.42%) |
| Agoraphobia | 61 (13.01%) | 75 (12.48%) | 28 (9.49%) |
| Social Phobia | 64 (13.65%) | 64 (10.65%) | 44 (14.92%) |
| Specific Phobias | 91 (19.40%) | 127 (21.13%) | 46 (15.59%) |
| Chronic Fatigue Syndrome | 343 (73.13%) | 476 (79.20%) | 220 (74.58%) |

| MIR (N=480) | PANDA (N=652) | TREAD (N=361) | Overall (N=2858) |
|------------------------------|--------------------------------|--------------------------------|-----------------------------------|
| 31.1 (9.91) | 23.9 (10.3) | 32.1 (9.24) | 30.4 (10.5) |
| 30.0 [14.0, 58.0] | 23.0 [2.00, 54.0] | 31.0 [14.0, 57.0] | 30.0 [2.00, 60.0] |
| 332 (69.2%) | 384 (58.9%) | 239 (66.2%) | 1902 (66.6%) |
| 148 (30.8%) | 268 (41.1%) | 122 (33.8%) | 956 (33.4%) |
| 50.7 (13.2) | 39.7 (15.0) | 39.8 (12.6) | 42.5 (14.1) |
| 51.0 [19.0, 84.0] | 38.5 [18.0, 73.0] | 39.0 [18.0, 69.0] | 42.0 [18.0, 84.0] |
| 237 (49.4%) | 433 (66.4%) | 230 (63.7%) | 1641 (57.4%) |
| 102 (21.2%) | 73 (11.2%) | 48 (13.3%) | 532 (18.6%) |
| 141 (29.4%) | 146 (22.4%) | 83 (23.0%) | 685 (24.0%) |
| 95 (19.8%) | 230 (35.3%) | 87 (24.1%) | 609 (21.3%) |
| 135 (28.1%) | 220 (33.7%) | 104 (28.8%) | 670 (23.4%) |
| 150 (31.2%) | 145 (22.2%) | 102 (28.3%) | 590 (20.6%) |
| 100 (20.8%) | 57 (8.7%) | 68 (18.8%) | 388 (13.6%) |
| 0 (0%) | 0 (0%) | 0 (0%) | 601 (21.0%) |
| 469 (97.7%) | 579 (88.8%) | 336 (93.1%) | 2699 (94.4%) |
| 11 (2.3%) | 73 (11.2%) | 25 (6.9%) | 159 (5.6%) |
| 2.10 (0.97) | 1.43 (1.18) | 2.20 (1.17) | 2.09 (1.12) |
| 219 (45.63%) | 299 (45.86%) | 238 (65.93%) | 1664 (58.2%) |
| 62 (12.92%) | 52 (7.98%) | 50 (13.85%) | 419 (14.7%) |
| 45 (9.38%) | 42 (6.44%) | 14 (3.88%) | 235 (8.2%) |
| 81 (16.88%) | 42 (6.44%) | 35 (9.70%) | 322 (11.3%) |
| 58 (12.08%) | 68 (10.43%) | 52 (14.40%) | 350 (12.2%) |
| 62 (12.92%) | 98 (15.03%) | 61 (16.90%) | 485 (17%) |
| 311 (64.79%) | 288 (44.17%) | 257 (71.19%) | 1895 (66.3%) |

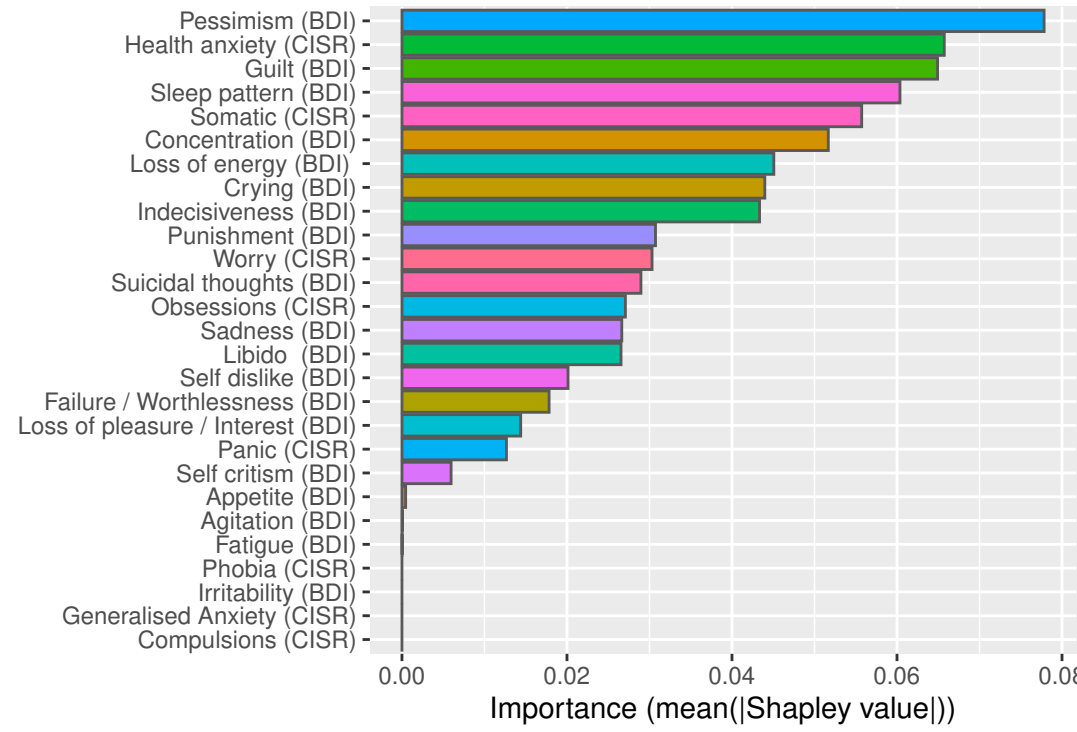
| | | PROMIS T-Score | | |
|-----------------------------|---------------|----------------|----------------|-------|
| | | RMSE | R ² | MAE |
| 3 to 4 months N=2646 | Items | 0.925 | 0.146 | 0.73 |
| | Sum scores | 0.926 | 0.143 | 0.73 |
| 6 to 8 months N=1297 | Items | 0.926 | 0.147 | 0.734 |
| | Sum scores | 0.924 | 0.146 | 0.735 |
| 9 to 12 months N=1110 | Items | 0.919 | 0.161 | 0.744 |
| | Sum scores | 0.935 | 0.126 | 0.753 |

Figure 1

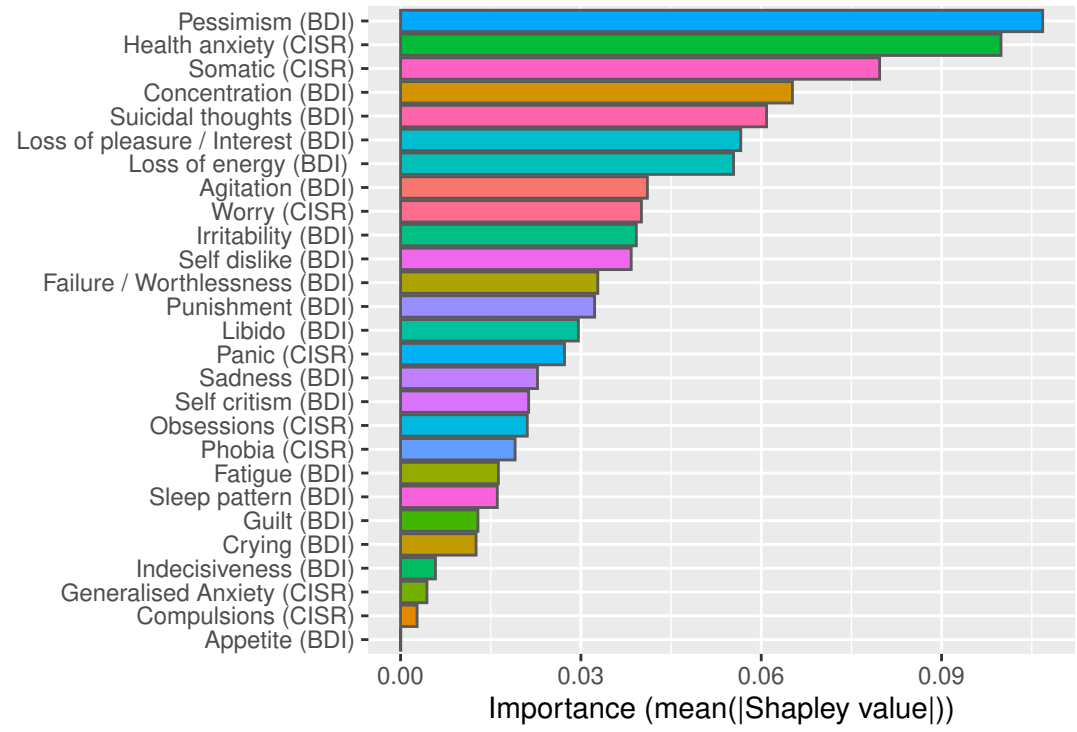
Outcome: 3 to 4 months (severity)

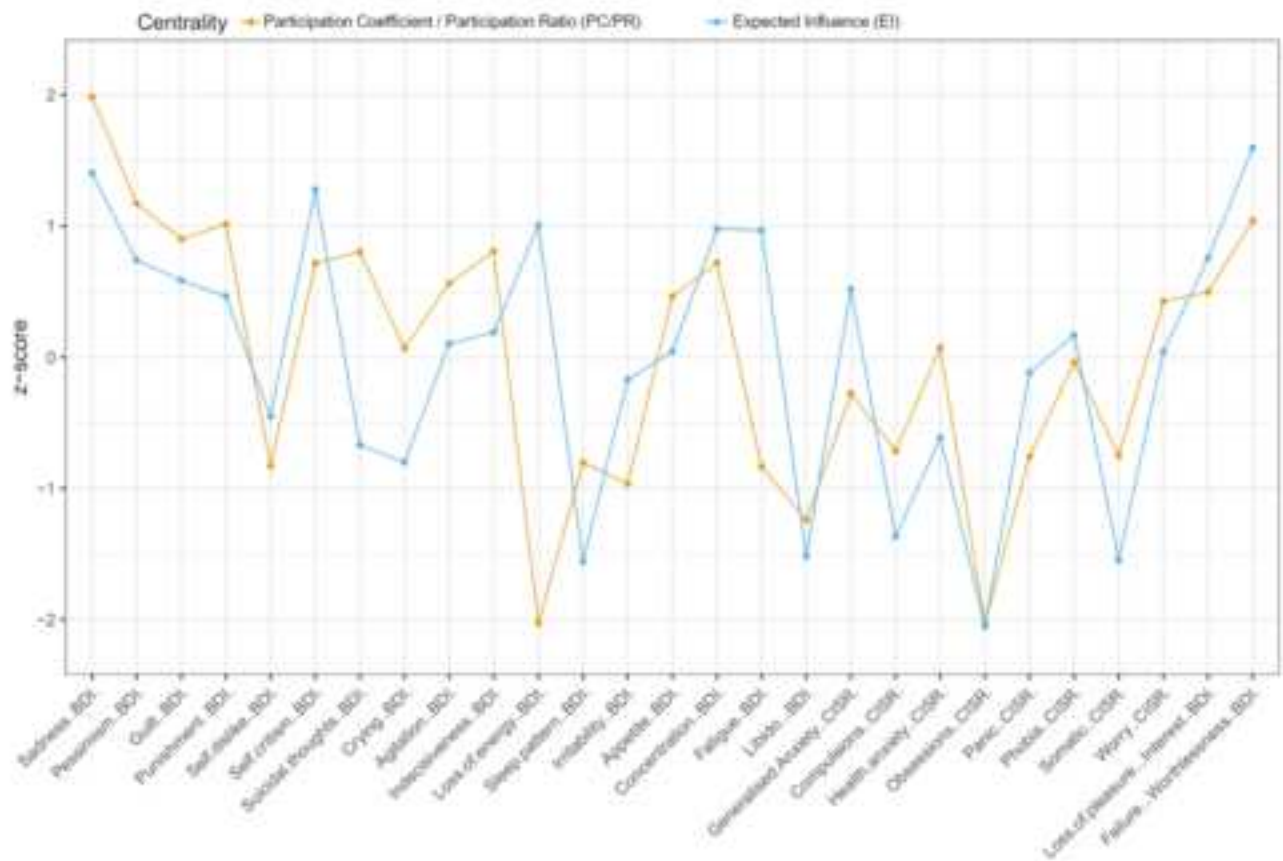
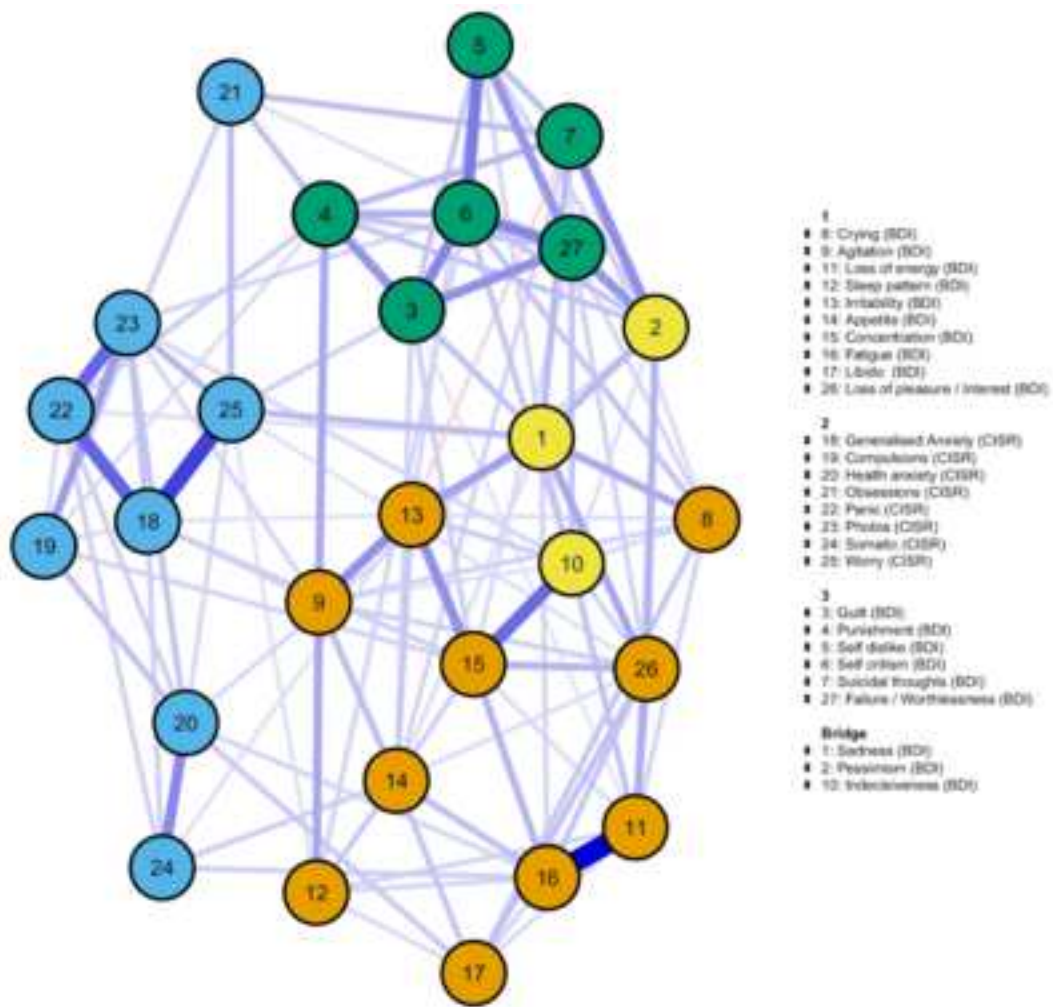


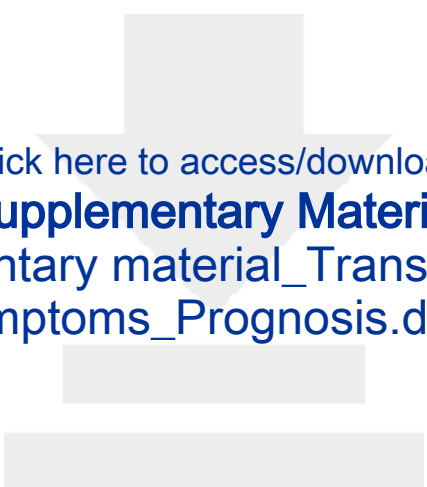
Outcome: 6 to 8 months (severity)



Outcome: 9 to 12 months (severity)







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