

This is a repository copy of *Effect of pragmatic versus explanatory interventions on medication adherence in people with cardiometabolic conditions: a systematic review and meta-analysis.*

White Rose Research Online URL for this paper: http://eprints.whiterose.ac.uk/163889/

Version: Published Version

Article:

Fitzpatrick, C., Gillies, C., Seidu, S. et al. (6 more authors) (2020) Effect of pragmatic versus explanatory interventions on medication adherence in people with cardiometabolic conditions: a systematic review and meta-analysis. BMJ Open, 10 (7). e036575. ISSN 2044-6055

https://doi.org/10.1136/bmjopen-2019-036575

Reuse

This article is distributed under the terms of the Creative Commons Attribution (CC BY) licence. This licence allows you to distribute, remix, tweak, and build upon the work, even commercially, as long as you credit the authors for the original work. More information and the full terms of the licence here: https://creativecommons.org/licenses/

Takedown

If you consider content in White Rose Research Online to be in breach of UK law, please notify us by emailing eprints@whiterose.ac.uk including the URL of the record and the reason for the withdrawal request.



BMJ Open Effect of pragmatic versus explanatory interventions on medication adherence in people with cardiometabolic conditions: a systematic review and meta-analysis

Claire Fitzpatrick,^{1,2} Clare Gillies,^{1,2} Samuel Seidu,^{1,2} Debasish Kar,³ Ekaterini Ioannidou,^{1,2} Melanie J Davies,^{1,2} Prashanth Patel,^{4,5} Pankaj Gupta,^{4,5} Kamlesh Khunti ⁽¹⁾ ^{1,2,6,7}

ABSTRACT

To cite: Fitzpatrick C, Gillies C, Seidu S, *et al.* Effect of pragmatic versus explanatory interventions on medication adherence in people with cardiometabolic conditions: a systematic review and meta-analysis. *BMJ Open* 2020;**10**:e036575. doi:10.1136/ bmjopen-2019-036575

Prepublication history and additional material for this paper are available online. To view these files, please visit the journal online (http://dx.doi. org/10.1136/bmjopen-2019-036575).

CF and CG contributed equally.

Received 23 December 2019 Revised 07 June 2020 Accepted 16 June 2020

Check for updates

© Author(s) (or their employer(s)) 2020. Re-use permitted under CC BY. Published by BMJ.

For numbered affiliations see end of article.

Correspondence to Professor Kamlesh Khunti; kk22@le.ac.uk

controlled trials (RCTs) of interventions aimed at increasing medication adherence in individuals with type 2 diabetes (T2DM) and/or cardiovascular disease (CVD). And, in a novel approach, to compare the intervention effect of studies which were categorised as being more pragmatic or more explanatory using the Pragmatic-Explanatory Continuum Indicator Summary-2 (PRECIS-2) tool, to identify whether study design affects outcomes. As explanatory trials are typically held under controlled conditions, findings from such trials may not be relatable to real-world clinical practice. In comparison, pragmatic trials are designed to replicate real-world conditions and therefore findings are more likely to represent those found if the intervention were to be implemented in routine care. Design Systematic review and meta-analysis. Data sources Ovid Medline, Ovid Embase, Web of Science and CINAHL from 1 January 2013 to 31 December 2018. Eligibility criteria for selecting studies RCTs lasting 23 months (90 days), involving \geq 200 patients in the analysis.

Objective To synthesise findings from randomised

with either established CVD and/or T2DM and which measured medication adherence. From 4403 citations, 103 proceeded to full text review. Studies published in any language other than English and conference abstracts were excluded.

Main outcome measure Change in medication adherence.

Results Of 4403 records identified, 34 studies were considered eligible, of which 28, including 30 861 participants, contained comparable outcome data for inclusion in the meta-analysis. Overall interventions were associated with an increase in medication adherence (OR 1.57 (95% CI: 1.33 to 1.84), p<0.001; standardised mean difference 0.24 (95% CI: -0.10 to 0.59) p=0.101). The effectiveness of interventions did not differ significantly between studies considered pragmatic versus explanatory (p=0.598), but did differ by intervention type, with studies that included a multifaceted rather than a single-faceted intervention having a more significant effect (p=0.010). The analysis used random effect models and used the revised Cochrane Risk of Bias Tool to assess study quality.

Strengths and limitations of this study

- In a novel approach, this systematic review compared whether study design (pragmatic vs explanatory as defined by the Pragmatic-Explanatory Continuum Indicator Summary-2 (PRECIS-2) tool) had an impact on intervention effect.
- The study selection was undertaken independently by two researchers to ensure that all relevant studies were included as well as preventing the risk of individual biases on study selection.
- The impact of study heterogeneity was explored using subgroup analyses.
- ► This review provides a contemporary update to the 2014 Cochrane Review on medication adherence.
- A potential limitation of this systematic review and meta-analysis is that a small number of studies had to be excluded as they did not categorise diabetes type.

Conclusions In this meta-analysis, interventions were associated with a significant increase in medication adherence. Overall multifaceted interventions which included an element of education alongside regular patient contact or follow-up showed the most promise. Effectiveness of interventions between pragmatic and explanatory trials was comparable, suggesting that findings can be transferred from idealised to real-word conditions.

PROSPERO registration number CRD42017059460.

INTRODUCTION

Prevalence of type 2 diabetes (T2DM) and cardiovascular diseases (CVDs) are increasing rapidly.¹ They have been identified as two of the most common cardiometabolic morbidities associated with multimorbidity¹ and are two of the leading causes of death worldwide.² Most people diagnosed with these conditions will likely have multimorbidity¹ (coexistence

of two or more chronic conditions), particularly those aged 65–84 years where prevalence is estimated at 65%.³

Management of multimorbidity is complex, typically relying on the coprescription of multiple drugs, which is strongly associated with medication non-adherence.⁴ The WHO estimates that 50% of individuals receiving chronic treatment are non-adherent (taking less than 80%) to prescribed medications.⁵ Medication non-adherence is considered the biggest cause of suboptimal clinical outcomes, accounting for approximately 57% of avoidable costs in relation to medication use.⁵

Despite the rise in cardiometabolic multimorbidity,^{6–8} interventions to increase medication adherence within this population are sparse. As such we decided to focus this review on T2DM and CVD as the most prevalent cluster cardiometabolic morbidities. Treatment nonadherence within these populations is well recognised and has become the focus of considerable research. While randomised controlled trials (RCTs) are widely accepted as a rigorous way of exploring the impact of interventions on specific health behavioural change outcomes, it has been identified that numerous components within a trial design can lead to biased interpretations of intervention effects.⁹ One such bias may be the controlled nature of these trials and the impact they impose on the cooperation of their participants and may prelude an action which does not necessarily represent what may occur in routine clinical practice.¹⁰ In pragmatic trials, the intervention is less strict and mimics usual practice as much as possible, thus lessening the unexpected reactions from the patients which lead to the biases.¹¹ As RCTs are generally expensive to conduct, it is important that their findings show real-world intervention effectiveness that is relevant to routine clinical practice. This study therefore aimed to not only identify interventions to increase medication adherence but also compare the effectiveness of interventions categorised as explanatory (undertaken in idealised settings) or pragmatic (undertaken in real-world settings). Described by Schwartz and Lellouch,¹² explanatory trial are those which confirm a physiological or clinical hypothesis; in contrast, pragmatic trails are those which inform clinical or policy decisions by evidencing the effect that adoption would have on routine care. As treatment effects of explanatory trials may be larger than those observed in pragmatic trials, traditional meta-analytic approaches may not account for this heterogeneity resulting in biased estimated treatment effects. While a handful of reviews in different research areas have now been published which retrospectively applied PRECIS-2 (Pragmatic-Explanatory Continuum Indicator Summary-2) to see if comparing pragmatic with explanatory trials altered review findings,¹³¹⁴ to the authors knowledge, this is the first review to incorporate PRECIS-2 scoring into the initial review process. While retrospective application provides an interesting comparison, whereby the overall intervention effects can be compared with explanatory and/or pragmatic intervention effects, pre- or post-use of PRECIS-2, it could lead

to an initial misinterpretation of findings. For example, a review containing a high number of explanatory trials may be much less applicable to routine clinical practice than one containing a greater number of pragmatic trials. Identification and comparison during the trial design stage would allow researchers to more easily identify how applicable findings would be to real- world clinical practice, rather than making an assumption based on a generalised outcome. In addition, this approach removes any risk of bias as the analysts are unaware of the results of the study. Using PRECIS-2 this review scored interventions from very explanatory to very pragmatic to explore whether the differences in characteristics between the study designs of these trials influenced intervention effectiveness estimates in a meta-analysis.

METHODS

This systematic review has been registered on PROS-PERO and was conducted in accordance with PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines.¹⁵

Data sources and searches

An extensive scoping search identified the great breath of research published within this topic area. It also identified a 2014 Cochrane Review on medication adherence which included studies published up to 11 January 2013. Due to factors relating to the focus of the review, changes in prescribed medication and the increasing use of mobile technology interventions in recent adherence research, we decided that this review would provide a contemporary update and therefore we refined our literature search to include studies published between 1 January 2013 to 31 December 2018. We searched Ovid Medline 1946 (Epub ahead of print, in process and other non-indexed citations), Ovid Embase 1974, Web of Science 1970 and Cumulative Index of Nursing and Allied Health Literature (EBSCO CINAHL-plus with full text). We searched "type 2 diabetes", "cardiovascular disease", "medication adherence" and "randomised control trial" using a combination of medical subject headings, keywords and synonyms with both English and American spellings. An example search strategy can be found in online supplementary etable 1.

Screening

Two investigators (CF and EI) independently screened all titles, abstracts and full text articles. Discrepancies were resolved through discussion with a third researcher (CG). Studies were assessed against five eligibility criteria to determine first, whether the study was an RCT; second, that patients were identified as having established CVD and/or T2DM; third, the trial measured medication adherence; fourth, study duration was \geq 3 months (90 days) and fifth, the study included >200 people in the analysis. Studies published in any language other than English and conference abstracts were excluded. Additional reasons for exclusion were where articles failed to specify diabetes type or where results were combined for individuals at risk of and with established CVD.

Data extraction

Data extraction was performed by CF and checked by CG. Where available the following data were recorded: study characteristics (authors, year, country, duration, sample size), participant characteristics (age, gender, disease), methods of assessment (self-report, pharmacy records, pill count), intervention type (online supplementary etable 2) and outcome measures of medication adherence. For trials reporting multiple follow-up, the final follow-up corresponding to study duration was used.

Risk of bias

Risk of bias assessment was undertaken using the revised Cochrane Risk of Bias Tool for RCTs.¹⁶¹⁷

Pragmatic-Explanatory Continuum Indicator Summary-2

Using Loudon *et al* (2015) for guidance,¹⁸ PRECIS-2 scoring to explore how pragmatic/explanatory different components of the included study designs was undertaken independently by CF and CG. Each of the nine domains was given a score from 1 (very explanatory 'ideal study conditions') to 5 (very pragmatic 'usual care conditions'). Results were compared and a consensus score was reached.

Some studies did not clearly explain all components making scoring difficult. To determine whether the treatment of unclear data affected the overall score, we conducted two classification processes. For the primary method we inputted a score of 3 and then calculated an average score by adding up the scores and dividing by 9. In the second, a score was not given for the missing domains and then an average was calculated based on available domain scores.

Data synthesis

Study data were reported as means and medians for continuous data and as proportions for categorical data. Twenty-two studies defined a cut point to determine adherence. Twenty used a binary outcome defining individuals as adherent or non-adherent; however, two categorised individuals based on a prespecified level of low, medium and high. Of those reporting levels of low, medium or high adherence, as has been done in previous studies,^{19 20} we combined medium and high levels and separated low level to form a binary outcome of adherence or non-adherence, respectively, therefore enabling OR of adherence to be calculated. In one study (Boyne et al, 2014), all participants in the intervention arm were adherent at follow-up, so a continuity correction of 0.5 was added to allow an OR to be calculated. Where an OR for medication adherence was reported in study results, this was used for the meta-analyses, rather than calculating an estimated OR from the raw numbers. Where adjusted ORs were reported, the OR adjusted for the most covariates was used in the meta-analyses.

A further six studies, plus one which was included in the previous analysis, reported adherence using a continuous scale which provided an overall group indication of adherence. For these studies, change in adherence for each study arm was calculated. As scores differed, we calculated standardised mean differences (SMDs) to combine study estimates in a meta-analysis. The SMD was calculated as the mean change in medication adherence in the intervention group minus mean change in the control group, divided by the pooled SD, using Cohen's method.²¹

For both meta-analysis models, funnel plots and Egger's test were carried out to assess for publication bias, and random effects meta-analyses were fitted to account for heterogeneity in study design. We assessed heterogeneity between studies by calculating the Higgins I² statistic, with an I² statistic >75% considered high heterogeneity.²² In one study (Boyne *et al*, 2014), all participants in the intervention arm were adherent at follow-up, so a continuity correction of 0.5 was added to allow an OR to be calculated.

We fitted meta-regression models assessing; study length, whether adherence was a primary or secondary outcome, mean age, percent male, disease (T2DM or CVD), and PRECIS-2 score, to explore the impact of study heterogeneity on the intervention effect. Three subgroup analyses were also carried out, whereby fitting a meta-regression model to compare the statistical difference between groups, separate meta-analyses were also run for each subgroup to enable the pooled estimate for each subgroup to be calculated, and hence a more explicit comparison to be made. The first compared intervention effects of studies that included a self-reported measure of adherence, to those with an objective adherence measure. The second compared pragmatic with explanatory studies (pragmatic studies were those with an average PRECIS-2 score >3 and explanatory a score of \leq 3) and the third compared outcomes of interventions which were identified as multifaceted against those with a singular intervention component.

Where studies had obtained both pharmacy and selfreport data, pharmacy data were used in preference in the meta-analysis. In the case of a few studies which did not report overall level of adherence and instead listed individual drug adherence, we made the pragmatic decision to monitor the adherence effect of the first drug listed by the author.

All analyses were performed using Stata 15 with the METAN, METAREG and METABIAS commands. Results are reported with 95% CIs with a p value <0.05 considered statistically significant. For the few studies not included within the meta-analysis, descriptive synthesis of intervention effects is discussed.

RESULTS Study selection

Following deduplication we screened 4403 titles and abstracts which yielded 103 potentially relevant studies. Thirty-three met the inclusion criteria with an additional study²³ identified from reference list and forward citation

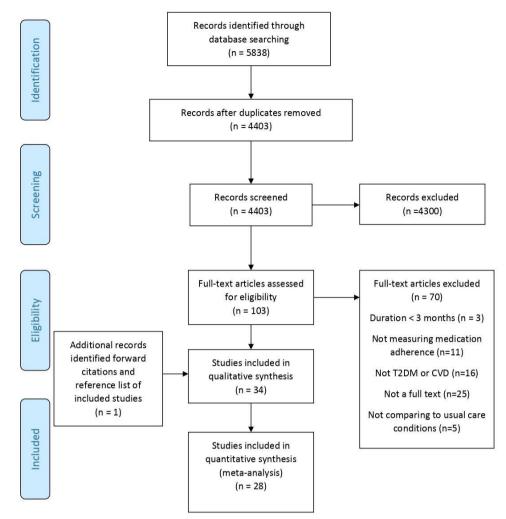


Figure 1 PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) flow chart of included studies. CVD, cardiovascular disease; T2DM, type 2 diabetes mellitus.

screening (figure 1).²³ Of the 34 studies, 28 including 30 861 participants contained comparable outcome data for inclusion in the meta-analysis.

Characteristics of included studies

An overview of the characteristics of the included studies is available in table 1. A more detailed key characteristics of included studies is available in online supplementary etable 3. Studies were conducted in 17 countries, with one study²⁴ conducted across 4 countries. Twenty-three trials were conducted in high-income countries,^{23 25-46} seven in upper middle,^{47–53} two in lower middle,^{54 55} one in a low-income country and one across both upper middle and high income economies.⁵⁶ The proportion of males ranged from 27.9%³¹ to 94.0%.²³ Four trials did not report age^{24 26 52 56} and/or gender.^{24 26 56}

Medical conditions varied across studies; 12 studies were in participants with T2DM, 25 $^{28-31}$ 34 37 41 48 53 54 56 21 in participants with CVD 23 24 26 27 32 33 35 36 $^{38-40}$ $^{42-44}$ 46 47 $^{49-52}$ 55 and 1 in people with both. 45 Follow-up ranged from 3 to 36 months.

Medication adherence was assessed as a primary outcome in 19 studies. $^{24\,25\,27-29\,33\,35\,36\,39\,42\,44\,45\,48\,50-53\,55\,56}$ Methods of

measuring medication adherence varied between studies. Adherence was primarily assessed through self-report (n=23),²³⁻³² ³⁴ ³⁸ ⁴⁰ ⁴² ⁴⁴ ⁴⁷⁻⁵⁰ ⁵³⁻⁵⁶ followed by pharmacy data (n=10),^{28 35 36 41–43 45 46 50 53} pill count $(n=2)^{24 33}$ and electronic pill bottle opening (n=1).³⁹ Three studies did not report the method of medication adherence assessment used.^{37 51 52} Of those measuring adherence through self-report, the Morisky-Green Questionnaire⁵⁷ and the 8-item Morisky Medication Adherence Scale (MMAS-8)⁵⁸ (or translated versions) were the most frequently used. Four studies measured adherence using two different methods: self-report and either pharmacy or pill count data.^{24 28 50 53} Two of the studies combined the data to provide an overall level or adherence,^{24 28} whereas the other two reported adherence outcomes separately.^{50 53} Of the two studies reporting separate outcomes, both reported consistent findings and statistically significant improvements in adherence whether self-report or pharmacy data were used.

Interventions varied greatly between studies. We categorised and defined interventions (online supplementary etable 2) based on their most prominent feature into one

Table 1 Overview of characteristics of included studies									
First author, year	Sample size (analysed) Duration (data collection points)	Condition	Key intervention detail	Medication adherence measure					
Al-Haj Mohd, 2016 ²⁵	446 (446) 6 months (0 and 6 months)	T2DM	Multifaceted intervention strategy	Self-report (MMAS-8)					
Barker-Collo, 2015 ²⁶	386 (326) 12 months (0, 3, 6, 9 and 12 months)	Stroke/TIA	Behavioural/educational	Self-report with crosscheck of pharmacy data					
Boyne, 2014 ²⁷	382 (382) 12 months (0, 3, 6 and 12 months)	Heart failure	Telemonitoring/telemedicine	Self-report (European Heart Failure Self-care Behaviour Scale)					
Buhse, 2017 ²⁸	279 (279) 6 months (0 and 6 months)	T2DM	Behavioural/educational	Pharmacy data Self-report (interview)					
Caetano, 2018 ²⁹	709 (702) 6 months (0 and 6 months)	T2DM	Behavioural/educational	MAT scale					
Cao, 2017 ⁴⁷	236 (236) 90 days (0, 30 and 90 days)	Coronary heart disease	Collaborative care	Self-report (MMAS-8)					
Carrasquillo, 2017 ³⁰	300 (215) 12 months (0 and 12 months)	T2DM	Behavioural/educational	Self-report (MMAS-8)					
Castellano, 2014 ²⁴	695 (594) 9 months (0, 1, 4 and 9 months)	Myocardial infarction	Simplification of drug regimen	Pill count Self-report (Morisky-Green- Levine Adherence Scale)					
Chung, 2014 ⁴⁸	241 (241) 12 months (0, 4, 8 and 12 months)	T2DM	Multifaceted intervention strategy	Self-report—MMAS-8 (Revised Malaysian version)					
Crowley, 2013 ³¹	359 (329) 12 months (0, 3, 6, 9 and 12 months)	T2DM	Multifaceted intervention strategy	Self-report (Morisky-Green- Levine Adherence Scale)					
Du, 2016 ³²	979 (964) 36 months (0 and 36 months)	Percutaneous coronary intervention	Multifaceted intervention strategy	Self-report (Morisky-Green- Levine Adherence Scale)					
El-Touky, 2017 ³³	321 (276) 12 months (0, 1, 3, 6, 9 and 12 months)	Acute coronary syndrome	Behavioural/educational intervention	Pill count					
Graumlich, 2016 ³⁴	674 (674) 12 months (0, immediately postintervention, 3 and 6 months)	T2DM	Medication monitoring table	Self-report (PMAQ)					
Hedegaard, 2015 ³⁵	211 (203) 12 months (0, 3, 6, 9 and 12 months)	Stroke/TIA	Multifaceted intervention strategy	Pharmacy data-MPR					
Ho, 2014 ³⁶	253 (241) 12 months (0 and 12 months)	Acute coronary syndrome	Multifaceted intervention strategy	Pharmacy refill data					
Jeong, 2018 ³⁷	338 (338) 24 weeks (0 and 24 weeks)	T2DM	Telemonitoring/telemedicine	Does not report					
Jia, 2017 ⁴⁹	669 (669) 36 months (0, 1, 3, 6, 12 and 36 months)	Percutaneous coronary intervention	Intensified patient care	Self-report (Morisky-Green- Levine Adherence Scale)					
Kronish, 2014 ³⁸	600 (600) 6 months (0 and 6 months)	Stroke/TIA	Behavioural/educational intervention	Self-report (MMAS-8)					
Lin, 2017 ⁵⁰	288 (288) 18 months (0, 6, 12 and 18 months)	Coronary artery bypass grafting	Multifaceted intervention strategy	Self-report—MARS (5 item)					
Marin, 2015 ⁵⁶	467 (459) 12 months (0 and 12 months)	T2DM	Personalised medication management	Self-report (Morisky-Green- Levine Adherence Scale)					
Marquez-Contreras, 2018 ³⁹	726 (625) 18 months (0, 6 and 12 months)	Atrial fibrillation	Multifaceted intervention strategy	MEMs					

Continued

Table 1 Continue						
First author, year	Sample size (analysed) Duration (data collection points)	Condition	Key intervention detail	Medication adherence measure		
Meng, 2014 ²³	471 (425) 12 months (admission, discharge, 6 and 12 months)	Coronary heart disease	Behavioural/educational intervention	Self-report—MARS-D		
Meng, 2016 ⁴⁰	513 (449) 12 months (admission, discharge, 6 and 12 months)	Heart failure	Behavioural/educational intervention	Self-report—MARS-D		
Peng, 2014 ⁵¹	3821 (3330) 12 months (discharge, 6, 9 and 12 months)	Stroke/TIA	Behavioural/educational intervention	Does not report		
Pladevall, 2015 ⁴¹	1692 (1512) 18 months (0, 6, 12 and 18 months)	T2DM	Behavioural/educational intervention	Pharmacy data-PDC		
Rinfret, 2013 ⁴²	300 (300) 12 months (0 and 12 months)	Percutaneous coronary intervention	Intensified patient care	Pharmacy refill data		
Samtia, 2013 ⁵⁴	348 (348) 5 months (0 and 5 months)	T2DM	Behavioural/educational	Self-report		
Schou, 2014 ⁴³	921 (920) 13–72 months (every 1–3 months)	Heart failure	Multifaceted intervention strategy	Pharmacy data—PDC		
Schwalm, 2015 ⁴⁴	852 (852) 12 months (0, 3 and 12 months)	Myocardial infarction	Behavioural/educational	Self-report (Morisky-Green- Levine Adherence Scale)		
Su, 2017 ⁵²	1275 (1187) 12 months (0 and 12 months)	Stroke/TIA	Multifaceted intervention strategy	Does not report		
Vollmer, 2014 ⁴⁵	21 752 (21 752) 12 months (0 and 12 months)	T2DM and/or CVD	Telemonitoring/telemedicine	Pharmacy data-PDC		
Volpp, 2017 ⁴⁶	1509 (1503) 12 months (0 and 12 months)	Myocardial infarction	Multifaceted intervention strategy	Pharmacy data-PDC		
Xavier, 2016 ⁵⁵	805 (750) 12 months (0 and 12 months)	Acute coronary syndrome	Multifaceted intervention strategy	Pharmacy data—composite medical adherence score >80%		
Xin, 2015 ⁵³	240 (227) 12 months (0 and 12 months)	T2DM	Behavioural/educational	Prescription refill claims Self-report (Morisky-Green- Levine Adherence Scale)		

CVD, cardiovascular disease; MARS-D, Medication Adherence Report Scale (German version); MAT, Measure of Adherence to Treatments; MEMs, Medical Event Monitoring Systems; MMAS-8, Morisky Medication Adherence Scale (8-item); 0 months, baseline; MPR, medication possession ratio; PDC, proportion of days covered; PMAQ, Patient Medication Adherence Questionnaire; T2DM, type 2 diabetes mellitus; TIA, transient ischaemic attack.

of the following seven groups: (1) behavioural/educational, $^{23 26 28-30 33 34 38 40 41 44 51 53 54}$ (2) intensified patient care, $^{42 49}$ (3) collaborative care, 47 (4) simplification of drug regimen, 24 (5) personalised drug dispensing, 56 (6) telemonitoring/telemedicine $^{27 37 45}$ and (7) multifaceted intervention strategy. $^{25 31 32 35 36 39 43 46 48 50 52 55}$

PRECIS-2 scoring

Details of PRECIS-2 scoring for each of the included studies can be found in table 2. For the purpose of this study it was considered that each of the nine domains on the PRECIS-2 wheel had equal weighting. For visual clarity, PRECIS-2 results are presented as a shaded graph (table 2) where darker shades represent more pragmatic

components. We also inputted scores on the PRECIS-2 wheel (online supplementary efigures 1–3).

Of the 34 studies, 20 studies were identified as being more pragmatic, with an average score >3 (range 3.11– 4.11).²³ ²⁴ ²⁶ ²⁷ ²⁹ ³² ^{35–37} ^{40–49} ⁵² Of these, most (n=18) received an average score between 3 and 4, demonstrating a slightly more pragmatic intention on the PRECIS-2 continuum. Two of the studies scored ≥ 4 ,^{29 45} demonstrating a greater degree of pragmatism. The majority of these studies scored poorly (score of 1) in relation to pragmatism of the primary outcome as, based on the guidance by Loudon *et al*,¹⁸ medication adherence was not considered of obvious importance from the patients' perspective and was typically assessed by methods not

Table 2 Shaded graph to show PRECIS-2 scoring of included studies														
First author	Year	Eligibility	Recruitment	Setting	Organisation	Flexibility: delivery	Flexibility: adherence	Follow-up	o Primary outcome	Primary analysis	Total score	Average score	Amended total score	Amended average score
Al-Haj Mohd ²⁵	2016	4	4	5	1	2	3	2	1	5	27	3	24	3
Barker-Collo ²⁶	2015	5	5	5	1	1	3	1	5	5	31	3.44	28	3.5
Boyne ²⁷	2014	5	5	5	2	1	3	1	2	5	29	3.22	26	3.25
Buhse ²⁸	2017	1	4	5	1	2	3	1	1	5	23	2.56	20	2.5
Caetano 29	2018	4	5	5	4	5	5	3	1	5	37	4.11	34	4.25
Cao ⁴⁷	2017	2	4	5	1	2	4	2	5	5	30	3.33	30	3.33
Carrasquillo 30	2017	1	1	4	1	2	1	1	2	4	17	1.89	17	1.89
Castellano ²⁴	2014	4	5	5	5	5	2	2	1	5	34	3.78	34	3.78
Chung ⁴⁸	2014	2	5	5	4	1	5	2	1	5	30	3.33	30	3.33
Crowley ³¹	2013	1	1	5	2	2	3	2	2	4	22	2.44	19	2.38
Du ³²	2016	5	4	5	2	2	3	2	5	5	33	3.67	30	3.75
El-Touky ³³	2017	4	5	5	4	3	1	2	1	1	26	2.89	23	2.88
Graumlich 34	2016	1	2	5	1	1	4	2	2	5	23	2.56	23	2.56
Hedegaard 35	2015	4	4	5	2	2	3	4	1	5	30	3.33	27	3.38
Ho ³⁶	2014	4	4	5	2	1	3	5	1	5	30	3.33	27	3.38
Jeong ³⁷	2018	4	5	5	1	4	3	2	2	4	30	3.33	27	3.38
Jia ⁴⁹	2017	3	4	5	2	4	3	4	1	4	30	3.33	27	3.38
Kronish ³⁸	2014	4	1	4	1	2	3	1	2	5	23	2.56	20	2.5
Lin ⁵⁰	2017	1	2	4	1	2	3	1	1	5	20	2.22	17	2.13
Marin 56	2015	1	2	5	4	4	3	2	1	4	26	2.89	23	2.88
Marquez-Contreras 39	2018	2	1	5	2	2	1	3	1	2	19	2.11	19	2.11
Meng ²³	2014	4	5	5	2	4	3	2	2	4	31	3.44	28	3.5
Meng ⁴⁰	2016	2	5	4	2	4	3	2	4	4	30	3.33	27	3.38
Peng ⁵¹	2014	3	5	5	2	1	3	3	1	4	27	3	21	3
Pladevall ⁴¹	2015	2	2	5	1	5	3	4	2	5	29	3.22	26	3.25
Rinfret ⁴²	2013	4	4	5	2	4	3	4	1	5	32	3.56	29	3.63
Samtia 54	2013	2	4	4	4	3	3	1	2	4	27	3	21	3
Schou ⁴³	2014	1	5	5	4	4	4	4	1	4	32	3.56	32	3.56
Schwalm 44	2015	5	5	5	4	4	3	2	1	4	33	3.67	30	3.75
Su ⁵²	2017	4	5	5	2	2	3	2	1	4	28	3.11	25	3.13
Vollmer ⁴⁵	2014	4	5	5	4	4	3	5	1	5	36	4	33	4.13
Volpp ⁴⁶	2017	2	1	5	1	4	2	5	5	5	30	3.33	30	3.33
Xavier 55	2016	4	5	5	1	2	2	1	1	5	26	2.89	26	2.89
Xin ⁵³	2015	2	3	3	2	4	3	2	1	4	24	2.67	15	2.5

PRECIS-2, Pragmatic-Explanatory Continuum Indicator Summary-2.

7

Study	OR(95% CI)	% Weight
Pragmatic (PRECIS-2 score<3)		
Al Haj (2016)	2.00 (1.37, 2.91)	5.56
Barker Collo (2015)	1.06 (0.45, 2.49)	2.49
Boyne (2014)	→ 5.17 (0.24, 109.51)	0.27
Castellano (2014)	1.50 (1.11, 2.02)	6.25
Chung (2014)	2.11 (1.22, 3.66)	4.17
Du (2016)	2.04 (1.58, 2.64)	6.63
Hedegaard (2015)	0.66 (0.35, 1.26)	3.54
Ho (2014)	2.95 (1.46, 5.98)	3.19
lia (2017)	1.93 (1.42, 2.63)	6.19
Pladeval (2015)	1.05 (0.82, 1.35)	6.69
Schwalm (2015)	1.03 (0.77, 1.36)	6.39
Su (2017)	2.14 (1.75, 2.46)	7.30
/ollmer (2014)	1.16 (1.06, 1.26)	7.76
\diamond	1.53 (1.23, 1.89)	66.44
Explanatory (PRECIS-2 score>=3)		
Buhse (2017)	1.20 (0.50, 2.60)	2.62
Crowley (2013)	4.40 (1.80, 10.60)	2.37
l Touky (2017)	1.89 (0.63, 5.67)	1.72
Graumlich (2016)	1.13 (0.99, 1.29)	7.54
(ronish (2014)	1.12 (0.80, 1.56)	5.95
Peng (2014)	3.47 (0.83, 14.49)	1.12
Samtia (2013)	1.27 (0.83, 1.96)	5.09
(avier (2016)	2.69 (1.36, 5.34)	3.31
(in (2015)	2.98 (1.64, 5.41)	3.84
\diamond	1.69 (1.24, 2.31)	33.56
Dverall (I2 = 80.5%, p < 0.001)	1.57 (1.33, 1.84)	100.00
	1 20	
	s intervention	

Figure 2 Forest plot of pooled ORs for medication adherence, stratified by Pragmatic-ExplanatoryContinuum Indicator Summary-2 (PRECIS-2) score.

used routinely in primary care. Eleven studies were identified as being more explanatory receiving an average score of <3 (range 1.89–2.89).^{28 30 31 33 34 38 39 50 53 55 56} Of these, most (n=10) received an average score between 2 and 3 demonstrating that they favour a more explanatory intention on the continuum. Finally, three studies were given a score of $3^{25 51 54}$ suggesting that these were equally pragmatic and explanatory.

The domains within which trials scored a more pragmatic rating were recruitment, setting of the intervention and the primary analysis which received average scores of 3.7, 4.8 and 4.4, respectively. All other domains appeared more explanatory with average scores of 2.2, 2.8, 2.9, 2.4 and 1.8 for the domains of organisation, flexible delivery, flexible adherence, follow-up and primary outcome, respectively. The eligibility domain was the only one to receive a score of 3.

When results were compared based on the way missing data were scored, no changes to any studies overall categorisation was observed.

META-ANALYSIS

Effect of interventions

Data from 22 studies were pooled in a meta-analysis (figure 2). Irrespective of intervention type or duration,

interventions significantly improved medication adherence when compared with control (OR 1.57 (95% CI: 1.33 to 1.84), p<0.001). For the seven studies where adherence was reported as a continuous variable, the pooled SMD was estimated as 0.24 ((95% CI: -0.10 to 0.59), p=0.101) (online supplementary efigure 4), in favour of the intervention group, but was not statistically significant.

Both meta-analyses showed statistically significant heterogeneity between studies (p<0.001) with high I^2 values of 80.5% and 95.4% for the meta-analyses of the ORs and SMDs, respectively, where the I^2 value represents the percentage of variability due to heterogeneity rather than sampling error.

We investigated the heterogeneity by conducting metaregression analyses (online supplementary etable 4). No statistically significant association was found between the intervention effects and study length, mean study age, male per cent, disease (CVD or T2DM), whether it was a primary or secondary outcome or PRECIS-2 score. Subgroup analyses identified no differences in the intervention effect between self-reported and objective measures of adherence, and between PRECIS-2 score (online supplementary etable 5). Subgroup analyses comparing multifaceted versus singular-faceted interventions showed that multifaceted interventions led to

Study	OR (95% CI)	% Weight
Multifaceted Intervention		
Al Haj (2016)	2.00 (1.37, 2.91)	5.56
Chung (2014)	- 2.11 (1.22, 3.66)	4.17
Crowley (2013)	4.40 (1.80, 10.60)	2.37
Du (2016)	2.04 (1.58, 2.64)	6.63
Hedegaard (2015)	0.66 (0.35, 1.26)	3.54
Ho (2014)	2.95 (1.46, 5.98)	3.19
Su (2017)	2.14 (1.75, 2.46)	7.30
Xavier (2016)	2.69 (1.36, 5.34)	3.31
\diamond	2.05 (1.62, 2.59)	36.07
Single-faceted Intervention		
Barker Collo (2015)	1.06 (0.45, 2.49)	2.49
Boyne (2014)	◆ 5.17 (0.24, 109.51)	0.27
Buhse (2017)	1.20 (0.50, 2.60)	2.62
Castellano (2014)	1.50 (1.11, 2.02)	6.25
El Touky (2017)	1.89 (0.63, 5.67)	1.72
Graumlich (2016)	1.13 (0.99, 1.29)	7.54
Jia (2017)	1.93 (1.42, 2.63)	6.19
Kronish (2014)	1.12 (0.80, 1.56)	5.95
Peng (2014)	➡ 3.47 (0.83, 14.49)	1.12
Pladeval (2015)	1.05 (0.82, 1.35)	6.69
Samtia (2013)	1.27 (0.83, 1.96)	5.09
Schwalm (2015)	1.03 (0.77, 1.36)	6.39
Vollmer (2014)	1.16 (1.06, 1.26)	7.76
Xin (2015)	2.98 (1.64, 5.41)	3.84
♦	1.28 (1.13, 1.46)	63.93
Overall (I2 = 80.5%, p < 0.001)	1.57 (1.33, 1.84)	100.00
0.5 1 2 Favours control Fav	5 20 ours intervention	

Figure 3 Forest plot of ORs for adherence, stratified by complexity of the intervention.

a statistically significant improvement in odds of adherence (p=0.010) (figure 3; online supplementary figure S3, etable 5). The funnel plot and Egger's test (p=0.184) showed that no indication of publication bias was present for the meta-analysis of SMDs (online supplementary efigure 5). For the meta-analysis of ORs, Egger's test for publication bias was statistically significant (p=0.041), with the funnel plot asymmetry indicating a slightly uneven presence of small studies showing a favourable intervention effect (online supplementary efigure 6).

Descriptive synthesis

Six studies did not contain comparable data for inclusion in the meta-analysis.^{37 39 42 43 46 56} Of those, four studies showed a positive intervention effect on rates of medication adherence,^{37 39 42 56} while two showed no statistically significant intervention effects.^{43 46}

Assessment of risk of bias

Table 3 displays trial specific risk of bias assessment. Few studies were deemed to be of fair quality with low risk of bias and none were considered to be free of bias. Twenty-seven trials provided information about adequate sequence generation and 17 regarding allocation concealment. Incomplete outcome data was a concern in four trials and selective reporting was primarily deemed unclear (n=19) due to the number of trials without a published protocol. Blinding of participants and personnel was the main domain scoring poorly for risk of bias; however, the nature of these types of trials means it is often impractical to do so.

DISCUSSION

Medication adherence is a key component of chronic disease care yet many patients fail to follow prescribing guidelines.⁵ This review identified 34 trials that evaluated the impact of interventions aimed at improving medication adherence in individuals with either T2DM and/ or established CVD. Overall interventions significantly increased medication adherence.

To our knowledge this is the first review to systematically compare and synthesise medication adherence outcomes based on the PRECIS-2 classification of trials. Our review showed that interventions improve medication adherence and the PRECIS-2 classification did not appear to affect the outcome when compared in subgroup and meta-regression analyses. This finding differs from previous research where explanatory trials have reported significantly larger effect sizes than pragmatic trials.^{13 14} This suggests that findings from these interventions are representation of real-world clinical practice.

Table 3 Revised Cochrane risk of bias of included studies

First author, year	Adequate random sequence generation	Allocation concealment	Blinding of participants	Blinding of personnel	Blinding of outcome assessors	Incomplete outcome data	Selective reporting	Other
Al-Haj Mohd ²⁵	+	+	-	-	?	+	?	?
Barker-Collo et al ²⁶	+	+	_	_	+	_	+	+
Boyne ²⁷	+	?	?	?	?	+	?	?
Buhse ²⁸	+	+	_	-	+	+	-	+
Caetano ²⁹	+	?	-	-	-	+	?	?
Cao ⁴⁷	+	+	-	-	+	+	?	?
Carrasquillo ³⁰	+	+	-	-	+	+	+	?
Castellano ²⁴	+	+	-	-	-	+	+	+
Chung ⁴⁸	?	?	-	-	-	+	?	?
Crowley ³¹	+	+	-	-	-	+	+	?
Du ³²	+	+	-	-	?	+	?	?
El-Touky ³³	?	?	?	?	?	-	?	?
Graumlich ³⁴	+	+	-	-	-	+	+	+
Hedegaard ³⁵	+	+	-	-	+	+	?	?
Ho ³⁶	+	+	?	?	+	+	+	+
Jeong ³⁷	?	?	-	-	-	+	-	?
Jia ⁴⁹	?	+	?	?	?	-	?	?
Kronish ³⁸	+	+	_	-	+	+	?	?
Lin ⁵⁰	+	?	?	?	+	+	?	?
Marin ⁵⁶	+	?	?	?	?	+	-	-
Marquez-Contreras	+	+	?	?	?	-	?	?
Meng ²³	-	-	-	-	-	+	?	?
Meng ⁴⁰	+	+	+	+	?	+	+	?
Peng ⁵¹	+	?	?	?	?	?	+	?
Pladevall ⁴¹	+	?	-	-	+	+	?	-
Rinfret ⁴²	+	+	-	-	+	+	?	?
Samtia 54	?	?	?	?	?	+	?	-
Schou ⁴³	+	+	-	-	-	+	+	-
Schwalm ⁴⁴	+	-	?	?	+	+	+	-
Su et al ⁵²	+	?	?	?	?	-	?	?
Vollmer ⁴⁵	+	?	-	-	-	+	?	?
Volpp ⁴⁶	+	?	-	-	+	+	-	?
Xavier ⁵⁵	+	?	-	-	-	+	+	?
Xin ⁵³	?	?	-	-	+	+	?	?

Key: + low risk of bias, -high risk of bias, ? unclear risk of bias.

Using the PRECIS-2 tool retrospectively posed a number of challenges. Missing data in published manuscripts is a known problem when conducting systematic reviews, and posed additional problems for this review as manuscripts often did not include the level of detail required to accurately score certain domains (online supplementary etable 6). As has been identified in previous research,⁵⁹ the most challenging domains to score were those relating to the flexibility of delivery and flexibility of participant adherence. Most studies focused on detail the content of the intervention not its delivery and adherence. To define a study as having a more pragmatic or

10

explanatory tendencies, we averaged domain scores to >3 or \leq 3, respectively. While considered the best method, we were aware of problems which could arise as a result of combining data in this way. Combining scores could result in two or more studies receiving the same or similar scores even though individual domain scores were very different. It also meant that each domain was given equal weighting in the overall score of pragmatism. While a known limitation to the authors, no published guidance on how best to report the data could be found.

We restricted our review to articles published since 2013 to provide a contemporary update following a large Cochrane review published 2014 by Nieuwlaat et al.60 Similarly to the 2014 review, our findings suggest that interventions were diverse in nature and the majority were complex, involving multiple different components. Of the number of 'successful' interventions, multifaceted interventions which included an element of education alongside regular patient contact showed the most promise, suggesting that frequent engagement with the healthcare team may trigger behavioural change or act as a reminder to undertake the behaviour. While promising, considerations need to be made as to the ability of such interventions to be upscaled and implemented. Multifaceted interventions can be expensive and therefore their cost utility needs to be explored prior to such interventions becoming embedded in clinical care pathways.

As the number of people with access to mobile technologies has increased in the past decade, particularly in the over 60 age group, interventions which rely on frequent patient contact are becoming increasingly plausible and contact via either calls or SMS provide both a pragmatic and cheap alternative to face-to-face healthcare professional contact.⁶¹ A systematic review by Changizi and Kaveh looking into the effectiveness of mobile health in the elderly showed that it can be used effectively and is widely accepted as a source of health literacy, particularly SMS messages due to their low requirement for technological competency.⁶²

Our review highlighted a number of limitations to medication adherence research conducted to date. First, the lack of gold standard method of measuring adherence makes it difficult to pool and compare outcomes. The most frequently used method within this review was self-report, particularly the tools developed by Morisky and colleagues.^{57 58} These tools are quick, easy and costeffective to administer making them ideal for use within a clinical care setting. The tools are validated and have shown moderate comparability with other indirect methods of medication adherence; however, their use has been associated with overestimation of true treatment adherence as they carry the potential for recall bias.⁶³

Even among the studies using the Morisky scales methods of scoring and presenting data varied. For example, in the case of the Morisky-Green Scale, some used a Likert scoring system,³¹ some classed answers of no to all questions as indicative of good adherence,⁴⁴ whereas others classed good adherence as an answer of

yes to all.⁵³ For the MMAS-8, studies such as Cao *et al*⁴⁷ reported mean patient scores, with higher scores representing greater adherence. In contrast, Al-Haj Mohd *et al*²⁵ report the proportion of people categorised as low medium or highly adherent based on a score of <6, 6–7.9 or >8, respectively. Discrepancies in reporting make it difficult to compare outcomes and therefore there is a need for standardisation. The best method to measure medication adherence across disease populations, and the best approach to reporting said results, therefore continues to be an area requiring further exploration.

We also acknowledge that our findings may be limited by our strict inclusion criteria. All papers reviewed were written in English and published since 2013 and therefore results may not represent non-English and older research. Finally the lack of quality of included studies, particularly in relation to participant blinding and reporting bias could compromise the integrity of review findings. In addition, there was some indication of publication bias for the meta-analysis of ORs (p=0.041). This suggest that some smaller studies showing a negative result might not have been published; therefore, the results of this metaanalysis need to be interpreted with some caution.

CONCLUSION

This systematic review showed that interventions had a significant effect on improving medication adherence in populations with T2DM and/or CVD. Multifaceted interventions which included either regular patient contact or an element of education had the most significant effect. There is, however, a need to compare more standardised interventions and assess these using more uniform methods of measuring medication adherence to enable studies to be more realistically compared.

With regard to trial design, recently there has been a focus on designing trials that are pragmatic and therefore more representative of 'real life'. The findings from this review suggest that the effectiveness of interventions between pragmatic and explanatory trials was comparable, suggesting that findings can be transferred from idealised to real-word conditions. There is, however, a need for further guidance to be developed to assist researchers in characterising and scoring studies.

Author affiliations

¹Diabetes Research Centre, University of Leicester, Leicester, UK

²Leicester Diabetes Centre, University Hospitals of Leicester NHS Trust, Leicester General Hospital, Leicester, UK

³Department of Oncology and Metabolism, University of Sheffield, Sheffield, UK ⁴Department of Cardiovascular Sciences, University of Leicester, Leicester, UK ⁵Department of Chemical Pathology and Metabolic Diseases, University Hospitals of Leicester NHS Trust, Leicester, UK

⁶NIHR CLAHRC East Midlands, Leicester, UK

⁷NIHR ARC East Midlands, Leicester, UK

Twitter Samuel Seidu @sis11@le.ac.uk

Acknowledgements The authors acknowledge the financial support received from the National Institute for Health Research Collaboration for Leadership in Applied Health Research and Care—East Midlands (NIHR CLAHRC—EM), NIHR Applied

Contributors CF developed the study protocol, carried out the scientific literature search, screened and extracted data, undertook PRECIS-2 scoring, quality assessed studies, interpreted the results and developed the first draft of the manuscript. CG quality assessed studies, interpreted the results, undertook data checking, PRECIS-2 scoring, performed all statistical analysis and contributed to the development of the manuscript. SS assisted with the design of the review and provided intellectual content to the drafted manuscript. DK assisted with the design of the review, applied for funding and assisted with protocol development. El screened articles for inclusion in the review. MJD provided intellectual content to the drafted manuscript. PP and PG assisted with the design of the review. KK had the concept of the study, applied for funding, assisted with protocol development and provided intellectual oversight on the development of the manuscript. KK accepts full responsibility for the conduct of the study, has access to the data and controlled the decision to publish. KK as the corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted.

Funding The authors received funding from the NIHR CLAHRC—EM and the Leicester Biomedical Research Centre.

Disclaimer The views expressed are those of the authors and not necessarily those of the National Health Service, the NIHR or the Department of Health.

Competing interests MJD reports grants from Novo Nordisk, grants from Sanofi-Aventis, grants from Lilly, grants from Boehringer Ingelheim, grants from Janssen, personal fees from Novo Nordisk, personal fees from Sanofi-Aventis, personal fees from Lilly, personal fees from Merck Sharp & Dohme, personal fees from Boehringer Ingelheim, personal fees from AstraZeneca, personal fees from Janssen, personal fees from Servier, personal fees from Mitsubishi Tanabe Pharma, personal fees from Takeda Pharmaceuticals International, outside the submitted work. KK reports personal fees from Amgen, personal fees from AstraZeneca, personal fees from Bayer, personal fees from NAPP, personal fees from Lilly, personal fees from Merck Sharp & Dohme, personal fees from Novartis, personal fees from Novo Nordisk, personal fees from Roche, personal fees from Berlin-Chemie AG/Menarini Group, personal fees from Sanofi-Aventis, personal fees from Servier, personal fees from Boehringer Ingelheim, grants from Pfizer, grants from Boehringer Ingelheim, grants from AstraZeneca, grants from Novartis, grants from Novo Nordisk, grants from Sanofi-Aventis, grants from Lilly, grants from Merck Sharp & Dohme, grants from Servier, outside the submitted work.

Patient and public involvement Patient and public involvement (PPI) was undertaken prior to the development of the review to obtain an understanding of patient reasons for non-adherence. A number of themes were identified and included side effects of drugs and poor explanations from healthcare professionals about prescribed drugs. Following the completion of the review, alongside the views and opinions expressed during the PPI, the study team aim to develop a toolbox and education programme to support medication adherence. A second focus group is planned comprising health professionals, with the objective of gathering their perspectives on how to promote the 'toolbox' for general utilisation in clinical practice and how best to disseminate and mainstream the toolbox of interventions.

Patient consent for publication Not required.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data are available on reasonable request. Data are available on reasonable request from the corresponding author.

Open access This is an open access article distributed in accordance with the Creative Commons Attribution 4.0 Unported (CC BY 4.0) license, which permits others to copy, redistribute, remix, transform and build upon this work for any purpose, provided the original work is properly cited, a link to the licence is given, and indication of whether changes were made. See: https://creativecommons.org/ licenses/by/4.0/.

ORCID iD

Kamlesh Khunti http://orcid.org/0000-0003-2343-7099

REFERENCES

- 1 Guthrie B, Payne K, Alderson P, *et al.* Adapting clinical guidelines to take account of multimorbidity. *BMJ* 2012;345:e6341.
- 2 World Health Organisation. *Global health estimates 2016: deaths by cause, age, sex, by country, and by region, 2000-2016.* Geneva, 2018.

- 3 Barnett K, Mercer SW, Norbury M, et al. Epidemiology of multimorbidity and implications for health care, research, and medical education: a cross-sectional study. Lancet 2012;380:37–43.
- 4 Marengoni A, Onder G. Guidelines, polypharmacy, and drug-drug interactions in patients with multimorbidity. *BMJ* 2015;350:h1059.
- 5 Sabaté E. Adherence to long-term therapies: evidence for action. World Health Organization, 2003.
- 6 Glynn LG. Multimorbidity: another key issue for cardiovascular medicine. *Lancet* 2009;374:1421–2.
- 7 Tinetti ME, Fried TR, Boyd CM. Designing health care for the most common chronic condition--multimorbidity. *JAMA* 2012;307:2493-4.
- 8 Weiss CO, Boyd CM, Yu Q, *et al.* Patterns of prevalent major chronic disease among older adults in the United States. *JAMA* 2007;298:1158–62.
- 9 McCambridge J, Kypri K, Elbourne D. In randomization we trust? there are overlooked problems in experimenting with people in behavioral intervention trials. *J Clin Epidemiol* 2014;67:247–53.
- 10 Godin G, Sheeran P, Conner M, et al. Asking questions changes behavior: mere measurement effects on frequency of blood donation. *Health Psychol* 2008;27:179–84.
- McCambridge J, Kypri K, Elbourne D. Research participation effects: a skeleton in the methodological cupboard. *J Clin Epidemiol* 2014;67:845–9.
- 12 Schwartz D, Lellouch J. Explanatory and pragmatic attitudes in therapeutical trials. J Chronic Dis 1967;20:637–48.
- 13 Sajobi TT, Li G, Awosoga O, et al. A comparison of meta-analytic methods for synthesizing evidence from explanatory and pragmatic trials. Syst Rev 2018;7:19.
- 14 Yoong SL, Wolfenden L, Clinton-McHarg T, et al. Exploring the pragmatic and explanatory study design on outcomes of systematic reviews of public health interventions: a case study on obesity prevention trials. J Public Health 2014;36:170–6.
- 15 Moher D, Liberati A, Tetzlaff J, et al. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. Ann Intern Med 2009;151:264–9.
- 16 Sterne JAC, Savović J, Page MJ, et al. Rob 2: a revised tool for assessing risk of bias in randomised trials. BMJ 2019;366:I4898.
- 17 Higgins JPT, Altman DG, Gøtzsche PC, et al. The Cochrane collaboration's tool for assessing risk of bias in randomised trials. BMJ 2011;343:d5928.
- 18 Loudon K, Treweek S, Sullivan F, et al. The PRECIS-2 tool: designing trials that are fit for purpose. BMJ 2015;350:h2147.
- 19 Khayyat SM, Khayyat SMS, Hyat Alhazmi RS, et al. Predictors of medication adherence and blood pressure control among Saudi hypertensive patients attending primary care clinics: a crosssectional study. *PLoS One* 2017;12:e0171255.
- 20 Krousel-Wood M, Islam T, Muntner P, et al. Association of depression with antihypertensive medication adherence in older adults: crosssectional and longitudinal findings from CoSMO. Ann Behav Med 2010;40:248–57.
- 21 Cohen J. Statistical power analysis for the behavioral sciences. Routledge, 2013.
- 22 Higgins JPT, Thompson SG, Deeks JJ, *et al.* Measuring inconsistency in meta-analyses. *BMJ* 2003;327:557–60.
- 23 Meng K, Seekatz B, Haug G, et al. Evaluation of a standardized patient education program for inpatient cardiac rehabilitation: impact on illness knowledge and self-management behaviors up to 1 year. Health Educ Res 2014;29:235–46.
- 24 Castellano JM, Sanz G, Peñalvo JL, *et al*. A polypill strategy to improve adherence: results from the focus project. *J Am Coll Cardiol* 2014;64:2071–82.
- 25 Al-Haj Mohd MMM, Phung H, Sun J, et al. Improving adherence to medication in adults with diabetes in the United Arab Emirates. BMC Public Health 2016;16:857.
- 26 Barker-Collo S, Krishnamurthi R, Witt E, *et al.* Improving adherence to secondary stroke prevention strategies through motivational interviewing: randomized controlled trial. *Stroke* 2015;46:3451–8.
- 27 Boyne JJJ, Vrijhoef HJM, Spreeuwenberg M, et al. Effects of tailored telemonitoring on heart failure patients' knowledge, self-care, self-efficacy and adherence: a randomized controlled trial. Eur J Cardiovasc Nurs 2014;13:243–52.
- 28 Buhse S, Kuniss N, Liethmann K, *et al.* An informed shared decision making programme for patients with type 2 diabetes in primary care: cluster randomised controlled trial. *Diabetologia* 2017;1:S347.
- 29 Caetano IRCES, Santiago LM, Marques M. Impact of written information on control and adherence in type 2 diabetes. *Rev Assoc Med Bras* 2018;64:140–7.
- 30 Carrasquillo O, Lebron C, Alonzo Y, et al. Effect of a community health worker intervention among Latinos with poorly controlled type 2 diabetes: the Miami healthy heart initiative randomized clinical trial. *JAMA Intern Med* 2017;177:948–54.

- 31 Crowley MJ, Powers BJ, Olsen MK, *et al*. The cholesterol, hypertension, and glucose education (change) study: results from a randomized controlled trial in African Americans with diabetes. *Am Heart J* 2013;166:179–86.
- 32 Du L, Dong P, Jia J, et al. Impacts of intensive follow-up on the long-term prognosis of percutaneous coronary intervention in acute coronary syndrome patients - a single center prospective randomized controlled study in a Chinese population. Eur J Prev Cardiol 2016;23:1077–85.
- 33 El-Toukhy H, Omar A, Abou Samra M. Effect of acute coronary syndrome patients' education on adherence to dual antiplatelet therapy adherence to dual antiplatelet therapy. *J Saudi Heart Assoc* 2017;29:252–8.
- 34 Graumlich JF, Wang H, Madison A, et al. Effects of a patient-provider, collaborative, Medication-Planning tool: a randomized, controlled trial. J Diabetes Res 2016;2016:2129838
- 35 Hedegaard U, Kjeldsen LJ, Pottegard A, et al. A multifaceted pharmacist intervention to support medication adherence after stroke and transient ischemic attack. Int J Clin Pharm 2015;37:187.
- 36 Ho PM, Lambert-Kerzner A, Carey EP, et al. Multifaceted intervention to improve medication adherence and secondary prevention measures after acute coronary syndrome hospital discharge: a randomized clinical trial. JAMA Intern Med 2014;174:186–93.
- 37 Jeong JY, Jeon J-H, Bae K-H, et al. Smart care based on Telemonitoring and telemedicine for type 2 diabetes care: multi-center randomized controlled trial. *Telemed J E Health* 2018;24:604–13.
- 38 Kronish IM, Goldfinger JZ, Negron R, et al. Effect of peer education on stroke prevention: the prevent recurrence of all inner-city strokes through education randomized controlled trial. Stroke 2014;45:3330–6.
- 39 Márquez-Contreras E, Martell-Claros N, Márquez-Rivero S, et al. Strategies for improving dabigatran adherence for stroke prevention in patients with non-valvular atrial fibrillation: education and drug intake reminders (FACILITA study). Curr Med Res Opin 2018;34:1301–8.
- 40 Meng K, Musekamp G, Schuler M, et al. The impact of a selfmanagement patient education program for patients with chronic heart failure undergoing inpatient cardiac rehabilitation. *Patient Educ Couns* 2016;99:1190–7.
- 41 Pladevall M, Divine G, Wells KE, et al. A randomized controlled trial to provide adherence information and motivational interviewing to improve diabetes and lipid control. *Diabetes Educ* 2015;41:136–46.
- 42 Rinfret S, Rodés-Cabau J, Bagur R, et al. Telephone contact to improve adherence to dual antiplatelet therapy after drug-eluting stent implantation. *Heart* 2013;99:562–9.
- 43 Schou M, Gislason G, Videbaek L, et al. Effect of extended followup in a specialized heart failure clinic on adherence to guideline recommended therapy: NorthStar adherence study. Eur J Heart Fail 2014;16:1249–55.
- 44 Schwalm J-D, Ivers NM, Natarajan MK, et al. Cluster randomized controlled trial of delayed educational reminders for long-term medication adherence in ST-elevation myocardial infarction (DERLA-STEMI). Am Heart J 2015;170:903–13.
- 45 Vollmer WM, Owen-Smith AA, Tom JO, et al. Improving adherence to cardiovascular disease medications with information technology. Am J Manag Care 2014;20:sp502–10.
- 46 Volpp KG, Troxel AB, Mehta SJ, et al. Effect of electronic reminders, financial incentives, and social support on outcomes after myocardial

infarction: the HeartStrong randomized clinical trial. *JAMA Intern Med* 2017;177:1093–101.

- 47 Cao X-Y, Tian L, Chen L, et al. Effects of a hospital-community partnership transitional program in patients with coronary heart disease in Chengdu, China: a randomized controlled trial. Jpn J Nurs Sci 2017;14:320–31.
- 48 Chung WW, Chua SS, Lai PSM, et al. Effects of a pharmaceutical care model on medication adherence and glycemic control of people with type 2 diabetes. *Patient Prefer Adherence* 2014;8:1185–94.
- 49 Jia J-J, Dong P-S, Du L-J, *et al.* Impact of Physician-Coordinated intensive follow-up on long-term medical costs in patients with unstable angina undergoing percutaneous coronary intervention. *Acta Cardiol Sin* 2017;33:173–81.
- 50 Lin C-Y, Yaseri M, Pakpour AH, et al. Can a multifaceted intervention including motivational interviewing improve medication adherence, quality of life, and mortality rates in older patients undergoing coronary artery bypass surgery? A multicenter, randomized controlled trial with 18-month follow-up. *Drugs Aging* 2017;34:143–56.
- 51 Peng B, Ni J, Anderson CS, et al. Implementation of a structured guideline-based program for the secondary prevention of ischemic stroke in China. Stroke 2014;45:515–9.
- 52 Su Q, Li C, Long F, *et al.* Effects of a health promotion program on medication adherence to antiplatelet therapy among ischemic stroke patients in Hainan Province, China. *Vascular* 2017;25:242–8.
- 53 Xin C, Xia Z, Jiang C, et al. Effect of pharmaceutical care on medication adherence of patients newly prescribed insulin therapy: a randomized controlled study. *Patient Prefer Adherence* 2015;9:797–802.
- 54 Samtia AM, Rasool MF, Ranjha NM, et al. A multifactorial intervention to enhance adherence to medications and disease-related knowledge in type 2 diabetic patients in southern Punjab, Pakistan. *Trop J Pharm Res* 2013;12:851–6.
- 55 Xavier D, Gupta R, Kamath D, *et al.* Community health worker-based intervention for adherence to drugs and lifestyle change after acute coronary syndrome: a multicentre, open, randomised controlled trial. *Lancet Diabetes Endocrinol* 2016;4:244–53.
- 56 Marin GH, Risso P, Sbatella D, *et al*. Treatment adherence by personalizing the drug dispensing for diabetic patients in social vulnerable situation. *Quality in Primary Care* 2015;23:93–6.
- 57 Morisky DE, Green LW, Levine DM. Concurrent and predictive validity of a self-reported measure of medication adherence. *Med Care* 1986;24:67–74.
- 58 Morisky DE, Ang A, Krousel-Wood M, et al. Predictive validity of a medication adherence measure in an outpatient setting. J Clin Hypertens 2008;10:348–54.
- 59 Johnson KE, Neta G, Dember LM, *et al*. Use of Precis ratings in the National Institutes of health (NIH) health care systems research Collaboratory. *Trials* 2016;17:32.
- 60 Nieuwlaat R, Wilczynski N, Navarro T, et al. Interventions for enhancing medication adherence. *Cochrane Database Syst Rev* 2014;11:CD000011.
- 61 Zhou J, Rau P-LP, Salvendy G. Age-Related difference in the use of mobile phones. *Univers Access Inf Soc* 2014;13:401–13.
- 62 Changizi M, Kaveh MH. Effectiveness of the mHealth technology in improvement of healthy behaviors in an elderly population-a systematic review. *Mhealth* 2017;3:51.
- 63 Stirratt MJ, Dunbar-Jacob J, Crane HM, et al. Self-Report measures of medication adherence behavior: recommendations on optimal use. *Transl Behav Med* 2015;5:470–82.