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## signaling as a driver for refractive error

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### Abstract

Refractive errors, including myopia, are the most frequent eye disorders worldwide and an increasingly common cause of blindness. This genome-wide association meta-analysis in 160,420 participants and replication in 95,505 participants, increased the established independent signals from 37 to 161 and revealed high genetic correlation between Europeans and Asians ( $>0.78$ ). Expression experiments and comprehensive *in silico* analyses identified retinal cell physiology and light processing as prominent mechanisms, and functional contributions to refractive error development in all cell types of the neurosensory retina, retinal pigment epithelium, vascular endothelium and extracellular matrix. Newly identified genes elicited novel mechanisms such as rod and cone bipolar synaptic neurotransmission, anterior segment morphology, and angiogenesis. Thirty-one loci resided in or near regions transcribing small RNAs, suggesting a role for post-transcriptional regulation. Our results support the notion that refractive errors are caused by a light-dependent retina-to-sclera signaling cascade, and delineate potential pathobiological molecular drivers.

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#### Competing interests

N.A.F., N.E., J.Y.T., and the 23andMe Research Team are current or former employees of 23andMe, Inc., and hold stock or stock options in 23andMe. J.B. J. is a patent holder with Biocompatibles UK Ltd. (Franham, Surrey, UK) (Title: Treatment of eye diseases using encapsulated cells encoding and secreting neuroprotective factor and/or anti-angiogenic factor; Patent number: 20120263794), and patent application with University of Heidelberg (Heidelberg, Germany) (Title: Agents for use in the therapeutic or prophylactic treatment of myopia or hyperopia; European Patent Number: 3 070 101). The other authors declare no competing financial interests.

#### Author contributions

M.S.T., V.J.M.V., S.M., J.A.G., A.I.I.G., R.W., P.G.H., A.I.I.G., and E.M.v.L. performed the analyses. C.C.W.K., V.J.M.V., M.S.T., R.W., J.A.G., and S.M. drafted the manuscript, and C.J.H., P.G.H., A.P.K., C.M.v.D., D.S., E.M.v.L., J.E.B.W., J.Y.T., N.A.F., Q.F., S.M.S., and V.V. critically reviewed the manuscript. A.N., A.P.K., A.T., C.B., C.Gi., C.L.S., C.Y.C., G.Bi., G.C., I.R., J.E.B.W., J.E.H., J.S.Ri., J.W., J.X., K.M.W., K.Y., M.P.C., M.S.H., M.S.T., N.A.F., N.E., P.C., P.Gh., P.K.J., Q.F., R.Ho., R.L.S., R.P.I., R.W., T.H., T.H.S.A., T.Z., V.V., W.Y.S., W.Z., X.L.S., Y.C.H., Y.S., and Y.Y.T. performed data analysis for the individual studies; and A.D.P., A.G.U., A.T., A.W.H., B.E.K.K., C.C.W.K., C.D., C.Gr., C.H., C.J.H., C.W., C.Y.C., D.A.M., F.R., G.Be., H.M.H., J.A.G., J.B.J., J.E.B.W., J.E.C., J.F.W., J.H.L., J.R.V., J.S.Ra., J.S.Ri., J.Y.T., K.Y., M.A.M.S., N.G.M., N.P., O.Po., O.Pa., O.T.R., P.Gu., P.J.F., P.M., P.N.B., R.K., S.K.I., S.M.S., T.L., T.M., W.Z., Y.C.H., and Y.X.W. contributed to data assembly. A.A.B.B., A.W., C.Gr., D.S., K.N.W., S.W.J.T., and T.Y. performed expression experiments, and M.S.T., A.A.B.B., P.J.v.d.S., and R.Ha. performed *in silico* pathway analyses. C.C.W.K. and C.J.H. conceived and designed the outline of the current report, and jointly with A.M., A.H., A.W.H., C.D., C.H., C.J.H., C.M.v.D., C.W., C.Y.C., D.A.M., D.S., E.S.T., F.M., G.Bi., I.R., J.A.G., J.B.J., J.E.B.W., J.E.C., J.F.W., J.H.L., J.R.V., J.Y.T., N.A., N.A.F., N.P., O.Pa., O.T.R., P.J.F., P.N.B., S.K.I., S.M.S., T.L., T.Y.W., T.Y., V.V., Y.X.W., and Y.Y.T. supervised conduction of experiments and analyses.

## INTRODUCTION

Refractive errors are common optical aberrations determined by mismatches in the focusing power of the cornea, lens and axial length of the eye. Their distribution is rapidly shifting towards myopia, or nearsightedness, all over the world. The myopia boom is particularly prominent in urban East Asia where up to 95% of twenty-year-olds in cities such as Seoul and Singapore have this refractive error<sup>1-4</sup>. Myopia prevalence is also rising throughout Western Europe and the USA, affecting ~50% of young adults in these regions<sup>5,6</sup>. While refractive errors can be optically corrected, even at moderate values they carry significant risk of ocular complications with high economic burden<sup>7-9</sup>. One in three individuals with high myopia (-6 diopters or worse) will develop irreversible visual impairment or blindness, mostly due to myopic macular degeneration, retinal detachment, or glaucoma<sup>10,11</sup>. At the other extreme, high hyperopia predisposes to strabismus, amblyopia and angle-closure glaucoma<sup>10,12</sup>.

Refractive errors result from a complex interplay of lifestyle and genetic factors. The most established lifestyle factors for myopia are high education, lack of outdoor exposure, and excessive near work<sup>3</sup>. Recent research has identified many genetic variants for refractive errors, myopia, and axial length<sup>13-25</sup>. Two large studies, the international Consortium for Refractive Error and Myopia (CREAM)<sup>26</sup> and the personal genomics company 23andMe, Inc.<sup>17,27</sup> have provided the most comprehensive results.<sup>28</sup>

Given that only 3.6% of the variance of the refractive error trait was explained by the identified genetic variants<sup>26</sup>, we presumed a high missing heritability. We therefore combined CREAM and 23andMe, and expanded the study sample to 160,420 individuals from a mixed ancestry population with quantitative information on refraction for a genome-wide association (GWAS) meta-analysis. Index variants were tested for replication in an independent cohort consisting of 95,505 individuals from the UK Biobank. We conducted systematic comparisons to assess differences in genetic inheritance and distribution of risk variants between Europeans and Asians. Polygenic risk analyses were performed to evaluate the contribution of the identified variants to the risk of myopia and hyperopia. Finally, we integrated expression data and bioinformatics on the identified genes to gain insight into the possible mechanisms underlying the genetic associations.

## RESULTS

### Susceptibility loci for refractive error

We performed a GWAS meta-analysis on adult untransformed spherical equivalent (SphE) using summary statistics from 37 studies from CREAM and on age of diagnosis of myopia (AODM) from two cohorts from 23andMe (Supplementary Figure 1, Supplementary Table 1a)<sup>26,27</sup>. Analyses were based on ~11 million genetic variants (SNPs, insertions and deletions) genotyped or imputed to 1000 Genomes Project Phase I reference panel (version 3, March 2012 release<sup>29</sup>) that passed extensive quality control (Supplementary Figures 2-4, Supplementary Table 1b).

Meta-analyses were conducted in three stages: *Stage 1* CREAM (CREAM-EUR,  $N=44,192$ ; CREAM-ASN,  $N=11,935$ ); *Stage 2* 23andMe ( $N=104,293$ ; Online Methods); *Stage 3* joint meta-analysis of Stage 1 and 2. As CREAM and 23andMe applied different phenotype measures, we used signed Z-scores as the mean per-allele effect size and assigned equal weights to CREAM and 23andMe. We identified 7,967 genome-wide significant genetic variants clustering in 140 loci (Figure 1a,b; Supplementary Figure 5–6, Supplementary Table 2-5, Supplementary Data 1-2), replicating all 37 previously discovered loci and finding 104 novel loci. We applied genomic control at each stage and checked for population stratification using LD score regression<sup>30</sup> (Stage 1-2 inflation factors ( $\lambda_{GC}$ )  $<1.1$  and LD score regression intercepts ( $LDSC_{intercept}$ ) 0.892-1.023; Supplementary Table 6; Supplementary Figure 6-7). At Stage 3, we observed a genomic inflation ( $\lambda_{GC}=1.129$ ; Supplementary Figure 6), probably due to true polygenicity rather than population stratification or cryptic relatedness<sup>31</sup>.  $LDSC_{intercept}$  remained undetermined due to mixed ancestry.

To detect the presence of multiple independent signals at the discovered loci, a stepwise conditional analysis was performed with GCTA-COJO<sup>32</sup> on summary statistics from all European cohorts ( $N=148,485$ ) using the Rotterdam Study I-III (RS I-III) as a reference panel for LD structure ( $N_{RSI-III}=10,775$ ). This analysis yielded 27 additional independent variants, resulting in a total of 167 loci (Supplementary Table 2).

We advanced these loci for replication in a GWAS of refractive error carried out by the UK Biobank Eye & Vision (UKEV) Consortium ( $N=95,505$ )<sup>33</sup> (Online Methods). Six out of the 167 variants were not considered for replication analysis. One of these five variants (rs3138141, *RDH5*) was identified previously and therefore still considered as a refractive error risk variant<sup>26,27</sup>. The remaining 161 genetic variants were tested for replication. 86% (138/161) of the candidate variants replicated significantly: 104 (65%) replicated surpassing genome-wide significance and 34 replicated surpassing Bonferroni *correction* ( $P < 3.0 \times 10^{-4}$ ; 21.1%); another 12 showed nominal evidence for replication ( $0.05 < P < 3.0 \times 10^{-4}$ ; 7.5%) and only 11 (7%) did not replicate at all (Table 1, Supplementary Table 2).

As CREAM and 23andMe employed different phenotypic outcomes, we evaluated consistency of genotypic effects by comparing marker-wise additive genetic effect sizes (in units diopters per risk allele variant) for SphE from CREAM-EUR against those (in units log(HR) per risk allele variant) for AODM from 23andMe. All variants strongly associated with either outcome ( $P < 0.001$ ) were concordant in direction-of-effect, and had highly correlated effect sizes (Figure 2a,b; Supplementary Figure 8). For these variants a 10% decrease in log(HR) for AODM, indicating an earlier age-at-myopia onset, was associated with a decrease of 0.15 diopters in SphE. A quantitative analysis for all common SNPs ( $MAF > 0.01$ ; HapMap3) using LD score regression yielded a genetic correlation of 0.93 (95% CI 0.86-0.99;  $P = 2.1 \times 10^{-159}$ ), confirming that effect sizes for both phenotypic outcomes were closely related.

### Gene annotation of susceptibility loci

We annotated all genetic variants with WANNVAR using the University of California Santa Cruz (UCSC) Known Gene database<sup>34,35</sup>. The identified 139 genetic loci were

annotated to 208 genes and known transcribed RNA genes (Table 1, Supplementary Table 2, Online Methods). The physical positions of the lead genetic variants relative to protein-coding genes are shown in Figure 1c. 86% of the identified variants were either intragenic or less than 50 kb from the 5' or 3' end of the transcription start site. We found seven exonic variants (Supplementary Table 7) of which two had MAF  $\geq 0.05$ : rs5442 (*GNB3*) and rs17400325 (*PDE11A*). The index SNP in the *GNB3* locus with MAF 0.05 in Europeans is a highly conserved missense variant (G272S) predicted to be damaging by PolyPhen-2<sup>36</sup> and SIFT<sup>37</sup>. *PDE11A* is presumed to play a role in tumorigenesis, brain function, and inflammation<sup>38</sup>. The index SNP in the *PDE11A* locus with MAF 0.03 in Europeans is also a highly conserved missense variant (Y727C); this variant was predicted to be damaging by PolyPhen<sup>36</sup>, SIFT<sup>39</sup> and align GVGD<sup>40,41</sup>. The other exonic variants, rs1064583 (*COL10A1*), rs807037 (*KAZALD1*), rs1550094 (*PRSS56*), rs35337422 (*RD3L*), and rs6420484 (*TSPAN10*) were not predicted to be damaging.

The most significant variant (Stage 3; rs12193446,  $P=4.21 \times 10^{-84}$ ) resides on chromosome 6 within a non-coding RNA, *BC035400*, in an intron of the *LAMA2* gene. This locus had been identified previously, but our current fine mapping redefined the most associated variant. The function and potential downstream target sites for BC035400 are currently unknown. The previously most strongly associated variant, rs524952 on chromosome 15 near *GJD2*, was the second most significant variant ( $P=2.28 \times 10^{-65}$ ).

### Post-GWAS analyses

We performed two gene-based tests, fastBAT<sup>42</sup> and EUGENE<sup>43</sup>, and applied a functional enrichment approach using fgwas<sup>44</sup> (Online Methods). With fastBAT, we identified 13 genes at  $P < 2.0 \times 10^{-6}$ , one of which (*CHD7*) had been identified previously<sup>26,27</sup>. Using EUGENE, we found 7 genes at  $P < 2.0 \times 10^{-6}$  after incorporation of blood eQTLs. With fgwas, we identified 6 loci, which could be annotated to 9 genes, at posterior probability  $> 0.9$ . Two genes (*HMGN4* and *TLX1*) showed significant associations in two or more approaches. Taken together, these post-GWAS approaches resulted in a total of 22 additional candidate loci for refractive error, annotated to 25 genes (Supplementary Table 8). This increases the overall number of significant genetic associations to 161 candidate loci.

### Polygenic risk scores

We calculated polygenic risk scores (PGRS)<sup>45</sup> per individual at various  $P$  thresholds (Online Methods) for Rotterdam Study I-III (RS I-III;  $N=10,792$ ) after recalculating  $P$  and  $Z$ -scores of variants from Stage 3 excluding RS I-III. We found the highest fraction of phenotypic variance (7.8%) explained with 7,307 variants at  $P$  value threshold 0.005 (Supplementary Table 9). A PGRS based on these variants distinguished well between individuals with hyperopia and myopia at the lower and higher deciles (Figure 3); those in the highest decile had a 40-fold increased risk of myopia. When the PGRS was stratified for the median age ( $< 63$  or  $> 63+$  yrs), we found a significant difference in the variance explained ( $< 63$  yrs 8.9%;  $63+$  yrs 7.4%;  $P=0.0038$ ). The variance explained by PGRS was not significantly different between males and females (8.3% vs 7.5%;  $P=0.13$ ). The predictive value (area under the receiver operating characteristic curve, AUC) of the PGRS for myopia versus

hyperopia adjusted for age and gender was 0.77 (95% CI=0.75–0.79), a 10% increase compared to previous estimations<sup>46</sup>.

### Trans-ethnic comparison of genotypic effects

To explore potential ancestry differences in the identified refractive error loci, we calculated the heritability explained by common genetic variants (SNP- $h^2$ ) for Europeans and Asians using LD score regression<sup>47</sup>. SNP- $h^2$  was 0.214 (95% CI 0.185- 0.243) and 0.172 (95% CI 0.154- 0.190) in the European samples (CREAM-EUR and 23andMe), while it was only 0.053 (95% CI -0.025- 0.131) in the Asian sample (CREAM-EAS). Next, we estimated the genetic correlation between Europeans and Asians by comparing variant effect size for common variants using Popcorn<sup>48</sup> (Online Methods). Two genetic correlation metrics were calculated: First, a genetic effect correlation ( $\rho_{ge}$ ) that quantifies the correlation in SNP effect sizes between Europeans and Asians without taking into account ancestry-related differences in allele frequency; and second, a genetic impact correlation ( $\rho_{gi}$ ) that estimates the correlation in variance-normalized SNP effect sizes between the two ancestry groups (Table 2). Estimates of  $\rho_{ge}$  were high between Europeans and Asians, but significantly different from 1 (0.79 and 0.80, respectively at  $P < 1.9 \times 10^{-6}$ ; Table 2), indicating a clear genetic overlap but a difference in per allele effect size. Estimates of  $\rho_{gi}$  were similarly high ( $>0.8$ ), but not significantly different from 1 for the correlation between CREAM-EUR and CREAM-ASN ( $P=0.065$ ), indicating that the genetic impact of these alleles may still be similar.

### In silico pathway analysis

We used an array of bioinformatics tools to investigate potential functions and pathways of the associated genes. We first employed DEPICT<sup>49</sup> to perform a gene set enrichment analysis, a tissue type enrichment, and a gene prioritization analysis, on all variants with  $P < 1.00 \times 10^{-5}$  from Stage 3. The gene set enrichment analysis resulted in 66 reconstituted gene sets, of which 55 (83%) were eye-related. To reduce redundancies between pathways, we clustered the significant pathways into 13 meta gene sets (false discovery rate (FDR)  $< 5\%$  and a  $P < 0.05$ ) (Supplementary Note 2, Figure 4, Supplementary Table 10). The most significant gene set was the ‘abnormal photoreceptor inner segment morphology’ (MP:0003730;  $P=1.79 \times 10^{-7}$ ). The eye-related meta gene sets consisted of the ‘thin retinal outer nuclear layer’ (MP:0008515; 27 (55%) gene sets), ‘detection of light stimulus’ (GO:0009583; 13 (24%) gene sets), ‘nonmotile primary cilium’ (GO:0031513; 4 (6%) gene sets), and ‘abnormal anterior eye segment morphology’ (MP:0005193; 4 (6%) gene sets). The first three meta gene sets had a Pearson’s correlation  $> 0.6$ . Interestingly, *RGR*, *RPIL1*, *RORB* and *GNB3* were present in all of these meta gene sets. Retina was the most significant tissue of expression according to the tissue enrichment analysis ( $P=1.11 \times 10^{-4}$ , FDR  $< 0.01$ ). From the gene prioritization according to DEPICT, 7 genes were highlighted as the most likely causal genes at  $P < 7.62 \times 10^{-6}$  and FDR  $< 0.05$ : *ANO2*, *RPIL1*, *GNB3*, *EDN2*, *RORB* and *CABP4*.

Next, we performed a canonical pathway analysis on all genes annotated to the variants of Stage 3 using Ingenuity Pathway Analysis (See URLs). All genes were run against the IPA database incorporating functional biological evidence on genomic and proteomic expression

based on regulation or binding studies. IPA identified “Glutamate Receptor Signaling” with central player *NF-κB* gene as the most significant pathway after correction for multiple testing (ratio of the number of molecules 8.8% and Fisher’s Exact test  $P=1.56\times 10^{-4}$ ; Supplementary Figure 9).

### From disease-associated loci to biological mechanisms

We adapted the scoring scheme designed by Fritsche et al.<sup>50</sup> to highlight genes for which there is biological plausibility for a role in eye growth<sup>50</sup>. We used 10 equally rated categories (Online Methods; Figure 5; Supplementary Table 11; Supplementary Note 2). One-hundred-and-nine index variants replicated in two or more individual cohorts; we found evidence for seven genetic variants with eQTL effects in multiple tissue types; nine exonic variants, of which seven predicted protein-alterations (Supplementary Table 7); 31 RNA genes, five located in the 3’ or 5’ UTR (Supplementary Table 12, Supplementary Figure 10), 84 genes resulting in an ocular phenotype in humans (Supplementary Table 13) and 36 in mice (Supplementary Table 14); 172/212 (81%) genes expressed in human ocular tissue (Supplementary Note 2, Supplementary Table 15); 41 genes identified by DEPICT at  $P < 5.4\times 10^{-4}$  and  $FDR < 0.05$  and 45 genes contributed to the most significant canonical pathways of IPA. Notably, 48 of the associated genes encode known drug targets (Supplementary Table 16).

The gene with the highest biological plausibility score (score=8) was *GNB3*, a highly conserved gene encoding a guanine nucleotide-binding protein expressed in rod and cone photoreceptors and ON-bipolar cells<sup>51</sup>. *GNB3* participates in signal transduction through G-protein-coupled receptors and enhances the temporal accuracy of phototransduction and ON-center signaling in the retina<sup>51</sup>. As described above, the index SNP harbors a missense variant associated with refractive errors. Non-synonymous mutations within *GNB3* are known to cause syndromic congenital stationary night blindness<sup>52</sup> in humans, progressive retinopathy and globe enlargement in chickens<sup>51</sup>, and abnormal development of the photoreceptor-bipolar synapse in knock-out mice<sup>53,54</sup>.

Other genes highly ranked (score=7) include *CYP26A1*, *GRIA4*, *RDH5*, *RORB*, and *RGR*, all previously associated with refractive error, and one newly identified gene: *EFEMP1*. *EFEMP1* encodes a member of the fibulin family of extracellular matrix glycoproteins, and is found pan-ocularly including in the inner nuclear layer and Bruch’s membrane. Mutations in this gene lead to specific macular dystrophies<sup>55</sup>, while variants have also been shown to co-segregate with primary open-angle glaucoma<sup>56</sup> and associate with optic disc cup area<sup>57</sup>.

Several other genes are noteworthy for their function. *CABP4*, a calcium-binding protein expressed in cone and rod photoreceptor cells, mediates  $Ca^{2+}$ -influx and glutamate release in the photoreceptor-bipolar synapse<sup>58</sup>. Mutations in this gene have been described in congenital cone-rod synaptic disorder<sup>59</sup>, a retinal dystrophy associated with nystagmus, photophobia, and, remarkably, high hyperopia. *KCNMA1* encodes pore-forming alpha subunits of  $Ca^{2+}$ -activated  $K^+$  (BK) channels. These channels regulate synaptic transmission exclusively in the rod pathway<sup>60</sup>. *ANO2* is a  $Ca^{2+}$ -activated  $Cl^-$  channel recently reported to regulate retinal pigment epithelial (RPE) cell volume in a light-dependent manner<sup>64</sup>. *EDN2* is a potent vasoconstrictor that binds to two G-protein-coupled receptors, *EDNRA*, which

resides on bipolar dendrites, and *EDNRB*, which is present on Mueller and horizontal cells. Both receptors are also present on choroidal vessels<sup>65</sup>, implying that the choroid as well as retinal cells are target sites for this gene. *RPIL1* is expressed in cone and rod photoreceptors where it is involved in the maintenance of microtubules in the connecting cilium<sup>66</sup>. Mutations in this gene cause dominant macular dystrophy and retinitis pigmentosa<sup>67</sup>. We replicated two genes known to cause myopia in family studies. *FBNI* harbors mutations causing with Marfan (OMIM #154700) and Weil Marchesani (OMIM #608328) syndrome; *PTPRR* was one of the candidates in the MYP3 locus, which was found by linkage in families with high myopia<sup>68</sup>.

The location of rs7449443 ( $P=3.58\times 10^{-8}$ ) is notable as it resides in between *DRD1* and *FLJ16171*. *DRD1* encodes dopamine receptor 1 and is known to modulate dopamine receptor 2-mediated events<sup>69,70</sup>. The dopamine pathway has been implicated in myopia pathogenesis in many studies<sup>69,71</sup>. SNPs in and near other genes involved in the dopamine pathway (dopamine receptors, synthesis, degradation, and transporters)<sup>72-74</sup> did not reveal genome-wide significant associations (Supplementary Note 3, Supplementary Table 17; Supplementary Figure 11).

There were 31 genetic variants in or near DNA structures transcribing RNA genes (non coding RNA, linc RNAs, tRNAs, snoRNAs, rRNAs). Notably, five were in the transcription region and thirteen were in the vicinity ( $>0$  kb and  $\leq 50$  kb) of start or end of the RNA gene transcription region. They received low scores, since many have no reported function or disease association to date (Figure 5, Supplementary Figure 10, Supplementary Table 12). Our ranking of genes based on functional information existing in the public domain does not necessarily represent the true order of importance for refractive error pathogenesis. The observation that genes with strong statistical association were distributed over all scores supports this concept. Nevertheless, this list may help to select genes for subsequent functional studies.

Finally, integration of all our findings, supported by literature, allowed us to annotate a large number of genes to ocular cell types (Figure 6). Remarkably, all cell types of the retina harbored refractive error genes, as well as the RPE, vascular endothelium, and extracellular matrix.

### Genetic pleiotropy

We performed a GWAS catalogue look up using FUMA to investigate overlap of genes with other common traits (Supplementary Figure 12)<sup>75</sup>. Refractive error and hyperopia were replicated significantly after correcting for multiple testing (adjusted  $P$  value= $1.44\times 10^{-52}$  and  $9.34\times 10^{-9}$ , respectively). We found significant overlap with 74 other traits, of which height (adjusted  $P$  value= $1.11\times 10^{-10}$ ), obesity (adjusted  $P=1.38\times 10^{-10}$ ), and BMI (adjusted  $P=4.05\times 10^{-7}$ ) were most important. Ocular diseases significantly associated were glaucoma (optic cup area, intraocular pressure; adjusted  $P=2.69\times 10^{-5}$  and  $3.01\times 10^{-5}$ , respectively) and age-related macular degeneration (adjusted  $P$  value= $1.27\times 10^{-3}$ ).

## DISCUSSION

Myopia may become the leading cause of world blindness in the near future, a grim outlook for which current counteractions are still insufficient<sup>11,76</sup>. To improve understanding of the genetic landscape and biology of refractive error, we conducted a large GWAS meta-analysis in 160,420 participants of mixed ancestry and replicated in 95,505 participants. This led to the identification of 139 independent susceptibility loci by single variant analysis and 22 additional loci through post-GWAS methods, a four-fold increase in refractive error genes. The majority of annotated genes were found to be expressed in the human posterior segment of the eye. Using in silico analysis, we identified significant biological pathways, of which retinal cell physiology, light processing, and, specifically, glutamate receptor signaling were the most prominent mechanisms. Our integrated bio-informatic approach highlighted known ocular functionality for many genes.

To ensure robustness of our genetic associations, we included studies of various designs and populations, sought replication in an independent cohort of significant sample size, and stringently accounted for population stratification by performing genomic control at all stages of the meta-analysis<sup>77</sup>. We combined studies with outcomes based on actual refractive error measurements as well as on self-reported age-of-myopia-onset, and found the direction-of-effect of the associated variants, as well as their effect size, to be remarkably consistent. Combining two different outcome measures may appear unconventional, but age of onset and refractive error have been shown to be very tightly correlated<sup>11,28,78,79</sup>. Moreover, the high genetic correlation (93%) of common SNPs between the two phenotypes underscores their similarity. Most compelling evidence was provided by replication of 86% of the discovered variants in the independent UKEV which also used conventional refractive error measurements. This robustness indicates that both phenotypic outcomes can be used to capture a shared source of genetic variation. In addition, we found trans-ethnic replication of significant loci, and a high correlation of genetic effects of common variants in the Europeans and Asians. Our findings support a largely shared genetic predisposition to refractive error and myopia in the two ethnicities, although ancestry-specific allelic effects may exist. The low heritability estimate in Asians may, in part, be explained by the low representation of this ethnicity in our study sample. Alternatively, it may imply that environmental factors explain a greater proportion of the phenotypic risk and recent rise in myopia prevalence in this ancestry group<sup>80</sup>.

Limitations of our study were the possibility of false negative findings due to genomic control, and underrepresentation of studies with Asian ancestry. Heterogeneity of observed effect estimates was large for several associated variants, but not unexpected, given the large number of collaborating studies with varying methodology.

Although neurotransmission was previously suggested pathway<sup>26,27</sup>, our current pathway analyses provide more in-depth insights into the retinal circuitry driving refractive error. DEPICT identified ‘thin retinal outer nuclear layer’, ‘detection of light stimulus’, and ‘nonmotile primary cilium’ as the most important meta-gene sets. These are the main characteristics of photoreceptors, which are located in the outer retina and contain cilia. These photosensitive cells drive the phototransduction cascade in response to light, which in

turn induces visual information processing. IPA pointed towards glutamate receptor signaling as the most significant pathway. Glutamate is released by photoreceptors and determines conductance of retinal signaling to the ON and OFF bipolar cells<sup>81</sup>. Our functional gene look ups provide evidence that rod (*CLU*) as well as cone (*GNB3*) bipolar cells play a role. Taken together, these findings strongly suggest that light response and light processing in the retina are initiating factors leading to refractive error.

The genetic association with light-dependent pathways may also link to the well-established protective effect of outdoor exposure on myopia. We found suggestive evidence for a genetic association with *DRD1*. The dopaminergic pathway has been studied extensively in animal models for its role in controlling eye growth in response to light<sup>69,71,82–91</sup>. *DRD1* was found to be a mediator in this process, as bright light increased DRD1 activity in the bipolar ON-pathway, and diminished form-deprivation myopia in mice. Blockage of DRD1 reversed this inhibitory effect<sup>92</sup>. We did not find evidence for direct involvement of other genes in the dopamine pathway, but *GNB3* may be an indirect modifier as it is a downstream signaling molecule of dopamine and has been shown to influence availability of the dopamine transporter DAT<sup>93</sup>. Although a promising target for therapy, further evidence of *DRD1* in human myopiagenesis is warranted.

Novel pathways elicited by the newly identified genes are anterior segment morphology (*TCF7L2*, *VIPR2*, *MAF*) and angiogenesis (*FLT1*). In addition, the high number of variants residing near small RNA genes suggests that post-transcriptional regulation is an important mechanism, as these RNAs are known to play a distinct and central regulatory role in cells<sup>94</sup>. These findings will serve as leads for future studies performing detailed mapping of cellular networks, and functional studies on genes implicated in ocular phenotypes, harboring protein-altering variants, and proven drug targets.

Our evaluation of shared genetics between refractive error and other disease-relevant phenotypes highlighted overlap with anthropometric traits such as height, obesity, and body mass index. This could give valuable additional clues as to the phenotypic outcomes of perturbations of some of the networks identified.

Our genetic observations add credence to the current notion that refractive errors are caused by a retina-to-sclera signaling cascade that induces scleral remodeling in response to light stimuli. The concept of this cascade originates from various animal models showing that form deprivation, retinal defocus and contrast, ambient light, and wavelength can influence eye growth in young animals<sup>95–97</sup>. Cell-specific moieties in this putative signaling cascade in humans were largely unknown, although animal models implicated GABA, dopamine, all-trans-retinoic acid and TGF- $\beta$ <sup>69,91,98,99</sup>. Our study provides a large number of new molecular candidates for this cascade, and clearly shows that a wide range of neuronal cell types in the retina, the RPE, the vascular endothelium, as well as components of the extracellular matrix are implicated. The many interprotein relationships exemplify the complexity of eye growth, and provide a challenge to develop strategies to prevent pathological eye elongation.

In conclusion, by using a cross-ancestry design in the largest study population on common refractive errors to date, we uncovered numerous novel loci and pathways involved in eye growth. Our multi-disciplinary approach incorporating GWAS data with *in silico* analyses and expression experiments provides an example for the design of future genetic studies for complex traits. Additional genetic insights into refractive errors will be gained by increasing sample size and genotyping depth, by performing family studies to identify rare alleles of large effects, and by evaluating population extremes. Our list of plausible genes and pathways provide a plethora of data for future studies focusing on gene-environment interaction, and on translation of GWAS findings into starting points for therapy.

## ONLINE METHODS

### Ethics Statement

All human research was approved by the relevant institutional review boards and conducted according to the Declaration of Helsinki. All CREAM participants provided written informed consent; all 23andMe applicants provided informed consent online, and answered surveys according to 23andMe's human subjects protocol, which was reviewed and approved by Ethical & Independent Review Services, an AAHRPP-accredited institutional review board. The UK Biobank received ethical approval from the National Health Service National Research Ethics Service (reference 11/NW/0382).

### Study data

The study populations were participants of the Consortium for Refractive Error and Myopia (CREAM) comprising of 41,793 individuals with European ancestry from 26 cohorts (CREAM-EUR) and 11,935 individuals with Asian ancestry from 8 studies (CREAM-ASN); and customers of the 23andMe genetic testing company who gave informed consent for inclusion in research studies consisting of 104,293 individuals (2 cohorts of individuals with European ancestry,  $N=12,128$  and  $N=92,165$ , respectively). All participants included in this analysis from CREAM and 23andMe were aged 25 years or older. Participants with conditions that could alter refraction, such as cataract surgery, laser refractive procedures, retinal detachment surgery, keratoconus as well as ocular or systemic syndromes were excluded from the analyses. Recruitment and ascertainment strategies varied per study (Supplementary Table 1a,b, and Supplementary Note 4). Refractive error represented by measurements of refraction and analyzed as spherical equivalent (SphE =spherical refractive error + 1/2 cylinder refractive error) was the outcome variable for CREAM; myopic refractive error represented by self-reported age of diagnosis of myopia (AODM) for 23andMe<sup>27</sup>.

### Genotype calling and imputation

Samples were genotyped on different platforms and study specific quality control measures of the genotyped variants were implemented before association analysis (Supplementary Table 1b). Genotypes were imputed using the appropriate ancestry-matched reference panel for all cohorts from the 1000 Genomes Project (Phase I version 3, March 2012 release) with either minimac<sup>100</sup> or IMPUTE<sup>101</sup>. The metrics for pre-imputation quality control varied amongst studies, but genotype call rate thresholds were set at high level ( $\geq 0.95$  for both

CREAM and 23andMe). These metrics were similar to our previous GWAS analyses<sup>26,27</sup>; details per cohort can be found in Supplementary Table 1b.

### GWAS per study

For each CREAM cohort, a single marker analysis for the SphE (in diopters) phenotype was carried out using linear regression adjusting for age, sex and up to the first five principal components. All non-family-based cohorts removed one of each pair of relatives (after detection using either GCTA or IBS/IBD analysis). In family-based cohorts, a score test-based association was used to adjust for within-family relatedness<sup>102</sup>. For the 23andMe participants, Cox proportional hazards analysis testing AODM as the dependent variable were performed as previously described<sup>27</sup>, with  $P$  calculated using a likelihood ratio test for the single marker genotype term. We used an additive SNP allelic effect model for all analyses.

### Centralized quality control per study

After individual GWAS, all studies underwent a second round of quality control (QC). Quantile-quantile, effect allele frequency,  $P$ - $Z$  test, standard error – sample size, and genomic control inflation factor plots were generated for each individual cohort using EasyQC<sup>103</sup> (Supplementary Figure 2.1 (Supplementary Figure 2.1.1 and 2.1.2), 2.2 (Supplementary Figure 2.2.1 – 2.2.5), 2.3 (Supplementary Figure 2.3.1 and 2.3.2)). All analytical issues discovered during this QC step were resolved per individual cohort.

### GWAS meta-analyses

The GWAS meta-analyses were performed in three stages (Supplementary Figure 1). In *Stage 1*, European (CREAM-EUR,  $N=44,192$ ) and Asian (CREAM-ASN,  $N=11,935$ ) participants from the CREAM cohort were meta-analysed separately. Subsequently, all CREAM cohorts (CREAM-ALL) were meta-analysed. Variants with MAF  $< 1\%$  or imputation quality score  $< 0.3$  (info metric of IMPUTE) or  $R_{sq} < 0.3$  (minimac) were excluded. A fixed effects inverse variance-weighted meta-analysis was performed using METAL<sup>104</sup>. 1,063 variants clustering in 24 loci (Supplementary Table 2) were genome-wide significant ( $P=5.0 \times 10^{-8}$ ). All 37 loci that were previously found by CREAM and 23andMe using genotype data imputed to the HapMap II reference panel were replicated ( $p_{Bonferroni} 1.85 \times 10^{-3}$ ), and 36 of the 37 were genome-wide significant (Supplementary Table 2)<sup>26,27</sup>. In *Stage 2*, a meta-analysis of the two 23andMe cohorts ( $N_{23andMe\_v2}=12,128$ ;  $N_{23andMe\_v3}=92,165$ ) was performed, using similar filtering but a lower MAF threshold ( $< 0.5\%$ ). A total of 5,205 genome-wide significant variants clustered in 112 loci (Supplementary Table 2).

In *Stage 3*, CREAM-ALL and 23andMe samples were combined using a fixed effects meta-analysis based on  $P$  value and direction of effect. In all stages, each genetic variant had to be represented by at least half of the entire study population and at least represented by 13 cohorts in CREAM and one cohort in 23andMe. For SNPs with high heterogeneity (at  $P < 0.05$ ), we also performed a random effects meta-analysis using METASOFT<sup>50</sup>. We chose a different weighting scheme due to the differences in effect size scaling; 23andMe used a less accurate phenotype variable (AODM); i.e. the *effective* sample size of the 23andMe was

approximately equivalent to the *effective* sample size of CREAM-ALL (Figure 2b), thus weighting by  $(1/\sqrt{n_{\text{effective}}})$  yielded a final weighting ratio of 1:1<sup>105</sup>. Genome-wide statistical significance was defined at  $P < 5.0 \times 10^{-8}$ <sup>106</sup>.

All three meta-analysis stages were performed under genomic control. Study specific and meta-analysis lambda ( $\lambda$ ) estimates are shown in Supplementary Figure 6; to check for confounding biases (e.g. cryptic relatedness and population stratification), LD score intercepts from LD score regressions per ancestry were constructed (Supplementary Figure 7)<sup>30</sup>. To check the robustness of signals, we performed a conventional random effects models using METASOFT, fixed effects models weighted on sample size and on weights estimated from standard error per allele tested using METAL (Supplementary Table 2 and Supplementary Table 3).

Manhattan (modified version of package ‘qqman’), regional, box, and forest plots were made using R version 3.2.3 and LocusZoom<sup>107</sup>. An overview of the Hardy Weinberg  $P$  of all index variants per cohort can be found in Supplementary Table 4. The comparison between refractive error and age-of-onset was performed using the LDSC program<sup>30</sup>.

### Population stratification and heritability calculations

Each study assessed the degree of genetic admixture and stratification in their study participants through the use of principal components. Homogeneity of participants was assured by removal of all individuals whose ancestry did not match the prevailing ancestral group. We used genomic inflation factors to control for admixture and stratification, and performed genomic-controlled meta-analysis to account for the effects of any residual heterogeneity. To further distinguish between inflation from a true polygenic signal and population stratification, we examined the relationship between test statistics and linkage disequilibrium (LD) with LDSC. CREAM-EUR, CREAM-ASN and 23andMe were evaluated separately; variants not present in HapMap3 and MAF < 1% were excluded. SNP heritability estimates were calculated using LDSC for the same set of genetic variants.

### Locus definition and annotation

All study effect size estimates were oriented to the positive strand of the NCBI Build 37 reference sequence of the human genome. The index variant of a locus was defined as the variant with the lowest  $P$  in a region spanning a 100 kb window of the most outer genome-wide significant variant of that same region. We annotated all index variants using the web-based version of ANNOVAR<sup>108</sup> based on UCSC Known Gene Database<sup>35</sup>. For variants within the coding sequence or 5′ or 3′ untranslated regions of a gene, that gene was assigned to the index variant (note that this led to more than 1 gene being assigned to variants located within the transcription units of multiple, overlapping genes). For variants in intergenic regions, the nearest 5′ gene and the nearest 3′ gene were assigned to the variant. Index variants were annotated to functional RNA elements when described as such in the UCSC Known Gene Database. We used conservation (PhyloP<sup>109</sup>) and prediction tools (SIFT<sup>39</sup>, Mutation Taster<sup>110</sup>, align GVG<sup>40,41</sup>, PolyPhen-2<sup>36</sup>) to predict the pathogenicity of protein-altering exonic variants.

### Conditional signal analysis

We performed conditional analysis to identify additional independent signals nearby the index variant at each locus, using GCTA-COJO<sup>32</sup>. We transformed the Z-scores of the summary statistics to beta's using the following formula:

Standard Error =  $\sqrt{1/2 * N * MAF(1 - MAF)}$ . We performed the GCTA-COJO analysis<sup>32</sup>, utilizing summary-level statistics from the meta-analysis on all cohorts. Linkage disequilibrium (LD) between variants was estimated from the Rotterdam Study I-III.

### Replication in UK Biobank

The UK Biobank Eye & Vision (UKEV) Consortium performed a GWAS of refractive error in 95,505 participants of European ancestry aged 37-73 year with no history of eye disorders<sup>33</sup>. Refractive error was measured using an autorefractor; SphE was calculated per eye and averaged between the two eyes. To account for relatedness a mixed model analysis with BOLT-LMM was used<sup>111</sup>, including age, gender, genotyping array, and the first 10 principal components as covariates. Analysis was restricted to markers present on the HRC reference panel<sup>112</sup>. We performed lookups for all independent genetic variants identified in our Stage 3 meta-analysis and conditional analysis. For 16 variants not present in UKEV, we performed lookups for a surrogate variant in high LD ( $r^2 > 0.8$ ). When more than one potential surrogate variant was available, the variant in strongest LD with the index variant was selected. Six variants were not available for replication: one variant (rs188159083) was not present on the array nor was a surrogate available in UKEV and five variants showed evidence of departure from HWE (HWE exact test  $P < 3.0 \times 10^{-4}$ ).

### Post-GWAS analyses

We performed two gene-based tests to identify additional significant genes not found in the single variant analysis. First, we applied the gene-based test implemented in fastBAT<sup>42</sup> to the per-variant summary statistics of the meta-analysis of all European cohorts (23andMe and CREAM-EUR). We used the default parameters (all variants in or within 50kb of a gene) and focused on variants with a gene-based  $P < 2 \times 10^{-6}$  (Bonferroni correction based on 25,000 genes) and the per-variant  $P > 5 \times 10^{-8}$ . Secondly, we applied another gene-based test in EUGENE<sup>43</sup> which only includes variants which are eQTLs (GTEx, blood<sup>113</sup>). EUGENE tests an hypothesis predicated on eQTLs as key drivers of the association signal. eQTLs within 50kb of a gene were included in the test. Genes with EUGENE  $P < 2 \times 10^{-6}$  (and not found in the single variant analysis) were considered to be significant. Finally, we used functional annotation information from genome-wide significant loci to reweight results using fgwas (version 0.3.64<sup>44</sup>). Fgwas incorporates functional annotation (e.g. DNase I-hypersensitive sites in various tissues and 3'UTR regions) to reweight data from GWAS, and uses a Bayesian model to calculate a posterior probability of association. This approach is able to identify risk loci that otherwise might not reach the genome-wide significance threshold in standard GWAS. Details about this approach can be found in Supplementary Note 5.

## Refractive errors and myopia risk prediction

To assess the risk of the entire range of refractive errors, we computed polygenic risk scores (PGRS) for the population-based Rotterdam Studies (RS) I, RS-II and RS-III using the  $P$  and  $Z$  scores from a meta-analysis on CREAM-ALL and 23andMe, excluding the RS I-III cohorts. Only variants with high imputation quality (IMPUTE info score  $> 0.5$  or minimac  $R_{sq} > 0.8$ ) and  $MAF > 1\%$  were considered.  $P$ -based clumping was performed with PLINK<sup>114</sup>, using an  $r^2$  threshold of 0.2 and a physical distance threshold of 500 kb, excluding the MHC region. This resulted in a total of 243,938 variants. For each individual in RS-I, RS-II and RS-III ( $N = 10,792$ ), PGRS were calculated using the `-score` command in PLINK across strata of  $P$  thresholds:  $5.0 \times 10^{-8}$ ,  $5.0 \times 10^{-7}$ ,  $5.0 \times 10^{-6}$ ,  $5.0 \times 10^{-5}$ ,  $5.0 \times 10^{-4}$ , 0.005, 0.01, 0.05, 0.1, 0.5, 0.8 and 1.0. The proportion of variance explained by each PGRS model was calculated as the difference in the  $R^2$  between two regression models; one where SphE was regressed on age, sex, the first five principal components, and the other also including the PGRS as an additional covariate. Subsequently, AUCs were calculated for myopia (SphE  $\leq -3$  SD) versus hyperopia (SphE  $\geq +3$  SD).

## Genetic correlation between ancestries

We used Popcorn<sup>48</sup> to investigate ancestry-related differences in the genetic architecture of refractive error and myopia. Popcorn takes summary GWAS statistics from two populations and LD information from ancestry-matched reference panels, and computes genetic correlations by implementing a weighted likelihood function that accounts for the inflation of  $Z$  scores due to LD. Pairwise analyses were carried out using the GWAS summary statistics from 23andMe ( $N = 104,292$ ), CREAM-EUR ( $N = 44,192$ ) and CREAM-EAS ( $N = 9,826$ ) meta-analyses. Only SNPs with  $MAF \geq 5\%$  were included, resulting in a final set of 3,625,602 SNPs for analyses involving 23andMe and 3,642,928 SNPs for the CREAM-EUR versus CREAM-EAS analysis. Reference panels were constructed using genotype data from 503 European and 504 East Asian individuals sequenced as part of the 1000 Genomes Project (release 2013-05-02 downloaded from: <ftp.1000genomes.ebi.ac.uk>). The reference panel VCF files were filtered using PLINK<sup>114</sup> to remove indels, strand-ambiguous variants, variants without an “rs” id prefix, and variants located in the MHC region on chromosome 6 (chr6:25,000,000-33,500,000; Build 37).

## Analysis between phenotypes

To evaluate consistency of genotypic effects across studies that employed different phenotype definitions, we compared effect sizes from GWAS studies of either SphE or AODM in Europeans, i.e. CREAM-EUR ( $N = 44,192$ ) or 23andMe ( $N = 104,293$ ) respectively. Marker-wise additive genetic effect sizes (in units diopters per copy of the risk allele) for SphE were compared against those (in units  $\log(HR)$  per copy of the risk allele) for AODM. Data was visualised using R. Genetic correlation between the two phenotypes SphE and AODM was calculated using LD score regression. This analysis included all common SNPs ( $MAF > 0.01$ ) present in HapMap3.

## Evidence for functional involvement

In order to rank genes according to biological plausibility, we scored annotated genes based on our own findings and published reports for a potential functional role in refractive error. Points were assigned for each gene on the basis of 10 categories (details on the methodology per category are provided in Supplementary Note 4): internal replication of index genetic variants in the individual cohort GWAS analyses through Bonferroni corrections (CREAM-ASN, CREAM-EUR and 23andMe;  $p_{\text{Bonferroni}} 1.19 \times 10^{-4}$ ), evidence for eQTL using the FUMA<sup>32</sup> and extensive look-ups in GtEx, evidence of expression in the eye in developmental and adult ocular tissues, presence of an eye phenotype in knock-out mice (MGI and IMPC database), presence of an eye phenotype in humans (OMIM; see URLs, DisGeNET<sup>115</sup>), location in a functional region of a gene (wANNOVAR; see URLs), presence of the gene in a significant enriched functional pathway with false discovery rate  $< 0.05$  (DEPICT<sup>49</sup>), presence of the gene in the gene priority analysis of DEPICT with false discovery rate  $< 0.05$  and the presence of the gene in the canonical pathway analysis of Ingenuity Pathway Analysis (IPA; See URLs). Furthermore, we performed a systematic search for each gene to assess its potential as a drug target (SuperTarget<sup>116</sup>, STITCH<sup>117</sup>, DrugBank<sup>118</sup>, PharmaGkb<sup>119</sup>). All information derived from this study and literature were used to annotate genes to retinal cell types.

## Genetic pleiotropy

To investigate overlap of genes with other common traits, we performed a look-up in the GWAS catalog using FUMA. Multiple testing correction (i.e. Benjamini-Hochberg) was performed. Traits were significantly associated when adjusted  $P \leq 0.05$  and the number of genes that overlap with the GWAS catalog gene sets was  $\geq 2$ .

**Data availability statement**—The summary statistics of the Stage 3 meta-analysis are included in Supplementary Data 3 of this published article. In order to protect the privacy of the participants in our cohorts, further summary statistics of Stage 1 (CREAM) and Stage 2 (23andMe) will be available upon request. Please contact [c.c.w.klaver@erasmusmc.nl](mailto:c.c.w.klaver@erasmusmc.nl) (CREAM) and/or [apply.research@23andMe.com](mailto:apply.research@23andMe.com) (23andMe) for more information and to access the data.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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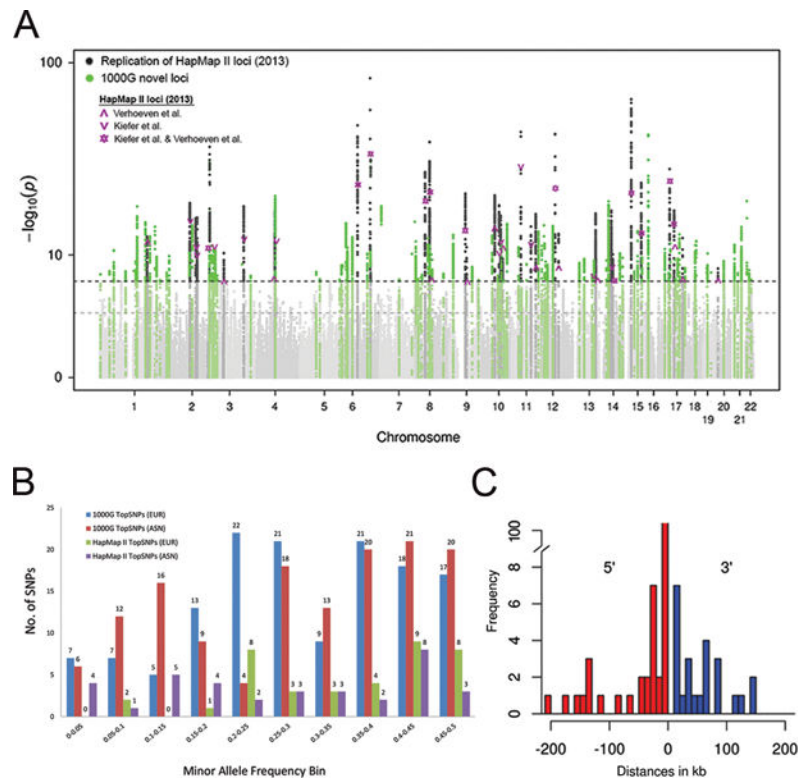
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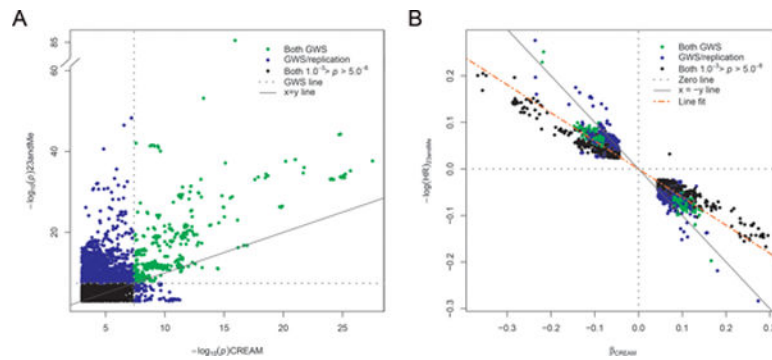
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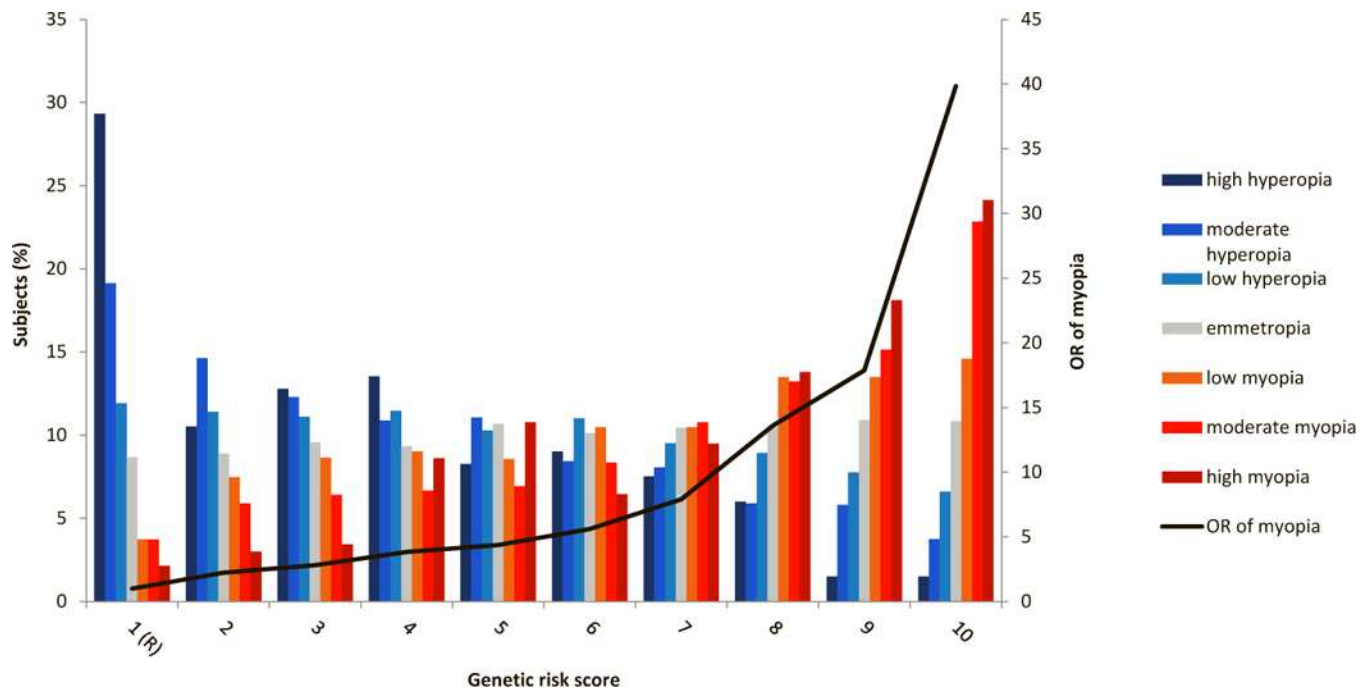


**Figure 1. GWAS meta-analysis identifies 140 loci for refractive error (Stage 3)**

(a) We conducted a meta-analysis of genome-wide single-variant analyses for >10 million variants in 160,420 participants of CREAM and 23andMe (Stage 3). Shown is the Manhattan plot depicting  $P$  for association, highlighting new ( $P < 5 \times 10^{-8}$  for the first time; green) and known (dark grey) refractive error loci previously found using HapMap II imputations from Kiefer et al.<sup>27</sup> and Verhoeven et al.<sup>26</sup> (Table 1). The horizontal lines indicate suggestive significance ( $P = 1 \times 10^{-5}$ ) or genome-wide significance ( $P = 5 \times 10^{-8}$ ). (b) We compared the minor allele frequencies of the 140 discovered index variants based on 1000G (blue: Europeans; red: Asians) to the minor allele frequencies of the previously found genetic variants based on HapMap II (green: Europeans; purple: Asians). Observed are an increase in genetic variants found across all minor allele frequency bins increase, including the lower minor allele frequency bins. (c) We annotated the 167 loci to genes using wANNOVAR. Shown are the distances between index variants from the nearest gene and its gene on the 5' and/or 3' site. The majority of index variants (84%) were at a distance of less than 50 kb up- or downstream from the annotated gene.

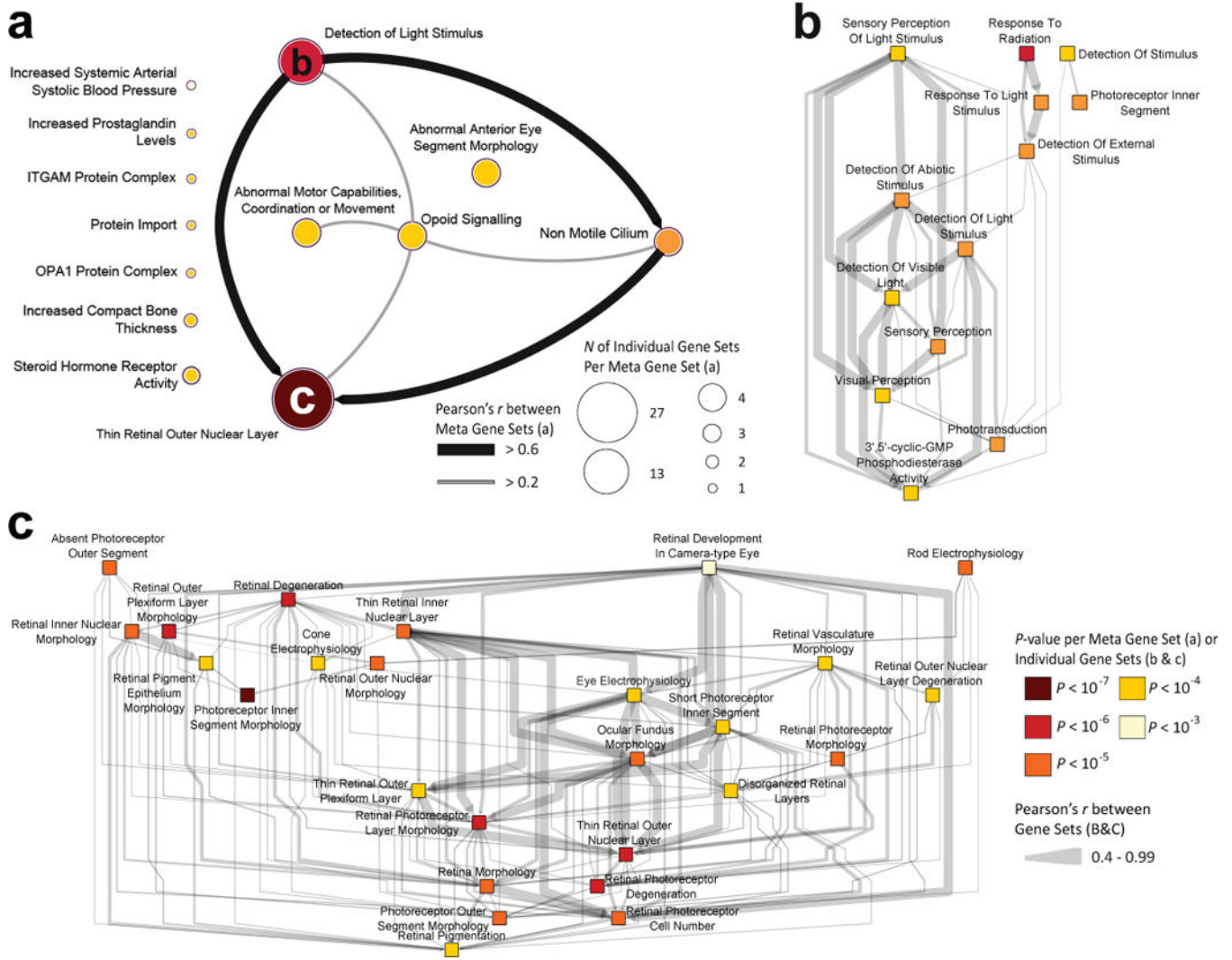


**Figure 2. Correlation of statistical significance and effect size of SNPs based on spherical equivalent (SphE) in diopters and age of diagnosis of myopia (AODM) in years**  
**(a)** *P* comparison of all genetic variants with  $P < 1.0 \times 10^{-3}$  ( $n=7249$ ) between CREAM meta-analysis (Stage 1) and 23andMe (Stage 2) meta-analysis. Shown is the overlap (red) and the difference (green) in *P* signals per cohort for genetic variants. Green genetic variants are only genome wide significant in either CREAM or 23andMe. Blue: genetic variants with *P* between  $5.0 \times 10^{-8}$  and  $1.0 \times 10^{-3}$  in both CREAM and 23andMe. **(b)** Comparison of effects (SphE and logHR of AODM in years;  $P < 1.0 \times 10^{-3}$ ;  $n=7249$ ) between CREAM and 23andMe. Same color code was applied as in (a). The effects were concordant in their direction of effect on refractive error. We performed a simple linear regression between the effects of CREAM and 23andMe; the regression slope is -0.15 diopters per logHR of AODM in years.



**Figure 3. Risk of refractive error per decile of polygenic risk score (Rotterdam Study I-III, N=10,792)**

Distribution of refractive error in subjects from Rotterdam Study I-III ( $N=10,792$ ) as a function of the optimal polygenic risk score (including 7,303 variants at  $P \leq 0.005$  explaining 7.8% of the variance of SphE; Supplementary Table 9). Mean OR of myopia (black line) was calculated per polygenic risk score category using the lowest category as a reference. High myopia (SphE  $\leq -6$  D), moderate myopia (SphE  $> -6$  D &  $\leq -3$  D), low myopia (SphE  $> -3$  D &  $< -1.5$  D), emmetropia (SphE  $\geq -1.5$  D and  $\leq 1.5$  D), low hyperopia (SphE  $> 1.5$  D &  $< 3$  D), moderate hyperopia (SphE  $\geq 3$  D &  $< 6$  D), high hyperopia (SphE  $\geq 6$  D).



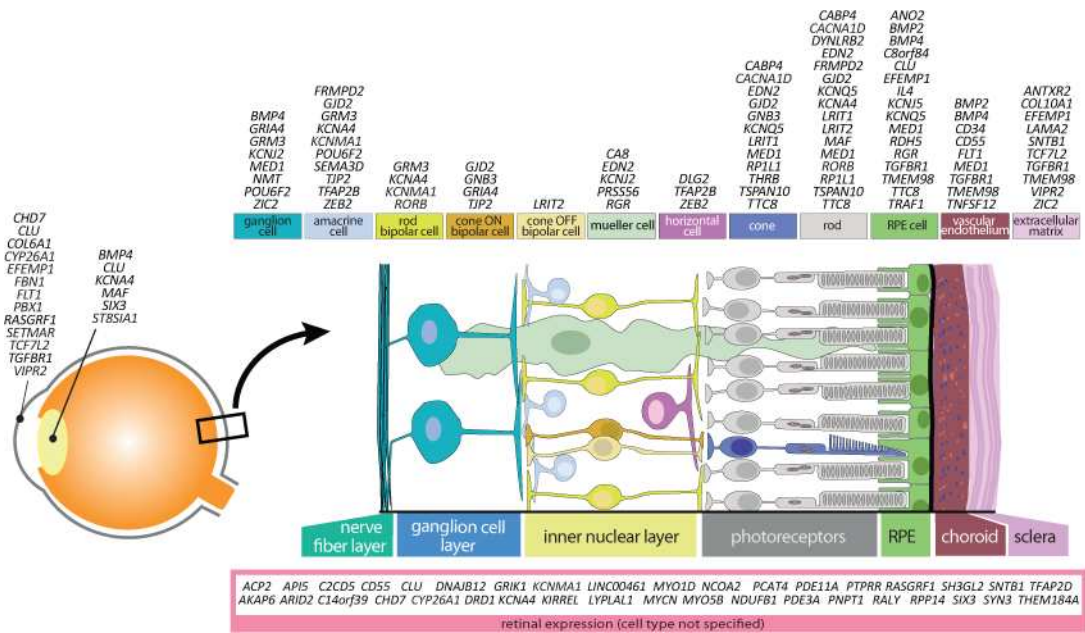
**Figure 4. Visualization of the DEPICT gene-set enrichment analysis based on loci associated with refractive error and the correlation between the (meta)gene sets**  
**(a)** Shown are the 66 significantly enriched reconstituted gene sets clustered into thirteen meta gene sets based on the gene set enrichment analysis of DEPICT (pairwise Pearson correlations;  $P < 0.05$ ). All genetic variants with a  $P < 1 \times 10^{-5}$  in the GWAS meta-analysis of stage 3 ( $n=21,073$ ) and an  $FDR < 0.05$  were considered. **(b)** Visualization of the interconnectivity between gene sets ( $n=13$ ; pairwise Pearson correlations;  $P < 0.05$ ) of the meta gene set ‘Detection of Light Stimulus’ (GO:0009583). **(c)** Visualization of the interconnectivity between gene sets ( $n=27$ ; pairwise Pearson correlations;  $P < 0.05$ ) of the largest meta gene set ‘Thin Retinal Outer Nuclear Layer’ (MP:0008515). In all panels, (meta)gene sets are represented by nodes colored according to statistical significance, and similarities between them are indicated by edges scaled according to their correlation; Pearson’s  $r \geq 0.2$  are shown in panel **(a)** and Pearson’s  $r \geq 0.4$  are shown in panel **(b,c)**.

Locus	Locus Name	Gene Priority Score	Annotation*				Expression			Biology		Pathways			Known Drug Target
			Σ	1	1	1	1	1	1	1	1	1	1	1	
GNB3	GNB3	8													
RDH5	BLOC1S1-RDH5, RDH5	7													
CYP26A1	CYP26A1, MYOF	7													
EFEMP1	EFEMP1, PNPT1	7													
GRIA4	GRIA4	7													
RGR	RGR	7													
RORB	RORB	7													
TJP2	TJP2	6													
PRSS56	PRSS56	6													
CABP4	CABP4	6													
FBN1	FBN1	6													
GJD2	GJD2, GOLGA8B	6													
KCNJ2	BC039327, KCNJ2	6													
KCNMA1	KCNMA1	6													
MAF	DYNLRB2, MAF	6													
RCBTB1	RCBTB1	6													
ST18	FAM150A, ST18	6													
TCF7L2	TCF7L2	6													
ZEB2	ZEB2	6													

**Figure 5. Genes ranked according to biological and statistical evidence**

Genes were ranked (orange) based on 10 equal categories which can be divided in four categories: internal replication of genetic variant in more than two cohorts (purple; CREAM-EUR, CREAM-ASN and/or 23andMe), annotation (light blue; genetic variant harboring an exonic protein altering variant or non-protein altering variant, genetic variant residing in a 5' or 3' UTR region of a gene or transcribing an RNA structure), expression (yellow; eQTL, expression in adult human ocular tissue, expression in developing ocular tissue), biology (dark yellow; ocular phenotype in mice, ocular phenotype in humans), pathways (green; DEPICT gene-set enrichment, DEPICT gene prioritization analysis and canonical pathway analysis of IPA). We assessed genes harboring drug targets (salmon red), but did not assign a scoring point to this category.

\*Only one point can be assigned in the category 'ANNOTATION', even though it has four columns (i.e. a genetic variant is located in only 1 of these four categories).



**Figure 6. Schematic representation of the human eye, retinal cell types, and functional sites of associated genes**

We assessed gene expression sites and/or functional target cells in the eye for all genes using our expression data and literature and data present in the public domain. The genes appear to be distributed across virtually all cell types in the neurosensory retina, in the RPE, vascular endothelium and extracellular matrix; i.e., the route of the myopic retina-to-sclera signalling cascade.

Table 1

Results of the meta-analysis of CREAM and 23andMe for the previously-identified loci and a subset of the newly-identified loci, and replication in UK Biobank

**a Replication of the HapMap II index variants for refractive error per locus in the Stage 3 meta-analysis**

SNP	Chr	Position	Nearest Loci And Gene(s)	Effect Allele	Other Allele	EAF EUR	EAF ASN	Z-score	Direction	P value	HetISq	HetPVal	Sample Size (N)	HapMap II Discovery (2013)	Category	P value Replication
rs10500355	16	7459347	RBFOX1	A	T	0.354	0.133	-13.73	--	6.49E-43	9.1	2.93E-07	160,139	Kiefer et al. & Verhoeven et al.	I	2.50E-48
rs11145465	9	71766593	TJP2	A	C	0.212	NA	-9.55	--	1.35E-21	46.3	0.1722	153,174	Kiefer et al. & Verhoeven et al.	I	1.00E-10
rs11178469	12	71275137	PTPRR	T	C	0.752	0.638	-7.40	--	1.33E-13	0	0.6989	160,139	Verhoeven et al.	II (CREAM)	2.60E-04
rs11602008	11	40149305	LRRC4C	A	T	0.822	0.749	13.98	++	2.12E-44	22.5	1.56E-10	157,505	Kiefer et al.	II (23andMe)	2.90E-47
rs12193446	6	129820038	BC035400, LAMA2	A	G	0.906	NA	-19.43	--	4.21E-84	16.8	5.72E-15	150,269	Kiefer et al. & Verhoeven et al.	I	4.60E-106
rs1550094	2	233385396	CHRNA3, PRSS56	A	G	0.701	0.705	12.74	++	3.64E-37	26.3	0.003	159,422	Kiefer et al. & Verhoeven et al., Kiefer et al.	I	4.10E-59
rs1649068	10	60304864	BICC1	A	C	0.475	0.504	-9.44	--	3.77E-21	0	0.712	160,144	Verhoeven et al.	I	7.50E-11
rs17382981	10	94953258	CYP26A1, MYOF	T	C	0.417	0.190	-6.31	--	2.72E-10	67.9	0.077	155,332	Verhoeven et al.	II (CREAM)	4.10E-07
rs17428076	2	172851936	HAT1, METAP1D	C	G	0.768	0.854	-8.18	--	2.77E-16	0	0.003	160,151	Kiefer et al.	II (23andMe)	7.50E-08
rs1858001	1	207488004	C4BPA, CD55	C	G	0.676	0.415	7.28	++	3.45E-13	59.6	0.020	160,149	Verhoeven et al.	II (CREAM)	6.70E-20
rs1954761	11	105596885	GRIA4	T	C	0.371	0.377	-8.40	--	4.57E-17	0	0.911	160,122	Verhoeven et al.	I	1.20E-16
rs2155413	11	84634790	DLG2	A	C	0.482	0.655	-7.76	--	8.85E-15	0	2.99E-04	159,504	Kiefer et al.	II (23andMe)	1.10E-17
rs235770	20	6761765	BMP2	T	C	0.372	0.388	-5.93	--	3.11E-09	0	0.547	157,521	Verhoeven et al.	II (23andMe)	4.80E-11
rs2573081	2	178828507	PDE11A	C	G	0.524	0.538	8.21	++	2.18E-16	47.6	0.167	160,126	Kiefer et al.	II (23andMe)	1.60E-29
rs2753462	14	60850703	JB175233, C14orf39	C	G	0.296	0.568	-6.49	--	8.37E-11	73.9	0.050	157,352	Verhoeven et al.	II (CREAM)	2.00E-15
rs2855530	14	54421917	BMP4	C	G	0.507	0.474	-8.58	--	9.87E-18	41.7	0.190	160,092	Kiefer et al.	I	4.80E-22
rs2908972	17	11407259	SHISA6	A	T	0.415	0.484	-11.13	--	9.46E-29	23	0.254	160,123	Kiefer et al. & Verhoeven et al.	I	6.10E-29
rs3138141	12	56115778	BLOC1S1-RDH5, RDH5	A	C	0.214	0.147	13.80	++	2.46E-43	3.2	5.05E-07	157,531	Kiefer et al. & Verhoeven et al.	I	2.30E-56
rs4687586	3	53837971	CACNA1D	C	G	0.691	NA	-6.55	--	5.86E-11	0	0.605	150,217	Verhoeven et al.	III	1.60E-08
rs4793501	17	68718734	KCNJ2, BC039327	T	C	0.575	0.444	-7.21	--	5.53E-13	0	0.592	160,150	Verhoeven et al.	II (CREAM)	3.70E-12
rs524952	15	35005886	GOLGA8B, GJD2	A	T	0.475	0.507	-17.08	--	2.28E-65	67.2	0.015	160,150	Kiefer et al. & Verhoeven et al.	I	1.60E-103
rs56075542	2	146882415	BC040861, PABPC1P2	T	G	0.552	0.472	-8.99	--	2.39E-19	13.9	0.001	159,478	Kiefer et al.	II (23andMe)	1.30E-18
rs62070229	17	31227593	MYO1D, TMEM98	A	G	0.807	0.874	8.58	++	9.64E-18	0	0.416	156,570	Verhoeven et al.	I	1.30E-18
rs6495367	15	79375347	RASGRF1	A	G	0.408	0.399	-10.20	--	1.95E-24	0	0.667	160,144	Kiefer et al. & Verhoeven et al.	I	7.20E-37
rs7042950	9	77149837	RORB	A	G	0.732	0.392	6.80	++	1.07E-11	0	0.912	160,153	Verhoeven et al.	III	2.90E-18
rs72621438	8	60178580	SNORA51, CA8	C	G	0.642	0.609	-13.14	--	2.03E-39	38.4	0.006	160,128	Kiefer et al. & Verhoeven et al.	I	1.80E-49
rs745480	10	85986554	LRIT2, LRIT1	C	G	0.511	0.418	8.31	++	9.26E-17	67.3	0.081	159,504	Kiefer et al.	II (23andMe)	8.20E-18

**a Replication of the HapMap II index variants for refractive error per locus in the Stage 3 meta-analysis**

SNP	Chr	Position	Nearest Loci And Gene(s)	Effect Allele	Other Allele	EAF EUR	EAF ASN	Z-score	Direction	P value	HetISq	HetPVal	Sample Size (N)	HapMap II Discovery (2013)	Category	P value Replication
rs7624084	3	141093285	ZBTB38	T	C	0.568	0.633	-8.81	--	1.24E-18	18.5	0.018	160,151	Kiefer et al.	II (23andMe)	6.50E-17
rs7662551	4	80537638	LOC100506035, PCAT4	A	G	0.723	0.558	8.53	++	1.47E-17	19.4	0.265	160,147	Verhoeven et al.	I	6.00E-12
rs7692381	4	81903049	C4orf22, BMP3	A	G	0.763	0.630	9.40	++	5.55E-21	0	0.013	160,134	Kiefer et al.	I	7.50E-13
rs7744813	6	73643289	KCNQ5	A	C	0.591	0.602	-14.56	--	5.43E-48	35	0.001	160,091	Kiefer et al. & Verhoeven et al.	I	1.00E-75
rs7829127	8	40726394	ZMAT4	A	G	0.792	0.897	-10.91	--	1.02E-27	15.9	2.77E-04	160,132	Kiefer et al. & Verhoeven et al.	II (23andMe)	3.10E-22
rs7895108	10	79061458	KCNMA1	T	G	0.351	0.118	-8.87	--	7.56E-19	32.8	0.021	160,140	Kiefer et al.	II (23andMe)	1.10E-27
rs79266634	16	7309047	RBFOX1	C	G	0.093	0.115	-5.93	--	3.00E-09	0	0.561	156,268	Kiefer et al. & Verhoeven et al.	III	1.50E-08
rs837323	13	101175664	PCCA	T	C	0.512	0.762	6.32	++	2.65E-10	35.6	0.213	160,142	Verhoeven et al.	II (23andMe)	5.30E-16
rs9517964	13	100717833	ZIC2,PCCA	T	C	0.589	0.786	8.42	++	3.68E-17	0	0.020	160,121	Kiefer et al.	II (23andMe)	3.40E-20

**b Subset of the new loci harboring the smallest p-values for refractive error in the Stage 3 meta-analysis**

SNP	Chr	Position	Nearest Loci And Gene(s)	Effect Allele	Other Allele	EAF EUR	EAF ASN	Z-score	Direction	P value	HetISq	HetPVal	Sample Size (N)	Category	P value Replication
rs36024104	14	42294993	LRFN5	A	G	0.823	NA	9.09	++	9.86E-20	15.9	0.01414	152,585	II (23andMe)	2.20E-12
rs7456039	7	6901710	CCZ1B, LOC100131257	C	G	0.183	NA	8.82	++	1.18E-18	42.1	3.79E-12	121,337	II (23andMe)	6.50E-01
rs1556867	1	164213686	5S_rRNA, PBX1	T	C	0.264	0.494	-8.81	--	1.29E-18	71.1	0.06266	160,155	II (23andMe)	4.20E-17
rs12667032	7	154406581	DPP6	A	G	0.152	0.317	7.99	++	1.31E-15	82.3	1.02E-11	130,790	II (23andMe)	2.10E-01
rs2225986	1	200311910	LINC00862	A	T	0.381	0.169	-7.96	--	1.68E-15	40.2	0.196	160,152	II (23andMe)	7.50E-17
rs1207782	6	22059967	LINC00340	T	C	0.577	0.265	-7.92	--	2.47E-15	0	0.8946	160,149	I	4.90E-13
rs72826094	10	114801488	TCF7L2	A	T	0.799	0.838	7.88	++	3.20E-15	64.5	0.09323	156,825	II (23andMe)	4.90E-02
rs297593	2	157363743	GPD2	T	C	0.286	0.257	-7.82	--	5.45E-15	0	0.5285	159,461	II (23andMe)	7.80E-11
rs5442	12	6954864	GNB3	A	G	0.068	NA	-7.82	--	5.48E-15	8.8	0.03693	146,217	II (23andMe)	1.20E-33
rs10880855	12	46144855	ARID2	T	C	0.507	0.464	-7.78	--	7.35E-15	0	0.9683	160,144	I	4.80E-08
rs12405776	1	242431557	PLD5	T	C	0.220	0.521	7.75	++	9.52E-15	64.9	3.56E-10	153,784	II (23andMe)	1.50E-01
rs2150458	21	47377296	PCBP3, COL6A1	A	G	0.455	0.641	7.74	++	1.04E-14	55.7	0.1329	160,151	II (23andMe)	1.80E-13
rs12898755	15	63574641	APH1B	A	G	0.245	0.456	7.53	++	4.98E-14	7.9	0.2974	159,506	II (23andMe)	1.40E-16
rs7122817	11	117657679	DSCAML1	A	G	0.507	0.662	7.51	++	5.73E-14	73.8	0.05077	160,147	II (23andMe)	1.10E-10
rs10511652	9	18362865	SH3GL2, ADAMTSL1	A	G	0.416	0.445	7.36	++	1.91E-13	44.8	0.1782	160,149	II (23andMe)	3.50E-18
rs11101263	10	49414181	FRMPD2	T	C	0.258	0.105	-7.33	--	2.33E-13	0	0.3477	160,155	II (23andMe)	2.20E-13

**b Subset of the new loci harboring the smallest p-values for refractive error in the Stage 3 meta-analysis**

SNP	Chr	Position	Nearest Loci And Gene(s)	Effect Allele	Other Allele	EAF EUR	EAF ASN	Z-score	Direction	P value	HetISq	HetPVal	Sample Size (N)	Category	P value Replication
rs11118367	1	219790221	LYPLAL1	T	C	0.482	0.630	-7.29	--	3.16E-13	0	0.8576	160,141	III	1.20E-13
rs9395623	6	50757699	TFAP2D, TFAP2B	A	T	0.315	0.381	7.25	++	4.16E-13	0	0.9579	160,151	III	2.20E-10
rs284816	8	53362145	ST18, FAM150A	A	G	0.163	0.198	-7.21	--	5.52E-13	0	0.9242	160,140	III	1.60E-08
rs12965607	18	47391025	MYO5B	T	G	0.857	0.923	7.07	++	1.52E-12	20.8	0.01674	157,604	II (23andMe)	8.10E-16
rs7747	4	80827062	ANTXR2	T	C	0.202	0.093	7.03	++	2.05E-12	5.4	0.01267	150,327	II (23andMe)	7.70E-16
rs12451582	17	54734643	NOG, C17orf67	A	G	0.369	0.308	7.02	++	2.22E-12	0	0.5925	160,155	II (23andMe)	8.80E-18
rs80253120	17	14138507	CDRT15	T	C	0.626	0.723	6.97	++	3.25E-12	58.6	0.12	156,054	II (23andMe)	7.20E-11
22:23069851:1	22	23069851	DKFZp667J0810	ATG	A	0.084	0.1582	6.95	+-	3.56E-12	98.5	4.80E-16	120,481	II (23andMe)	9.30E-01
rs7968679	12	9313304	PZP	A	G	0.700	0.894	6.95	++	3.65E-12	0	0.01951	160,076	II (23andMe)	4.20E-10
rs11202736	10	90142203	RNLS	A	T	0.717	0.762	-6.92	--	4.53E-12	0	0.4007	160,150	II (23andMe)	9.40E-07
rs11088317	21	16574122	NRIP1, USP25	T	C	0.287	0.299	-6.90	--	5.38E-12	72.5	0.05657	160,116	II (23andMe)	6.50E-06
rs10853531	18	42824449	SLC14A2	A	G	0.200	0.182	6.88	++	5.89E-12	0	0.6755	160,104	III	2.60E-10
rs72655575	8	60556509	SNORA51, CA8	A	C	0.201	0.124	6.87	++	6.54E-12	0	0.8811	156,566	I	7.10E-07
rs12998513	2	242879499	CXXC11, AK097934	A	G	0.880	0.676	-6.86	+-	7.15E-12	65.2	4.51E-14	117,611	II (23andMe)	7.80E-01
rs1790165	11	131928971	NTM	A	C	0.411	0.283	6.85	++	7.17E-12	0	0.003708	160,131	II (23andMe)	1.80E-10
rs511217	11	30029948	METTL15, KCNA4	A	T	0.738	0.729	-6.79	--	1.10E-11	0	0.3626	160,143	II (23andMe)	1.40E-17
rs1150687	6	28162469	ZNF192P1, TRNA_Ser	T	C	0.619	0.504	6.78	++	1.17E-11	56.2	0.131	159,448	II (23andMe)	3.10E-10
rs56055503	16	80532694	MAF, DYNLRB2	A	G	0.751	0.539	-6.72	--	1.83E-11	0	0.8407	160,145	II (23andMe)	8.00E-06
rs9681162	3	8194734	AK124857, LMCD1-AS1	T	C	0.680	0.437	6.70	++	2.10E-11	63	0.1002	160,152	II (23andMe)	6.30E-13
rs11589487	1	61342229	AK097193, BC030753	A	G	0.445	0.089	6.67	++	2.64E-11	34.6	0.2163	160,143	II (23andMe)	2.20E-10

We identified 140 loci for refractive error with genome-wide significance ( $P < 5 \times 10^{-8}$ ) on the basis the meta-analyses of the genome-wide single-variant linear regressions performed in 160,420 participants of mixed ancestries (CREAM-ASN, CREAM-EUR and 23andMe). Shown are the replication of the previously found loci from HapMap II and a subset of the new loci harboring the smallest  $P$  values. For each locus, represented by an index variant (the variant with smallest p-value in that locus), Effect Allele, Other Allele, effect allele frequencies per ancestry (EAF AZN and EAF EUR), effect size (Z-score), direction of the effect (Direction), the  $P$  value, heterogeneity  $I$  square (HetISq), heterogeneity  $P$  value (HetPval), Sample Size (N), Category (1 = both GWS in Stage 1 and 2, 2 = one of two cohorts (CREAM or 23andMe) GWS, 3 = both not GWS in Stage 1 or 2) and  $P$  value of the replication in UK Biobank are shown (Full table: Supplementary Table 2). Chr, chromosome; EAF, effect allele frequency; ASN, Asian; EUR, European; GWS, genome wide significant.

**Table 2**

Genetic correlation for refractive error between Europeans and East Asians

Sample 1	Sample 2	Genetic effect correlation ( $pg_e$ ) <sup>a</sup>	Standard error $pg_e$	P value $pg_e$	Genetic impact correlation ( $pg_i$ ) <sup>a</sup>	Standard error $pg_i$	P value $pg_i$
EUR CREAM	EAS CREAM	0.804	0.041	1.83E-06	0.888	0.061	0.065
EUR 23andMe	EAS CREAM	0.788	0.041	2.48E-07	0.865	0.054	0.014

Abbreviations: EUR, European; EAS, East Asian.

<sup>a</sup>P-value relates to a test of the null hypothesis that  $pg_e=1$  or  $pg_i=1$ .

We calculated the genetic correlation of effect ( $pg_e$ ) and impact ( $pg_i$ ) using Popcorn to compare the genetic associations between Europeans (CREAM-EUR, N= 44,192; 23andMe, N=104,292) and East Asians (CREAM-ASN, N= 9,826). Reference panels for Popcorn were constructed using genotype data for 503 EUR and 504 EAS individuals sequenced as part of the 1000 Genomes Project. SNPs used had a MAF of at least 5% in both populations, resulting in a final set of 3,625,602 SNPs for the analyses using the 23andMe GWAS sample and 3,642,928 SNPs for those using the CREAM-EUR sample. These findings support a largely common genetic predisposition to refractive error and myopia in Europeans and Asians, although ancestry-specific risk alleles may exist.