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1 **Affective instability in those with and without mental disorders: a case control study**

2

3 ^{1,2} * **Steven Marwaha**, ³**Charlotte Price**, ⁴**Jan Scott**, ⁵**Scott Weich**, ¹**Aimee Cairns**,
4 ⁶**Jeremy Dale**, ¹**Catherine Winsper**, ²**Matthew R. Broome**

5

6 ¹Mental Health and Wellbeing, Division of Health Sciences, University of Warwick

7 CV47AL, UK

8 ² Institute for Mental Health, University of Birmingham, Edgbaston, Birmingham, B15 2TT,

9 UK

10

11 * Corresponding author

12

13 ³Operational Research and Management Sciences Group, Warwick Business School,

14 University of Warwick, CV4 7AL, UK

15

16 ⁴Institute of Neuroscience, Newcastle University, Newcastle upon Tyne, UK and

17 IOPPN, Kings College London, UK

18

19 ⁵ Mental Health Research Unit, School of Health and Related Research (ScHARR),

20 University of Sheffield, Regent Court, 30 Regent Street, Sheffield S1 4DA, UK

21

22 ⁶ Department of Primary Care, Warwick Medical School, University of Warwick, CV47AL,

23 UK

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26 **Word count: 4292**

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29 **1. Introduction**

30 Affective Instability (AI) is a transdiagnostic symptom (Broome et al., 2015b; Henry et al.,
31 2001b). It has been defined as rapid oscillations of intense affect, with difficulty regulating
32 these or their behavioural consequences (Marwaha, 2013). Multiple strands of evidence have
33 associated AI with suicidal thinking (Palmier-Claus et al., 2012; Yen et al., 2004), health
34 service use (Marwaha et al., 2013c), new onset of depression (Marwaha et al., 2015),
35 psychotic symptoms (Marwaha et al., 2013a), onset of bipolar disorder and increasing time to
36 recovery (Howes et al., 2011; Stange et al., 2016). It is also independently linked to greater
37 medication use and detention under mental health legislation (Patel et al., 2015). It is
38 associated with childhood trauma experiences (including abuse) and it is suggested that it
39 may partly explain the connection between these and psychiatric disorders (Aas et al., 2016;
40 Marwaha et al., 2016; Moffa et al., 2017). The estimated prevalence of AI in the general
41 population is 14%, with levels being higher in younger people and women (Marwaha et al.,
42 2013c; Patel et al., 2015). We (Broome et al., 2015a), and others (Harrison et al., 2017) have
43 suggested that trans-diagnostic investigation of AI is compatible with the NIMH Research
44 Domain Criteria project (Insel, 2014), a framework for understanding mental disorders by
45 study of dysfunction in individual psychological and biological systems.

46

47 Through systematically reviewing the literature we have previously identified AI as
48 having three core affect components: intensity, lability, and ability to control the oscillations
49 or their behavioural consequences (Marwaha et al., 2014). The review also identified that
50 comprehensive measurement of all three components is rarely undertaken. The current
51 literature is limited in part by theoretical and methodological heterogeneity in how AI is
52 understood and assessed. This means that studies of AI in different diagnostic groups cannot

53 be compared, and hence understanding whether AI is similar in different disorders and how it
54 contributes to outcomes such as functioning are hard to ascertain. As such, there is a
55 significant gap in understanding this clinical phenomenon. A second major shortcoming of
56 the current literature is that nearly all studies to date, apart from a few notable exceptions
57 (Ben-Zeev and Young, 2010; Ben-Zeev et al., 2009), have lacked comparisons with
58 individuals without mental disorder. This means it is unclear how far AI represents
59 psychopathology needing intervention or indeed whether it is a core aspect of abnormal
60 mental states, or is a feature of normal mental life.

61

62 To our knowledge, only three studies to date have compared AI in different diagnostic
63 groups using the same assessment procedures but limiting assessment of AI to two if its
64 domains. Henry et al. (2001a) examined AI using the Affect Lability Scale (ALS) and Affect
65 Intensity Measure (AIM) in out-patients with Borderline Personality Disorder (BPD) (N=29),
66 bipolar disorder: type II without BPD (N=14), BPD and bipolar disorder: type II (N=12), and
67 no BPD or bipolar disorder but other personality disorders (N=93). Lability scores were
68 significantly ($p < 0.05$) higher in BPD, whilst bipolar patients tended ($p = 0.06$) to have higher
69 lability scores than other personality disorders. No differences in affect intensity were
70 observed.

71

72 In a similar study those with bipolar disorder had significantly higher scores on the
73 euthymia-elation subscale of the ALS as well as significantly higher scores on affect
74 intensity, whereas those with BPD experienced more shifts between anxiety-depression,
75 euthymia-anger and significantly fewer shifts between euthymia-elation and depression-
76 elation (Reich et al., 2012). Most recently Richard-Lepouriel et al (Richard-Lepouriel et al.,
77 2016) compared ALS and AIM scores in people with bipolar disorder, attention deficit

78 hyperactivity disorder (ADHD) and controls (dentistry students and doctors). Those with
79 ADHD and bipolar disorder scored higher on ALS than controls, with AIM scores being
80 highest for ADHD. Whilst affective lability appears to be higher in BPD, and people with
81 bipolar disorder, results for affective intensity are discrepant between studies with some
82 evidence that affective intensity may be higher in people with mental disorders than in
83 controls.

84

85 Given the paucity of previous research, we aimed to expand the diagnostic groups in
86 which AI is examined (given the suggestion that AI is transdiagnostic), compare these
87 “cases” with psychologically “healthy controls”, assess AI more comprehensively, and test
88 whether AI is independently linked to functioning within a clinical population.

89

90 **2. Aims**

91

92 We aimed to answer the following research questions:

93

- 94 1. Does affective instability differ between clinical cases and controls and between
95 diagnostic groups?
- 96 2. To what extent are measures of affective lability, intensity and ability to control affect
97 correlated in a trans-diagnostic clinical sample?
- 98 3. Is affective instability associated with functioning in a clinical population independent
99 of diagnosis?

100 **3. Methods**

101 We undertook a case-control study among users of secondary care mental health services
102 (cases) and primary care attenders without evidence of current mental disorder (controls).

103 Ethical approval was obtained from the Coventry and Warwickshire Ethics Committee, UK.

104 Participant consent and data collection was completed by an experienced researcher with a
105 psychology background.

106

107 *3.1 Participants*

108

109 Individuals with a range of diagnoses were recruited from secondary care mental
110 health services within Coventry and Warwickshire, UK through convenience sampling. The
111 aim was to include individuals who were representative of the ‘typical’ case mix of these
112 services, so participants were recruited in out-patient departments, day hospitals, community
113 mental health teams and a specialist personality disorder service. Inclusion criteria were: a]
114 aged 18-65 years; b] capacity to give informed consent; c] the primary reason for attending
115 the mental health service was for management of a clinical diagnosis of BPD, bipolar
116 affective disorder, major depressive episode (moderate or severe depressive episode) or non-
117 affective psychosis as reported by a Consultant Psychiatrist. The researcher confirmed the
118 diagnosis with the Psychiatrist using ICD-10 criteria.

119

120 Exclusion criteria were: a] an acute illness episode (sufficient to require urgent or
121 inpatient care) according to the patient’s Consultant Psychiatrist; b] unable or unwilling to
122 complete the assessments (e.g. individuals with a clinically assessed learning disability, with
123 insufficient command of the English language to understand and complete questionnaires); or
124 c] individuals with a primary ICD-10 diagnosis of dependency to drugs or alcohol (to avoid
125 confounding by drug or alcohol misuse).

126

127 Control participants were recruited from primary care (general practitioner surgeries).
128 Physicians asked patients if they were interested, a researcher in the waiting room then
129 consented the patients and completed the battery of questionnaires. Exclusion criteria for the
130 control group were: a] presence of a current mental disorder (including common mental

131 disorders such as depression or anxiety disorders); b] dependency on substances or alcohol;
132 c] previous diagnosis of BPD, bipolar disorder, or non-affective psychosis, according to their
133 primary care records.

134

135 *3.2 Materials*

136

137 Details were collected on participants' diagnosis (for cases) and confirmed by their
138 Consultant Psychiatrist. Details on duration of illness (cases only) and current medications
139 were identified by a researcher, from clinical records. Medications were grouped as anti-
140 psychotic, anti-depressant, anti-anxiety, mood stabiliser, anti-depressant/mood-stabiliser or
141 'other' (medication not directly related to the patient's psychiatric diagnosis).

142 Sociodemographic information was collected on age (years), gender, marital status
143 (married/cohabiting, single, separated/widowed), employment (employed, unemployed,
144 other), ethnicity (White British, other), and highest education level (None/GCSE, A Level,
145 Degree/higher degree) (see Table 1).

146

147 *3.3 Assessment scales*

148

149 Participants were asked to complete four questionnaires relating to affective / mood
150 instability.

- 151 1. The Affective Lability Scale - short form (ALS-18) (Harvey et al., 1989), is a
152 reliable and valid measure comprising 18 items coded 0-3. Overall score is
153 obtained by taking the mean of the scores for each item as indicated by the scale
154 developers. Three subscales can be derived; 'anxiety-depression', 'depression-
155 elation', and 'anger'.
- 156 2. The Affective Intensity Measure-20 (AIM) (Larsen et al., 1986a) was used to
157 measure affect intensity. The AIM contains 20 items coded 1 to 6. A person's

158 overall score is obtained by taking the mean of the scores for each item. The AIM
159 has good internal consistency, test-retest reliability and construct validity (Larsen
160 et al., 1986b).

- 161 3. The Affective Control Scale (ACS) (Williams et al., 1997) comprises 42 items,
162 coded 1 to 7 (with some items requiring reverse scoring); it has good
163 psychometric properties including construct validity. A higher ACS score
164 indicates reduced ability to control affect. Four subscales can be derived; ‘anger’,
165 ‘positive affect’, ‘depressed mood’, and ‘anxiety’.
- 166 4. Mood fluctuation rate: Because of the lack of a previously well validated scale for
167 fluctuation rate (Marwaha, 2013) we used a new bespoke schedule for this study.
168 Mood fluctuation rate was assessed using a question from the Structured Clinical
169 Interview for DSM Disorders (SCID). It asks the respondent to state how often
170 they experienced a sudden marked shift in mood. Study participants rated the
171 number of significant mood fluctuations they experienced over a week.
172 Respondents were asked to consider this for each one of the weeks in the last
173 month prior to assessment, and possible responses were 0, 1-3, 4-7, or >7 mood
174 changes over each week.

175

176 Functioning was measured using the WHO Disability Assessment Schedule 2.0 – 12
177 item version (WHODAS; Üstün, 2010). This contains 12 items each coded 0 to 4. To obtain
178 a person’s final score, the simple version entails summing the scores from each of the 12
179 items, scores range from 0-48. For consistency in comparing with the other scores above, the
180 mean rather than the sum was used in the current study.

181

182 *3.4 Data analysis*

183

184 Descriptive statistics including means/medians with standard deviations/interquartile ranges,

185 or frequencies with percentages where relevant, were used to investigate participant

186 demographics and characteristics of AI in the different diagnostic groups and controls. There

187 are no clear rules about the acceptable fraction of missing data to justify imputation. As such,

188 we decided on 10%, as a level that would allow imputation, thus enabling us to use as much

189 of the data as possible, whilst also retaining reliability and accuracy (Steyerberg, 2008). As

190 such scores were imputed if the patient had less than 10% missing items. This translates as:

191 AIM: Up to 2 missing values, ALS-18: Up to 2 missing values, ACS: Up to 4 missing values,

192 WHODAS: 1 missing value.

193

194 Two sample t-tests were used to compare means between the cases and controls after

195 verifying that relevant assumptions were valid. Proportions were compared using chi-squared

196 tests. General linear models (GLMs) were used to compare the mean lability (ALS), intensity

197 (AIM), and subjective ability to control affect (ACS) outputs across cases (different

198 diagnostic groups) and the control group. Adjustment was made for age, sex and educational

199 level if necessary. Model assumptions were checked and, in the case of an overall significant

200 difference in mean score across the diagnosis groups, pairwise post-hoc comparisons of

201 adjusted mean scores were performed with a Bonferroni correction. To investigate how far

202 the different aspects of affective instability correlate with each other, the linear association

203 between each pair of measurement scales for the full sample and for the cases only was

204 assessed using Pearson's product moment correlation. Association between each

205 measurement scale and the mood fluctuation rate was assessed using Spearman's rank

206 correlation.

207
208 Multiple regression was used to examine the association between affective instability
209 and general assessment of functioning as measured by the WHODAS (Üstün, 2010) in
210 clinical cases, adjusting for diagnosis and other patient characteristics. A purposeful selection
211 approach was used to fit the model. Manual backward elimination was first used to remove
212 variables based on Wald statistics using $p = 0.05$ as the cut-point for removal. Removed
213 variables were then re-entered into the model one-by-one to check their significance.
214 Variables initially considered in the model included: (a) socio-demographics: age, sex,
215 ethnicity, education level, marital status, (employment status was not considered in the model
216 since the WHODAS incorporates this parameter in ratings), (b) illness characteristics:
217 diagnosis, duration of illness, total number of medications, (c) AI measures: mood fluctuation
218 rate and the mean scores for the ACS, ALS-18 and AIM. All analyses were conducted in
219 IBM SPSS Statistics 24.

220

221 **4. Results**

222

223 4.1 Participant characteristics

224

225 The initial dataset comprised 101 participants, but 9 individuals were excluded due to
226 missing data (3 bipolar, 2 major depression, and 2 controls). Hence the final sample (N=94)
227 comprised of 69 cases and 25 controls.

228

229 Table 1 describes the socio-demographic characteristics of included participants by
230 group (case versus control), and diagnostic subgroups (bipolar disorder (n=11), BPD (n=12),
231 psychosis (n=21), and major depression (n=25)). There was a significant difference between
232 cases and controls in mean age ($p=0.001$), employment status ($p=0.001$) and marital status
233 ($p<0.001$). Age was controlled for during regression analysis as AI is influenced by this

234 (Marwaha et al., 2013b). Duration of illness was recorded in the dataset for 67 out of the 69
235 cases and was positively skewed with the sample having been ill for a median duration of 36
236 months (interquartile range (IQR) 15-156 months). Across the diagnostic groups, participants
237 with depression reported the longest duration of illness (median 120 months, IQR 12-258),
238 followed by participants with BPD (median 36 months, IQR 24-120), psychosis (median 27
239 months, IQR 20.5-111), and bipolar disorder (median 24 months, IQR 9-36), respectively.

240

241 *Table 1 about here*

242

243 The commonest class of medication prescribed was anti-depressants, and most patients
244 reported being prescribed one (25%) or two (25%) medications. Seventeen percent of cases
245 (N=12) were not taking any medication. We did not explore, type of medications and their
246 impact on our results because of the lack of a robust typology of the effectiveness of
247 medications indicated for affective instability (Lieb et al., 2010).

248

249 4.2 Comparison of questionnaire scores between cases and controls

250 Differences in the unadjusted mean scores between cases and the controls for all
251 measures are presented in Table 2. Age and sex were found not to be significant across the
252 sample in the general linear model (GLM) for the ACS, AIM, and WHODAS scores; whilst
253 there was trend towards significance for the effect of age on ALS score ($P = 0.068$).

254

255 *Table 2 about here*

256

257 Affect lability (ALS): When adjusted for age, a statistically overall significant
258 difference was observed in mean ALS-18 scores between cases and controls ($F(4,88) =$

259 7.195, $p < 0.001$). Post-hoc pairwise comparisons of mean scores revealed significantly lower
260 mean ALS-18 scores for the control group compared to each diagnosis group but no
261 significant differences between diagnoses.

262

263 Affect intensity (AIM): There was little difference in the mean AIM scores between
264 groups, with slightly higher mean scores found for controls compared to cases. These
265 differences were not statistically significant ($p = 0.867$).

266

267 Ability to control affect (ACS): An overall significant difference was found between
268 mean ACS scores across the different diagnostic groups, including controls ($F(4,89) =$
269 14.520 , $p < 0.001$). Post-hoc pairwise comparisons of the mean scores revealed significantly
270 higher mean ACS scores (meaning lower control) for each diagnostic group compared to
271 controls ($p < 0.05$). A significant difference was also found between the mean scores in
272 borderline personality disorder patients and patients with non-affective psychosis ($p = 0.010$).

273

274 Mood fluctuations in the last week: Table 2 shows the number of participants (i.e.
275 frequency with percentage) who reported each number of mood fluctuations over the past
276 week prior to assessment. This revealed that cases tended to have more changes in their mood
277 state than controls, although no overall differences were found in rate of mood fluctuation
278 between groups ($p=0.310$). Those with major depression reported the greatest number of
279 mood fluctuations in the last week, followed by non-affective psychosis, borderline
280 personality disorder and then bipolar disorder.

281

282 *4.3 Correlations between different components of AI, mood fluctuation rate and functioning*

283 Correlations are shown in table 3. Strong positive correlations were found between
284 the ALS and the ACS in the full and cases only analysis. Weak to moderate correlations were
285 found between the AIM and the ALS. When assessing the association between each
286 measurement scale and mood fluctuation rate ‘last week’, moderate to strong positive
287 correlations were found between mood fluctuation and ALS and ACS. There was a weak
288 correlation between AIM and mood fluctuation rate. All correlations were weaker when
289 focusing on the cases only.

290

291 *Table 3 about here*

292

293 *4.4 AI and functioning*

294

295 In the clinical sample, an overall significant difference was observed between mean
296 WHODAS scores across the different diagnosis groups, $F(4,89) = 11.454$, $p < 0.001$ ($p <$
297 0.05 for bipolar disorder). Post-hoc pairwise comparisons revealed significantly lower mean
298 WHODAS scores for the control group compared to each diagnosis group, as might be
299 expected, but differences between diagnostic groups were not significant.

300

301 A multiple regression model investigating factors associated with the WHODAS
302 score, demonstrated that both ALS-18 and ACS scores were significantly associated with
303 current level of functioning. After correcting for multicollinearity, ALS-18 score was retained
304 in the final model ($\beta=0.845$, $p<0.001$), along with the total number of medications
305 ($\beta=0.107$, $p<0.046$). All other variables considered, including diagnosis, were not
306 significantly associated with WHODAS score in the final model.

307

308 **5. Discussion**

309

310 *5.1 Main findings*

311
312 This is the first study, to our knowledge, that has comprehensively assessed the core
313 components of affective instability in a trans-diagnostic clinical population and compared
314 clinical cases with a control group without mental disorder. We found only affective lability
315 and affective control is significantly different in people with a range of mental disorders in
316 comparison to those without. No differences were observed between people with and without
317 mental disorder in the intensity of affect experienced or the rate of mood fluctuation in the
318 last week. Two of the three components of affective instability (lability and intensity) did not
319 differ significantly between individuals with different psychiatric diagnoses, although ability
320 to control affect was significantly different in individuals with BPD in comparison to non-
321 affective psychosis. Whilst the small numbers within each diagnostic group mean that
322 interpretation can only be exploratory, contrary to expectation, we found that the greatest
323 number of mood changes in a week was experienced by people with major depression,
324 followed by non-affective psychosis, BPD and then bipolar disorder.

325
326 In terms of the affective instability construct, the strongest inter-correlation was found
327 between lability and ability to control affect, with much weaker (modest) correlations
328 between affective intensity and ability to control affect (or lability and control). Finally, only
329 affective lability, but not affective intensity, ability to control affect or mood fluctuation rate
330 was associated within functioning independent of diagnosis and other important confounders.

331

332 *5.2 Limitations*

333
334 Our sample size was relatively modest (just under 100). This limited the statistical
335 power of our analyses and increased the risk that our results might be due to type II error.
336 This means that comparisons of affective instability between diagnostic groups in particular,

337 should be considered entirely exploratory, and other interpretations tentative. Another caveat
338 to comparisons between diagnoses is that we did not complete inter-rater reliability
339 assessments. However, this is the largest study to date exploring our questions.

340

341 Our observations related to affective instability are limited to the four mental
342 disorders that we sampled. We cannot therefore generalize our findings to other disorders,
343 where affective instability is known to be important such as OCD (Bowen et al., 2015) or
344 ADHD (Asherson et al., 2007). Furthermore, we could not take into account the contribution
345 made by mental or physical comorbidities in our sample. However, given our sample of cases
346 were those in contact with secondary mental health services there are likely to be high levels
347 of comorbidity. Therefore, it is possible that high levels of affective lability and problems
348 with affective control are linked to comorbidity and this should be the focus of future studies.
349 In our regression modelling we were not able to control for some factors known to impact
350 functioning such as cognition, illness severity, premorbid functioning and depressive
351 symptoms.

352

353 The cases sampled were not in an acute illness episode and it is conceivable that this
354 biased estimate of group difference towards the null, that is, there is no difference between
355 the cases and controls on affective instability measures. Affect intensity (and possibly
356 instability) might vary with illness acuity, which might explain why differences between
357 cases and controls in the present study were smaller than those reported in an in-patient
358 sample (Henry et al., 2008; Reich et al., 2012). Whilst we did not assess illness severity, we
359 adjusted for illness duration and number of medications, both of which might be expected to
360 be associated with illness severity. More specifically, we also did not assess current mood
361 state using standardised measures and therefore do not know how far the severity of current

362 mood (e.g depth of depression) could have impacted on our results. There is little current
363 evidence on how far AI changes, as mood becomes lower or more elated to guide how this
364 could have influenced our main findings. Indeed, in bipolar disorder, AI is found in both
365 euthymic and periods of acute illness (Harvey, 2008). We explored whether AI is different
366 between cases and controls. Future studies should also aim to explain the differences between
367 affective instability in people with mental disorders and without.

368

369 We used assessment measures which require recall of affective experiences. These
370 may be prone to bias, particularly when compared to ecological momentary assessments
371 (EMA) (Broome et al., 2015b). How accurately people with mental disorders recall their
372 affective experiences might differ depending on diagnosis. The ratings themselves at an
373 individual level may also be dependent on an initial calibration to understand what is meant
374 by a “marked” shift in mood (Holmes et al., 2016). Therefore, paradoxically individuals with
375 fewer mood fluctuations may better report retrospective fluctuations as they would have
376 stood out in their experience, whilst those with more frequent fluctuations may only report
377 “marked” ones, as small fluctuations were perhaps normalised by their experience. This is
378 one potential explanation of why depressed patients reported more fluctuations than other
379 groups, though this was not statistically significant. The question used to assess mood
380 fluctuation didn’t specify type of affect and therefore could have excluded swings in anger
381 and irritability, which have been shown to differentiate between diagnosis (Tsanas et al.,
382 2016). We also recognise that current mood state may have impacted on assessment.

383

384 Whilst momentary assessment of psychopathology appears feasible using
385 smartphones (Tsanas et al., 2016), it is as yet unclear whether retrospective affective
386 assessments and EMA relate to the same underlying psychological or biological processes,

387 especially as the former will be subject to important cognitive processes (e.g contextual
388 processing), which control how mood is experienced (Dubad et al., 2018). There is also the
389 issue of how far individuals recognize and name affective states in the same way.

390
391 *5.3 Theoretical and clinical implications*
392
393

394 Our findings only partly validated our original definition of affective instability as a
395 trans-diagnostic parameter incorporating affect lability, ability to control and intensity
396 (Marwaha et al., 2014). Affect lability and the ability to control these were indeed found to
397 occur at higher levels than in controls and at similar levels across the different diagnostic
398 groups. Scores on both measures were also relatively strongly correlated with each other re-
399 enforcing the notion that they are facets of the same or similar underlying latent construct.
400 Affective intensity was only relatively weakly associated with other affective instability
401 measures. Replication in a much larger sample is required to understand how far this pattern
402 holds true. In the current study affective intensity was no different between cases and controls
403 or between the cases themselves consistent with previous literature (Henry et al., 2001b).
404 Whilst caution is necessary in interpretation, this does suggest that intensity of affect may not
405 be a feature that may help delineate the boundaries of “normal” or “abnormal” affective
406 experience, or at least in the way that it was measured here. Again, a study with a larger
407 sample size is required.

408
409 Mood fluctuation rate (as measured by our bespoke instrument) showed some
410 concurrent validity with two measures of affective instability, and surprisingly, fluctuation
411 rate was no different between cases and controls. This may be a function of our sample size,
412 but this finding should prompt larger studies, with more comprehensive fluctuation change

413 assessments to investigate this area. Crucially, these studies need to include people without
414 mental disorders as controls.

415

416 We used a comprehensive way to measure affective instability in people with
417 different diagnoses and the current results as well as previous research provides some
418 counterbalance to the notion that affective instability is specific to or more severe in people
419 with bipolar disorder or borderline personality disorder. The challenge now is to understand
420 whether more subtle differences exist that may be clinically useful, such as whether a
421 particular valence change is more or less common in different disorders (Reich et al., 2012)
422 or whether richer, digitally captured mood data is helpful in differentiating disorders. Current
423 evidence indicates clinicians do not use diagnostic criteria effectively to distinguish disorders
424 such as BPD and bipolar disorder in which affective instability symptoms are seen to overlap
425 (Saunders et al., 2015). Further research into common and uncommon valence changes in the
426 disorders, perhaps incorporating digital mood monitoring, may help to resolve this clinical
427 difficulty.

428

429 Finally, we demonstrate that affective instability independently adversely impacts
430 functioning in people with mental disorders, and this is independent of diagnosis. The
431 measure of functioning that we used suggests the impact could be on multiple domains
432 including learning new tasks, joining in community activities, day to day work and
433 maintaining friendships. We have previously found that interpersonal conflict is part of the
434 pathway from affective instability and incident depression (Marwaha et al., 2015) and the
435 current study is also consistent with other work highlighting the impact of affective instability
436 on functioning in bipolar and transdiagnostically (Patel et al., 2015; Strejilevich et al., 2013).

437 We extend these previous findings by identifying that affective lability, as opposed to other
438 aspects of AI such as ability to control affect or intensity, has the greatest impact.

439

440 As such affective lability has the potential for being a therapeutic target that could
441 improve functional outcomes in mental disorders. Pharmacological interventions that are
442 widely used (e.g mood stabilising antipsychotics) and emotional regulation training (Berking
443 et al., 2008) need more robust trial evidence, but could have a significant impact on distress
444 and outcomes.

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Table 1: Descriptive statistics by group; case versus control and diagnosis subgroups (n = 94)

Characteristic	Control (n = 25)	Case (n = 69)	Cases by diagnosis				Total (n = 94)	Case versus control	
			Bipolar (n = 11)	Borderline (n = 12)	Psychosis (n = 21)	Depression (n = 25)		P value	Test
Age (years); mean (SD)	48.5 (10.8)	38.2 (12.8)	35.6 (13.3)	33.9 (11.0)	34.9 (9.8)	44.2 (14.1)	41.0 (13.1)	0.001	t test
Male; n (%)	9 (36.0%)	36 (52.2%)	6 (54.5%)	2 (16.7%)	12 (57.1%)	16 (64.0%)	45 (47.9%)	0.165	Chi square test
Employment; n (%):									
– Employed	19 (76.0%)	26 (37.7%)	5 (45.5%)	4 (33.3%)	6 (28.6%)	11 (44.0%)	45 (47.9%)	0.001	Chi Square test (employed vs unemployed, n = 90)
– Unemployed	5 (20.0%)	40 (58.0%)	6 (54.5%)	8 (66.7%)	14 (66.7%)	12 (48.0%)	45 (47.9%)		
– Other	1 (4.0%)	3 (4.3%)	0 (0.0%)	0 (0.0%)	1 (4.8%)	2 (8.0%)	4 (4.3%)		
Ethnicity; n (%)									Chi square test
– White British	18 (72.0%)	57 (82.6%)	9 (81.8%)	10 (83.3%)	14 (66.7%)	24 (96.0%)	75 (79.8%)	0.259	
– Other	7 (28.0%)	12 (17.4%)	2 (18.2%)	2 (16.7%)	7 (33.3%)	1 (4.0%)	19 (20.2%)		
Education; n (%)									Chi square test (n = 89)
– None/GCSE	7 (28.0%)	35 (50.7%)	5 (45.5%)	5 (41.7%)	12 (57.1%)	13 (52.0%)	42 (44.7%)	0.061	
– A level	3 (12.0%)	16 (23.2%)	4(36.4%)	2 (16.7%)	5 (23.8%)	5 (20.0%)	19 (20.2%)		
– Degree/higher degree	11 (44.0%)	17 (24.6%)	2 (18.2%)	5 (41.7%)	3 (14.3%)	7 (28.0%)	28 (29.8%)		
– [Missing]	4 (16.0%)	1 (1.4%)	-	-	1 (4.8%)	-	5 (5.3%)		

Marital status; n (%)									Chi square test
– Married/cohabiting	20 (80.0%)	22 (31.9%)	1 (9.1%)	6 (50.0%)	5 (23.8%)	10 (40.0%)	42 (44.7%)	< 0.001	
– Single/separated/ widowed	5 (20.0%)	47 (68.1%)	10 (90.9%)	6 (50.0%)	16 (76.2%)	15 (60.0%)	52 (55.3%)		

Table 2: Unadjusted mean scores (with standard deviation) and count (%) for each measurement scale and subscales

Measurement scale	Controls (n = 25)	Cases (n = 69)				All (n = 94)	Case versus control (General linear models, F test)
		Bipolar (n = 11)	Borderline (n = 12)	Psychosis (n = 21)	Depression (n = 25)		
ACS (scale 1-7)	3.36 (0.56)	4.39 (0.66)	5.14 (0.58)	4.23 (0.86)	4.52 (0.86)	4.21 (0.92)	p<0.001
ACS: Anger	3.06 (0.60)	4.31 (0.83)	5.21 (0.96)	4.07 (0.89)	4.08 (1.15)	3.97 (1.11)	
ACS: Positive-affect	3.48 (0.63)	4.17 (0.73)	4.12 (1.02)	4.05 (0.89)	3.87 (1.03)	3.88 (0.89)	
ACS: Depressed	3.36 (0.75)	4.96 (0.88)	5.92 (0.72)	4.40 (1.01)	5.38 (1.04)	4.64 (1.27)	
ACS: Anxiety	3.42 (0.57)	4.31 (0.84)	5.63 (0.60)	4.42 (1.23)	4.93 (0.95)	4.43 (1.13)	
ALS-18 (scale 0-3)	0.64 (0.58)	1.47 (0.62)	1.66 (0.49)	1.53 (0.67)	1.50 (0.62)	1.29 (0.71)	p<0.001 ¹
ALS-18: Anxiety/Depression	0.55 (0.65)	1.49 (0.69)	2.23 (0.79)	1.60 (1.01)	1.88 (0.87)	1.46 (1.00)	
ALS-18: Depression/Elation	0.86 (0.70)	1.66 (0.57)	1.45 (0.58)	1.65 (0.73)	1.52 (0.53)	1.38 (0.70)	
ALS-18: Anger	0.38 (0.48)	1.16 (0.95)	1.42 (0.96)	1.27 (0.88)	1.08 (1.03)	0.99 (0.92)	
AIM (scale 1-6)	3.50 (0.48)	3.45 (0.37)	3.37 (0.44)	3.42 (0.56)	3.37 (0.39)	3.42 (0.45)	p=0.867
Number of mood fluctuations reported in the last week							P=0.310
0	13 (52%)	2 (18.2%)	0 (0%)	5(23.8%)	2 (8%)	22 (23.4%)	
1-3	10 (40%)	5 (45.5%)	3 (25%)	4 (19%)	8 (32%)	30 (31.9%)	
4-7	2 (8%)	3 (27.3%)	7 (58.3%)	7 (33.3%)	7 (28%)	26 (27.7%)	
>7	0 (0%)	1 (9.1%)	2 (16.7%)	5 (23.8%)	8 (32%)	16 (17.0%)	
WHODAS ² (scale 0-4)	0.54 (0.11)	1.43 (0.61)	1.83 (0.69)	1.75 (1.00)	1.89 (0.91)	1.44 (0.96)	p<0.001

¹ Adjusted for age.

Table 3: Correlation coefficients between each pair of measurement scales

Full -sample (N=94)					
		AIM	ALS-18	ACS	Mood fluctuation (last week)
AIM		1	0.210	0.188	0.12
ALS-18			1	0.776	0.61
ACS				1	0.53
Mood fluctuation (last week)					1
Cases only (N=69)					
		AIM	ALS-18	ACS	Mood fluctuation (last week)
AIM		1	0.322	0.265	0.157
ALS-18			1	0.666	0.45
ACS				1	0.29
Mood fluctuation (last week)					1