White Rose Research Online

eprints@whiterose.ac.uk
https://eprints.whiterose.ac.uk
Universities of Leeds, Sheffield and York

Deposited via The University of Leeds.

White Rose Research Online URL for this paper:
https://eprints.whiterose.ac.uk/id/eprint/124363/

Version: Accepted Version

**Article:**
Ullah, S, Akhtar, P and Zaefarian, G (2018) Dealing with endogeneity bias: The generalized method of moments (GMM) for panel data. Industrial Marketing Management, 71. pp. 69-78. ISSN: 0019-8501

https://doi.org/10.1016/j.indmarman.2017.11.010

© 2017 Elsevier Inc. This manuscript version is made available under the CC-BY-NC-ND 4.0 license http://creativecommons.org/licenses/by-nc-nd/4.0/

**Reuse**
This article is distributed under the terms of the Creative Commons Attribution-NonCommercial-NoDerivs (CC BY-NC-ND) licence. This licence only allows you to download this work and share it with others as long as you credit the authors, but you can't change the article in any way or use it commercially. More information and the full terms of the licence here: https://creativecommons.org/licenses/

**Takedown**
If you consider content in White Rose Research Online to be in breach of UK law, please notify us by emailing eprints@whiterose.ac.uk including the URL of the record and the reason for the withdrawal request.

UNIVERSITY OF LEEDS
University of Sheffield
UNIVERSITY of York

# Dealing with endogeneity bias: The generalized method of moments for panel data

**Subhan Ullah***
Accounting and Finance,
Faculty of Business, Law and Politics,
University of Hull
United Kingdom

**Pervaiz Akhtar**
Management Systems, Logistics Institute,
Faculty of Business, Law and Politics,
University of Hull
United Kingdom
IESEG School of Management, France.

**Ghasem Zaefarian**
Leeds University Business School,
University of Leeds,
United Kingdom

*** Corresponding author: Dr Subhan Ullah**, Lecturer in Accounting and Finance, Faculty of Business, Law and Politics, Hull University Business School.

Email: subhan.ullah@hull.ac.uk

**Acknowledgement**

The authors would like to thank Adam Lindgreen and Anthony Di Benedetto for encouraging us to write about the topic of how to deal with endogeneity bias in panel data as well as Peter Naudé for his suggestions and help with the manuscript.

# Dealing with endogeneity bias: The generalized method of moments for panel data

**Abstract:**

Endogeneity bias can lead to inconsistent estimates and incorrect inferences, which may provide misleading conclusions and inappropriate theoretical interpretations. Sometimes such bias can even lead to coefficients having the wrong sign. Although this is a long-standing issue, it is now emerging in marking and management science, with high-ranked journals increasingly exploring the issue. In this paper, we methodologically demonstrate how to detect and deal with endogeneity issues in panel data. For illustration purposes, we used a dataset consisting of observations over a 15-year period (i.e., 2002 to 2016) from 101 UK listed companies and examined the direct effect of R&D expenditures, corporate governance, and firms' characteristics on performance. Due to endogeneity bias, the result of our analyses indicates significant differences in findings reported under the ordinary least square (OLS) approach, fixed effects and the generalized method of moments (GMM) estimations. We also provide generic STATA commands that can be utilized by marketing researchers in implementing a GMM model that better controls for the three sources of endogeneity, namely, unobserved heterogeneity, simultaneity and dynamic endogeneity.

**Keyword:** endogeneity bias; generalized method of moments; methodological issues; panel data.

## 1. Introduction

Endogeneity in regression models refers to the condition in which an explanatory (endogenous, e.g., research and development expenditure) variable correlates with the error term, or if two error terms correlate when dealing with structural equation modelling. Endogeneity bias can therefore cause inconsistent estimates (i.e., not tend to be the true value as sample size increases), which potentially leads to wrong inferences, misleading conclusions and incorrect theoretical interpretations. Ketokivi and McIntosh (2017) even stated that researchers might not get the correct sign of coefficients if endogeneity bias exits. Research suggests approximately 90% of papers published in premier journals have not adequately addressed endogeneity bias (e.g., Antonakis, Bendahan, Jacquart, & Lalive, 2010; Hamilton & Nickerson, 2003). Based on a study of over 100 articles in top journals, it is claimed that "*researchers fail to address at least 66% and up to 90% of design and estimation conditions that make causal claims invalid*" (Antonakis et al., 2010, p. 1086).

Despite recent methodological advances and the relevant literature in econometrics/psychology, other social science disciplines (e.g., marketing, operations management, international business and supply chain management) have largely produced inconsistent estimates due to not addressing endogeneity biases. However, marketing (e.g., Journal of Marketing, Journal of Marketing Research, and more recently Industrial Marketing Management) and operations management (e.g., Journal of Operations Management) journals have started to take it more seriously, and asked authors to fully address endogeneity in their studies (e.g. Ketokivi & McIntosh, 2017; Reeb, Sakakibara, & Mahmood,2012; Zaefarian et al., 2017). Researchers are responding to this call; for example, the Industrial Marketing Management journal has seen an increase number of authors addressing endogeneity bias in their studies published in 2017 (a total of 6 papers to be exact) compared to that of the previous year (only 1 paper). The reviewers associated with these journals have also played their part in directing researchers to address

such methodological complications. Nonetheless, many researchers in management disciplines are not yet fully aware of endogeneity, its sources, and relevant remedies (Antonakis, Bendahan, & Lalive, 2014; Guide & Ketokivi, 2015; Zaefarian et al., 2017).

Importantly, endogeneity bias can have different origins, and different methods exist to address them. For example, the dynamic generalized method of moments model (GMM) is used to address panel data (i.e., dynamic endogeneity bias) and two-stage least squares (2SLS)/three-stage least squares (3SLS) are often used for survey data. Some researchers have recently provided reviews to understand the key endogeneity concepts and relevant techniques (e.g., see Zaefarian et al., 2017). However, a step-by-step procedure on how to execute these techniques for a particular research problem is still missing. We therefore provide a succinct overview of the key endogeneity sources and solutions, and comprehensively demonstrate the GMM method using a case study of a panel dataset consisting of 15 years of observations. Specifically, this study explores how the dynamic nature of investment in R&D expenditures together with corporate governance affect firm performance. To better illustrate how endogeneity bias may cause incorrect estimates, we examine our proposed model using three different approaches, namely, ordinary least square (OLS), fixed effects, and the generalized method of moments (GMM). Practically, our main aim and contribution is to provide a comprehensive procedure for researchers to produce consistent estimates and to draw valid inferences when dealing with panel data.

In addition, panel data is used far less frequently in business-to-business than in the business-to-consumer marketing domain, and this article could provide a starting point as to how industrial marketing and management researchers can utilize such datasets to provide insights for business practitioners. For instance, the research and development expenditure and its relationship to firm financial performance in industrial marketing can be explored by using

panel datasets that are available from different databases (e.g., DataStream), unfortunately researchers are often unaware of such resources).

## 2. Sources of endogeneity

The error term in endogeneity bias is unobservable, so there is no direct way to statistically test that an endogenous variable is correlated with the error term. Also, exogenous variables are probably never exogenous precisely (Ketokivi & McIntosh, 2017). It is therefore almost impossible to statistically ensure that an endogeneity problem can be completely resolved (Roberts & Whited, 2012). That is why such dilemmas do not ask for solutions, they require better choices (Ketokivi & McIntosh, 2017). For choices, researchers need to understand the sources of the problem and then take reasonable actions to reduce the negative impact in order to deal with endogeneity effectively. As there are no direct tests for endogeneity, the choices of indirect tests and precautionary measures can help to guide relevant insights and conclusions (Ketokivi & McIntosh, 2017). Endogeneity encompasses common-method variance, measurement errors, omitted variables/selections and simultaneity. It is important to address them theoretically (e.g., extensively reviewing literature and providing comprehensive research designs that could help to apply appropriate statistical tools) as well as empirically (e.g. using statistical techniques to ensure that data is rigorously investigated) (Antonakis et al., 2010; Ketokivi & McIntosh, 2017).

### *2.1 Common-method variance and its remedies*

Common-method variance (CMV) is related to measurement methods. CMV is problematic due to its interlinks with the sources of measurement errors, These sources can come from common-rater effects (e.g., only collecting information from similar respondents), common measurement content (e.g., time, location and a single-medium used to collect data), common-item context or item characteristics (e.g., wording, length and clarity), scale types, respondents,

response formats and the general content (Malhotra, Kim, & Patil, 2006; Podsakoff, MacKenzie, Lee, & Podsakoff, 2003). Research suggests that the difference between the amounts of variance accounted is 24% when CMV is controlled (i.e., 35%) versus when it is not controlled (i.e., 11%). Thus, CMV can have a substantial effect on the relationships between measures or constructs (Podsakoff et al., 2003).

A series of steps can be taken to minimize the CMV bias. Theoretically, one can use research to develop a systematic questionnaire and measures (items) to form the constructs, which can be further refined statistically using exploratory factor analysis and reliability measures. It is good practice to avoid unfamiliar words, double-barrelled questions and technical words and to keep items simple, specific and concise. The items could be further grouped with different construct items (i.e., not in conceptual dimensions) (Tourangeau, Rips, & Rasinski, 2000). Some researchers also suggest to avoid adding (many) negatively-worded items because of a lack of confidence in respondents' ability to fully understand them, as highlighted by Podsakoff et al. (2003). Researchers often have to delete such items because their loadings are not strong enough to meet the minimum criterion. In addition, respondents should be informed of the anonymity of the survey - individuals and companies should not be identified and only aggregate data used. Moreover, to avoid a single-informant bias, data could be collected from multiple informants. For example, a survey data collection may involve multiple management positions such as chief executive officers, managing directors, project managers, marketing managers, senior operations managers and team leaders (e.g., Akhtar, Tse, Khan, & Rao-Nicholson, 2016).

In order to test for CMV, researchers commonly use Harman's one-factor test (Malhotra et al., 2006; Podsakoff et al., 2003). In this method, the analysis produced from multiple factors (based on eigenvalues greater than 1 and scree plot observations) with reasonable variances is compared to a single factor solution or other combinations. However, this test is insensitive,

and as such it is insufficient test to rule out the potential existence of common method bias (Podsakoff et al., 2003). Although all statistical approaches to control for CMV bias have their particular advantages and disadvantages (Malhotra et al., 2006; Podsakoff et al., 2003), it is also useful to use the marker variable technique (e.g., the number of languages that respondents speak, as a marker variable) proposed by Lindell and Whitney (2001), which is a good alternative to assess the CMV bias. Additionally, the latent factor approach can be used for assessing CMV (see Malhotra et al., 2006). It is a good practice to use these mentioned multiple remedies to minimize possible concerns. This leads to employ different methods and one can follow a rigorous statistical procedure by using these techniques to deal with CMV.

### *2.2 Measurement errors*

Measurement error is a common problem in marketing, management, business and other social science research. This is because the constructs of interest cannot be measured perfectly, as researchers can do in natural sciences. Consequently, the estimates are inconsistent and the error affects other variables involved (Antonakis et al., 2010; Antonakis et al., 2014; DeShon, 1998).

Although structural equation modelling analysis (e.g., maximum likelihood estimate) does correct for the biasing effects of measurement errors (Frone, Russell, & Cooper, 1994) or correct for the small amount of measurement errors (DeShon, 1998), researchers still need to control for measurement errors when they use a single indicator approach, that is parcelling using multi-item scales (DeShon, 1998). For example, if researchers use parcelling (averaging the relevant items) for environmental and financial (performance) constructs, they should be corrected for the random measurement error by constraining the relevant random error variance equal to the product of the variance multiply by one minus the reliability. The relevant loadings (i.e., SD * square-root of alpha) for the parcels are also fixed (Antonakis et al., 2014; Bollen,

1989; DeShon, 1998). By controlling for the errors, besides minor changes in significance levels, researchers can find that the difference between the amounts of variance accounted is improved when measurement error is controlled versus when it is not controlled. This indicates the importance of correcting measurement errors and following a methodological rigorous approach.

### *2.3 Omitted bias and simultaneity*

Omitted bias results in various forms (See for details Antonakis et al., 2010; Antonakis et al., 2014). One possibility of omitted bias can be when researchers test the validity of a construct without including other important variables/constructs. For instance, one measures operational performance of supply chain partners without including social performance among them. In this regard, Anonakis and Dietz (2011), Antonakis et al. (2014), and Ketoviki and McIntosh (2017) all emphasize the comprehensive integration of theories. It is also useful to use the multiple dimensions of constructs, which themselves may consist of sub-constructs (e.g., service quality and product quality forming operational construct; trust in and satisfaction with supply chain partners representing social construct) built based on extensive theories. After this, statistical analysis such as exploratory factor analysis, reliability and validity tests can be applied to further refine them.

The problem of simultaneity occurs when two variables simultaneously affect/cause each other and have reciprocal feedback loops (non-recursive models). Though the problem is easy to understand, it is complicated to resolve statistically, particularly when a study involves multiple constructs. The problem may be addressed using instrumental variables. Zaefarian et al. (2017) provide details of how to use instrumental variables by utilizing two-stage or three stage least square approaches. However, finding instrumental variables for a number of constructs is not easy, sometimes even it is impossible (Antonakis et al., 2010; Antonakis et al., 2014). The

fundamental problem for such models in structural models is the identification of the models due to numerous complications such as not having sufficient data, high correlations of instrumental variables and deficient order and rank conditions (Bentler & Chou, 1987; Martens & Haase, 2006). It is a good academic practice to acknowledge that alternative models may exist, which are equally possible on logical grounds. Researchers should use strong theoretical arguments to build the directions of hypotheses and models. Interested readers may also want to consult available studies on the topic (e.g., Bentler & Chou, 1987; Martens & Haase, 2006).

Cumulatively, by investigating the datasets in detail and addressing different types of endogeneity bias from research designs to statistical analysis, researchers can provide evidence that they have made efforts to follow a rigorous theoretical procedure and comprehensively investigated datasets to comply with psychometric properties and 'urban legends'. As there is no hard and fast rule or a yes or no answer for such types of bias and urban rules, that is why some of the 'urban legends' (the cut-off criteria) have recently been criticized. For example, Lance, Butts, and Michels (2006) unambiguously debunk the 0.70 rule for reliability. This is not discussed further here, and we recommend readers to consult recent methodological advances in the area (e.g., Cortina, 2002; Lance, 2011; Lance et al., 2006).

Returning to the main focus of this study, a series of techniques to deal with endogeneity are succinctly provided by researchers (Guide & Ketokivi, 2015; Ketokivi & McIntosh, 2017; Zaefarian et al., 2017). Our purpose is to provide an in-depth step-by-step procedure as to how the general dynamic generalized method of moments (GMM) model can be used to address endogeneity in panel data. For other types of data and understanding concepts and relevant techniques (e.g., cross-sectional and structural equation modelling), researchers can consult recent studies (e.g., Ketokivi & McIntosh, 2017; Zaefarian et al., 2017). A systematic procedure for handling endogeneity in panel data is provided in the next section.

### *2.4 GMM models and dealing with endogeneity in panel data*

The emergence of GMM models in marketing research when using panel data can resolve some of unanswered questions raised in the recent literature when discussing econometric techniques. Arellano and Bond (1991) and Blundell and Bond (1998) developed the generalized method of moments model, which can be used for dynamic panel data. In dynamic panel data, the cause and effect relationship for underlying phenomena is generally dynamic over time. For example, it may not be the current year's marketing expenses that are affecting performance, but rather the previous year's expenses that could be playing a significant role.

To capture this, dynamic panel data estimation techniques use lags of the dependent variables as explanatory variables. Lagged values of the dependent variables are therefore used as instruments to control this endogenous relationship. These instruments are often called 'internal instruments' as they are used from the existing econometric model (Roodman, 2009). The GMM model, which is generally used for panel data, provides consistent results in the presence of different sources of endogeneity, namely "*unobserved heterogeneity, simultaneity and dynamic endogeneity*" (Wintoki, Linck, & Netter, 2012, p. 588). Traditionally, researchers (Schultz et al., 2010; Wintoki et al., 2012) have used two lags of the dependent variables and they argue that two lags are sufficient to capture the persistence of the dependent variable (say for example firm performance).

The GMM model removes endogeneity by "*internally transforming the data*" – transformation refers to a statistical process where a variable's past value is subtracted from its present value (Roodman, 2009, p. 86). In this way, the number of observations is reduced and this process (internal transformation) enhances the efficiency of the GMM model (Wooldridge, 2012). Furthermore, two kinds of transformation methods, known as first-difference transformation (one-step GMM) and second-order transformation (two-step GMM), can also be used as GMM

estimators. However, the first-difference transformation (one-step GMM) has some limitations. For instance, if a variable's recent value is missing, then the first-difference transformation (where a variable's past value is deducted from its current value) could result in the loss of too many observations (Roodman, 2009). To avoid potential data loss owing to the internal transformation problem with the first-step GMM, Arellano and Bover (1995) recommended the use of a second order transformation (two-step GMM).

The second-order transformation (two-step GMM) applies '*forward orthogonal deviations'*, which means that instead of subtracting the previous observations of a variable from its current value, the two-step GMM model subtracts the average of all future available observations of a particular variable (Roodman, 2009, p. 86). Using a two-step GMM model, researchers can prevent unnecessary data loss. Therefore, in the case of a balanced panel dataset, a two-step GMM model provides more efficient and consistent estimates for the involved coefficients (Arellano and Bover 1995

Following Arellano and Bond (1991) and Blundell and Bond (1998), we use the following general dynamic generalized method of moments model (GMM):

$$P_{it} = \partial P_{i,t-1} + RDSALES_{it} + G\beta_{it} + Xn_{it} + \mu_{it} + \varepsilon_{it} \quad (1)$$

Where: $P_{it}$ denotes firm performance, which in our example is measured in terms of return on assets (ROA), $P_{i,t-1}$ is a one period lag operator (previous year firm performance); $RDSALES_{it}$ represents our main explanatory variable, R&D expenditures; $G\beta_{it}$ represents corporate governance variables; $Xn_{it}$ represents control variables over the time period; $\mu_{it}$ is firm-specific fixed effects; and $\varepsilon_{it}$ represents the error term.

In this paper, our methodical demonstration and discussion are relating to the impact of R&D expenditures, firm-level governance and financial characteristics on the return on assets (ROA)

as an indicator of firm's financial performance. The birth of a corporate form of business organization resulted in a separation of ownership and control (Berle and Means, 1932), with managers having sufficient control in running the day-to-day affairs of an organization. It is unlikely that managers would always act in the best interest of the owners. This phenomenon was first pointed out by Adam Smith in 1776 and was formally presented in the form of agency theory by Jensen and Meckling (1976). According to Jensen and Meckling (1976), the opportunistic behavior by managers would cause a conflict of interests between owners and managers, thereby negatively affecting a firm's financial performance. Jensen and Meckling (1976) also suggested that incentive mechanisms (executive compensation) and effective control (corporate governance) mechanisms are likely to re-align the interests of owners and managers. The board of directors thus play an important role in aligning the interests of owners and managers.

We include a number of firm-level governance (monitoring) mechanisms and incentives paid to senior directors to test how these mechanisms affect firms' performance, in addition to other firms' specific characteristics, such assize, sales growth and R&D expenditures. The literature uses several proxies for strong internal governance that enhances firms' performance (Beiner, Drobetz & Schmid, 2006). Some attributes of board of directors that represent strong governance and monitoring include: smaller board size, highest percentage of non-executive directors, appropriate executive compensation, a highest percentage of gender diversity, and a higher number of board meetings. We therefore include these proxy measures to demonstrate the application of GMM in panel data research.

Roodman (2009) presented assumptions that need to be fulfilled when employing GMM estimations, namely (a) some regressors may be endogenously determined; (b) the nature of the relationship is dynamic, implying that current performance is affected by previous ones; (c) the idiosyncratic disturbances are uncorrelated across individual; (d) some regressors may not

necessarily be strictly exogenous; and finally, (e) the time periods in panel data, T, may be small. (i.e., "small T, large N."). The inclusion of lag performance variables changes the static nature of this econometric model to a dynamic panel data model. Two-step system GMM relies on internal instruments (lagged values, internal transformation) to address the different sources of endogeneity discussed in the literature review section.

## 3. Method

### *3.1 Data and Sample Selection Procedures*

Researchers in the *Journal of Marketing*, and *Journal of Marketing Research* have been using panel data for many years (see for example, McAlister, Srinivasan, & Kim, 2007; Nierop Bronnenberg, Paap, Wedel, & Franses. 2010; Ma, Ailawadi, Gauri, & Grewal, 2011; Sridhar, Germann, Kang, & Grewal, 2016; Han, Mittal & Zhang, 2017), while IMM researchers are still predominantly relying on the use of survey data. The emergence of business/financial databases (Bloomberg, Datastream, Thomson One, Compustat) provides more flexibility to marketing researchers to effectively utilize the marketing and product related data available in these databases. As panel datasets combine the characteristics of cross-sectional and time series data, we believe that the IMM community could benefit from the marketing related data available in these databases.

For example, corporate investment and financial data for US and non-US companies can be collected from Compustat. The Thompson Financial's Securities Data Corporation (SDC) database provide detailed information on different types of joint ventures. This database compiles information from different publicly available sources. Datastream includes global data on patents, brand value, number of consumer controversies directly linked to companies' products or services, product recall, eco-design products, energy footprint reduction, organic-product initiatives, among others. If IMM community could utilize such massive global

marketing datasets for listed companies, they could capture interesting insights about the on-going marketing management issues around the world. The companies associated with these databases put enormous efforts in collecting and reporting this data that is largely unexplored, particularly in the context of industrial marketing. Furthermore, using a panel data approach, researchers can use more observations and panel data has the capability to control for unobserved heterogeneity — a potential source of endogeneity discussed in the literature.

To demonstrate a step-by-step and rigorous procedure, we first collected the data of UK listed companies (FTSE ALL Share index) – available on the Datastream. The data included 15-year observations, ranging from 2002 to 2016. We included corporate governance data in our analysis, although this data is only available from 2002 onwards in Datastream. Our final sample includes data from the 101 UK listed companies, after excluding companies which were subsequently delisted and whose financial data was not available in the database. Where financial data was not available for the sample, we collected such data from the alternative databases (i.e., Bloomberg). Table 1 shows the industrial classification of our sample firms. The chosen 101 sample firms belong to 25 different industrial sectors, with over 1500 observations. Table 2 includes the definitions of these variables and relevant codes available for downloading data from the Datastream, which is a powerful database and it allows to scrutinise patterns, generate and test research questions and develop research on market trends.

---

Insert Table 1 about here

---

### *3.2 Variables and descriptive statistics*

Table 2 includes a list of variables used in this research. The explanatory variables are R&D expenditure, corporate governance variables including board size, the percentage of independent non-executive directors, the number of board meetings, gender diversity, CEO-Chairman duality, total senior executive compensation and the percentage of shares owned by

the single largest/biggest shareholder/owner. We included firm-level governance variables, as prior research shows that strong corporate governance (monitoring) have implications on the performance of firms (Beiner et al., 2006). The dependent variable in our model is firm performance, which is measured by return on assets (ROA). We also control for firm-specific characteristics including debt financing, firm size and sales growth.

---
Insert Table 2 about here
---

Table 3 presents descriptive statistics. The mean value for our key explanatory variable R&D expenditure is 4% of sales, with a maximum value of 50% of sales. This provides the insights about the level of R&D investment by the UK listed companies. The average board size of the UK listed companies is 9.07 and the maximum board size is 18. The average value for the percentage of non-executive directors (NEDs) in the UK companies is 65.26%, with a maximum number of 88.69% NEDs appointed by the UK companies. Except for smaller companies which are not constituent members of the FTSE 350 index in the UK, corporate regulations in the UK require at least 50% non-executive directors for companies listed on the London Stock Exchange (See regulation B.2.1, in the, Financial Reporting Council, 2016, p.11). A higher percentage of external non-executive directors implies strong monitoring at board levels, which could also be observed in corporate financial performance.

Number of board meetings is another monitoring mechanism, affecting corporate financial performance. The average number of board meetings (NBS) is 9 and the maximum value is 22. As companies are different in their size and structure, such variations in the number of board meetings are expected. The mean value for the percentage of women on the corporate boards of UK listed companies is 9% and the maximum value is 50%. This represents gender diversity that is another governance (monitoring) mechanism having implications for corporate financial performance. In this regard, the Lord Davies Report in the UK, entitled, '*Improving the Gender*

*Balance on British Boards'* sets out a 33% women quota for companies listed on the London Stock Exchange by the end of 2020 (Davies, 2011).

The average value for DUAL is 0.13, suggesting that 13% of the CEOs in UK listed companies hold the positions of CEO and Chairmanship simultaneously, while the regulations require splitting these role as one person may get too much powerful at the top. Hence, such dual positions may negatively affect corporate financial performance. The mean value for the logarithm of total senior executive compensation (TSEC) is 6.68 and the maximum value is 9.13. In terms of large shareholdings, the mean value for single largest blockholder ownership (SBO) is 14.06%. Generally, large investors are considered effective monitors compared to small and dispersed shareholders.

On average, 23% debt financing activities are carried out by UK companies during the reported period. Debt financing (LEVER) is also a monitoring mechanism as a lender of finance keep strict surveillance over their investee companies (Beiner et al., 2006), and hence we expect its implications on corporate performance. Firm size (SIZE) is measured by the natural logarithm of total assets and the mean value for the size is 14.28 and the maximum value is 19.71. The descriptive statistics indicate that the average sales growth of sample UK companies is 10% for the reporting period of 15 years, 2002-2016. Finally, the operating performance measure return on assets (ROA) shows that the mean ROA for the UK companies is 14% and the maximum value is 50%.

---

Insert Table 3 about here

---

## 4. Steps used in the estimation process (demonstration)

In the following section, we use a step-by-step procedure to demonstrate how GMM offers robust estimates compared to OLS and fixed-effects estimates. We first start with OLS analysis

and identify endogeneity issues by utilizing Durbin-Wu-Hausman test, followed by a fixed-effects model. The procedure then demonstrates that fixed-effects fail to capture dynamic endogeneity. The GMM model finally incorporates lagged-values of the dependent variable (previous year's financial performance). Thereby, the endogeneity concerns are addressed and the valid estimates are produced by using a rigorous GMM process.

***Step1 –Basic OLS analysis***

Owing to its wide usage in prior research, initially, an OLS analysis was carried out to examine the direct effect of our independent variables (i.e. R&D expenditures, corporate governance, and firms' characteristics) on our dependent variable, ROA, and the results are reported under model (1) column in Table 4. However, following Schultz et al. (2010) and Wintoki et al. (2012), before interpreting the results from OLS regression, a test for endogeneity was carried out to determine whether the results reported under the OLS models are consistent (see Step 2). Our main results in the OLS model shows that R&D has a positive effect on the performance of firms. This finding is consistent with the recent research of Ehie & Olibe (2010), who reported a positive relationship between investment in R&D expenditures and the market valuation of firms for a sample of 26,500 firm-years observations for a period of 18 years. In model (1), majority of the governance variables, BSIZE, TSEC, GD and SBO have a positive impact on the performance of firms, which suggests that strong firm-level (internal) corporate governance mechanisms can improve the performance of firms. This is in line with the assumptions of agency theory that strong firm-level monitoring mechanisms are likely to have positive implications on corporate performance (Beiner et al., 2006).

***Step 2– Detecting endogeneity bias***

The Durbin-Wu-Hausman test is commonly used to detect endogeneity of individual regressors. Theoretically, the explanatory variable on the right-hand side should be uncorrelated with the

error term, and this test determines whether the residuals (error term) are correlated with the explanatory variable. A Durbin-Wu-Hausman test is thus used to detect endogeneity in the OLS regression. For the general understanding of IMM readers, we comprehensively demonstrate the procedures used to test the endogeneity and illustrate it by using an example of an explanatory variable, RDSALES. To do that, we follow the procedures suggested by Beiner et al. (2006), Schultz et al. (2010) and Wintoki et al. (2012). The following steps were implemented to carry out the Durbin–Wu–Hausman test.

a. To test whether an independent variable, for example, investment in R&D expenditures (e.g., RDSALES) is endogenous or exogenous, a regression was estimated on each independent variable with all other independent variables and control variables to predict the relevant residuals.

Example:

$$RDSALES_{it} = BSIZE_{it} + NEDs_{it} + NBM_{it} + GD_{it} + DUAL_{it} + TSEC_{it} + SBO_{it} + LEVER_{it} + SIZE_{it} + SG_{it} + \varepsilon_{it} \quad (2)$$

In our standard model, the dependent variable is ROA (see equation 1, $P_{it}$ ). To test for endogeneity/exogeneity of RDSALES variable, this variable is now included as a dependent variable rather than as an exploratory variable. This is the first step to conduct Durbin-Wu-Hausman test.

b. In the second step, the coefficients for the residuals were estimated to test whether the residuals (error terms, $\varepsilon_{it}$ ) are significant. The null hypothesis states that investment in R&D expenditures and corporate governance mechanisms are exogenous, implying they are uncorrelated with the residuals. For each individual explanatory variable in the model (e.g., RDSALES), we estimated residuals using the generic STATA command, '*predict new-*

*variable residuals'*. The residuals for RDSALES were then included in our basic OLS model, which takes the following form:

$$ROA_{it} = RDSALES_{it} + BSIZE_{it} + NEDs_{it} + NBM_{it} + GD_{it} + DUAL_{it} + TSEC_{it} + SBO_{it} + LEVER_{it} + SIZE_{it} + SG_{it} + \varepsilon_{it} \quad (3)$$

c. A significant test statistic of Durbin-Wu–Hausman test for an explanatory variable indicates that the variable is endogenous – the explanatory variable is correlated with the residuals (error term).[1]

If a single variable in the econometric specification is endogenous, obviously, researchers need to implement a superior estimation technique that provides consistent estimates than OLS. In the next step, we identify a list of endogenous variables and we discuss the implications of endogeneity issues, and suggest the use of fixed-effects estimation.

***Step 3 – Understanding the nature of endogenous variables***

The Durbin-Wu-Hausman test statistics shows that majority variables in Model 1 (e.g., BSIZE, NEDs, DUAL, LEVER, SIZE, SG) are endogenously determined. From an econometrics perspective, and theoretically, it makes sense that some of the corporate governance mechanisms and investment in R&D could be endogenously determined. For example, a firm with poor performance in one year may change its board size or the percentage of non-executive directors in the following year (Beiner et al., 2006; Schultz et al., 2010; Wintoki et al., 2012). Similarly, poorly performing firms are likely to take more risks in the following years (Bromiley, 1991). In addition, firms with higher market valuations may choose to invest more in R&D expenditures in subsequent years (Gupta, Banerjee, & Onur, 2017). Overall, the results from Durbin-Wu Hausman test statistics suggest that endogeneity is a major problem in

[1] The procedures to carry out a Durbin-Wu-Hausman test in any version of the STATA is explained at https://www.stata.com/support/faqs/statistics/durbin-wu-hausman-test/

our OLS model. If only one variable is endogenous in a regression model, the results reported from OLS are inconsistent (Beiner et al., 2006; Schultz et al., 2010; Wintoki et al., 2012). This implies that the results reported from OLS are inconsistent because of the endogeneity issues.

**Step 4 –** ***Applying Fixed-effects estimation***

In Model 2 (Table 4), we employ a fixed-effects estimation technique, which can potentially control for unobservable heterogeneity under the assumption of strict exogeneity. Strict endogeneity means that a firm's current governance mechanisms, investment in R&D expenditures (independent variables) are not affected by any changes in a firm past and present financial performance (e.g., dependent variable ROA) (Schultz et al., 2010; Wintoki et al., 2012).

However, in reality, this assumption of strict exogeneity is violated because a firm's past/current performance may affect the current/future governance structure of a firm. In a fixed-effects model, firm-specific fixed effects are incorporated in the econometric model by either including a set of firm-specific indicator variables into the regression, or by internally transforming (differencing) to eliminate the time invariant components (Hamilton and Nickerson, 2003). Internal transformation (first-differencing) is a procedure in which each firm's relative observation in the period t-1 is subtracted from the corresponding observation in period t, for example $(\Delta RDSALES_{it} = RDSALES_{it} - RDSALES_{it-1})$. This process eliminates the time-invariant industry- and firm-level unmeasured variables from the right-hand side of the regression equation.

In qualitative terms, fixed effects models help in controlling unobserved heterogeneity, which is 'constant' over time and is also correlated with the explanatory variables. This constant/time invariant component of the model is usually removed from internal transformation. A general example of unobserved time-invariant individual effects is managerial capabilities and firm-

level institutional quality which cannot be captured in the econometric specification. The time invariant error term is treated as fixed effects (Gujarati, 1999). Furthermore, fixed-effects estimation is a static panel data model, which means it does not allow for the lag of the dependent variables (firm financial performance) to be included as an explanatory variable in the econometric model (Wooldridge, 2012). An ordinary least squares regression model fails to control for unobserved heterogeneity and the fixed-effects or random-effects models could potentially overcome this problem. However, fixed-effects estimation is employed to deal with endogeneity in circumstances where firm-specific characteristics (time invariant) are correlated with the explanatory variable (Wintoki et al., 2012).

In our example, the relationship between our explanatory variables and firm performance is dynamic — past realization of dependent variables (performance) may also affect current year performance. Simultaneity could potentially violate the strict exogeneity assumptions (Schultz et al. 2010). For example, in a fixed effect model, the relationship between R&D, corporate governance, control variables and operating performance of the firms may be determined simultaneously — investment in R&D and governance structure could be determined based on the expected corporate performance in the same year. Consequently, traditional fixed-effects and random-effects panel data static models would provide inconsistent and biased results (Wooldridge, 2012). As a GMM is more robust in dealing with these sources of endogeneity, we carried out a final check with the two-step system GMM. An important difference between fixed-effects and GMM is that fixed-effects estimation uses 'strict exogeneity' assumptions resulting into a static fixed-effects model in the following form:

*Performance (ROA) = f (RDSALES, Corporate governance, firm-level characteristics, and fixed effects).*[2]

[2] The following generic STATA command can be executed to run a fixed-effects model: *xtreg depvar indepvars, fe.*

We suggest that an appropriate econometric model should be a "dynamic" model of the form:

*Performance (ROA) = f (Past performance/lag ROA, RDSALES, Corporate governance, firm-level characteristics, and fixed effects).*

By including lagged values of past performance, we differentiated between a 'dynamic' and a 'static' panel data model. A depiction of the endogenous relationships is also presented in Figure 1.

---
Insert Figure 1 about here
---

Based on the existing literature about endogeneity, Figure 1 indicates how the impact of investment in R&D expenditure, corporate governance characteristics, and firm-level characteristics on the financial performance could be endogenously determined. Larger firms may choose to invest more in R&D expenditures. Similarly, firms with poor financial performance (ROA) in one year may also choose different governance arrangements in subsequent years. Furthermore, poor financial performance may also affect the likelihood/percentage of investment in R&D expenditures. Therefore, this complex and dynamic relationship could not be captured through 'static' OLS/fixed-effects models. In applying a fixed-effects model, time-invariant explanatory variables (e.g., the industry classification of a firm) are not included, and failing to estimate such time-constant variables, the fixed-effects estimator has been criticized for 'wasting' very relevant information from the econometric specification (Owusu-Gyapong, 1986)

In the next step, we explain how the two-step GMM system could address these endogeneity concerns.

'xtreg' fits regression models to panel data. 'Depvar' indicates the dependent variable (ROA in our case), indepvar represents independent variables, and 'fe' implies the fixed-effects option.

***Step 5 – Two-Step System of GMM and Comparisons***

In model 3 (Table 4), we use a dynamic panel data estimation to overcome the endogeneity issues arising from reverse causality.

Our two-step system GMM model is presented in the following equation:

$$ROA_{it} = ROA_{it-1} + ROA_{it-2} + RDSALES_{it} + BSIZE_{it} + NEDs_{it} + NBM_{it} + GD_{it} + DUAL_{it} + TSEC_{it} + SBO_{it} + LEVER_{it} + SIZE_{it} + SG_{it} + + \mu_{it} + \varepsilon_{it} \quad (4)$$

The definitions for all explanatory variables are presented in Table 2. $ROA_{it-1}$ indicates one lag of the dependent variable ROA (previous year performance), and $ROA_{it-2}$ denotes second lag of the dependent variable, representing performance in the year before previous year. These lags are included as explanatory variables in our GMM estimation.

The GMM model controls for endogeneity by internally transforming the data and by including lagged values of the dependent variable. In this way, the GMM model offers a superior estimation technique compared to the OLS model. In the next step, we report a revised analysis of the governance-performance relationship, using a GMM method. As the GMM model control for endogeneity and includes lagged values and applies internal transformation process, the results reported under the GMM could be significantly different than those reported in the OLS column (Table 4, Model 1). For instance, using an OLS approach, Schultz et al. (2010) find a significantly negative relationship between executive remuneration and the performance of Australian firms. However, after controlling for unobserved heterogeneity, simultaneity and dynamic endogeneity by using the GMM approach, Schultz et al. (2010) did not find any significant impact of executive remuneration on the performance of firms.

We find that the relationships of RDSALES, TSEC and SG with the operating performance of firms is consistent using OLS, fixed-effects and system GMM (no change in significance levels for all three models – see italic-values in Table 4). The impact of the remaining explanatory

variables changed significantly (or significant levels changed even some of them turned out to be significant or insignificant) when we use static panel data model (fixed effects) and dynamic panel model (GMM), which captured the lagged values of previous two years' financial performance. For example, with regard to control variable, firm size (SIZE) has a negative relation with ROA in all three models, which is consistent with the argument that larger firms have higher operating costs (Beiner et al. 2006). However, it showed the changes in the significance levels (the fixed model shows that it is significant at $p = 0.1$ while other two models show it is significant at $p = 0.01$). Similarly, the variable NEDs showed an insignificant relationship due to endogeneity in the OLS model, in fact it is significantly associated as demonstrated by the GMM model (with $p = 0.000213$***). The Duality (DUAL) shows a negative relationship in the OLS model but it is positively and significantly associated as demonstrated by other two models. Table 4 shows other similar examples highlighted with bold text.

---

Insert Table 4 about here

---

In short, the GMM controlled for different kinds of endogeneity by including previous financial performance (lagged values of the dependent variable ROA) as an explanatory variable in the model. The GMM model controlled for the three major sources of endogeneity: (i) unobserved heterogeneity; (ii) simultaneity; (iii) dynamic endogeneity. The emergence of the GMM technique could be considered as the new methodological development in business research, after it was recently employed by Wintoki et al. (2012) in governance research. The nature of the data (panel data) and the dynamic nature of the governance-performance relationship suggest that a GMM model offers more efficient and consistent estimates for the coefficients as compared with other estimation techniques.

When applying the generalized method of moments model, researchers need to apply two post-estimation tests to determine that an appropriate econometric model is applied. These tests are: (i) the Sargan test; and (ii) the Arellano-Bond test for first-order and second-order correlation. A critical assumption for the validity of GMM estimates requires that instruments are exogenous. In other words, the findings from GMM will not be valid if the instruments are endogenously determined. The Sargan test is used to determine whether the econometric model is valid or not, and whether the instruments are correctly specified or not. In other words, if the null hypothesis is rejected, the researcher needs to reconsider the model or the instruments used in the estimation process. The post-estimation (Sargan test) can be executed in STATA using *estat sargan* command. Subsequently, if the Sargan test turns out to be insignificant it implies that the instruments included in the econometric specifications are exogenous.

To examine the validity of a strong exogeneity assumption, the Arellano-Bond test for no auto-correlation (or no serial correlation) is used under the null hypothesis that the error terms of two different time periods are uncorrelated. In other words, it means that the lagged variables are not correlated with the error term in the governance-performance equation. To execute this post-estimation test in STATA, the user needs to employ the *estat abond* command. The values for these two post-estimation tests are reported in Table 4, which confirms the validity of the instruments (model) used in our estimation process.

## 5. Implementing GMM in STATA

This study uses STATA software to execute a generalized method of moments (GMM) model to deal with endogeneity, showing how this robust technique can control for different kinds of endogeneity issues and thus providing unbiased estimates. However, it may be complicated to implement in STATA, as Roodman (2009) argues that such analysis is complicated and can easily mislead researchers. "*Implementing them with a Stata command stuffs them into a black*

*box, creating the risk that users, not understanding the estimators' purpose, designs and limitations, will unwittingly misuse it*" (Roodman, 2009, p. 87). This is another reason that our step-by-step procedure can help researchers to better understand endogeneity sources and address them accordingly.

For implementing the GMM operations in STATA, users can apply a number of inbuilt commands e.g., *xtabond, xtdpd, and xtdpdsys* commands. The STATA syntax can be used to apply the Arellano–Bover/Blundell–Bond's dynamic panel data estimator. For a sample dataset with a dependent variable y and explanatory variables x1, x2, x3, and x4, Roodman (2009) developed the following *xtabond2*[3] STATA codes for dynamic panel data model

*xtabond2 y l.y x1 x2, x3, x4 gmm(y x1 x2, x3, x4 lag(a b)) noconstant twostep* (5)

Lagged values *(l.y)* in the above two step dynamic model are included as regressors. These lagged levels of the dependent variable in the Arellano–Bond's (1991) estimator are used as instruments to deal with endogeneity. Also, *lag(a b)* means the number of lags that researchers wish to include in the model. Owing to the internal transformation process, the numbers of observations are reduced by using the system GMM. In Appendix 1, we report generic STATA commands for implementing GMM in STATA and applying the post-estimation tests.

## 6. Conclusions

Endogeneity bias is an emerging issue in marketing and management science. Different sources of endogeneity in panel data could generate bias and inconsistent estimates. We used panel data for R&D expenditures, governance and financial performance of 101 UK listed companies over a period of 15 years from 2002 to 2016 as an example to empirically demonstrate how endogeneity can be addressed in panel data using generalized method of moments.

[3] STATA users can install this command in STATA using the following code: *ssc install xtabond2*

We explained step-by-step procedures that can be used to deal with endogeneity bias in panel data. Succinctly, OLS regression provides inconsistent results in the presence of any source of endogeneity bias. This is confirmed using Durbin-Wu Hausman test. To overcome unobserved heterogeneity and to capture firm-specific effects, we employed the fixed effects estimation approach, which removes unobserved heterogeneity by internally transforming the data. This means that a variable (e.g., RDSALES) value in previous year is subtracted from its value in current year: $\Delta RDSALES_{it} = RDSALES_{it} - RDSALES_{it-1}$.

However, fixed-effects model is also known as a '*static panel data model',* which implies that implementing this model does not allow to use lagged values of the dependent variable. The dynamic panel data GMM model further extends the fixed-effects model, and, in addition to the internal transformation process, the lagged values of the dependent variable are also included as instruments to control for dynamic endogeneity. Thereby, we systematically move from basic OLS estimation to more sophisticated econometric techniques to control the different sources of endogeneity bias. We also provided user-friendly generic STATA commands that can be used by marketing researchers when they deal with marketing related panel datasets.

Finally, we briefly introduced a number of potential databases (Bloomberg, Datastream, Thomson One and Compustat) that can be utilized by IMM researchers in extracting recent marketing related global datasets on advertising expenditures, patents, brand values, number of consumer controversies directly linked to companies' products or services, product recall, eco-design products, energy-footprint reduction and organic-product initiatives. While the quantitative side of the IMM community has predominantly relied on surveys, we believe that these databases would potentially overcome some of the data collection challenges faced by marketing researchers.

**References**


Akhtar, P., Tse, Y.K., Khan, Z., & Rao-Nicholson, R. 2016. Data-driven and adaptive leadership contributing to sustainability: Global agri-food supply chains connected with emerging markets. *International Journal of Production Economics*, 181(Part B), 392-401.

Antonakis, J., Bendahan, S., Jacquart, P., & Lalive, R. (2010). On making causal claims: A review and recommendations. *The Leadership Quarterly*, 21(6), 1086-1120.

Antonakis, J., Bendahan, S., & Lalive, R. (2014). Causality and endogeneity: Problems and solutions The Oxford handbook of leadership and organizations.

Antonakis, J., & Dietz, J. (2011). More on testing for validity instead of looking for it. *Personality and Individual Differences*, 50(3), 418-421.

Arellano, M. & Bond, S. (1991). Some tests of specification for panel data: Monte Carlo evidence and an application to employment equations. *Review of Economic Studies*, 58(2), 277–297.

Arellano, M. & Bover, O. (1995). Another look at the instrumental variable estimation of error-components models. *Journal of Econometrics*, 68(1), 29–51.

Beiner, S., Drobetz, W., Schmid, M.M. & Zimmermann, H. (2006). An integrated framework of corporate governance and firm valuation. *European Financial Management*, 12(2), 249-283.

Berle, A. A. & Means, G. C. (1932). *The Modern Corporation and Private Property*, New York, Macmillan Publishing Co.

Bentler, P. M., & Chou, C.P. (1987). Practical issues in structural modeling. *Sociological Methods & Research*, 16(1), 78-117.

Bollen, K. A. (1989). Structural equations with latent variables: John Wiley & Sons.

Blundell, R. & Bond, S. (1998). Initial conditions and moment restrictions in dynamic panel data models. *Journal of Econometrics*, 87(1), 115–143.

Bromiley, P. (1991). Testing a causal model of corporate risk taking and performance. *Academy of Management Journal*, 34(1), 37–59.

Cortina, J. M. (2002). Big things have small beginnings: An assortment of "minor" methodological misunderstandings. *Journal of Management*, 28(3), 339-362.

Davies, A. (2011). Women on boards February 2011. Department for Business, Innovations and Skills. URL: https://www.gov.uk/government/uploads/system/uploads/attachment_data/file/31480/11-745-women-on-boards.pdf

DeShon, R. P. (1998). A cautionary note on measurement error corrections in structural equation models. *Psychological Methods,* 3(4), 412-423.

Ehie, I.C. and Olibe, K., (2010). The effect of R&D investment on firm value: An examination of US manufacturing and service industries. *International Journal of Production Economics*, 128(1), pp.127-135.

Financial reporting Council. (2016). The UK corporate governance code. Available at URL: https://www.frc.org.uk/getattachment/ca7e94c4-b9a9-49e2-a824-ad76a322873c/UK-Corporate-Governance-Code-April-2016.pdf (accessed 28 September 2017)

Frone, M. R., Russell, M., & Cooper, M. L. (1994). Relationship between job and family satisfaction: Causal or noncausal covariation? *Journal of Management*, 20(3), 565-579.

Guide, V. D. R., & Ketokivi, M. (2015). Notes from the editors: Redefining some methodological criteria for the journal. *Journal of Operations Management* (37), v-viii.

Gujarati, D.N. (2009). Basic Econometrics. Tata McGraw-Hill Education.

Gupta, K., Banerjee, R. & Onur, I. (2017). The effects of R&D and competition on firm value: International evidence. *International Review of Economics & Finance*. 51,391–404

Hamilton, B. H., & Nickerson, J. A. (2003). Correcting for endogeneity in strategic management research. *Strategic organization*, 1(1), 51-78.

Han, K., Mittal, V. & Zhang, Y. (2017). Relative strategic emphasis and firm idiosyncratic risk: The moderating role of relative performance and demand instability. *Journal of Marketing*. 81(4), 25-44.

Jensen, M. C. & Meckling, W. H. (1976). Theory of the firm: Managerial behavior, agency costs and ownership structure. *Journal of Financial Economics*, 3(4), 305–360.

Ketokivi, M., & McIntosh, C. N. (2017). Addressing the endogeneity dilemma in operations management research: Theoretical, empirical, and pragmatic considerations. *Journal of Operations Management,* (52), 1-14.

Lance, C. E. (2011). More statistical and methodological myths and urban legends. Organizational *Research Methods*, 14(2), 279-286.

Lance, C. E., Butts, M. M., & Michels, L. C. (2006). The sources of four commonly reported cutoff criteria what did they really say? *Organizational Research Methods*, 9(2), 202-220.

Lindell, M. K., & Whitney, D. J. (2001). Accounting for common method variance in cross-sectional research designs. *Journal of Applied Psychology*, 86(1), 114-121.

Ma, Y., Ailawadi, K.L., Gauri, D.K. & Grewal, D. (2011). An empirical investigation of the impact of gasoline prices on grocery shopping behavior. *Journal of Marketing*, *75*(2), 18-35.

Malhotra, N. K., Kim, S. S., & Patil, A. (2006). Common method variance in IS research: A comparison of alternative approaches and a reanalysis of past research. *Management Science*, 52(12), 1865-1883.

Martens, M. P., & Haase, R. F. (2006). Advanced applications of structural equation modeling in counseling psychology research. *The Counseling Psychologist*, 34(6), 878-911.

McAlister, L., Srinivasan, R., and Kim, M. (2007) Advertising, research and development, and systematic risk of the firm. *Journal of Marketing*, 71(1), 35-48.

Nierop, V.E., Bronnenberg, B., Paap, R., Wedel, M. and Franses, P.H. (2010). Retrieving unobserved consideration sets from household panel data. *Journal of Marketing Research*, 47(1), 63-74.

Owusu-Gyapong, A. (1986). Alternative estimating techniques for panel data on strike activity. *Review of Economics and Statistics*, 68(3), 526-531.

Podsakoff, P. M., MacKenzie, S. B., Lee, J.-Y., & Podsakoff, N. P. (2003). Common method biases in behavioral research: a critical review of the literature and recommended remedies. *Journal of Applied Psychology*, 88(5), 879-903.

Reeb, D., Sakakibara, M., & Mahmood, I. P. (2012). From the editors: Endogeneity in international business research. *Journal of International Business Studies*, 43(3), 211-218.

Roberts, M. R., & Whited, T. M. (2012). Endogeneity in empirical corporate finance. SSRN paper. Simon School Working Paper No. FR 11-29.

Roodman, D. (2009). How to do xtabond2: An introduction to difference and system GMM in Stata. *The Stata Journal*, Vol. 9, No. 1, pp. 86–136.

Schultz, E. L., Tan, D. T. & Walsh, K. D. (2010) Endogeneity and the corporate governance-performance relation, *Australian Journal of Management,* 35(2), 145–163.

Sridhar, S., Germann, F., Kang, C. & Grewal, R., (2016). Relating online, regional, and national advertising to firm value. *Journal of Marketing*, *80*(4), 39-55.

Tourangeau, R., Rips, L. J., & Rasinski, K. (2000). The psychology of survey response: Cambridge University Press.

Wintoki, M. B., Linck, J. S. & Netter, J. M. (2012). Endogeneity and the dynamics of internal corporate governance. *Journal of Financial Economics*, (105)3, 581–606.

Wooldridge, J.M. (2012). Introductory Econometrics: A Modern Approach, (Boston: Cengage Learning).

Zaefarian, G., Kadile, V., Henneberg, S. C., & Leischnig, A. (2017). Endogeneity bias in marketing research: Problems, causes and remedies. *Industrial Marketing Management*, (65), 39-46.

Figure1. Conceptual framework

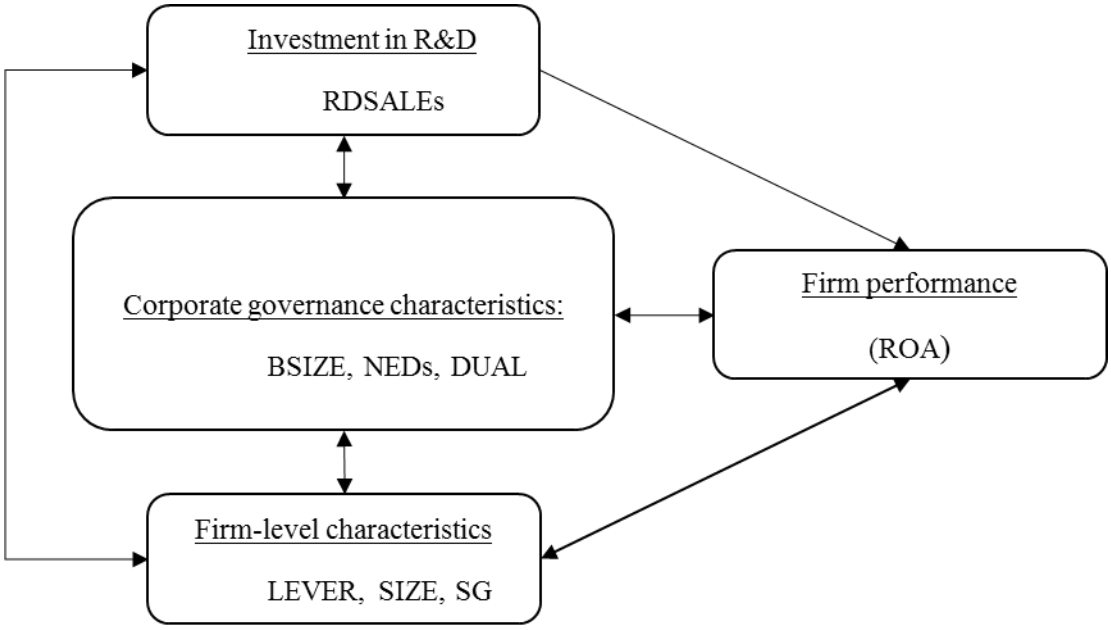

**Table 1 Industrial Classification**

| Name of industry | No. of companies |
|---|---|
| Aerospace | 4 |
| Auto Parts | 1 |
| Biotechnology | 2 |
| Beverages | 4 |
| Business Support System | 7 |
| Utilities | 6 |
| Containers and Packages | 3 |
| Defense | 4 |
| Industrial Parts/Machinery | 9 |
| Electrical/Electronic Equipment | 8 |
| Exploration and Production | 1 |
| Telecommunications | 5 |
| Food Products | 8 |
| Mining | 4 |
| Construction | 6 |
| Oil and Gas | 3 |
| Media | 2 |
| Pharmaceuticals | 6 |
| Household products | 4 |
| Paper | 1 |
| Software | 5 |
| Chemicals | 5 |
| Retailors | 1 |
| Tobacco | 1 |
| Transport Services | 1 |
| Total | 101 |

***Table 2 Definitions of variables***

| ***Variable name*** | ***Definition*** | ***Dtatastream Codes*** |
|---|---|---|
| R&D Expenditures (RDSALES) | Research and development expenditures divided by sales. | 119/104 |
| ***Governance Variables*** | | |
| Board size (BSIZE) | The total number of board members at the end of the fiscal year. | CGBSDP060 |
| Non-executive directors (NEDs) | Percentage of non-executive board members. | CGBSO06V |
| Number of board meetings (NBM) | The number of board meetings during the year. | CGBFDP024 |
| Gender diversity (GD) | Percentage of women on the board of directors. | CGBSO17V |
| Duality (DUAL) | 1 if chairman and CEO are the same person, 0 otherwise. | CGBSO09V |
| Total senior executives' compensation (TSEC) | Logarithm of the total compensation paid to all senior executives (if total aggregate is reported by the company). | CGCPDP054 |
| Single largest/biggest shareholder/owner (SBO) | The percentage ownership of the single largest/biggest owner (by voting power) having shares ownership ≥ 3%. | CGSRDP045 |
| ***Firm-specific Characteristics/Control*** | | |
| Debt financing (LEVER) | A firm's total debt divided by its total assets. | WC03255/DWTA |
| Firm size (SIZE) | Natural l logarithm of a firm's total assets at the end of a financial year. | DWTA |
| Sales growth (SG) | Current year's sales minus previous year's sales divided by previous years sales | 104-104/104 |
| ***Operating performance-Dependent variable*** | | |
| Return on assets (ROA) | Operating income divided by total assets at the end of a financial year. | WC01250/DWTA |

**Table 3 Descriptive Statistics**

| Variable | Observations | Mean | Standard deviation | Min | Max |
|---|---|---|---|---|---|
| RDSALES | 1515 | 0.04 | 0.08 | 0.00 | 0.50 |
| BSIZE | 1515 | 9.07 | 2.50 | 5.00 | 18.00 |
| NEDs | 1515 | 65.26 | 9.99 | 22.22 | 88.69 |
| NBM | 1515 | 8.75 | 2.51 | 2.00 | 26.00 |
| GD | 1514 | 9.18 | 10.46 | 0.00 | 50.00 |
| DUAL | 1515 | 0.13 | 0.34 | 0.00 | 1.00 |
| TSEC | 1515 | 6.68 | 0.45 | 3.44 | 9.13 |
| SBO | 1027 | 14.06 | 11.44 | 3.00 | 72.30 |
| LEVER | 1515 | 0.23 | 0.17 | 0.00 | 0.89 |
| SIZE | 1515 | 14.28 | 1.88 | 8.93 | 19.71 |
| SG | 1515 | 0.10 | 0.39 | -0.89 | 4.86 |
| ROA | 1513 | 0.14 | 0.10 | -0.22 | 0.50 |

**Table 4 Empirical Results Reported for OLS, Fixed-effects and System GMM**

| VARIABLES | Model(1) OLS | Model(2) Fixed effects | Model(3) System GMM |
|---|---|---|---|
| L.ROA | | | 0.594*** |
| | | | (0.0176) |
| L2.ROA | | | 0.0706*** |
| | | | (0.0130) |
| RDSALES | *0.376*** | *0.140*** | *0.193*** |
| | (0.0413) | (0.0529) | (0.0302) |
| BSIZE | **0.00636*** ** | **-0.00263*** | **0.000626*** |
| | (0.00182) | (0.00137) | (0.000341) |
| NEDs | **0.000355** | **0.000444*** | **0.000213*** ** |
| | (0.000312) | (0.000236) | (4.56e-05) |
| DUAL | **-0.00884** | **0.0165** ** | **0.00891*** ** |
| | (0.00928) | (0.00760) | (0.00340) |
| TSEC | *0.0650*** | *0.0331*** | *0.0325*** |
| | (0.00963) | (0.00617) | (0.00394) |
| LEVER | **0.137*** ** | **-0.0364*** | **-0.00839** |
| | (0.0191) | (0.0187) | (0.00769) |
| NBM | **-0.00263** ** | **-0.00127*** | **-0.00154*** ** |
| | (0.00106) | (0.000710) | (0.000161) |
| GD | **0.00108*** ** | **0.000341** | **-0.000121** |
| | (0.000313) | (0.000215) | (7.62e-05) |
| SBO | **0.000606** ** | **-0.00144*** ** | **0.000816*** ** |
| | (0.000270) | (0.000346) | (0.000109) |
| SIZE | **-0.0163*** ** | **-0.00855*** | **-0.00676*** ** |
| | (0.00306) | (0.00480) | (0.00171) |
| SG | *0.0525*** | *0.0246*** | *0.0496*** |
| | (0.0115) | (0.00616) | (0.00330) |
| Constant | **-0.181*** ** | **0.0729** | **-0.0972*** ** |
| | (0.0499) | (0.0685) | (0.0171) |
| Observations | 1,026 | 1,026 | 1,006 |
| R-squared | 0.207 | 0.089 | |
| Number of firms | 101 | 101 | 101 |
| AR(1) | | | 0.0022 |
| AR(2) | | | 0.1411 |
| Sargan test statistics | | | 86.10505 |

Standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1.

**Appendix 1 Generic STATA Commands for Implementing GMM in STATA and Applying Post-estimation Tests**

| | |
|---|---|
| Fixed effects | xtreg depvar indepvars, fe |
| Installing xtabond2 in STATA | ssc install xtabond2 |
| Applying Two-step GMM using xtabond2 command | xtabond2 y l.y x1 x2, x3, x4 gmm(y x1 x2, x3, x4 lag(a b)) noconstant twostep |
| Sargan test | estat sargan |
| The Arellano-Bond test for first-order and second-order correlation | estat abond |