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# 1 Genome-wide association analyses identify new risk variants 2 and the genetic architecture of amyotrophic lateral sclerosis

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262 **To elucidate the genetic architecture of amyotrophic lateral sclerosis (ALS) and find**  
263 **associated loci, we assembled a custom imputation reference panel from whole genome-**  
264 **sequenced ALS patients and matched controls ( $N = 1,861$ ). Through imputation and**  
265 **mixed-model association analysis in 12,577 cases and 23,475 controls, combined with**  
266 **2,579 cases and 2,767 controls in an independent replication cohort, we fine mapped a**  
267 **novel locus on chromosome 21 and identified *C21orf2* as an ALS risk gene. In addition,**  
268 **we identified *MOBP* and *SCFD1* as novel associated risk loci. We established evidence**  
269 **for ALS being a complex genetic trait with a polygenic architecture. Furthermore, we**  
270 **estimated the SNP-based heritability at 8.5%, with a distinct and important role for low**  
271 **frequency (1–10%) variants. This study motivates the interrogation of larger sample**  
272 **sizes with full genome coverage to identify rare causal variants that underpin ALS risk.**

273

274 ALS is a fatal neurodegenerative disease that affects 1 in 400 people, death occurring within  
275 three to five years<sup>1</sup>. Twin-based studies estimate heritability to be around 65% and 5–10% of  
276 ALS patients have a positive family history<sup>1,2</sup>. Both are indicative of an important genetic  
277 component in ALS etiology. Following the initial discovery of the *C9orf72* locus in GWASs<sup>3–</sup>  
278 <sup>5</sup>, the identification of the pathogenic hexanucleotide repeat expansion in this locus  
279 revolutionized the field of ALS genetics and biology<sup>6,7</sup>. The majority of ALS heritability,  
280 however, remains unexplained and only two additional risk loci have been identified robustly  
281 since<sup>3,8</sup>.

282

283 To discover new genetic risk loci and elucidate the genetic architecture of ALS, we genotyped  
284 7,763 new cases and 4,669 controls and additionally collected existing genotype data of  
285 published GWAS in ALS. In total, we analyzed 14,791 cases and 26,898 controls from 41  
286 cohorts (**Supplementary Table 1, Supplementary Methods**). We combined these cohorts  
287 based on genotyping platform and nationality to form 27 case-control strata. In total 12,577  
288 cases and 23,475 controls passed quality control (Online methods, **Supplementary Tables 2–**  
289 **5**).

290

291 For imputation purposes we obtained high-coverage (~43.7X) whole genome sequencing data  
292 from 1,246 ALS patients and 615 controls from The Netherlands (Online methods, and  
293 **Supplementary Fig. 1**). After quality control, we constructed a reference panel including  
294 18,741,510 single nucleotide variants. Imputing this custom reference panel into Dutch ALS  
295 cases increased imputation accuracy of low-frequency genetic variation (minor allele

296 frequency, MAF 0.5–10%) considerably compared to commonly used reference panels: the  
297 1000 Genomes Project phase 1 (1000GP)<sup>9</sup> and Genome of The Netherlands (GoNL)<sup>10</sup> (**Fig.**  
298 **1a**). The improvement was also observed when this reference panel was used to impute into  
299 ALS cases from the UK (**Fig. 1b**). To benefit from the global diversity of haplotypes, the  
300 custom and 1000GP panels were combined, which further improved imputation. Given these  
301 results, we used the merged reference panel for imputation of all strata in our study.

302  
303 In total we imputed 8,697,640 variants passing quality control in the 27 strata and separately  
304 tested these for association with ALS risk by logistic regression. Results were then included  
305 in an inverse-variance weighted fixed effects meta-analysis, which revealed 4 loci at genome-  
306 wide significance ( $p < 5 \times 10^{-8}$ ) (**Fig. 2a**). The previously reported *C9orf72* (rs3849943)<sup>3–5,8</sup>,  
307 *UNC13A* (rs12608932)<sup>3,5</sup> and *SARM1* (rs35714695)<sup>8</sup> loci all reached genome-wide  
308 significance, as did a novel association for a non-synonymous variant in *C21orf2*  
309 (rs75087725,  $p = 8.7 \times 10^{-11}$ , **Supplementary Tables 8 and 10–13**). Interestingly, this variant  
310 was present on only 10 haplotypes in the 1000GP reference panel (MAF = 1.3%), while our  
311 custom reference panel included 62 haplotypes carrying the minor allele (MAF = 1.7%). As a  
312 result, more strata passed quality control for this variant by passing the allele frequency  
313 threshold of 1% (**Supplementary Table 9**). This demonstrates the benefit of the merged  
314 reference panel with ALS-specific content, which improved imputation and resulted in a  
315 genome-wide significant association.

316  
317 Linear mixed models (LMM) can improve power while controlling for sample structure<sup>11</sup>,  
318 particularly in our study that included a large number of imperfectly balanced strata. Even  
319 though LMM for ascertained case-control data has a potential small loss of power<sup>11</sup>, we  
320 judged the advantage of combining all strata while controlling the false positive rate, to be  
321 more important and therefore jointly analyzed all strata in a LMM to identify additional risk  
322 loci. There was no overall inflation of the linear mixed model's test statistic compared to the  
323 meta-analysis (**Supplementary Fig. 2**). We observed modest inflation in the QQ-plot ( $\lambda_{GC} =$   
324 1.12,  $\lambda_{1000} = 1.01$ , **Supplementary Fig. 3**). LD score regression yielded an intercept of 1.10  
325 (standard error  $7.8 \times 10^{-3}$ ). While the LD score regression intercept can indicate residual  
326 population stratification, which is fully corrected for in a LMM, the intercept can also reflect  
327 a distinct genetic architecture where most causal variants are rare, or a non-infinitesimal  
328 architecture<sup>12</sup>. The linear mixed model identified all four genome-wide significant

329 associations from the meta-analysis. Furthermore, three additional loci that included the  
330 *MOBP* gene on 3p22.1 (rs616147), *SCFDI* on 14q12 (rs10139154) and a long non-coding  
331 RNA on 8p23.2 (rs7813314) were associated at genome-wide significance (**Fig. 2b, Table 1,**  
332 **Supplementary Tables 14–16**). Interestingly, the SNPs in the *MOBP* locus have been  
333 reported in a GWAS on progressive supranuclear palsy (PSP)<sup>13</sup> and as a modifier for survival  
334 in frontotemporal dementia (FTD)<sup>14</sup>. The putative pleiotropic effect of variants within this  
335 locus suggests a shared neurodegenerative pathway between ALS, FTD and PSP. We also  
336 found rs74654358 at 12q14.2 in the *TBKI* gene approximating genome-wide significance  
337 (MAF = 4.9%, OR = 1.21 for A allele,  $p = 6.6 \times 10^{-8}$ ). This gene was recently identified as an  
338 ALS risk gene through exome sequencing<sup>15,16</sup>.

339

340 In the replication phase, we genotyped the newly discovered associated SNPs in nine  
341 independent replication cohorts, totaling 2,579 cases and 2,767 controls. In these cohorts we  
342 replicated the signals for the *C2Iorf2*, *MOBP* and *SCFDI* loci, with lower p-values in the  
343 combined analysis than the discovery phase (combined p-value =  $3.08 \times 10^{-10}$ ,  $p = 4.19 \times 10^{-10}$   
344 and  $p = 3.45 \times 10^{-8}$  for rs75087725, rs616147 and rs10139154 respectively, **Table 1,**  
345 **Supplementary Fig. 4**)<sup>17</sup>. The combined signal for rs7813314 was less significant due to an  
346 opposite effect between the discovery and replication phase, indicating non-replication.  
347 Although replication yielded similar effect estimates for rs10139154 compared to the  
348 discovery phase, this was not statistically significant ( $p = 0.09$ ) in the replication phase alone.  
349 This reflects the limited sample size of our replication phase, which is inherent to the low  
350 prevalence of ALS and warrants even larger sample sizes to replicate this signal robustly.

351

352 There was no evidence for residual association within each locus after conditioning on the top  
353 SNP, indicating that all risk loci are independent signals. Apart from the *C9orf72*, *UNC13A*  
354 and *SARM1* loci, we found no evidence for associations previously described in smaller  
355 GWAS (**Supplementary Table 17**).

356

357 The associated low-frequency non-synonymous SNP in *C2Iorf2* suggested that this gene  
358 could directly be involved in ALS risk. Indeed, we found no evidence that linkage  
359 disequilibrium of sequenced variants beyond *C2Iorf2* explained the association within this  
360 locus (**Supplementary Fig. 5**). In addition, we investigated the burden of rare coding  
361 mutations in a set of whole genome sequenced cases (N = 2,562) and controls (N = 1,138).  
362 After quality control these variants were tested using a pooled association test for rare variants

363 corrected for population structure (T5 and T1 for 5% and 1% allele frequency,  
364 **Supplementary methods**). This revealed an excess of non-synonymous and loss-of-function  
365 mutations in *C21orf2* among ALS cases that persists after conditioning on rs75087725 ( $p_{T5} =$   
366  $9.2 \times 10^{-5}$ ,  $p_{T1} = 0.01$ , **Supplementary Fig. 6**), which further supports that *C21orf2*  
367 contributes to ALS risk.

368

369 In an effort to fine-map the other loci to susceptibility genes, we searched for SNPs in these  
370 loci with *cis*-eQTL effects observed in brain and other tissues (**Supplementary methods**,  
371 **Supplementary Table 18**)<sup>18</sup>. There was overlap with previously identified brain *cis*-eQTLs  
372 for five regions (**Supplementary Fig. 7, Supplementary Table 19, Supplementary Data**  
373 **Set 1**). Interestingly, within the *C9orf72* locus we found that proxies of rs3849943 (LD  $r^2 =$   
374  $0.21 - 0.56$ ) had a brain *cis*-eQTL effect on *C9orf72* only (minimal  $p = 5.27 \times 10^{-7}$ ), which  
375 harbors the hexanucleotide repeat expansion that drives this GWAS signal. Additionally, we  
376 found that rs12608932 and its proxies within the *UNC13A* locus had exon-level *cis*-eQTL  
377 effect on *KCNN1* in frontal cortex ( $p = 1.15 \times 10^{-3}$ )<sup>19</sup>. Another overlap was observed in the  
378 *SARM1* locus where rs35714695 and its proxies had the strongest exon-level *cis*-eQTL effect  
379 on *POLDIP2* in multiple brain tissues ( $p = 2.32 \times 10^{-3}$ ). Within the *SCFD1* locus rs10139154  
380 and proxies had a *cis*-eQTL effect on *SCFD1* in cerebellar tissue ( $p = 7.71 \times 10^{-4}$ ). For the  
381 *MOBP* locus, rs1768208 and proxies had a *cis*-eQTL effect on *RPSA* ( $p = 7.71 \times 10^{-4}$ ).

382

383 To describe the genetic architecture of ALS, we calculated polygenic scores that can be used  
384 to predict phenotypes for traits with a polygenic architecture<sup>20</sup>. We calculated the SNP effects  
385 using a linear mixed model in 18 of the 27 strata and subsequently assessed their predictive  
386 ability in the other 9 independent strata. The analysis revealed that a significant, albeit  
387 modest, proportion of the phenotypic variance could be explained by all SNPs (Nagelkerke  $r^2$   
388  $= 0.44\%$ ,  $r^2 = 0.15\%$  on the liability scale,  $p = 2.7 \times 10^{-10}$ , **Supplementary Fig. 8**). This  
389 finding adds to the existing evidence that ALS is a complex genetic trait with a polygenic  
390 architecture. To further quantify the contribution of common SNPs to ALS risk, we estimated  
391 the SNP-based heritability using three approaches, all assuming a population baseline risk of  
392  $0.25\%$ <sup>21</sup>. The variance explained by all SNPs using GCTA-REML estimated heritability at  
393  $8.5\%$  (SE  $0.5\%$ ). Haseman-Elston regression yielded a very similar  $7.9\%$  and LD score  
394 regression estimated the SNP-based heritability at  $8.2\%$  (SE  $0.5\%$ ). The heritability estimates  
395 per chromosome were strongly correlated with chromosome length ( $p = 4.9 \times 10^{-4}$ ,  $r^2 = 0.46$ ,  
396 **Fig. 3a**), which again is indicative of the polygenic architecture of ALS.

397

398 We found that the genome-wide significant loci only explained 0.2% of the heritability and  
399 thus the bulk of the heritability (8.3%, SE 0.3%) was captured in SNPs below genome-wide  
400 significance. This implies that many genetic risk variants have yet to be discovered.

401 Understanding where these unidentified risk variants remain across the allele frequency  
402 spectrum will inform designing future studies to identify these variants. We, therefore,  
403 estimated heritability partitioned by minor allele frequency. Furthermore, we contrasted this  
404 to common polygenic traits studied in GWASs such as schizophrenia. We observed a clear  
405 trend that indicated that most variance is explained by low-frequency SNPs (**Fig. 3b**).

406 Exclusion of the *C9orf72* locus, which harbors the rare pathogenic repeat expansion, and the  
407 other genome-wide significant loci did not affect this trend (**Supplementary fig. 9**). This  
408 architecture is different from that expected for common polygenic traits and reflects a  
409 polygenic rare-variant architecture observed in simulations<sup>22</sup>.

410

411 To gain better insight into the biological pathways that explain the associated loci found in  
412 this study we looked for enriched pathways using DEPICT<sup>23</sup>. This revealed SNAP receptor  
413 (SNARE) activity as the only enriched category (FDR < 0.05, **Supplementary Fig. 10**).  
414 SNARE complexes play a central role in neurotransmitter release and synaptic function<sup>24</sup>,  
415 which are both perturbed in ALS<sup>25</sup>.

416

417 Although the biological role of *C21orf2*, a conserved leucine-rich repeat protein, remains  
418 poorly characterized, it is part of the ciliome and is required for the formation and/or  
419 maintenance of primary cilia<sup>26</sup>. Defects in primary cilia are associated with various  
420 neurological disorders and cilia numbers are decreased in G93A *SOD1* mice, a well-  
421 characterized ALS model<sup>27</sup>. *C21orf2* has also been localized to mitochondria in immune  
422 cells<sup>28</sup> and is part of the interactome of the protein product of *NEK1*, which has previously  
423 been associated with ALS<sup>15</sup>. Both proteins appear to be involved in DNA repair  
424 mechanisms<sup>29</sup>. Although future studies are needed to dissect the function of *C21orf2* in ALS  
425 pathophysiology it is tempting to speculate that defects in *C21orf2* lead to primary cilium  
426 and/or mitochondrial dysfunction or inefficient DNA repair and thereby adult onset disease.  
427 The other associated loci will require more extensive studies to fine-map causal variants. The  
428 *SARM1* gene has been suggested as a susceptibility gene for ALS, mainly because of its role  
429 in Wallerian degeneration and interaction with *UNC13A*<sup>8,30</sup>. Although these are indeed  
430 interesting observations, the brain *cis*-eQTL effect on *POLDIP2* suggests that *POLDIP2* and

431 not *SARM1* could in fact be the causal gene within this locus. Similarly, *KCNN1*, which  
432 encodes a neuronal potassium channel involved in neuronal excitability, could be the causal  
433 gene either through a direct eQTL effect or rare variants in LD with the associated SNP in  
434 *UNC13A*.

435

436 In conclusion, we identified a key role for rare variation in ALS and discovered SNPs in  
437 novel complex loci. Our study therefore informs future study design in ALS genetics: the  
438 combination of larger sample sizes, full genome coverage and targeted genome editing  
439 experiments, leveraged together to fine map novel loci, identify rare causal variants and  
440 thereby elucidate the biology of ALS.

441 **ACCESSION CODES**

442 NIH Genome-Wide Association Studies of Amyotrophic Lateral Sclerosis (phs000101.v3.p1),  
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445 (PD) (phs000126.v1.p1), Genome-Wide Association Study of Parkinson Disease: Genes and  
446 Environment (phs000196.v1.p1)

447

448 **DATA ACCESS**

449 The GWAS summary statistics and sequenced variants are publicly available through the  
450 Project MinE data browser: <http://databrowser.projectmine.com>

451

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455

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483 analyses. U.V., L.F., W.v.R. and J.H.V. performed eQTL analyses. W.v.R., A.S., A.A.-C.,  
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486

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559

560 **FIGURE LEGENDS**

561 **Figure 1. Imputation accuracy comparison.** The aggregate  $r^2$  value between imputed and  
562 sequenced genotypes on chromosome 20 using different reference panels for imputation.  
563 Allele frequencies are calculated from the Dutch samples included in the Genome of the  
564 Netherlands cohort. The highest imputation accuracy was achieved when imputing from the  
565 merged custom and 1000GP panels. This difference is most pronounced for low frequency  
566 (0.5–10%) alleles in both ALS cases from The Netherlands (a) and United Kingdom (b).

567

568 **Figure 2. Meta-analysis and linear mixed model associations.** (a) Manhattan plot for meta-  
569 analysis results. This yielded four genome-wide significant associations highlighted with  
570 names indicating the closest gene. The associated SNP in *C21orf2* is a non-synonymous  
571 variant not found in previous GWAS. (b) Manhattan plot for linear mixed model results. This  
572 association analysis yielded three additional loci reaching genome-wide significance (*MOBP*,  
573 *LOC101927815* and *SCFD1*). SNPs in the previously identified ALS risk gene *TBK1*  
574 approached genome-wide significance ( $p = 6.6 \times 10^{-8}$ ). Since the *C21orf2* SNP was removed

575 from a Swedish stratum because of a  $MAF < 1\%$ , this SNP was tested separately, but is  
 576 presented here together with all other SNPs with a  $MAF > 1\%$  in every stratum. Here,  
 577 *LOC101927815* is colored grey because the association for this locus could not be replicated.

578  
 579 **Figure 3. Partitioned heritability.** (a) The heritability estimates per chromosome were  
 580 strongly correlated with chromosome length ( $p = 4.9 \times 10^{-4}$ ). (b) For ALS there was a clear  
 581 trend where more heritability was explained within the lower allele frequency bins. This  
 582 effect was still observed when, for a fair comparison between ALS and a previous study  
 583 partitioning heritability for schizophrenia (SCZ) using identical methods<sup>22</sup>, SNPs present in  
 584 HapMap3 (HM3) were included. The pattern for ALS resembles that observed in a rare  
 585 variant model simulation performed in this study. Error bars reflect standard errors.

586

587 **TABLES**588 **Table 1. Discovery and replication of novel genome-wide significant loci.**

SNP	Discovery					Replication				Combined	
	$MAF_{cases}$	$MAF_{controls}$	OR	$P_{meta}$	$P_{LMM}$	$MAF_{cases}$	$MAF_{controls}$	OR	$P$	$P_{combined}$	$I^2$
rs75087725	0.02	0.01	1.45	$8.65 \times 10^{-11}$	$2.65 \times 10^{-9}$	0.02	0.01	1.65	$3.89 \times 10^{-3}$	$3.08 \times 10^{-10}$	0.00*
rs616147	0.30	0.28	1.10	$4.14 \times 10^{-5}$	$1.43 \times 10^{-8}$	0.31	0.28	1.13	$2.35 \times 10^{-3}$	$4.19 \times 10^{-10}$	0.00*
rs10139154	0.34	0.31	1.09	$1.92 \times 10^{-5}$	$4.95 \times 10^{-8}$	0.33	0.31	1.06	$9.55 \times 10^{-2}$	$3.45 \times 10^{-8}$	0.05*
rs7813314	0.09	0.10	0.87	$7.46 \times 10^{-7}$	$3.14 \times 10^{-8}$	0.12	0.10	1.17	$7.75 \times 10^{-3}$	$1.05 \times 10^{-5}$	0.80**

589

590 **Table 1. Discovery and replication of novel genome-wide significant loci.** Genome-wide  
 591 significant loci from the discovery phase including 12,557 cases and 23,475 controls were  
 592 directly genotyped and tested for association in the replication phase including 2,579 cases  
 593 and 2,767 controls. The three top associated SNPs in the *MOBP* (rs616147), *SCFD1*  
 594 (rs10139154) and *C21orf2* (rs75087725) loci replicated with associations in identical  
 595 directions as in the discovery phase and an association in the combined analysis that exceeded  
 596 the discovery phase. \* Cochran's Q test:  $p > 0.1$ , \*\* Cochran's Q test:  $p = 4.0 \times 10^{-6}$ , Chr =  
 597 chromosome; SNP = single nucleotide polymorphism, MAF = minor allele frequency, OR =  
 598 odds ratio,  $P_{meta}$  = meta-analysis p-value,  $P_{LMM}$  = linear mixed model p-value,  $P_{combined}$  = meta-  
 599 analysis of discovery linear mixed model and associations from replication phase.

600

601 **AUTHOR INFORMATION**

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605

## 606 **ONLINE METHODS**

607 Software packages used, their version, web source, and references are described in the  
608 **Supplementary Table 20**.

609

610 **GWAS discovery phase and quality control.** Details on the acquired genotype data from  
611 previously published GWAS are described in **Supplementary Table 1**. Methods for case and  
612 control ascertainment for each cohort are described in the **Supplementary methods**. All  
613 cases and controls gave written informed consent and the relevant institutional review boards  
614 approved this study. To obtain genotype data for newly genotyped individuals, genomic DNA  
615 was hybridized to the Illumina OmniExpress array according to manufacturer's protocol.

616 Subsequent quality control included:

- 617 1) Removing low quality SNPs and individuals from each cohort.
- 618 2) Combining unbalanced cohorts based on nationality and genotyping platform to form  
619 case-control strata.
- 620 3) Removing low quality SNPs, related individuals and population outliers per stratum.
- 621 4) Calculate genomic inflation factors per stratum.

622 More details are described in the **Supplementary methods**. The number of SNPs and  
623 individuals failing each QC step per cohort and stratum are displayed in **Supplementary**  
624 **Tables 2–5**.

625

626 **Whole genome sequencing (custom reference panel).** Individuals were whole genome  
627 sequenced on the Illumina HiSeq 2500 platform using PCR free library preparation and 100bp  
628 paired-end sequencing yielding a minimum 35X coverage. Reads were aligned to the hg19  
629 human genome build and after variant calling (Isaac variant caller) additional SNV and  
630 sample quality control was performed (**Supplementary methods**). Individuals in our custom  
631 reference panel were also included in the GWAS in strata sNL2, sNL3 and sNL4.

632

633 **Merging reference panels.** All high quality calls in the custom reference panel were phased  
634 using SHAPEIT2 software. After checking strand and allele inconsistencies, both the 1000  
635 Genomes Project (1000GP) reference panel (release 05-21-2011)<sup>31</sup> and custom reference

636 panel were imputed up to the union of their variants as described previously<sup>32</sup>. Those variants  
637 with inconsistent allele frequencies between the two panels were removed.

638

639 **Imputation accuracy performance.** To assess the imputation accuracy between different  
640 reference panels, 109 unrelated ALS cases of Dutch ancestry sequenced by Complete  
641 Genomics and 67 ALS cases from the UK sequenced by Illumina were selected as a test  
642 panel. All variants not present on the Illumina Omni1M array were masked and the SNVs on  
643 chromosome 20 were subsequently imputed back using four different reference panels  
644 (1000GP, GoNL, custom panel and merged panel). Concordance between the imputed alleles  
645 and sequenced alleles was assessed within each allele frequency bin where allele frequencies  
646 are calculated from the Dutch samples included in the Genome of the Netherlands cohort.

647

648 **GWAS imputation.** Pre-phasing was performed per stratum using SHAPEIT2 with the  
649 1000GP phase 1 (release 05-21-2011) haplotypes<sup>31</sup> as a reference panel. Subsequently, strata  
650 were imputed up to the merged reference panel in 5 megabase chunks using IMPUTE2.  
651 Imputed variants with a MAF < 1% or INFO score < 0.3 were excluded from further analysis.  
652 Variants with allele frequency differences between strata, defined as deviating > 10SD from  
653 the normalized mean allele frequency difference between those strata and an absolute  
654 difference > 5%, were excluded, since they are likely to represent sequencing or genotyping  
655 artifacts. Imputation concordance scores for cases and controls were compared to assess  
656 biases in imputation accuracy (**Supplementary Table 6**).

657

658 **Meta-analysis.** Logistic regression was performed on imputed genotype dosages under an  
659 additive model using SNPTTEST software. Based on scree plots, 1 to 4 principal components  
660 were included per stratum. These results were then combined in an inverse-variance weighted  
661 fixed effect meta-analysis using METAL. No marked heterogeneity across strata was  
662 observed as the Cochran's Q test statistics did not deviate from the null-distribution ( $\lambda =$   
663 0.96). Therefore, no SNPs were removed due to excessive heterogeneity. The genomic  
664 inflation factor was calculated and the quantile-quantile plot is provided in **Supplementary**  
665 **Fig. 3a**.

666

667 **Linear mixed model.** All strata were combined including SNPs that passed quality control in  
668 every stratum. Subsequently the genetic relationship matrices (GRM) were calculated per

669 chromosome including all SNPs using the Genome-Wide Complex Trait Analysis (GCTA)  
670 software package. Each SNP was then tested in a linear mixed model including a GRM  
671 composed of all chromosomes excluding the target chromosome (leave one chromosome out,  
672 LOCO). The genomic inflation factor was calculated and the quantile-quantile plot is  
673 provided as **Supplementary Fig. 3b**.

674

675 **Replication.** For the replication phase independent ALS cases and controls from Australia,  
676 Belgium, France, Germany, Ireland, Italy, The Netherlands and Turkey that were not used in  
677 the discovery phase were included. A pre-designed TaqMan genotyping assay was used to  
678 replicate rs75087725 and rs616147. Sanger sequencing was performed to replicate  
679 rs10139154 and rs7813314 (**Supplementary methods, Supplementary Table 7**). All  
680 genotypes were tested in a logistic regression per country and subsequently meta-analyzed.

681

682 **Rare variant analysis in *C21orf2*.** The burden of non-synonymous rare variants in *C21orf2*  
683 was assessed in whole genome sequencing data obtained from ALS cases and controls from  
684 The Netherlands, Belgium, Ireland, United Kingdom and the United States. After quality  
685 control the burden of non-synonymous and loss-of-function mutations in *C21orf2* were tested  
686 for association per country and subsequently meta-analyzed. More details are provided in the  
687 **Supplementary methods**.

688

689 **Polygenic risk scores.** To assess the predictive accuracy of polygenic risk scores in an  
690 independent dataset SNP weights were assigned based on the linear mixed model (GCTA-  
691 LOCO) analysis in 18/27 strata. SNPs in high LD ( $r^2 > 0.5$ ) within a 250 kb window were  
692 clumped. Subsequently, polygenic risk scores for cases and controls in the 9 independent  
693 strata were calculated based on their genotype dosages using PLINK v1.9. To obtain the  
694 Nagelkerke  $R^2$  and corresponding p-values these scores were then regressed on their true  
695 phenotype in a logistic regression where (based on scree plots) the first three PCs, sex and  
696 stratum were included as covariates.

697

698 **SNP-based heritability estimates. *GCTA-REML*.** GRMs were calculated using GCTA  
699 software including genotype dosages passing quality control in all strata. Based on the  
700 diagonal of the GRM individuals representing subpopulations that contain an abundance of  
701 rare alleles (diagonal values mean  $\pm$  2SD) were removed (**Supplementary Fig. 14a**). Pairs  
702 where relatedness (off-diagonal) exceeded 0.05 were removed as well (**Supplementary Fig.**

703 **14b**). The eigenvectors for the first 10 PCs were included as fixed effects to account for more  
704 subtle population structure. The prevalence of ALS was defined as the life-time morbid risk  
705 for ALS (i.e. 1/400)<sup>19</sup>. To estimate the SNP-based heritability for all non-genome-wide  
706 significant SNPs, genotypes for the SNPs reaching genome-wide significance were modeled  
707 as fixed effect. The variance explained by the GRM therefore reflects the SNP-based  
708 heritability of all non-genome-wide significant SNPs. SNP-based heritability partitioned by  
709 chromosome or MAF was calculated by including multiple GRMs, calculated on SNPs from  
710 each chromosome or within the respective frequency bin, in one model.

711 *Haseman-Elston regression*. The Phenotype correlation - Genotype correlation (PCGC)  
712 regression software package was used to calculate heritability based on the Haseman-Elston  
713 regression including the eigenvectors for the first 10 PCs as covariates. The prevalence was  
714 again defined as the life-time morbid risk (1/400).

715 *LD score regression*. Summary statistics from GCTA-LOCO and LD scores calculated from  
716 European individuals in 1000GP were used for LD score regression. Strongly associated  
717 SNPs ( $p < 5 \times 10^{-8}$ ) and variants not in HapMap3 were excluded. Considering adequate  
718 correction for population structure and distant relatedness in the linear mixed model, the  
719 intercept was constrained to 1.0<sup>12</sup>.

720 **Biological pathway analysis (DEPICT)**. Functional interpretation of associated GWAS loci  
721 was carried out using DEPICT, using locus definition based on 1000GP phase 1 data. This  
722 method prioritizes genes in the affected loci, predicts involved pathways, biological processes  
723 and tissues, using gene co-regulation data from 77,840 expression arrays. Three separate  
724 analyses were performed for GWAS loci reaching  $p = 10^{-4}$ ,  $p = 10^{-5}$  or  $p = 10^{-6}$ . One thousand  
725 permutations were used for adjusting the nominal enrichment p-values for biases and  
726 additionally 200 permutations were used for FDR calculation.

727

## 728 REFERENCES FOR METHODS

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730 and array data to create an improved 1000 Genomes Project haplotype reference panel. *Nat.*  
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