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Country-Level Cost-Effectiveness Thresholds: Initial Estimates and the Need for Further Research

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Country-level cost-effectiveness thresholds: initial estimates and the need for further research

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Abstract

Healthcare systems in low- and middle-income countries (LMICs) face considerable population healthcare needs with markedly fewer resources than those in developed countries. The way in which available resources are allocated across competing priorities is crucial in affecting how much health is generated overall, who receives healthcare interventions and who goes without. Cost-effectiveness analysis (CEA) is one tool that can assist policy-makers in resource allocation. The central concern in CEA is whether the health gains offered by an intervention are large enough relative to its costs to warrant adoption. This requires some notion of the value that must be realized by an intervention, which is most frequently represented using a cost-effectiveness threshold (CET).

CETs should be based on estimates of the forgone benefit associated with alternative priorities which consequently cannot be implemented as a result of the commitment of resources to an alternative. For most health care systems these opportunity costs fall predominantly on health as a result of fixed budgets or constraints on health systems' abilities to increase expenditures. However, many CEAs to inform decisions in LMICs have used aspirational expressions of value, such as the World Health Organization's (WHO) recommended CETs (of 1-3 times GDP per capita in a country) which are not based upon opportunity costs. In contrast, we estimate CETs for a number of countries based upon recent empirical estimates of foregone benefit (from the English NHS) and international income elasticities of the value of health. The resulting CETs are much lower than those previously posited by WHO. There is no intention to provide definitive CETs; rather, the study is intended to provoke further research in this area of crucial policy importance and outlines how more robust estimates of CETs could be generated.

1. Introduction

Policy-makers in all healthcare systems face difficult choices about which interventions, programmes or activities (hereinafter referred to solely as 'interventions') should be funded from limited available resources. The tools of economic evaluation offer a variety of means to assist policymakers in the process of prioritization. A common approach is that of incremental cost-effectiveness analysis (CEA) which is based upon the comparative assessment of costs and health benefits. CEA seeks to identify which interventions offer health benefits large enough relative to costs to warrant adoption and those that do not.[1]

Results in CEA are often represented in the form of an incremental cost-effectiveness ratio (ICER); the ratio of incremental costs to incremental health benefits of one intervention compared to another mutually exclusive alternative ($\Delta costs / \Delta health benefits$).[1] Health benefits are often represented in the form of quality adjusted life years (QALYs); and so the ICER gives the 'cost per QALY gain' associated with an intervention. To determine whether an intervention's ICER is likely to represent a 'good buy' requires comparison to some benchmark of value. The benchmark most frequently used is the cost-effectiveness threshold (CET). If an ICER
CET, an intervention's health benefits are deemed large enough in comparison to costs (i.e. it is not cost-effective).

The choice of CET is, therefore, crucial in determining value from healthcare interventions when applying incremental CEA. A key concept in all economic analysis is that of 'opportunity costs'.[2] All healthcare systems face resource constraints so if resources are committed to the funding of one intervention these are not available to fund and deliver others. Opportunity costs are the health benefits forgone due to the commitment of scarce resources to particular investments, or if additional resources are made available, the health that could have been gained by investing these resources elsewhere in the healthcare system. For CEA to be consistent with population health improvement there is a need for chosen CETs to reflect opportunity costs. However, the estimation of opportunity costs for different kinds of decisions remains a major challenge.

2. Understanding and estimating cost-effectiveness thresholds

Recent methods research has brought greater clarity as to normative basis for the use of CETs and how they can be appropriately estimated so as to improve population health.[3, 4] A clear distinction needs to be made between two related, but separate, concepts of opportunity costs: (i) opportunity costs in terms of health forgone – due to the commitment of limited collective healthcare resources to particular interventions; (ii) opportunity costs in terms of forgone consumption (the 'consumption value of health') when additional costs fall on consumption opportunities outside health care rather than on health care expenditure. The first, (i), can also be termed the 'shadow price of the budget constraint'. It is an issue of 'fact', resulting from limits in the overall collective budget available for healthcare or constraints on health system's abilities to increase expenditure. The second, (ii), is an issue of 'value' and depends upon how individuals value health as compared to other forms of consumption.

For economic evaluation of an intervention, it is important to consider what type of opportunity costs would result from investment in new activities. If opportunity costs result in the form of health forgone (e.g. through displacement of other health generating interventions funded from a collective healthcare budget), then the CET should reflect this – let's denote this 'k' (the ratio of the cost reduction associated with displacing interventions to the QALYs lost). If opportunity costs are in terms of other forms of consumption (e.g. if the intervention relies on other sources of funding such as out-of-pocket contributions), the CET should reflect the consumption value of health - let's denote this 'v'.

Observing that an estimate of the consumption value of health is higher than the amount of healthcare resource required to improve health (v > k) suggests that the healthcare system is not meeting individual preferences, because individuals would be willing to give up more of the resources available to them to improve their own health than the healthcare system would require. Simply assuming that v is equal to k would be inappropriate since there are good reasons to expect v>k not least the welfare losses associated with socially acceptable ways to finance healthcare systems (i.e. society would not increase health expenditure to the point where v=k).

For incremental CEA to suitably inform the allocation of healthcare expenditures, for which the primarily purpose is generally regarded as being the generation of 'health' from limited collective healthcare resources, CETs reflecting the opportunity costs of healthcare spending (k) will always be required if there are any restrictions on the growth in health care expenditure. There are at least two ways in which k could be determined.

Firstly, a threshold can be chosen based upon what it is felt a health system *should* be willing to pay for health – for which distinction can be made between how much individuals would be willing to pay to improve health (v: the consumption value of health) and simply aspirational notions of how the world should be (or how much other individuals should be willing to spend to improve their health, without consideration of other calls on their resources).

Secondly, the relationship between changes in healthcare expenditure and health outcomes can be estimated (i.e. the marginal productivity of the healthcare system in generating health) as a direct measure of what is likely to be displaced when a cost-escalating intervention is adopted or what could be gained elsewhere if cost-savings are made or additional resources are made available.

The first of the approaches tends to dominate in practice in all healthcare systems. For instance, a value of US\$50,000 has been commonly applied in the United States.[5] Similarly, for low and middle income countries, the World Health Organization (WHO) has recommended thresholds of 1

to 3 times gross domestic product (GDP) per capita for use in low and middle income countries.[6] These values – from the US and from WHO – are not based upon assessment of health opportunity costs resulting from resource constraints. Instead they can be seen as statements of 'value'; what it is felt health systems *should* be willing to pay for health gains (the US estimate motivated more by individuals' expressions of willingness to pay for health gains; the WHO approach based on aspiration). However, their use is not necessarily consistent with population health improvement as they do not reflect the opportunity cost that is imposed on the health care system. They can therefore be seen as inappropriate measures of k that risk reducing, rather than increasing, population health.

The second approach is an empirical means to estimate k based upon assessment of health benefits forgone when funding is committed to particular interventions. Estimating k in this way provides a basis for making resource allocation decisions that increase population health. However, there is a paucity of estimates available in any jurisdiction using these approaches. One notable exception is Claxton et al.[4] who used local level programme expenditure data, in a range of disease areas, to estimate the relationship between changes in healthcare expenditure and health outcomes in the English National Health Service.

By exploiting the variation in expenditure and mortality outcomes, Claxton et al. estimate the relationship between changes in spending and mortality while accounting for endogeneity. With additional information about age and gender of the patient population these mortality effects were expressed as a cost per life year threshold (£25,241 per life year). These life year effects were adjusted for quality of life with additional information about quality of life norms by age and gender, as well as the quality of life impacts of different types of disease. By using the effect of expenditure on the mortality and life year burden of disease as a surrogate for the effects on a more complete measure of health burden (i.e. that includes quality of life burden) a cost per QALY threshold was estimated of UK£12,936 (US\$20,212) per QALY.[4]

3. Estimating cost-effectiveness thresholds for other countries

There is a lack of empirically based estimates of k (based on opportunity costs in terms of health) for jurisdictions other than the UK despite growing recognition that such estimates are required to suitably inform resource allocation decisions.[7, 8] The Claxton et al. estimate of k is based upon estimates of the marginal productivity of healthcare spending in just one jurisdiction.[4] In principle, a similar approach could be adopted to estimate the relationship between healthcare spending and health outcomes internationally, using countries as units of analysis, to determine k in a wide range of settings.

To date, however, cross-country evidence on the productivity of healthcare spending has focused on answering the question "does healthcare spending improve health outcomes?" Recent research adjusting for potential reverse causality in this relationship (e.g. governments may spend more when health outcomes are worse) suggests that the answer to this question is yes.[9] However, the available literature does not focus on how the effect of healthcare spending on health outcomes varies according to the level of healthcare spending or country income. The analyses available do suggest that the marginal productivity of healthcare spending diminishes with healthcare spending or country income.[10-12] This indicates that the threshold should increase with country income or healthcare spending rises. However, there is little information about the rate at which this increase occurs¹. Understanding this would require careful exploration of the functional form of the relationship between healthcare spending and health outcomes.

However, there is a body of literature that estimates v (the consumption value of health) in different countries. Some of this literature is based upon stated preference elicitation of individuals' willingness-to-pay (WTP) for morbidity adjusted life years (e.g. QALYs)[13, 14], but a larger body of work estimates the 'value of a statistical life' (VSL) from mortality reductions (e.g. through estimating compensating differentials for on the job risk exposure in labour markets).[15, 16] Moreover, the VSL literature also examines how the VSL varies across jurisdictions as a function of national per capita income (i.e. the elasticity of the VSL with respect to income, ε) (see below). This potentially provides information about the income elasticity of v if we can assume that the income elasticity of the VSL is equal to the income elasticity of the value of a life year (e.g. QALY). For this to be the case across countries, a VSL must convert to the same number of QALYs across countries (the plausibility of this assumption is examined in the Discussion).

Understanding the income elasticity of v across countries of different income levels raises an interesting prospect. If a similar income elasticity of k exists as for v, income elasticities of the VSL can be applied to the Claxton et al. estimate of k for the English NHS to provide estimates of k in a wide range of jurisdictions. We follow this approach to provide estimates of k for application in different countries based upon their per capita income levels using the cost-effectiveness threshold for the NHS, per capita income in the UK and the elasticity of VSL with respect to income. (Figure 1). The approach is based upon two core assumptions: (1) that the discrepancy between v and k is constant across countries in percentage terms (i.e. that 'underfunding' of healthcare through collectively pooled resources relative to individual preferences over consumption and health is

¹ Typically constant elasticity production functions are modelled. These impose significant constraints on the way the threshold can vary with income or health care spending. Namely, values of the elasticity of health with respect to health care spending or income must be <1.0 for the threshold to be increasing in health care spending or income. However, values of the elasticity <1.0 also imply that the threshold increases at a diminishing rate with health care spending or income.

constant with respect to GDP p.c.); (2) that the UK is 'typical' of other countries with respect to the values of v and k. These assumptions are examined in the Discussion.

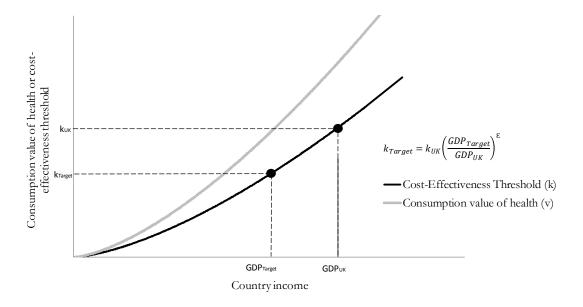


Figure 1: Method for inferring country-specific cost-effectiveness thresholds from UK threshold

4. What is known about the income elasticity of the value of health?

The relationship between the value of a statistical life (VSL) and per capita income at the level of jurisdictions (e.g. of countries) is investigated in a small but emerging literature.[16-18] The literature has evolved out of a longer standing body of work which has examined the relationship between income and health valuation at the level of individuals (i.e. 'within' countries).[15, 17]

Of central interest in both these bodies of work (within-group, at the individual level, and betweengroup, at the level of jurisdictions) is the income elasticity of the value of health and whether this is less than 1 (so health is a necessity good) or greater than 1 (so health is a luxury good); and whether the elasticity is likely to change for different levels of income.[17]

Initial research conducted primarily in higher income countries amongst individuals, and most often in the United States, suggested elasticities in the range 0.4-0.6.[15, 16] These estimates came mainly from cross-sectional studies looking at wage-risk premiums. However, the estimates have been described as 'nonsensical' when extrapolated to lower income countries since the corresponding VSL would be beyond the ranges deemed plausible in such settings.[16]

The methods to estimate income elasticities of VSL have, therefore, been more carefully scrutinized in more recent years. In particular, cross sectional ('within group') estimates from earlier studies have been contrasted with longitudinal or cohort ('between group') studies (which typically estimate elasticities>1; even within countries) and reasons for inconsistencies explored.[16, 17, 19] For instance, Aldy and Smyth (2014) use a life-cycle model applied to US data of the consumption and labour supply choices faced by individuals with uncertain life expectancy and wage income to explain this discrepancy.[19] They argue that cross-sectional studies are more likely to capture changes in realised income, whereas longitudinal or across cohort studies capture the impact of permanent income which is more informative when translating VSL estimates across countries. Estimates of elasticity with respect to realised income are lower as realised income is more variable. Similarly, Getzen (2000) concludes that whereas 'within group' health may be deemed a necessity, analyses of expenditures 'between group' indicate it is almost certainly a luxury good at the national level.[17]

The recent consensus then is that the income elasticity of VSL to transfer estimates across countries should be >1.[16, 18] We choose to apply a range of elasticities (1.0, 1.5 and 2.0) to reflect uncertainty in the literature. Based upon Milligan et al (2014), a function is also applied of an elasticity of 2.5 for countries with gross domestic product (GDP) per capita < US\$10,725 (2005 price year) purchasing power parity (PPP) and of 0.7 for countries with per capita incomes greater than this level².[18]

² In line with the recommendations in Milligan et al. 2014 the elasticities from this study are applied to 2013 PPP-adjusted GDP, deflated to reflect 2005 international dollars. The resulting threshold values are then inflated to reflect 2013 international dollars.

5. Results: country level estimates of cost-effectiveness thresholds

The best estimate of the UK CET is £12,936 per QALY (US\$18,609 purchasing power parity (PPP) adjusted³). Gross domestic product (GDP) per capita estimates were obtained from the World Bank dataset for 2013.⁴ In line with the literature on the value of a statistical life we apply elasticities to countries' GDP per capita adjusted for purchasing power parity (see for example Milligan (2014)).[18] CETs are reported in 2013 PPP adjusted US dollar values. Values without PPP adjustment are also provided for specific countries⁵.

Predicted CETs across country income levels are shown in Figure 2 for a range of income elasticities for the VSL. Higher income elasticities imply lower CETs in countries with lower GDP p.c. compared to the UK, and higher CETs in countries with higher GDP p.c. compared to the UK. The impact of alternative choices of elasticity is larger as the discrepancy between the GDP of the country of interest and UK GDP widens. Results for a selection of specific countries are shown in Table 1.

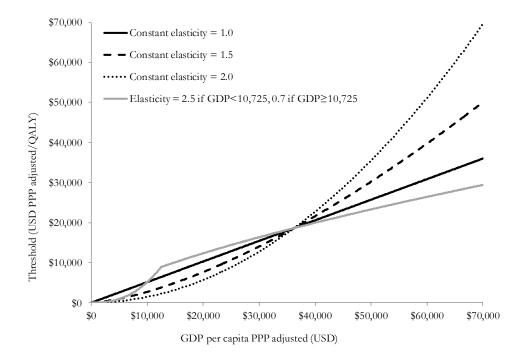


Figure 2: Predicted cost-effectiveness threshold (k) values by country income

³ This analysis uses the PPP conversion factor: <u>http://data.worldbank.org/indicator/PA.NUS.PPP</u>

⁴ This analysis uses PPP adjusted GDP per capita: <u>http://data.worldbank.org/indicator/NY.GDP.PCAP.PP.CD</u>

⁵ This analysis uses data on the ratio of the PPP conversion factor to the market exchange rate to remove the PPP adjustment but retain the presentation in dollars: <u>http://data.worldbank.org/indicator/PA.NUS.PPPC.RF</u>. Results are also presented in comparison to non-PPP adjusted GDP: <u>http://data.worldbank.org/indicator/NY.GDP.PCAP.CD</u>.

| | DDD adjusto | d (2013 USD) | Actual values | (2012) | |
|------------------------|-------------|------------------------------|---------------|------------------------------|------------------------|
| Country/classification | GDP p.c. | Threshold range ⁺ | GDP p.c. | Threshold range ⁺ | Threshold as % GDP |
| Malawi | 780 | 9 - 401 | 226 | 3 - 116 | 1% - 51% |
| Indonesia | 9,559 | 1,298 - 4,914 | 3,475 | 472 - 1,786 | 14% - 51% |
| Chile | 21,911 | 6,819 - 13,141 | 15,732 | 4,896 - 9,436 | 31% - 60% |
| Kazakhstan | 23,206 | 7,648 - 13,675 | 13,610 | 4,485 - 8,018 | 33% - 59% |
| United kingdom | 36,197 | 18,609 - 18,609 | 41,787 | 20,223 - 20,223 | 48% - 48% [#] |
| Canada | 43,247 | 21,051 - 26,564 | 51,958 | 25,292 - 31,915 | 49% - 61% |
| United states | 53,143 | 24,283 - 40,112 | 53,042 | 24,283 - 40,112 | 46% - 75% |
| Norway | 65,461 | 28,057 - 60,862 | 100,819 | 43,211 - 93,736 | 43% - 93% |
| Low/middle income* | 1,045 | 16 - 537 | | | 1% - 51% |
| Middle/high income* | 12,746 | 2,307 - 9,028 | Not available | | 18% - 71% |

Table 1: Example results for a range of countries and the World Bank income classification cut-offs

* We have assumed Gross National Income per capita to be the same as PPP adjusted GDP. These values relate to the income cut-offs for low to middle income and middle to high income countries as defined by the World Bank. ⁺ Reflects range of values obtained when using elasticity estimates of 1.0, 1.5, 2.0 and 2.5 for GDP less than \$10,725 and 0.7 for GDP greater than \$10,725. [#] For the UK the World Bank ratio of PPP conversion factor to market exchange rate did not correspond to the ratio of reported actual GDP to reported PPP-adjusted GDP. The threshold as a % GDP value for the UK therefore depends on whether PPP-adjusted or actual data are used (51% and 48% respectively).

The estimated CETs are substantially lower than those currently used by decision-making agencies and international organizations. Compared to a threshold of US\$50,000 per QALY that has been conventionally applied in the US, our approach estimates a CET in the range US\$24,283-US\$40,112 per QALY. Even more starkly, in comparison to thresholds of 1-3 times GDP per capita recommended by the WHO, the upper bounds of our estimates indicate values of 51% GDP p.c. for countries classified as low income and 71% GDP p.c. for those classified as middle income. US dollar CET values with and without PPP adjustment are provided in the appendix for all countries for which data was available from the World Bank database for 2013. Values without adjustment for PPP can be converted to local currency using standard exchange rates.

6. Discussion

Policymakers in all countries face difficult decisions about how to allocate scarce healthcare resources to maximize health gains for their populations. Cost-effectiveness analysis offers a means by which to compare the costs and health gains from alternative interventions as a basis to inform investment. For the results of CEA to align with population health improvement requires that health gains from recommended interventions exceed the opportunity costs of health forgone when resources are committed to those interventions. The health opportunity costs should be reflected in the cost effectiveness threshold (k). Health opportunity costs can be contrasted to the consumption value of health (v) based upon individuals' expressions of the value of health gains.

In this paper we provide estimates of cost-effectiveness thresholds (the 'k's) in a number of countries based upon assessment of the likely marginal productivity of their healthcare systems. This has required (i) an estimate of the marginal productivity in the English NHS from Claxton et al.; (ii) estimates of the GDP p.c. of the UK and other countries of interest; and (iii) estimates of the elasticity of the value of health ('v') with respect to countries' per capita income (proxied by income elasticities of the value of a statistical life (VSL)). We estimate thresholds far lower than those conventionally applied; which implies that CEAs are likely to be recommending interventions which lead to reductions in population health.

The results presented rely upon some core assumptions if they are to be reasonable extrapolations to reflect the marginal productivity of other (non-UK) healthcare systems: (1) that the discrepancy between the consumption value of health (v) and cost-effectiveness threshold for health (k) is constant in percentage terms across countries; (2) that the UK is 'typical' of other countries with respect to the values of v and k; and (3) that the income elasticity of VSL equals the income elasticity of the consumption value of a QALY.

Two reasons why the proportional discrepancy between k and v may differ across country income levels include: (i) non-constant ratios between individual and societal valuations of health (e.g. due to differences in understanding of health production functions and the trade-off associated with an implied v); (ii) constraints on the size of the health care budget differing for reasons other than societal valuations of healthcare spending (e.g. differences in tax generating abilities due to technical/administrative reasons). The ratio of v/k represents the value of collective compared to private resource or the shadow price of public expenditure.

The assumption that the UK is 'typical' of countries with respect to health production possibilities in part relies upon the UK healthcare system (the NHS) being of 'typical' efficiency in converting health sector resources to health gains for a country of its income level. There is some limited evidence the UK NHS may incur less waste than other health sectors (as indicated for example by relatively high use of generic drugs [20]) or be more efficient due to bodies such as the National Institute for Health and Care Excellence. If the UK is more productive, this means our analysis over-estimates the marginal productivity of other healthcare systems and under-estimates k. However, if the UK is less wasteful than other countries, the UK estimate of k can be considered relevant to a higher income level than UK GDP (as waste is simply lost income), this suggests our analysis may over-estimate the marginal productivity of other healthcare systems and under- or over-estimate k. However, if this were the case, it also implies that perhaps other countries could seek efficiency gains by reassessing what interventions they currently provide and at what cost, rather than evaluating new interventions more leniently.

For the elasticity of the VSL to equal the elasticity of the consumption value of health (v), a statistical life should provide the same units of morbidity-adjusted health (e.g. QALYs) across countries. This could be questioned if lives-saved were expected to generate very different remaining morbidity-adjusted life expectancies. Although life expectancy at birth varies considerably across countries, remaining life expectancy differs much less due to differences in age demographics (e.g. Hammitt and Robinson, 2011, find remaining life expectancies between 34 and 45 years in countries with widely varying per capita incomes as a result of much older populations in countries with higher life expectancies at birth[16]). Again, although quality of life is likely to be increasing in income, older populations would also be expected to have higher levels of morbidity so differences in QALYs gained may also be small.

Therefore, although our results are embedded with many assumptions, it is not immediately clear whether these are likely to lead to our estimates of k being positively or negatively biased.

These results should, however, only be regarded as a first attempt to inform this area of crucial policy importance. The estimates can be used to inform policy decisions, particularly in the absence of other estimates based upon an assessment of health opportunity costs; but they should be applied cautiously. They can in some ways be regarded as a provocation to inspire more research. What research could then be undertaken to better inform thresholds for use within countries?

Aside from the relatively elaborate analysis of Claxton et al. being repeated in every country, which would likely not be possible due to data constraints in many jurisdictions, econometric analyses of policy reforms and other natural experiments could inform the marginal productivity of health sectors. Another approach could be to explore the cost-effectiveness of interventions currently provided within a country and those falling outside of the budget envelope. In this way, policymakers can undergo a process of 'threshold seeking'[21] and become more informed about k as the number of CEAs in their jurisdiction increases. One example of a study using an approach similar to this is from Malawi - the country with the lowest per capita income from those presented in this paper – and suggests a threshold of no more (and perhaps less) than US\$150 in that country.[22] Encouragingly, this is slightly above the range (US\$3-US\$116) that we estimate.

As has already been highlighted, an analysis of cross-country data with countries as units of analysis could be used to inform international estimates of marginal productivity of healthcare. Moreno-Serra and Smith (2014), adopting an approach similar to this but focused on mortality, estimate a marginal cost of only \$145 per life year saved from public healthcare expenditure in a typical low and middle income country.[10] Another alternative could be a hybrid-type approach whereby a small number of within-country empirical estimates of thresholds are produced which are then extrapolated across income levels to provide estimates for a greater number of countries.

All such analyses are likely to have important roles in the quest for greater allocative efficiency, leading to overall population health gains, in health sectors across the world.

7. Conclusion

The cost-effectiveness thresholds presented in this paper are much lower than those previously applied in many countries (e.g. using WHO recommended cost-effectiveness thresholds of 1-3 times GDP per capita). The results have potentially profound consequences for resource allocation. It is hoped they will provoke further research on this topic.

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Appendix: All country values

The table below presents the PPP-adjusted and non-adjusted range of threshold values for each country for which PPP-adjusted GDP was reported in the World Bank database. In some cases data was not available to remove the PPP-adjustment and only PPP-adjusted threshold values are reported.

| Country | Cost-effectiveness threshold range | | |
|--------------------------|------------------------------------|-----------------|--|
| | (USD, PPP adjusted) | (actual USD) | |
| Afghanistan | 56 - 1,023 | 19 - 349 | |
| Albania | 1,563 - 5,816 | 702 - 2,612 | |
| Algeria | 2,514 - 9,300 | 1,012 - 3,743 | |
| Angola | 807 - 3,875 | 603 - 2,897 | |
| Antigua and Barbuda | 6,250 - 12,750 | 3,965 - 8,090 | |
| Armenia | 858 - 3,997 | 387 - 1,801 | |
| Australia | 21,153 - 26,938 | 32,771 - 41,732 | |
| Austria | 21,355 - 27,684 | 23,727 - 30,759 | |
| Azerbaijan | 4,172 - 11,085 | 1,901 - 5,051 | |
| Bahrain | 21,245 - 27,277 | 11,962 - 15,358 | |
| Bangladesh | 93 - 1,315 | 30 - 427 | |
| Belarus | 4,407 - 11,297 | 1,895 - 4,857 | |
| Belgium | 20,060 - 23,111 | 22,570 - 26,003 | |
| Belize | 1,012 - 4,340 | 584 - 2,503 | |
| Benin | 46 - 921 | 20 - 414 | |
| Bhutan | 835 - 3,943 | 267 - 1,258 | |
| Bolivia | 534 - 3,151 | 250 - 1,474 | |
| Bosnia and Herzegovina | 1,318 - 4,952 | 644 - 2,421 | |
| Botswana | 3,490 - 10,419 | 1,621 - 4,839 | |
| Brazil | 3,210 - 10,122 | 2,393 - 7,544 | |
| Brunei Darussalam | 29,901 - 73,137 | 16,065 - 39,294 | |
| Bulgaria | 3,609 - 10,541 | 1,720 - 5,025 | |
| Burkina Faso | 38 - 840 | 17 - 379 | |
| Burundi | 8 - 396 | 3 - 137 | |
| Cabo Verde | 584 - 3,297 | 343 - 1,935 | |
| Cambodia | 131 - 1,564 | 44 - 518 | |
| Cameroon | 104 - 1,394 | 49 - 654 | |
| Canada | 21,051 - 26,564 | 25,292 - 31,915 | |
| Central African Republic | 5 - 310 | 3 - 171 | |
| Chad | 61 - 1,070 | 31 - 540 | |
| Chile | 6,819 - 13,141 | 4,896 - 9,436 | |
| China | 2,013 - 7,957 | 1,151 - 4,550 | |
| Colombia | 2,174 - 8,754 | 1,370 - 5,518 | |
| Comoros | 35 - 801 | 19 - 452 | |
| Congo, Dem. Rep. | 8 - 384 | 5 - 230 | |
| Congo, Rep. | 489 - 3,016 | 264 - 1,628 | |
| Costa Rica | 2,733 - 9,574 | 2,006 - 7,027 | |
| | | | |

| Country | Cost-effectiveness threshold range | | |
|----------------------|------------------------------------|-----------------|--|
| | (USD, PPP adjusted) | (actual USD) | |
| Cote d'Ivoire | 129 - 1,548 | 61 - 737 | |
| Croatia | 6,206 - 12,720 | 3,953 - 8,101 | |
| Cyprus | 12,318 - 16,130 | 11,020 - 14,430 | |
| Czech Republic | 10,620 - 15,322 | 7,325 - 10,569 | |
| Denmark | 20,888 - 25,974 | 28,767 - 35,771 | |
| Djibouti | 128 - 1,541 | 71 - 857 | |
| Dominica | 1,429 - 5,205 | 991 - 3,611 | |
| Dominican Republic | 1,943 - 7,618 | 937 - 3,675 | |
| Ecuador | 1,557 - 5,788 | 858 - 3,191 | |
| Egypt, Arab Rep. | 1,745 - 6,669 | 522 - 1,993 | |
| El Salvador | 856 - 3,991 | 422 - 1,967 | |
| Equatorial Guinea | 16,150 - 17,717 | 9,843 - 10,798 | |
| Eritrea | 20 - 615 | 9 - 280 | |
| Estonia | 8,912 - 14,418 | 6,574 - 10,636 | |
| Ethiopia | 26 - 696 | 10 - 255 | |
| Fiji | 897 - 4,086 | 507 - 2,307 | |
| Finland | 19,334 - 20,781 | 23,867 - 25,653 | |
| France | 18,861 - 19,347 | 21,168 - 21,713 | |
| Gabon | 5,268 - 12,018 | 3,164 - 7,218 | |
| Gambia, The | 39 - 857 | 12 - 252 | |
| Georgia | 729 - 3,683 | 366 - 1,850 | |
| Germany | 21,080 - 26,668 | 21,933 - 27,747 | |
| Ghana | 224 - 2,043 | 104 - 951 | |
| Greece | 9,345 - 14,658 | 7,982 - 12,520 | |
| Grenada | 1,878 - 7,302 | 1,272 - 4,948 | |
| Guatemala | 756 - 3,750 | 360 - 1,788 | |
| Guinea | 22 - 645 | 9 - 269 | |
| Guinea-Bissau | 22 - 639 | 9 - 256 | |
| Guyana | 610 - 3,368 | 348 - 1,924 | |
| Haiti | 41 - 875 | 20 - 421 | |
| Honduras | 299 - 2,360 | 149 - 1,177 | |
| Hong Kong SAR, China | 24,302 - 40,202 | 17,409 - 28,801 | |
| Hungary | 7,434 - 13,540 | 4,268 - 7,773 | |
| Iceland | 19,942 - 22,720 | 22,567 - 25,712 | |
| India | 416 - 2,781 | 115 - 770 | |
| Indonesia | 1,298 - 4,914 | 472 - 1,786 | |
| Iran, Islamic Rep. | 3,450 - 10,378 | 1,054 - 3,171 | |
| Iraq | 3,276 - 10,194 | 1,504 - 4,679 | |
| Ireland | 21,071 - 26,634 | 23,063 - 29,153 | |
| Israel | 15,243 - 17,366 | 16,821 - 19,163 | |
| Italy | 16,712 - 17,928 | 16,867 - 18,094 | |
| Jamaica | 1,122 - 4,570 | 668 - 2,719 | |
| Japan | 18,651 - 18,731 | 19,769 - 19,854 | |

| Country | Cost-effectiveness threshold range | | |
|-----------------------|------------------------------------|------------------|--|
| | (USD, PPP adjusted) | (actual USD) | |
| Jordan | 1,971 - 7,757 | 872 - 3,432 | |
| Kazakhstan | 7,648 - 13,675 | 4,485 - 8,018 | |
| Kenya | 73 - 1,164 | 32 - 519 | |
| Kiribati | 49 - 954 | 43 - 848 | |
| Korea, Rep. | 15,598 - 17,505 | 12,227 - 13,722 | |
| Kosovo | 1,085 - 4,493 | 473 - 1,961 | |
| Kyrgyz Republic | 147 - 1,651 | 58 - 649 | |
| Lao PDR | 329 - 2,474 | 113 - 852 | |
| Latvia | 7,532 - 13,602 | 5,133 - 9,270 | |
| Lebanon | 4,187 - 11,098 | 2,420 - 6,416 | |
| Lesotho | 95 - 1,329 | 41 - 581 | |
| Liberia | 11 - 451 | 6 - 234 | |
| Libya | 6,503 - 12,927 | 3,697 - 7,349 | |
| Lithuania | 9,175 - 14,565 | 5,598 - 8,886 | |
| Luxembourg | 35,195 - 117,072 | 43,092 - 143,342 | |
| Macao SAR, China | 48,116 - 288,671 | 30,832 - 184,977 | |
| Macedonia, FYR | 1,978 - 7,791 | 824 - 3,246 | |
| Madagascar | 28 - 717 | 9 - 235 | |
| Malawi | 9 - 401 | 3 - 116 | |
| Malaysia | 7,709 - 13,712 | 3,481 - 6,192 | |
| Maldives | 1,929 - 7,550 | 1,103 - 4,318 | |
| Mali | 38 - 844 | 17 - 368 | |
| Malta | 12,965 - 16,419 | 10,138 - 12,838 | |
| Marshall Islands | 196 - 1,908 | 182 - 1,774 | |
| Mauritania | 131 - 1,564 | 46 - 550 | |
| Mauritius | 4,202 - 11,112 | 2,248 - 5,945 | |
| Mexico | 3,850 - 10,780 | 2,410 - 6,749 | |
| Micronesia, Fed. Sts. | 180 - 1,829 | 162 - 1,646 | |
| Moldova | 310 - 2,400 | 148 - 1,151 | |
| Mongolia | 1,264 - 4,849 | 543 - 2,085 | |
| Montenegro | 2,912 - 9,786 | 1,464 - 4,921 | |
| Morocco | 736 - 3,702 | 316 - 1,590 | |
| Mozambique | 16 - 537 | 8 - 294 | |
| Namibia | 1,332 - 4,979 | 791 - 2,958 | |
| Nepal | 72 - 1,154 | 22 - 357 | |
| Netherlands | 21,104 - 26,757 | 23,153 - 29,354 | |
| New Zealand | 17,226 - 18,117 | 20,555 - 21,619 | |
| Nicaragua | 297 - 2,350 | 118 - 937 | |
| Niger | 12 - 469 | 5 - 213 | |
| Nigeria | 446 - 2,880 | 239 - 1,545 | |
| Norway | 28,057 - 60,862 | 43,211 - 93,736 | |
| Oman | 21,322 - 27,562 | , 00,.00 | |
| Pakistan | 314 - 2,416 | 87 - 669 | |

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| Country | Cost-effectiveness threshold range | | |
|--------------------------------|------------------------------------|------------------|--|
| | (USD, PPP adjusted) | (actual USD) | |
| Palau | 3,235 - 10,149 | 2,531 - 7,940 | |
| Panama | 5,352 - 12,083 | 3,042 - 6,869 | |
| Papua New Guinea | 92 - 1,305 | 75 - 1,073 | |
| Paraguay | 919 - 4,135 | 484 - 2,179 | |
| Peru | 1,969 - 7,747 | 1,114 - 4,383 | |
| Philippines | 606 - 3,358 | 256 - 1,421 | |
| Poland | 7,694 - 13,703 | 4,440 - 7,908 | |
| Portugal | 9,527 - 14,756 | 7,738 - 11,985 | |
| Puerto Rico | 17,145 - 18,088 | 14,075 - 14,849 | |
| Qatar | 45,558 - 246,565 | 31,105 - 168,345 | |
| Romania | 4,932 - 11,746 | 2,467 - 5,875 | |
| Russian Federation | 8,263 - 14,046 | 5,007 - 8,511 | |
| Rwanda | 30 - 746 | 13 - 323 | |
| Samoa | 363 - 2,598 | 265 - 1,897 | |
| Sao Tome and Principe | 125 - 1,527 | 68 - 827 | |
| Saudi Arabia | 24,484 - 41,080 | 11,799 - 19,797 | |
| Senegal | 73 - 1,166 | 34 - 544 | |
| Serbia | 2,175 - 8,760 | 1,061 - 4,275 | |
| Seychelles | 8,310 - 14,074 | 5,470 - 9,265 | |
| Sierra Leone | 53 - 990 | 23 - 435 | |
| Singapore | 31,889 - 88,068 | 22,342 - 61,701 | |
| Slovak Republic | 9,686 - 14,841 | 6,561 - 10,053 | |
| Slovenia | 11,374 - 15,690 | 9,135 - 12,603 | |
| Solomon Islands | 61 - 1,063 | 57 - 1,004 | |
| South Africa | 2,221 - 8,909 | 1,175 - 4,714 | |
| South Sudan | 77 - 1,198 | 40 - 617 | |
| Spain | 14,638 - 17,124 | 13,277 - 15,531 | |
| Sri Lanka | 1,346 - 5,005 | 453 - 1,686 | |
| St. Kitts and Nevis | 6,222 - 12,731 | 4,110 - 8,409 | |
| St. Lucia | 1,584 - 5,914 | 1,107 - 4,133 | |
| St. Vincent and the Grenadines | 1,615 - 6,058 | 998 - 3,746 | |
| Sudan | 162 - 1,734 | 84 - 901 | |
| Suriname | 3,740 - 10,672 | 2,286 - 6,525 | |
| Swaziland | 634 - 3,436 | 288 - 1,559 | |
| Sweden | 21,148 - 26,917 | 28,306 - 36,028 | |
| Switzerland | 24,450 - 40,914 | 36,661 - 61,348 | |
| Tajikistan | 90 - 1,291 | 37 - 533 | |
| Tanzania | 45 - 912 | 18 - 357 | |
| Thailand | 2,941 - 9,820 | 1,181 - 3,943 | |
| Timor-Leste | 71 - 1,153 | | |
| Togo | 27 - 715 | 13 - 327 | |
| Tonga | 399 - 2,726 | 333 - 2,275 | |
| Trinidad and Tobago | 13,159 - 16,503 | 7,941 - 9,959 | |

| Country | Cost-effectiveness threshold range | | |
|----------------|------------------------------------|-----------------|--|
| - | (USD, PPP adjusted) | (actual USD) | |
| Tunisia | 1,747 - 6,680 | 678 - 2,592 | |
| Turkey | 5,114 - 11,895 | 2,950 - 6,861 | |
| Turkmenistan | 2,784 - 9,635 | 1,588 - 5,495 | |
| Tuvalu | 188 - 1,870 | 200 - 1,991 | |
| Uganda | 28 - 725 | 11 - 293 | |
| Ukraine | 1,097 - 4,518 | 487 - 2,005 | |
| United Kingdom | 18,609 - 18,609 | 20,223 - 20,223 | |
| United States | 24,283 - 40,112 | 24,283 - 40,112 | |
| Uruguay | 5,450 - 12,160 | 4,548 - 10,147 | |
| Uzbekistan | 379 - 2,656 | 138 - 965 | |
| Vanuatu | 127 - 1,538 | 139 - 1,685 | |
| Venezuela, RB | 4,701 - 11,553 | 3,724 - 9,151 | |
| Vietnam | 398 - 2,721 | 144 - 982 | |
| Yemen, Rep. | 223 - 2,035 | 83 - 757 | |
| Zambia | 144 - 1,635 | 68 - 768 | |
| Zimbabwe | 41 - 874 | 21 - 455 | |