



This is a repository copy of *A test based on panel refreshments for panel conditioning in stated utility measures*.

White Rose Research Online URL for this paper:
<http://eprints.whiterose.ac.uk/99204/>

Version: Accepted Version

Article:

Van Landeghem, B. (2014) A test based on panel refreshments for panel conditioning in stated utility measures. *Economics Letters*, 124 (2). pp. 236-238. ISSN 0165-1765

<https://doi.org/10.1016/j.econlet.2014.05.024>

© 2014 Elsevier. This is an author produced version of a paper subsequently published in *Economics Letters*. Uploaded in accordance with the publisher's self-archiving policy. Article available under the terms of the CC-BY-NC-ND licence (<https://creativecommons.org/licenses/by-nc-nd/4.0/>)

Reuse

This article is distributed under the terms of the Creative Commons Attribution-NonCommercial-NoDerivs (CC BY-NC-ND) licence. This licence only allows you to download this work and share it with others as long as you credit the authors, but you can't change the article in any way or use it commercially. More information and the full terms of the licence here: <https://creativecommons.org/licenses/>

Takedown

If you consider content in White Rose Research Online to be in breach of UK law, please notify us by emailing eprints@whiterose.ac.uk including the URL of the record and the reason for the withdrawal request.



eprints@whiterose.ac.uk
<https://eprints.whiterose.ac.uk/>

A Test Based on Panel Refreshments for Panel Conditioning in Stated Utility Measures

(Forthcoming in Economics Letters)

Bert Van Landeghem

University of Sheffield, Maastricht University, and IZA

b.vanlandeghem@sheffield.ac.uk

April 2014

Abstract

Adapting a methodology proposed in Das et al. (2011), this paper uses panel refreshments as a natural experiment to determine whether trends in stated utility measures observed in panel data are genuine or rather caused by measurement issues.

Keywords: Stated Utility; Panel Attrition; Panel Conditioning; Time Trends.

JEL Codes: C18, D60

1 Introduction

It is increasingly common practice in economics to use ‘stated utility’ measures, which are derived from survey questions asking people about their satisfaction with life, income, or health. While such measures have passed important external validity checks, in some circumstances measurement and methodological issues need to be taken seriously when interpreting empirical results (Benjamin et al., 2013; Conti and Pudney, 2011; Heffetz and Rabin, 2013).

Many researchers have noticed a sharp drop in stated utility measures in the first rounds of a commonly used panel containing such data, the German Socioeconomic Panel. It is important to find out whether such a drop reflects a genuine decline in stated utility in the society the dataset represents, or whether it is rather due to panel conditioning or a panel effect, a change in the way questions measuring stated utility are answered the more experienced one is in answering such questions. This paper aims to help make progress in solving this puzzle inspired by an identification strategy proposed in Das et al. (2011). The contribution of the paper is first of all empirical since it studies a different type of data than Das et al. (2011). Second, the long time dimension of the panel data allows us to get insights into the dynamics of panel conditioning, and to deal more easily with the consequences of a relatively large attrition rate.

2 Data

The analysis will use panel refreshers (newcomers in the panel) as a natural experiment to obtain identification. To ensure a higher degree of external validity, two different nationally representative datasets are being used, the German Socioeconomic Panel (German SOEP) and the Swiss Household Panel (SHP).

As for the German SOEP, the original West German sample started in 1984, and the original East German sample in 1990. Comparable refreshment samples (for both West and East) were introduced in 1998, 2000 and 2006. The Swiss Household Panel (SHP) started in 1999 and a refreshment sample was introduced in 2004. A robustness check with the SHP seems particularly useful since the mode of interview (telephone) is the same across all interviews while there are several small nonrandom variations in the face-to-face set-up in the German SOEP.

Both datasets contain an identical measure of stated utility in every round from the first round onwards for the German SOEP, and the second round onwards for the SHP, which is derived from the following survey question:

“On a scale from 0 (completely dissatisfied) to 10 (completely satisfied): How satisfied are you with your life, all things considered?”

In addition to refreshers stemming from the introduction of a refreshment sample (“induced refreshers”), a second group of “natural refreshers” enter steadily each year, and can stem from three sources. First, members from interviewed households will reach the eligible age to enter the panel. Second, if a new member eligible for the survey moves into the household, the enumerator will attempt to interview this new member as well. Third, if a household member leaves the household, the aim is to follow the respondent and, at the same time, try to interview other members eligible for the survey in that person’s new household. Generally, first-time respondents not stemming from a refreshment sample account for around 3% of the respondents.

3 Estimation and Results

Das et al. (2011) illustrate with dichotomous variables that one can still calculate upper and lower bounds of the panel effect if no assumptions about the attrition process are

to be made. In this context, however, bounds seem not informative due to the much larger scale of the dependent variable and due to a relatively high attrition rate. Many strategies are available that will, under different identifying assumptions, lead to point estimates. In this case, since the datasets used in this paper have refreshment samples introduced well before the last calendar year for which data are available, one can restrict the newcomers to those who will stay in the panel for several additional years to make the two samples more comparable and to obtain insights in the sign of the attrition bias.

Columns 1, 3, and 5 in Table 1, show average differences between new respondents and more experienced respondents for each calendar year in which a refreshment sample was introduced. The other three columns show a replication of this analysis when restricting the sample of newcomers to those who will at least stay for three additional consecutive years. To make first-time respondents and more experienced respondents as comparable as possible, the differences are conditional on socioeconomic variables that are likely not to be prone to panel conditioning themselves. One important covariate is a dummy taking one if the respondent entered the panel as a natural refresher, since the latter entered the panel through a different sampling procedure than the others.

In all cases, scores in the calendar year in which a refreshment sample is started are statistically and substantially higher for the refreshment sample than for the more experienced sample. When the correction for attrition biases is applied, the results are even more pronounced. The strategy to correct for attrition will, however, only lead us to correct conclusions about the sign of the attrition bias if the attrition processes in the first years of the panel are similar for the different samples, or if at least the sign of the attrition bias for the stated utility measure is similar for the different samples. This is investigated by running pooled OLS regressions for each of the samples, with stated utility as the dependent variable, and a selectivity dummy, a set of controls and time dummies as independent variables. Following Kapteyn et al. (2005), the selectivity dummy equals one when the individual will be participating in the next round of the

Table 1: Conditional Differences in Stated utility Between New Respondents and More Experienced Respondents for Years with Substantial Panel Refreshments

	West Germany		East Germany		Switzerland	
	ur	r	ur	r	ur	r
1998	0.505*** (0.051)	0.595*** (0.059)	0.525*** (0.090)	0.636*** (0.107)		
2000	0.339*** (0.031)	0.421*** (0.034)	0.399*** (0.051)	0.448 (0.054)		
2004					0.184*** (0.035)	0.215*** (0.046)
2006	0.491*** (0.041)	0.567*** (0.052)	0.742*** (0.075)	0.760*** (0.092)		

Robust standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Source: German Socioeconomic Panel and Swiss Household Panel.

ur = unrestricted, and **r** = restricted, which means that the newcomers are restricted to those who will at least remain in the panel for three additional consecutive years.

All regressions include age and labour force status dummies, education dummies, marital status dummies and a dummy for having entered the panel as natural refresher.

panel, and 0 otherwise. The coefficients on the selectivity dummies can inform us about the sign of the bias caused by attrition, if we are willing to make the assumption which Das et al. (2011) call “stationary attrition bias”: the difference of stated utility between the entire population and those who will not drop out remains stable over time. The regression results (available upon request) show substantial and positive coefficients on all the selectivity dummies, which are significant at conventional significance levels in all but one of the samples.

There might be remaining concerns that the results are driven by the natural refreshers that have been continuously entering the panel, and that adding a dummy to control for this is not sufficient. Hence, Table 2 compares the subsequent German SOEP’s refreshment samples with each other, as well as the SHP’s main and refreshment sample, in the case when the natural refreshers are included, and in the case when they are omitted. Even though there is some variation in the coefficients’ size, the results show that the evidence for panel conditioning remains when comparing the refreshment samples with each other, and do not change much after removing the natural refreshers from the analysis.

Finally, since the data include newcomers in every year, and since many rounds of data are available, there is an opportunity for a further extension to get more precise insights into the size and dynamics of panel conditioning. A subsample of individuals is selected who are interviewed in at least five consecutive rounds. The stated utility measure is regressed on dummies indicating whether one is asked the stated utility question for the first, second, . . . or fifth time¹ as well as on a set of covariates.

Results in Table 3 show us that the panel effect is not entirely established between being interviewed for the first and second time but that it accumulates over the different survey rounds. In West and East Germany, there is a substantial panel effect from the second to the fifth interview, which is also statistically significant at any conventional

¹Hence, the baseline category is having answered six times or more.

Table 2: Conditional Differences in Stated Utility Between New Respondents and Respondents Stemming from the Previous Substantial Refreshment: Analyses With and Without Natural Refreshers

	Germany	Switzerland	Germany	Switzerland
	Including Natural Refreshers		Excluding Natural Refreshers	
new 2000	0.093** (0.047)		0.088* (0.049)	
new 2004		0.215*** (0.046)		0.216*** (0.046)
new 2006	0.486*** (0.048)		0.489*** (0.049)	

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Source: German Socioeconomic Panel and Swiss Household Panel.
Controls are similar as in Table 1, and the correction for panel attrition has been applied.

significance level.² The pattern for Switzerland is in line with that for Germany, but no panel effect is measured after the third interview.

The coefficient on the dummy for being interviewed for the fifth time can be interpreted as the negative of a residual panel effect, that is, which will be established over all the interviews after the fifth has taken place.³ For West and East Germany, this coefficient still has a substantial magnitude of 0.18 and 0.29, respectively. It is rather speculative why in the Swiss data the cumulation path is somewhat shorter than in the German data, but it might be due to a different mode of interview or due to the fact that the key question was only asked from the second year onwards in the SHP, when respondents had already some experience with answering surveys.

²A panel effect for interview n is calculated as the difference in coefficients between the n th and $n - 1$ th interview.

³Contrary to the path of panel effects from the first to the fifth interviews, the estimate for a residual panel effect might be slightly biased due to panel attrition. Indeed, some respondents will no longer be interviewed after the fifth interview, while others will remain in the panel for many years.

Table 3: The Path of Panel Effects over Interviews: Results for West Germany, East Germany, and Switzerland

	West Germany	East Germany	Switzerland
1st interview	0.488*** (0.023)	0.463*** (0.040)	0.231*** (0.059)
2nd interview	0.379*** (0.022)	0.345*** (0.037)	0.105** (0.043)
3rd interview	0.288*** (0.021)	0.317*** (0.035)	0.072* (0.039)
4th interview	0.205*** (0.020)	0.261*** (0.033)	0.061* (0.036)
5th interview	0.181*** (0.020)	0.241*** (0.032)	0.050 (0.032)
Observations	211,847	74,727	50,206
R-squared	0.047	0.075	0.050

Robust standard errors in parentheses

*** p<0.01, ** p<0.05, * p<0.1

Source: German Socioeconomic Panel and Swiss Household Panel.

All regressions include marital status dummies, labour force status dummies, education dummies, age and time dummies, and a dummy for having entered the panel as natural refresher.

4 Conclusion

This paper has used panel refreshers as a source of identification to show that a time trend of stated utility measures one often observes in panel data can be attributed to panel conditioning, that is, answering questions differently the longer one has been in the panel. Reasons for this might be numerous and it is likely that an interplay of factors is at work, e.g. cognitive biases as described in Kahneman and Krueger (2006) might change, or one might interact differently with the interviewer (Chadi, 2013). There seems to be some variation in panel effects across the different samples in the analysis. Since the stated utility question is only asked from the second wave onwards in the SHP, the smaller panel effect in Swiss data in Table 3 might reveal that survey participation as a whole rather than having repeatedly answered the stated utility question causes the panel effect. Interaction effects of the panel effect with other variables (such as survey design characteristics and macroeconomic shocks), as well as sampling error might be other explanations for variations in the size of the estimated panel effect.

As it seems that panel conditioning can be identified without strong arbitrary assumptions, as panel attrition seems to bias the estimated effect towards zero rather than the reverse, and as the phenomenon does not seem to be bound to one particular dataset, the results might imply important consequences for interpreting and designing studies that look at a society's (or specific cohort's) well-being over time.

Acknowledgments

Acknowledgments

I thank Eric Bonsang, Sarah Brown, Gabriella Conti, Arnaud Dupuy, Richard Easterlin, Arjan Non, Andrew Oswald, Jennifer Roberts, Katrin Sommerfeld, Raphael Studer, and

Arthur van Soest for valuable comments, discussions and advice. I am also grateful to seminar and conference participants in Berlin, Bonn, Cambridge, Freiburg, Maastricht, Malaga, Philadelphia and Milan. This project was partly funded by a Rubicon grant from the Dutch Scientific Organization (NWO).

References

Benjamin, D., O. Heffetz, M. Kimball and A. Rees-Jones (2013) Can Marginal Rates of Substitution Be Inferred from Happiness Data? Evidence from Residency Choices. *American Economic Review*, *Forthcoming*.

Chadi, A. (2013) The Role of Interviewer Encounters in Panel Responses on Life Satisfaction. *Economics Letters* 121, 550-554.

Conti, G. and S. Pudney (2011) Survey Design and the Analysis of Satisfaction. *Review of Economics and Statistics* 93, 1087-1093.

Das, J., A. van Soest and V. Toepoel (2011) Nonparametric Tests of Panel Conditioning and Attrition Bias in Panel Surveys. *Sociological Methods and Research* 40, 32-56.

Heffetz, O. and M. Rabin (2013) Conclusions Regarding Cross-Group Differences in Happiness Depend on Difficulty of Reaching Respondents. *American Economic Review* 103, 3001-3021.

Kahneman, D. and A. Krueger (2006) Developments in the Measurement of Subjective Well-being. *Journal of Economic Perspectives* 20, 3-24.

Kapteyn, A., R. Alessie and A. Lusardi (2005) Explaining the Wealth Holdings of Different Cohorts: Productivity Growth and Social Security. *European Economic Review* 49, 1361-1391.