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Appendix S3: Model Selection and Wood's Test Statistic

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1 Procedures

Fit models predicting the demographic rate using body size, current environment, and the day by treatment. Fit models that contain all subsets of predictors. Find the best subset of predictors using AICc.

2 Preliminaries

```
library(reshape)
library(gamm4)
library(bbmle) #for AICtab
library(plyr)
```

Read in data that was already organized.

```
load("../Individual soil mite data.Rdata")
```

Set the maximum number of knots for the splines.

```
ke=10 #environment
kd=10 #days
kl=5 #length
```

3 Growth

```
temp=na.omit(subset(Indiv, stage!="egg")[
  ,c("10.s", "124", "density.s", "supply.s",
    "day", "trt", "10", "pop.ID", "stage")])
temp=transform(temp, 124.s=scale(124))
```

```
gm_full=gamm4(124~s(10.s, k=kl)+s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  random=~(1|pop.ID), data=temp, REML=FALSE)
gm_eb=gamm4(124~s(10.s, k=kl)+s(density.s, supply.s, k=ke),
  random=~(1|pop.ID), data=temp, REML=FALSE)
gm_et=gamm4(124~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  random=~(1|pop.ID), data=temp, REML=FALSE)
gm_bt=gamm4(124~s(10.s, k=kl)+trt+s(day, by=trt, k=kd),
  random=~(1|pop.ID), data=temp, REML=FALSE)
gm_b=gamm4(124~s(10.s, k=kl),
  random=~(1|pop.ID), data=temp, REML=FALSE)
gm_e=gamm4(124~s(density.s, supply.s, k=ke),
  random=~(1|pop.ID), data=temp, REML=FALSE)
```

```

gm_t=gamm4(l24~trt+s(day, by=trt, k=kd),
           random=~(1|pop.ID), data=temp, REML=FALSE)
gm_0=gamm4(l24~1,
           random=~(1|pop.ID), data=temp, REML=FALSE)

gm_sfull=gamm4(l24~stage+s(10.s, k=kl)+s(density.s, supply.s, k=ke)+
              trt+s(day, by=trt, k=kd), random=~(1|pop.ID), data=temp, REML=FALSE)

## Warning in optwrap(optimizer, devfun, getStart(start, rho$lower,
rho$pp), : convergence code 1 from bobyqa: bobyqa -- maximum number
of function evaluations exceeded

gm_seb=gamm4(l24~stage+s(10.s, k=kl)+s(density.s, supply.s, k=ke),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_set=gamm4(l24~stage+s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_sbt=gamm4(l24~stage+s(10.s, k=kl)+trt+s(day, by=trt, k=kd),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_sb=gamm4(l24~stage+s(10.s, k=kl),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_se=gamm4(l24~stage+s(density.s, supply.s, k=ke),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_st=gamm4(l24~stage+trt+s(day, by=trt, k=kd),
            random=~(1|pop.ID), data=temp, REML=FALSE)
gm_s0=gamm4(l24~stage,
            random=~(1|pop.ID), data=temp, REML=FALSE)

```

gm_sfull had a warning. It might be ok. So check out the components.

```

summary(gm_sfull$gam)

##
## Family: gaussian
## Link function: identity
##
## Formula:
## l24 ~ stage + s(10.s, k = kl) + s(density.s, supply.s, k = ke) +
##      trt + s(day, by = trt, k = kd)
##

```

```

## Parametric coefficients:
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.541814  0.003953 137.072 < 2e-16 ***
## stageproto   0.002755  0.002544   1.083  0.27892
## stagetrity   -0.003574  0.004140  -0.863  0.38814
## stagemale    -0.020074  0.004847  -4.142 3.58e-05 ***
## stagefemale  -0.016708  0.005102  -3.275  0.00107 **
## trtT2         -0.002651  0.003354  -0.790  0.42934
## trtT3         -0.005223  0.002083  -2.508  0.01221 *
## trtT4         0.001858  0.002069   0.898  0.36942
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##           edf Ref.df      F p-value
## s(l0.s)          3.918  3.918 5504.201 < 2e-16 ***
## s(density.s,supply.s) 5.828  5.828   3.063 0.00618 **
## s(day):trtT1      1.000  1.000   0.031 0.85984
## s(day):trtT2      1.000  1.000   2.863 0.09079 .
## s(day):trtT3      1.000  1.000   7.842 0.00515 **
## s(day):trtT4      3.770  3.770   6.225 0.00010 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.993
## lmer.REML = -11404  Scale est. = 0.00024229  n = 2087

summary(gm_sfull$mer)

## Linear mixed model fit by maximum likelihood ['lmerMod']
##
##           AIC          BIC    logLik deviance df.resid
## -11357.9 -11228.1  5702.0 -11403.9     2064
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.6189 -0.5956 -0.0353  0.5737  3.5610
##
## Random effects:
##  Groups   Name                Variance Std.Dev.

```

```

## pop.ID (Intercept) 1.532e-06 0.001238
## Xr.4 s(day):trtT4 3.813e-02 0.195262
## Xr.3 s(day):trtT3 0.000e+00 0.000000
## Xr.2 s(day):trtT2 0.000e+00 0.000000
## Xr.1 s(day):trtT1 0.000e+00 0.000000
## Xr.0 s(density.s,supply.s) 7.814e-03 0.088398
## Xr s(10.s) 1.134e+00 1.065099
## Residual 2.423e-04 0.015566
## Number of obs: 2087, groups:
## pop.ID, 8; Xr.4, 8; Xr.3, 8; Xr.2, 8; Xr.1, 8; Xr.0, 7; Xr, 3
##
## Fixed effects:
## Estimate Std. Error t value
## X(Intercept) 0.5418138 0.0039528 137.07
## Xstageproto 0.0027550 0.0025437 1.08
## Xstagetrityto -0.0035738 0.0041402 -0.86
## Xstagemale -0.0200745 0.0048468 -4.14
## Xstagefemale -0.0167082 0.0051020 -3.27
## XtrtT2 -0.0026515 0.0033543 -0.79
## XtrtT3 -0.0052233 0.0020825 -2.51
## XtrtT4 0.0018579 0.0020695 0.90
## Xs(10.s)Fx1 0.1918407 0.0030595 62.70
## Xs(density.s,supply.s)Fx1 -0.0029988 0.0011353 -2.64
## Xs(density.s,supply.s)Fx2 0.0019908 0.0009350 2.13
## Xs(day):trtT1Fx1 -0.0001325 0.0007501 -0.18
## Xs(day):trtT2Fx1 -0.0044768 0.0026458 -1.69
## Xs(day):trtT3Fx1 -0.0039615 0.0014146 -2.80
## Xs(day):trtT4Fx1 -0.0204536 0.0053623 -3.81
##
## Correlation of Fixed Effects:
## X(Int) Xstgpr Xstgtr Xstgml Xstgfm XtrtT2 XtrtT3 XtrtT4 X(0.)F
## Xstageproto -0.724
## Xstagetrityto -0.881 0.836
## Xstagemale -0.909 0.712 0.923
## Xstagefemale -0.904 0.688 0.888 0.968
## XtrtT2 -0.324 0.018 0.024 0.019 0.014
## XtrtT3 -0.338 0.012 0.020 0.022 0.026 0.711
## XtrtT4 -0.360 0.031 0.050 0.062 0.075 0.560 0.580
## Xs(10.s)Fx1 0.658 -0.674 -0.757 -0.645 -0.620 -0.076 -0.053 -0.105

```

```

## Xs(dn.,.)F1 -0.243  0.048  0.075  0.095  0.116  0.609  0.435  0.156 -0.089
## Xs(dn.,.)F2  0.092 -0.015 -0.025 -0.033 -0.048 -0.318 -0.253  0.083  0.047
## Xs(dy):T1F1 -0.020 -0.018 -0.009  0.006  0.021  0.060  0.039  0.022  0.008
## Xs(dy):T2F1 -0.211  0.008  0.014  0.018  0.017  0.809  0.513  0.299 -0.060
## Xs(dy):T3F1 -0.184 -0.025 -0.022 -0.016 -0.010  0.709  0.519  0.295 -0.034
## Xs(dy):T4F1 -0.109  0.033  0.037  0.033  0.034  0.259  0.196  0.120 -0.044
##
##          X(.,.)F1 X(.,.)F2 X():T1 X():T2 X():T3
## Xstageproto
## Xstagetrito
## Xstagemale
## Xstagefemal
## XtrtT2
## XtrtT3
## XtrtT4
## Xs(10.s)Fx1
## Xs(dn.,.)F1
## Xs(dn.,.)F2 -0.406
## Xs(dy):T1F1  0.062  -0.010
## Xs(dy):T2F1  0.657  -0.401  0.117
## Xs(dy):T3F1  0.655  -0.431  0.132  0.767
## Xs(dy):T4F1  0.255  -0.064  0.027  0.263  0.265
## convergence code: 1

```

Some of the random effects variance components associated with treatment splines (past environment variable) are 0. The maximum correlation of fixed effects is 0.96. This is weak evidence that we should possibly stick to the simpler model `gm_full`. So to be safe we omit `gm_sfull`.

```

AICctab(gm_full$mer, gm_eb$mer, gm_et$mer, gm_bt$mer,
        gm_b$mer, gm_e$mer, gm_t$mer, gm_0$mer,
        gm_seb$mer, gm_set$mer, gm_sbt$mer, gm_sb$mer,
        gm_se$mer, gm_st$mer, gm_s0$mer, nobs=nrow(temp))

##          dAICc  df
## gm_sbt$mer    0.0 20
## gm_seb$mer    5.7 12
## gm_sb$mer    25.8  9
## gm_bt$mer    74.8 16
## gm_full$mer   80.6 19
## gm_eb$mer    81.5  8

```

```
## gm_b$mer      111.7 5
## gm_st$mer     4906.6 18
## gm_set$mer    4911.3 21
## gm_se$mer     5109.8 10
## gm_s0$mer     5397.9 7
## gm_e$mer     10217.0 6
## gm_et$mer    10217.3 17
## gm_t$mer     10219.8 14
## gm_0$mer     10245.1 3
```

Now do Wood's test on the full model to see what's non-zero. Also check the best models's r^2 and population variance.

```
summary(gm_full$gam)

##
## Family: gaussian
## Link function: identity
##
## Formula:
## l24 ~ s(10.s, k = kl) + s(density.s, supply.s, k = ke) + trt +
##       s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.529256   0.001393 379.995  <2e-16 ***
## trtT2        0.002335   0.004437   0.526   0.5988
## trtT3       -0.001518   0.001855  -0.819   0.4131
## trtT4        0.003652   0.001977   1.847   0.0649 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df      F  p-value
## s(10.s)        3.880  3.880 69545.091 < 2e-16 ***
## s(density.s,supply.s) 5.773  5.773   4.372 0.000292 ***
## s(day):trtT1    1.000  1.000   0.052 0.820098
## s(day):trtT2    2.864  2.864   0.849 0.458313
## s(day):trtT3    1.000  1.000   0.068 0.793572
## s(day):trtT4    1.000  1.000   6.644 0.010014 *
```



```

## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.993
## lmer.REML = -11311  Scale est. = 0.00025413  n = 2087

summary(gm_sbt$gam)[["r.sq"]]

## [1] 0.9928647

summary(gm_sbt$mer)

## Linear mixed model fit by maximum likelihood ['lmerMod']
##
##      AIC      BIC  logLik deviance df.resid
## -11353.6 -11240.7  5696.8 -11393.6     2067
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.6166 -0.5955 -0.0504  0.5619  3.6141
##
## Random effects:
##  Groups   Name                Variance Std.Dev.
##  pop.ID   (Intercept)  1.477e-06 1.215e-03
##  Xr.3     s(day):trtT4  1.187e-01 3.445e-01
##  Xr.2     s(day):trtT3  0.000e+00 0.000e+00
##  Xr.1     s(day):trtT2  9.628e-16 3.103e-08
##  Xr.0     s(day):trtT1  0.000e+00 0.000e+00
##  Xr       s(10.s)       1.384e+00 1.176e+00
##  Residual                    2.440e-04 1.562e-02
## Number of obs: 2087, groups:
##  pop.ID, 8; Xr.3, 8; Xr.2, 8; Xr.1, 8; Xr.0, 8; Xr, 3
##
## Fixed effects:
##              Estimate Std. Error t value
## X(Intercept)    0.5386004  0.0036981  145.64
## Xstageproto     0.0032827  0.0025467   1.29
## Xstagetrityto  -0.0023875  0.0041290  -0.58
## Xstagefemale   -0.0185940  0.0048274  -3.85
## Xstagefemale   -0.0148305  0.0050769  -2.92

```

```

## XtrtT2          0.0034161  0.0018149   1.88
## XtrtT3          -0.0028136  0.0015385  -1.83
## XtrtT4          0.0039005  0.0015433   2.53
## Xs(10.s)Fx1    0.1904810  0.0030569  62.31
## Xs(day):trtT1Fx1 -0.0001667  0.0007093  -0.24
## Xs(day):trtT2Fx1  0.0002169  0.0013468   0.16
## Xs(day):trtT3Fx1 -0.0016907  0.0006548  -2.58
## Xs(day):trtT4Fx1 -0.0250289  0.0068840  -3.64
##
## Correlation of Fixed Effects:
##           X(Int) Xstgpr Xstgtr Xstgml Xstgfm XtrtT2 XtrtT3 XtrtT4 X(0.)F
## Xstageproto -0.755
## Xstageritro -0.915  0.834
## Xstagemale  -0.942  0.709  0.922
## Xstagefemal -0.933  0.685  0.886  0.967
## XtrtT2      -0.060 -0.049 -0.083 -0.120 -0.153
## XtrtT3      -0.167 -0.019 -0.032 -0.043 -0.050  0.438
## XtrtT4      -0.236  0.006  0.014  0.027  0.043  0.406  0.491
## Xs(10.s)Fx1  0.667 -0.672 -0.756 -0.642 -0.617  0.033  0.020 -0.066
## Xs(dy):T1F1 -0.029 -0.005  0.011  0.028  0.043 -0.013  0.000  0.015 -0.006
## Xs(dy):T2F1  0.098 -0.047 -0.075 -0.099 -0.129  0.432  0.012 -0.012  0.034
## Xs(dy):T3F1  0.156 -0.107 -0.141 -0.159 -0.177  0.042 -0.025 -0.012  0.075
## Xs(dy):T4F1 -0.030  0.031  0.033  0.028  0.029 -0.002 -0.001 -0.013 -0.031
##           X():T1 X():T2 X():T3
## Xstageproto
## Xstageritro
## Xstagemale
## Xstagefemal
## XtrtT2
## XtrtT3
## XtrtT4
## Xs(10.s)Fx1
## Xs(dy):T1F1
## Xs(dy):T2F1 -0.015
## Xs(dy):T3F1 -0.015  0.034
## Xs(dy):T4F1  0.000 -0.002 -0.004

```

The estimate of the standard deviation among populations is small relative to other estimates.

4 Probability of Initiating Transition in 24 Hours

Subset data to only include the juvenile survivors.

```
tempT=transform(subset(Indiv, (!stage %in% c("male","female") & (s24==1))),
  t.init.24=as.numeric(t24 %in% c("next stage", "quiescent")),
  t.done.24=as.numeric(t24 == "next stage"),
  stage=stage[, drop=TRUE]
)
temp=na.omit(tempT[
  ,c("t.init.24","l0.s","trt","day","density.s", "supply.s", "pop.ID", "stage")
```

```
t24m_full=gamm4(t.init.24~s(10.s, k=kl)+s(density.s, supply.s, k=ke)+
  trt+s(day, by=trt, k=kd), family=binomial, random=~(1|pop.ID), data=temp)
t24m_eb=gamm4(t.init.24~s(10.s, k=kl)+s(density.s, supply.s, k=ke),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_et=gamm4(t.init.24~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_bt=gamm4(t.init.24~s(10.s, k=kl)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_b=gamm4(t.init.24~s(10.s, k=kl),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_e=gamm4(t.init.24~s(density.s, supply.s, k=ke),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_t=gamm4(t.init.24~trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_0=gamm4(t.init.24~1,
  family=binomial, random=~(1|pop.ID), data=temp)

t24m_sfull=gamm4(t.init.24~stage+ s(10.s, k=kl)+s(density.s, supply.s, k=ke)+
  trt+s(day, by=trt, k=kd), family=binomial, random=~(1|pop.ID), data=temp)
t24m_seb=gamm4(t.init.24~stage+ s(10.s, k=kl)+s(density.s, supply.s, k=ke),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_set=gamm4(t.init.24~stage+ s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_sbt=gamm4(t.init.24~stage+ s(10.s, k=kl)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
t24m_sb=gamm4(t.init.24~stage+ s(10.s, k=kl),
  family=binomial, random=~(1|pop.ID), data=temp)
```

```

t24m_se=gamm4(t.init.24~stage+ s(density.s, supply.s, k=ke),
              family=binomial, random=~(1|pop.ID), data=temp)
t24m_st=gamm4(t.init.24~stage+ trt+s(day, by=trt, k=kd),
              family=binomial, random=~(1|pop.ID), data=temp)

## Warning in (function (fn, par, lower = rep.int(-Inf, n), upper
= rep.int(Inf, : failure to converge in 10000 evaluations

t24m_s0=gamm4(t.init.24~stage,
              family=binomial, random=~(1|pop.ID), data=temp)

```

t24m.st did not converge, indicating that it might be overparameterized, so we omit it from AIC comparison.

```

AICctab(t24m_full$mer, t24m_eb$mer, t24m_et$mer, t24m_bt$mer,
        t24m_b$mer, t24m_e$mer, t24m_t$mer, t24m_0$mer,
        t24m_sfull$mer, t24m_seb$mer, t24m_set$mer, t24m_sbt$mer,
        t24m_sb$mer, t24m_se$mer, t24m_s0$mer, nobs=nrow(temp))

##           dAICc df
## t24m_seb$mer    0.0  9
## t24m_sb$mer     8.2  6
## t24m_sbt$mer   13.6 17
## t24m_sfull$mer 14.7 20
## t24m_eb$mer   165.1  7
## t24m_full$mer 177.9 18
## t24m_bt$mer   187.0 15
## t24m_b$mer    217.0  4
## t24m_se$mer   242.9  7
## t24m_set$mer  258.7 18
## t24m_e$mer    273.3  5
## t24m_et$mer   289.0 16
## t24m_t$mer    295.6 13
## t24m_s0$mer   328.9  4
## t24m_0$mer    353.7  2

```

Now do Wood's test on the full model to see what's non-zero. Also check the best models's r^2 and population variance.

```

summary(t24m_sfull$gam)

##
## Family: binomial
## Link function: logit
##
## Formula:
## t.init.24 ~ stage + s(l0.s, k = kl) + s(density.s, supply.s,
##      k = ke) + trt + s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   4.6581     0.6320   7.370 1.70e-13 ***
## stageproto   -5.5702     0.6000  -9.284 < 2e-16 ***
## stagetrito  -13.9131     1.3623 -10.213 < 2e-16 ***
## trtT2         -1.1407     0.9877  -1.155   0.248
## trtT3          0.2351     0.3795   0.619   0.536
## trtT4         -1.8931     0.3789  -4.996 5.85e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df  Chi.sq p-value
## s(l0.s)                1      1 126.004 < 2e-16 ***
## s(density.s,supply.s)  2      2   4.748 0.09309 .
## s(day):trtT1           1      1   0.096 0.75653
## s(day):trtT2           1      1   0.175 0.67605
## s(day):trtT3           1      1   2.512 0.11300
## s(day):trtT4           1      1  10.591 0.00114 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.415
## glmer.ML = 528.15  Scale est. = 1          n = 995

summary(t24m_seb$gam)[["r.sq"]]

## [1] 0.451062

summary(t24m_seb$mer)

```

```

## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: binomial ( logit )
##
##      AIC      BIC   logLik deviance df.resid
##    554.1    598.2  -268.0   536.1     986
##
## Scaled residuals:
##   Min      1Q  Median      3Q      Max
## -2.653 -0.271 -0.087 -0.016  36.953
##
## Random effects:
##   Groups Name              Variance Std.Dev.
##   pop.ID (Intercept)        6.536e-03 0.08084
##   Xr.0  s(density.s,supply.s) 1.212e+03 34.81968
##   Xr    s(10.s)                9.167e+03 95.74623
## Number of obs: 995, groups:  pop.ID, 8; Xr.0, 7; Xr, 3
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## X(Intercept)      6.3526     0.9639   6.590 4.39e-11 ***
## Xstageproto      -8.7425     1.1986  -7.294 3.00e-13 ***
## Xstagetrity      -19.1178     2.2138  -8.636 < 2e-16 ***
## Xs(10.s)Fx1       5.1136     1.1491   4.450 8.58e-06 ***
## Xs(density.s,supply.s)Fx1 -0.5362     0.3048  -1.759 0.0785 .
## Xs(density.s,supply.s)Fx2  0.1934     0.2545   0.760 0.4474
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              X(Int) Xstgpr Xstgtr X(0.)F X(.,.)F1
## Xstageproto -0.882
## Xstagetrity -0.961  0.798
## Xs(10.s)Fx1  0.084 -0.412 -0.011
## Xs(dn.,.)F1  0.124 -0.072 -0.064  0.129
## Xs(dn.,.)F2  0.046 -0.029 -0.067  0.019  0.073

```

The estimate of the standard deviation among populations is small relative to other estimates.

5 Probability of Reproducing

Remove ones that were dead at 24hrs. Use females only.

```
Indivtmp=subset(Indiv, stage=="female" & !is.na(s24))
```

```
temp=na.omit(subset(Indivtmp, !is.na(10.s*as.numeric(trt)*f24) & (s24==1))[,c("f24", "10.s", "trt", "day", "density.s", "supply.s", "pop.ID")])
```

```
fbm_full=gamm4(as.numeric(f24>0)~s(10.s, k=kl)+s(density.s, supply.s, k=ke)+
  trt+s(day, by=trt, k=kd), family=binomial, random=~(1|pop.ID), data=temp)
fbm_eb=gamm4(as.numeric(f24>0)~s(10.s, k=kl)+s(density.s, supply.s, k=ke),
  family=binomial, random=~(1|pop.ID), data=temp)
fbm_et=gamm4(as.numeric(f24>0)~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
```

```
## Warning in (function (fn, par, lower = rep.int(-Inf, n), upper
= rep.int(Inf, : failure to converge in 10000 evaluations
```

```
fbm_bt=gamm4(as.numeric(f24>0)~s(10.s, k=kl)+trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
fbm_b=gamm4(as.numeric(f24>0)~s(10.s, k=kl),
  family=binomial, random=~(1|pop.ID), data=temp)
fbm_e=gamm4(as.numeric(f24>0)~s(density.s, supply.s, k=ke),
  family=binomial, random=~(1|pop.ID), data=temp)
fbm_t=gamm4(as.numeric(f24>0)~trt+s(day, by=trt, k=kd),
  family=binomial, random=~(1|pop.ID), data=temp)
fbm_0=gamm4(as.numeric(f24>0)~1,
  family=binomial, random=~(1|pop.ID), data=temp)
```

fbm_et did not converge, indicating that it might be overparameterized, so we omit it from AIC comparison.

```
AICctab(fbm_full$mer, fbm_eb$mer, fbm_bt$mer,
  fbm_b$mer, fbm_e$mer, fbm_t$mer, fbm_0$mer, nobs=nrow(temp))
```

```
##           dAICc df
## fbm_eb$mer    0.0 7
```

```

## fbm_full$mer    2.2 18
## fbm_bt$mer     18.0 15
## fbm_e$mer      20.5 5
## fbm_t$mer      28.8 13
## fbm_b$mer      82.0 4
## fbm_0$mer     169.2 2

summary(fbm_full$gam)

##
## Family: binomial
## Link function: logit
##
## Formula:
## as.numeric(f24 > 0) ~ s(l0.s, k = kl) + s(density.s, supply.s,
##      k = ke) + trt + s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)   2.1597     0.3729   5.792 6.96e-09 ***
## trtT2         -1.6526     0.7576  -2.181  0.0292 *
## trtT3          0.5145     0.5304   0.970  0.3320
## trtT4         -2.3255     0.4549  -5.112 3.19e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df Chi.sq  p-value
## s(l0.s)          1.000  1.000 23.763 1.09e-06 ***
## s(density.s,supply.s) 2.000  2.000 17.951 0.000126 ***
## s(day):trtT1      1.000  1.000  0.074 0.784928
## s(day):trtT2      1.000  1.000  0.149 0.699583
## s(day):trtT3      2.396  2.396 10.871 0.007447 **
## s(day):trtT4      1.000  1.000  0.001 0.976011
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.478
## glmer.ML = 395.08  Scale est. = 1          n = 568

```



```

summary(fbm_eb$gam)[["r.sq"]]

## [1] 0.3839028

summary(fbm_eb$mer)

## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: binomial ( logit )
##
##      AIC      BIC   logLik deviance df.resid
##  442.0   472.4  -214.0   428.0     561
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -6.7025 -0.2539  0.2499  0.3691  4.8874
##
## Random effects:
## Groups Name              Variance Std.Dev.
## pop.ID (Intercept)         0.877   0.9365
## Xr.0  s(density.s,supply.s) 277.581 16.6608
## Xr    s(10.s)                0.000   0.0000
## Number of obs: 568, groups: pop.ID, 8; Xr.0, 7; Xr, 3
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## X(Intercept)      1.4338    0.3613   3.968 7.24e-05 ***
## Xs(10.s)Fx1        0.8367    0.1760   4.755 1.99e-06 ***
## Xs(density.s,supply.s)Fx1 -0.9570    0.3438  -2.783 0.00538 **
## Xs(density.s,supply.s)Fx2  1.0313    0.3860   2.672 0.00754 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              X(Int) X(0.)F X(.,.)F1
## Xs(10.s)Fx1  0.107
## Xs(dn.,.)F1 -0.044 -0.104
## Xs(dn.,.)F2  0.024 -0.213  0.231

```

The estimate of the standard deviation among populations is in the same

range as other estimates. This could indicate that populations behaved differently. Figure S1 shows that the two populations within a treatment do not track each other perfectly, but fairly well.

6 Egg Counts, Given Reproduction

```
temp=na.omit(subset(Indivtmp, (f24>0) & (s24==1))[,
  c("f24", "l0.s", "trt", "day", "density.s", "supply.s", "pop.ID")])
```

```
fcm_full=gamm4(f24-1~s(l0.s,k=kl)+s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=k
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_eb=gamm4(f24-1~s(l0.s, k=kl)+s(density.s, supply.s, k=ke),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_et=gamm4(f24-1~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_bt=gamm4(f24-1~s(l0.s, k=kl)+trt+s(day, by=trt, k=kd),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_b=gamm4(f24-1~s(l0.s, k=kl),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_e=gamm4(f24-1~s(density.s, supply.s, k=ke),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_t=gamm4(f24-1~trt+s(day, by=trt, k=kd),
  family=poisson, random=~(1|pop.ID), data=temp)
fcm_0=gamm4(f24-1~1,
  family=poisson, random=~(1|pop.ID), data=temp)

AICctab(fcm_full$mer, fcm_eb$mer, fcm_et$mer, fcm_bt$mer,
  fcm_b$mer, fcm_e$mer, fcm_t$mer, fcm_0$mer, nobs=nrow(temp))
```

```
##          dAICc df
## fcm_eb$mer      0.0 7
## fcm_full$mer    8.5 18
## fcm_bt$mer     13.9 15
## fcm_t$mer     129.7 13
## fcm_et$mer     133.6 16
## fcm_e$mer     188.2 5
## fcm_b$mer     410.4 4
## fcm_0$mer     909.3 2
```

```

summary(fcm_full$gam)

##
## Family: poisson
## Link function: log
##
## Formula:
## f24 - 1 ~ s(l0.s, k = kl) + s(density.s, supply.s, k = ke) +
##      trt + s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  1.0088     0.1085   9.300 < 2e-16 ***
## trtT2        -0.3792     0.4363  -0.869  0.38470
## trtT3         0.5504     0.1778   3.096  0.00196 **
## trtT4        -0.3533     0.1942  -1.819  0.06889 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df  Chi.sq p-value
## s(l0.s)          3.182  3.182 149.766 < 2e-16 ***
## s(density.s,supply.s) 7.327  7.327 349.382 < 2e-16 ***
## s(day):trtT1      1.000  1.000   0.013 0.91027
## s(day):trtT2      1.000  1.000   0.361 0.54810
## s(day):trtT3      1.000  1.000   6.673 0.00979 **
## s(day):trtT4      1.000  1.000   1.210 0.27129
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) =  0.763
## glmer.ML = 677.08  Scale est. = 1          n = 409

summary(fcm_eb$gam)[["r.sq"]]

## [1] 0.7393331

summary(fcm_eb$mer)

## Generalized linear mixed model fit by maximum likelihood (Laplace

```

```

## Approximation) [glmerMod]
## Family: poisson ( log )
##
##      AIC      BIC   logLik deviance df.resid
## 1810.9   1839.0  -898.5  1796.9    402
##
## Scaled residuals:
##      Min      1Q  Median      3Q      Max
## -2.3432 -0.9055 -0.1298  0.6910  7.8782
##
## Random effects:
## Groups Name              Variance Std.Dev.
## pop.ID (Intercept)      0.02758  0.1661
## Xr.0  s(density.s,supply.s) 122.76138 11.0798
## Xr    s(10.s)              0.00000  0.0000
## Number of obs: 409, groups:  pop.ID, 8; Xr.0, 7; Xr, 3
##
## Fixed effects:
##              Estimate Std. Error z value Pr(>|z|)
## X(Intercept)      1.05810   0.07075  14.955 <2e-16 ***
## Xs(10.s)Fx1       0.39709   0.02835  14.007 <2e-16 ***
## Xs(density.s,supply.s)Fx1 0.73155   0.08451   8.657 <2e-16 ***
## Xs(density.s,supply.s)Fx2 -0.19052   0.08249  -2.310  0.0209 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##              X(Int) X(0.)F X(.,.)F1
## Xs(10.s)Fx1 -0.109
## Xs(dn.,.)F1 -0.167 -0.043
## Xs(dn.,.)F2  0.145  0.083 -0.569

```

The estimate of the standard deviation among populations is smaller than the other estimates.

7 Egg Size

Take the average within female as the response.

```

load("../Egg size from sampling MB.Rdata")
eggdat$day=as.numeric(eggdat$date.samp-min(eggdat$date.samp)+1)
temp=na.omit(ddply(eggdat, .(fem.ID), summarize,
  esize=mean(esize, na.omit=TRUE),
  l0.s=l0.s[1],
  pop.ID=pop.ID[1],
  density.s= density.s[1],
  supply.s= supply.s[1],
  day=day[1],
  trt=trt[1]))

esm_full=gamm4(esize~s(l0.s, k=kl)+s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_eb=gamm4(esize~s(l0.s, k=kl)+s(density.s, supply.s, k=ke),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_et=gamm4(esize~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_bt=gamm4(esize~s(l0.s, k=kl)+trt+s(day, by=trt, k=kd),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_b=gamm4(esize~s(l0.s, k=kl),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_e=gamm4(esize~s(density.s, supply.s, k=ke),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_t=gamm4(esize~trt+s(day, by=trt, k=kd),
  REML=FALSE, random=~(1|pop.ID), data=temp)
esm_0=gamm4(esize~1,
  REML=FALSE, random=~(1|pop.ID), data=temp)

AICctab(esm_full$mer, esm_eb$mer, esm_et$mer, esm_bt$mer,
  esm_b$mer, esm_e$mer, esm_t$mer, esm_0$mer, nobs=nrow(temp))

##           dAICc df
## esm_e$mer    0.0  6
## esm_eb$mer   1.6  8
## esm_t$mer    8.0 14
## esm_bt$mer  12.2 16
## esm_et$mer  12.5 17
## esm_b$mer   13.6  5
## esm_0$mer   14.9  3

```

```
## esm_full$mer 16.9 19

summary(esm_full$gam)

##
## Family: gaussian
## Link function: identity
##
## Formula:
## esize ~ s(l0.s, k = kl) + s(density.s, supply.s, k = ke) + trt +
##       s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.1772784  0.0006767 261.993  <2e-16 ***
## trtT2        -0.0048530  0.0027293  -1.778  0.0762 .
## trtT3        -0.0015562  0.0011223  -1.387  0.1663
## trtT4        -0.0008934  0.0011663  -0.766  0.4441
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df    F p-value
## s(l0.s)          1.000  1.000 0.002 0.96070
## s(density.s,supply.s) 2.000  2.000 0.925 0.39723
## s(day):trtT1     1.000  1.000 0.032 0.85724
## s(day):trtT2     1.000  1.000 2.730 0.09925 .
## s(day):trtT3     2.331  2.331 5.867 0.00205 **
## s(day):trtT4     1.000  1.000 7.355 0.00697 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## R-sq.(adj) = 0.0647
## lmer.REML = -2827.5  Scale est. = 5.8515e-05  n = 410

summary(esm_e$gam)[["r.sq"]]

## [1] 0.06183376

summary(esm_e$mer)
```

```
## Linear mixed model fit by maximum likelihood ['lmerMod']
##
##      AIC      BIC   logLik deviance df.resid
## -2804.7 -2780.6  1408.3  -2816.7     404
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -3.4629 -0.6682  0.0804  0.5928  2.7496
##
## Random effects:
##      Groups      Name                Variance Std.Dev.
## pop.ID      (Intercept)              0.000e+00 0.000000
## Xr          s(density.s,supply.s)  7.827e-04 0.027977
## Residual                                5.987e-05 0.007737
## Number of obs: 410, groups:  pop.ID, 8; Xr, 7
##
## Fixed effects:
##
##              Estimate Std. Error t value
## X(Intercept)      0.1764325  0.0003821  461.7
## Xs(density.s,supply.s)F1 -0.0013077  0.0005515   -2.4
## Xs(density.s,supply.s)F2  0.0013731  0.0006216    2.2
##
## Correlation of Fixed Effects:
##              X(Int) X(.,.)F1
## Xs(dn.,.)F1  0.000
## Xs(dn.,.)F2  0.000 -0.141
```

The estimate of the standard deviation among populations is 0.

8 Survival

```
temp=na.omit(Indiv[
  ,c("s24","l0.s","trt","day","density.s", "supply.s", "pop.ID", "stage")])
```

```
sm_full=gamm4(s24~s(l0.s, k=kl)+s(density.s, supply.s, k=ke)+
  trt+s(day, by=trt, k=kd), family=binomial, random=~(1|pop.ID), data=temp)
sm_eb=gamm4(s24~s(l0.s, k=kl)+s(density.s, supply.s, k=ke),
```

```

    family=binomial, random=~(1|pop.ID), data=temp)
sm_et=gamm4(s24~s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_bt=gamm4(s24~s(l0.s, k=kl)+trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_b=gamm4(s24~s(l0.s, k=kl),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_e=gamm4(s24~s(density.s, supply.s, k=ke),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_t=gamm4(s24~trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_0=gamm4(s24~1,
    family=binomial, random=~(1|pop.ID), data=temp)
sm_sfull=gamm4(s24~stage+ s(l0.s, k=kl)+s(density.s, supply.s, k=ke)+
    trt+s(day, by=trt, k=kd), family=binomial, random=~(1|pop.ID), data=temp)
sm_seb=gamm4(s24~stage+ s(l0.s, k=kl)+s(density.s, supply.s, k=ke),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_set=gamm4(s24~stage+ s(density.s, supply.s, k=ke)+trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_sbt=gamm4(s24~stage+ s(l0.s, k=kl)+trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_sb=gamm4(s24~stage+ s(l0.s, k=kl),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_se=gamm4(s24~stage+ s(density.s, supply.s, k=ke),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_st=gamm4(s24~stage+ trt+s(day, by=trt, k=kd),
    family=binomial, random=~(1|pop.ID), data=temp)
sm_s0=gamm4(s24~stage,
    family=binomial, random=~(1|pop.ID), data=temp)

AICctab(sm_full$mer, sm_eb$mer, sm_et$mer, sm_bt$mer,
    sm_b$mer, sm_e$mer, sm_t$mer, sm_0$mer,
    sm_sfull$mer, sm_seb$mer, sm_set$mer,
    sm_sbt$mer, sm_sb$mer, sm_se$mer, sm_st$mer, sm_s0$mer, nobs=nrow(temp))

##          dAICc df
## sm_se$mer    0.0  9
## sm_seb$mer   3.8 11
## sm_s0$mer    8.4  6

```



```

## sm_e$mer      9.7  5
## sm_eb$mer     11.3  7
## sm_sb$mer     11.5  8
## sm_0$mer      16.9  2
## sm_st$mer     17.5 17
## sm_set$mer    19.6 20
## sm_b$mer      19.9  4
## sm_sbt$mer    21.5 19
## sm_sfull$mer  23.4 22
## sm_t$mer      27.0 13
## sm_et$mer     29.0 16
## sm_bt$mer     29.3 15
## sm_full$mer   30.7 18

summary(sm_sfull$gam)

##
## Family: binomial
## Link function: logit
##
## Formula:
## s24 ~ stage + s(l0.s, k = kl) + s(density.s, supply.s, k = ke) +
##      trt + s(day, by = trt, k = kd)
##
## Parametric coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  2.66176    0.61444   4.332 1.48e-05 ***
## stageproto   0.18406    0.38759   0.475  0.635
## stagetrto    1.51998    0.62466   2.433  0.015 *
## stagemale    0.15889    0.71979   0.221  0.825
## stagefemale  0.31692    0.90874   0.349  0.727
## trtT2        0.28508    0.53668   0.531  0.595
## trtT3       -0.04604    0.33689  -0.137  0.891
## trtT4        0.06820    0.31447   0.217  0.828
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##              edf Ref.df Chi.sq p-value
## s(l0.s)      1      1  0.290  0.590

```

```

## s(density.s,supply.s)  2      2  4.056  0.132
## s(day):trtT1           1      1  1.141  0.285
## s(day):trtT2           1      1  1.041  0.308
## s(day):trtT3           1      1  0.264  0.608
## s(day):trtT4           1      1  0.000  0.988
##
## R-sq.(adj) =  0.0109
## glmer.ML = 877.19  Scale est. = 1          n = 2241

summary(sm_se$gam)[["r.sq"]]

## [1] 0.01224263

summary(sm_se$mer)

## Generalized linear mixed model fit by maximum likelihood (Laplace
## Approximation) [glmerMod]
## Family: binomial ( logit )
##
##      AIC      BIC   logLik deviance df.resid
##  898.9   950.3  -440.4   880.9   2232
##
## Scaled residuals:
##      Min       1Q   Median       3Q      Max
## -10.2986   0.1811   0.2215   0.2556   0.3666
##
## Random effects:
##  Groups Name              Variance Std.Dev.
##  pop.ID (Intercept)       5.856e-03 0.0765270
##  Xr      s(density.s,supply.s) 1.032e-08 0.0001016
## Number of obs: 2241, groups:  pop.ID, 8; Xr, 7
##
## Fixed effects:
##
##              Estimate Std. Error z value Pr(>|z|)
## X(Intercept)      2.92729    0.24776  11.815 < 2e-16 ***
## Xstageproto        0.10703    0.35420   0.302  0.762510
## Xstagetrito        1.29929    0.47865   2.714  0.006638 **
## Xstagemale        -0.20930    0.29623  -0.707  0.479848
## Xstagefemale       -0.14494    0.29592  -0.490  0.624274
## Xs(density.s,supply.s)Fx1  0.11831    0.12259   0.965  0.334499

```

```

## Xs(density.s,supply.s)Fx2 -0.33729    0.09433   -3.576 0.000349 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Correlation of Fixed Effects:
##           X(Int) Xstgpr Xstgtr Xstgml Xstgfm X(.,.)F1
## Xstageproto -0.675
## Xstagetrito -0.497  0.349
## Xstagemale  -0.813  0.565  0.417
## Xstagefemal -0.809  0.565  0.418  0.676
## Xs(dn.,.)F1  0.034  0.010  0.001  0.000  0.009
## Xs(dn.,.)F2 -0.119  0.006 -0.020  0.042  0.005  0.156

summary(sm_sfull$gam)

##
## Family: binomial
## Link function: logit
##
## Formula:
## s24 ~ stage + s(l0.s, k = kl) + s(density.s, supply.s, k = ke) +
##      trt + s(day, by = trt, k = kd)
##
## Parametric coefficients:
##           Estimate Std. Error z value Pr(>|z|)
## (Intercept)  2.66176    0.61444   4.332 1.48e-05 ***
## stageproto   0.18406    0.38759   0.475  0.635
## stagetrito   1.51998    0.62466   2.433  0.015 *
## stagemale    0.15889    0.71979   0.221  0.825
## stagefemale  0.31692    0.90874   0.349  0.727
## trtT2        0.28508    0.53668   0.531  0.595
## trtT3       -0.04604    0.33689  -0.137  0.891
## trtT4        0.06820    0.31447   0.217  0.828
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Approximate significance of smooth terms:
##           edf Ref.df Chi.sq p-value
## s(l0.s)           1      1  0.290  0.590
## s(density.s,supply.s) 2      2  4.056  0.132

```

```
## s(day):trtT1      1      1  1.141  0.285
## s(day):trtT2      1      1  1.041  0.308
## s(day):trtT3      1      1  0.264  0.608
## s(day):trtT4      1      1  0.000  0.988
##
## R-sq.(adj) =  0.0109
## glmer.ML = 877.19  Scale est. = 1          n = 2241
```

The estimate of the standard deviation among populations is smaller than the other estimates.

9 Version Info

```
sessionInfo()

## R version 3.2.1 (2015-06-18)
## Platform: x86_64-apple-darwin13.4.0 (64-bit)
## Running under: OS X 10.10.4 (Yosemite)
##
## locale:
## [1] en_US.UTF-8/en_US.UTF-8/en_US.UTF-8/C/en_US.UTF-8/en_US.UTF-8
##
## attached base packages:
## [1] stats4      stats      graphics  grDevices  utils      datasets  methods
## [8] base
##
## other attached packages:
## [1] gamm4_0.2-3  lme4_1.1-9  Matrix_1.2-1  plyr_1.8.3  bbmle_1.0.17
## [6] mgcv_1.8-6   nlme_3.1-120  ggplot2_1.0.1  reshape_0.8.5  knitr_1.11
##
## loaded via a namespace (and not attached):
## [1] Rcpp_0.12.0      magrittr_1.5      splines_3.2.1
## [4] MASS_7.3-40     munsell_0.4.2     colorspace_1.2-6
## [7] lattice_0.20-31 minqa_1.2.4       stringr_1.0.0
## [10] highr_0.5        tools_3.2.1       grid_3.2.1
## [13] gtable_0.1.2     digest_0.6.8      numDeriv_2014.2-1
## [16] nloptr_1.0.4     reshape2_1.4.1    formatR_1.2
## [19] codetools_0.2-11 evaluate_0.7.2     labeling_0.3
## [22] stringi_0.5-5    scales_0.2.5      proto_0.3-10
```