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1 Independent evaluation of melanoma 2 polygenic risk scores in UK and Australian 3 prospective cohorts

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27

28 **Running title:** Melanoma polygenic risk scores in UK and Australian prospective cohorts
29
30

31 **What's already known about this topic?**

- 32 • Melanoma risk has a substantial inherited (germline) genetic component.
- 33 • Polygenic risk scores (PRS) have the potential to improve risk stratification for
34 melanoma screening, but PRS-based risk prediction requires independent validation in
35 prospective cohorts, especially assessing calibration.
- 36 • A recent, large melanoma genome-wide association study has increased the number of
37 genetic variants independently associated with melanoma risk from 45 (in 21 genetic
38 regions) to 68 (in 54 genetic regions).

39 40 **What does this study add?**

- 41 • PRS enhanced melanoma risk prediction, with better performance of a PRS based on
42 genetic associations from a larger, more diverse meta-analysis.

- 43 • Population-average risks at age 50 years might be reached 10-20 years earlier by
44 individuals with the 20% highest PRS and 10-20 years later by those with 20% lowest
45 PRS.
46 • Re-calibration of 10-year absolute risks for application to specific populations may be
47 necessary, and will facilitate translation to clinical practice.
48

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64

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71 **Summary**

72 **Background:** Previous studies suggest polygenic risk scores (PRS) may improve melanoma
73 risk stratification. However, there has been limited independent validation of PRS-based risk
74 prediction, particularly assessment of calibration (comparing predicted to observed risks).
75

76 **Objectives:** To evaluate PRS-based melanoma risk prediction in prospective UK and
77 Australian cohorts with European ancestry.
78

79 **Methods:** We analysed invasive melanoma incidence in UK Biobank (UKB; n=395,647; 1,651
80 cases) and a case-cohort nested within the Melbourne Collaborative Cohort Study (MCCS,
81 Australia; n=4,765; 303 cases). Three PRS were evaluated: 68 SNPs at 54 loci from a 2020
82 meta-analysis (PRS68); 50 SNPs significant in the 2020 meta-analysis excluding UKB (PRS50);
83 45 SNPs at 21 loci known in 2018 (PRS45). 10-year melanoma risks were calculated from
84 population-level cancer registry data by age group and sex, with and without PRS
85 adjustment.
86

87 **Results:** Predicted absolute melanoma risks based on age and sex alone underestimated
88 melanoma incidence in UKB (ratio expected/observed cases E/O=0.65, 95% confidence
89 interval (95%CI) 0.62-0.68) and MCCS (E/O=0.63, 95%CI 0.56-0.72). For UKB, calibration was
90 improved by PRS-adjustment, e.g. PRS50-adjusted risks E/O=0.91, 95%CI 0.87-0.95.
91 Discriminative ability for PRS68- and PRS50-adjusted absolute risks was higher than for risks
92 based on age and sex alone (Δ AUC 0.07-0.10, $p < 0.0001$), and higher than for PRS45-
93 adjusted risks (Δ AUC 0.02-0.04, $p < 0.001$).
94

95 **Conclusions:** A PRS derived from a larger, more diverse meta-analysis improves risk
96 prediction compared to an earlier PRS, and might help tailor melanoma prevention and
97 early detection strategies to different risk levels. Re-calibration of absolute risks may be
98 necessary for application to specific populations.
99

100

101

102 **Introduction**

103 Polygenic risk scores (PRS) aggregate the effects of many genetic variants across the
104 genome into a single score aiming to reflect an individual's genetic risk of disease. PRS are
105 increasingly showing potential for risk-stratified cancer prevention, early detection and
106 screening.¹⁻⁴ This includes cutaneous melanoma, a relatively lethal form of skin cancer with
107 high and growing global burden (estimated >324,000 new invasive melanoma diagnoses and
108 >57,000 deaths worldwide in 2020, rising to >424,000 new invasive diagnoses and >84,000
109 deaths in 2040)⁵. While ~75% of melanomas globally have been attributed to excess
110 ultraviolet radiation exposure,⁶ multiple other risk factors are known, including number of
111 nevi (e.g. 3.63-fold relative risk for people with 1+ atypical nevi)⁷ and genetic predisposition
112 .⁸⁻¹⁰ Previous case-control studies have shown that incorporating melanoma PRS in risk
113 prediction models improves discriminatory accuracy compared to models using only
114 traditional risk factors.¹²

115

116 The lack of proper evaluation of model calibration using prospective cohorts has hindered
117 the integration of PRS into risk assessment for clinical and public health practice.^{13,14}
118 Calibration measures the agreement between the predicted risks from the model and the
119 observed risks in a population. A poorly calibrated model could lead to under- or over-
120 screening of individuals in a risk-stratified screening program, reducing the program's
121 benefits and cost-effectiveness. To date, calibration of PRS-based predicted absolute
122 melanoma risks has not been assessed.^{12,15} This is important because absolute risks (e.g. 10-
123 year risk) are commonly used for risk communication in clinical settings and to determine
124 eligibility in national cancer screening programs.¹⁶

125

126 Recently, a larger and more population-diverse melanoma genome-wide association study
127 (GWAS) increased the number of known susceptibility variants to 68 independent
128 associations located in 54 genetic regions (loci),¹⁰ more than doubling the numbers of
129 known risk loci compared with earlier GWAS and PRS (with 45 SNPs at 21 loci included in the
130 PRS in a 2018 study).¹² It is, therefore, also of interest to compare the risk prediction
131 performance of the previously proposed PRS with an analogous PRS based on the new
132 GWAS.

133

134 To address these research questions, we evaluated different PRS-based predicted
135 melanoma risks in prospective UK and Australian cohorts, assessing their association with
136 melanoma, calibration, and discriminative ability.

137

138 **Materials and methods**

139 **Study samples**

140 UK Biobank¹⁷ (UKB) is a prospective cohort study of ~500,000 participants recruited in 2006-
141 2010 ("baseline"). Deaths and invasive melanoma diagnoses, the primary outcome, were
142 ascertained using linkage to death and cancer registry records (see Appendix S1). Records
143 were censored for completeness on 31/03/2016 and 31/10/2015 for participants recruited
144 in England/Wales and Scotland, respectively. All participants were genotyped, with further
145 genotype imputation described previously.¹⁷

146

147 The prospective Melbourne Collaborative Cohort Study¹⁸ (MCCS) recruited 41,513
148 Melbourne residents in 1990-1994. To broaden the range of measured lifestyle exposures,

149 particularly Mediterranean diet, there was an increased effort to recruit Greek and Italian
150 migrants. A subcohort was defined by randomly selecting 4,710 participants from those who
151 attended the second follow-up wave in 2003-2007 (“baseline” for the current analysis).
152 Records of deaths and invasive melanoma diagnoses were ascertained using linkage to
153 death and cancer registries (see Appendix S1), and were censored for completeness on
154 30/06/2016. Subcohort members and all MCCS participants who developed invasive
155 melanoma within 10 years of the second follow-up visit and before 30/06/2016 were
156 genotyped, followed by imputation of additional genetic variants (see Appendix S1).

157

158 In both studies, we applied extensive quality control to genetic data, excluded non-
159 European participants and those with melanoma diagnosis before baseline (see Appendix
160 S1). We ascertained participants’ self-reported sex, ethnicity, ease of tanning (noting that
161 the “ethnic group(s)” question in MCCS has also been interpreted as country of birth), and
162 age at baseline. The first 20 genetic principal components (PCs) were calculated to account
163 for potential population stratification (i.e. confounding of genetic associations by ancestry).
164 For UKB, we also obtained information on skin and hair colour, education, and the
165 postcode-based Townsend deprivation index (derived from national census data on
166 unemployment, car ownership, home ownership, and overcrowding, see Appendix S1).¹⁹

167

168 All participants provided informed consent at recruitment; UKB was approved by the North
169 West Centre for Research Ethics Committee (11/NW/0382) and the MCCS by the Cancer
170 Council Victoria Human Research Ethics Committee.

171

172 **PRS**

173 We calculated three PRS based on 1) 68 SNPs at 54 loci from a 2020 meta-analysis¹⁰ (PRS68);
174 2) 50 SNPs at 42 loci significant in the 2020 meta-analysis after exclusion of UKB participants
175 to avoid overfitting and ensure a truly independent validation dataset (PRS50); and 3) 45
176 SNPs at 21 loci as analysed in a 2018 study¹² (PRS45). Table S1 shows all variants in the three
177 PRS, their odds ratios and allele frequencies. After quality control, the UKB data included all
178 variants for the three PRS, while the MCCS data included 66, 49, and 43 variants for PRS68,
179 PRS50 and PRS45, respectively. We followed a previously described approach²⁰ to
180 determine genotype-specific relative risks at each variant, then used genotype dosages and
181 a log-additive model to obtain a PRS-specific relative risk for each participant and each PRS
182 (see Appendix S1).

183

184 **Statistical methods**

185

186 ***Population-average and PRS-adjusted predicted absolute melanoma risks***

187 DevCan v6.7.8 was used to calculate sex-and-age-specific absolute 10-year risks of
188 melanoma, with 5-year groupings for age, adjusting for death as competing risk.²¹ The
189 calculation was performed using population-wide cancer incidence, and cancer and all-
190 cause mortality data from each of Victoria (Australia), Scotland, and England/Wales (see
191 Appendix S1). For Victoria, we used data for 2009-2013; for Scotland and England/Wales,
192 we performed separate calculations using data for 2001-2005, 2006-2010, and 2011-2015.
193 UKB participants were assigned absolute risks for Scotland if recruited in
194 Glasgow/Edinburgh, and risks for England/Wales otherwise. We rescaled the 10-year
195 absolute risks for each participant to account for different individual follow-up time (see

196 Appendix S1). PRS-adjusted predicted absolute melanoma risks were obtained by
197 multiplying the corresponding sex-and-age-specific absolute risk by the participants' PRS-
198 specific relative risk (separately for each PRS).

199

200 **Association between PRS and participant characteristics**

201 Associations between each PRS (log-transformed) and participant characteristics were
202 calculated using multivariable generalised linear regression models (R function "glm") that
203 included age, sex, ethnicity, ease of tanning, and 20 PCs, and univariable models separately
204 for age, sex, ethnicity, and ease of tanning. For UKB, we also evaluated associations of PRS
205 with skin and hair colour, education, and the Townsend deprivation index (using univariable
206 and multivariable models). To account for multiple testing, a Bonferroni multiple-testing
207 correction for 28 tests and the three PRS was applied (significance at $p < 0.0006$).

208

209 **Association between PRS and melanoma incidence**

210 Associations between the PRS (log-transformed) and melanoma incidence were evaluated
211 using Fine-Gray competing risks regression with death as competing risk, adjusting for age,
212 sex, ethnicity, ease of tanning, and 20 PCs. The function "crr" (R package "cmprsk") was
213 used to calculate subhazard ratios (SHR) and 95% confidence intervals (95% CIs). To facilitate
214 comparisons with other studies, we calculated SHRs per 1 standard deviation (sd) of each
215 PRS (for MCCS, sd was based on subcohort participants). Significance was defined at
216 $p < 0.017$, applying a Bonferroni multiple-testing correction for the three PRS. We compared
217 the fully-adjusted results to unadjusted and partially-adjusted results, and also tested the
218 association for PRS risk quintiles and deciles (see Appendix S1). For the MCCS, the results
219 were verified using a weighted Cox regression model designed for case-cohort studies (see
220 Appendix S1).

221

222 **Calibration**

223 Calibration of unadjusted (age-and-sex-specific) and PRS-adjusted predicted absolute risks
224 was evaluated by comparing the expected ("E") with the observed ("O") numbers of
225 melanoma cases ("E/O" ratio) for all participants and separately for males and females,
226 overall and for each risk quintile (see Appendix S1). Robust 95% CIs were used for the MCCS
227 case-cohort design.²² Calibration of risk quintiles was further evaluated after rescaling of
228 risks to achieve $E/O = 1$ overall (separately for males and females). For UKB, calibration by
229 Townsend deprivation index quartiles was also assessed.

230

231 **Discrimination**

232 The area under the receiver operating characteristic curve (AUC) was calculated for the PRS-
233 based relative risk, as well as the unadjusted (age-and-sex-specific) and PRS-adjusted
234 predicted absolute risks (see Appendix S1). The AUC represents the probability that the
235 predicted risk is higher for a randomly chosen case than a non-case, with 0.5 corresponding
236 to random ranking and 1.0 to perfect discrimination. P-values for AUC comparisons were
237 obtained using the Delong test for correlated curves (function "roc.test", R package
238 "pROC"). We also calculated the classification true and false positive rates, true and false
239 negative rates, sensitivity, specificity, positive and negative likelihood ratios at 5%, 10%, ...,
240 95% risk thresholds.

241

242 **R² on the liability scale**

243 We converted the AUC values for the PRS relative risk, as well as the unadjusted and PRS-
244 adjusted absolute risks, to the explained [variation \(R²\)](#) on the liability scale²³ (see Appendix
245 S1).

246

247 ***Estimated 10-year absolute risks by PRS quintile and age***

248 To illustrate risk-stratification, the predicted absolute 10-year melanoma risk by PRS50
249 quintile was estimated based on the association results in UKB and MCCS (see Appendix S1).

250

251 **Results**

252 After quality checks, we analysed genetic and phenotypic data from 395,647 UKB
253 participants, and separately, 4,765 MCCS participants (Table 1). Of these UKB participants,
254 1,651 (0.4%) were diagnosed with invasive melanoma during the follow-up period (median
255 7.2 years, interquartile range 6.5-7.8 years). The MCCS case-cohort included 303
256 participants diagnosed with melanoma in the 10 years following the baseline for this study.
257 Figure S1 shows the PRS distributions in cases and non-cases.

258

259 **Associations of melanoma PRS with participant characteristics including traditional 260 melanoma risk factors**

261 In both cohorts, all three PRS were strongly and consistently associated with self-reported
262 ease of tanning (Table S2). We also found multiple associations between the PRS and
263 genetic PCs (Table S2). In UKB, skin and hair colour were also highly associated with all three
264 PRS (Table S3). Self-reported ethnicity was associated with the PRS in univariable analyses,
265 but not in multivariable analyses that adjusted for genetic PCs (Appendix S1, Table S3).

266

267 **Association of melanoma PRS with melanoma incidence**

268 All three PRS were associated with melanoma incidence (Table 2), with generally stronger
269 associations in UKB than the MCCS (e.g. non-overlapping 95% CIs for PRS68 and PRS45 fully-
270 adjusted SHR estimates). Moreover, in both cohorts, the PRS45 SHR estimates were smaller
271 than, and outside the 95% CIs for PRS50 and PRS68 estimates; fully-adjusted SHRs per 1 sd of
272 PRS45, PRS50, and PRS68 were 1.61, 1.73, and 1.80 in UKB, and 1.25, 1.47 and 1.47 in
273 MCCS, respectively.

274

275 In UKB, associations between the PRS and melanoma incidence were robust to inclusion of
276 different covariates, and only minimally (3-4%) attenuated when additionally adjusting for
277 skin and hair colour. In the MCCS, associations between the PRS and melanoma incidence
278 were slightly more attenuated (12-14%) by the inclusion of traditional melanoma risk factors
279 and genetic PCs as covariates, but remained highly statistically significant.

280

281 We also considered the association of PRS quintiles and deciles with melanoma incidence
282 (Figure S2, Table S4). The results were consistent with the analysis of the PRS as a
283 continuous score, with stronger associations observed in UKB than in the MCCS (Appendix
284 S1).

285

286

287 **Calibration of absolute melanoma risks with and without PRS**

288 We found that absolute 10-year melanoma risks calculated based on population-level
289 cancer registry data have risen sharply in England/Wales over 2001-2015, especially for

290 males and older age groups, with some increases also observed in Scotland (Figure S3).
291 Nonetheless, we observed an under-prediction of melanoma incidence in UKB when using
292 sex-and-age-specific risks based on 2001-2005, 2006-2010 and 2011-2015 population-level
293 data, with the ratio of expected to observed cases E/O of 0.41 [95% CI 0.39-0.43], 0.55
294 [0.53-0.58] and 0.65 [0.62-0.68], respectively (Table S5). The 2011-2015 period corresponds
295 to the last five years of follow-up of UKB participants, so the absolute 10-year risks based on
296 this period were used for further analysis.

297
298 We found that PRS68- and PRS50-adjusted absolute risks still under-predicted melanoma
299 incidence in UKB ($E/O=0.89$ [95%CI 0.85-0.94] and 0.91 [0.87-0.95], respectively; Figure 1a),
300 but less so than the unadjusted risks based on age and sex alone. For PRS45-adjusted
301 absolute risks the E/O was 0.73 [95%CI 0.70-0.77]. The under-prediction of risks was similar
302 for males and females (Figures S4, S5), and stronger for participants with higher
303 socioeconomic status as quantified by the Townsend deprivation index (e.g. unadjusted
304 absolute risks: least deprived quartile $E/O=0.57$ [95%CI 0.52-0.62], most deprived quartile
305 $E/O=0.85$ [95%CI 0.76-0.95]; Figure S6).

306
307 The calibration results were similar in the MCCS, where the sex- and age-specific absolute
308 melanoma risks based on population-wide cancer registry data for Victoria, Australia, also
309 under-estimated melanoma incidence ($E/O=0.63$ [95%CI 0.56-0.72]; Figure 1b). The PRS68
310 and PRS50-adjusted absolute risks resulted in similarly under-predicted incidence (both
311 $E/O=0.66$ [95%CI 0.58-0.75]) and the under-prediction was more pronounced for PRS45
312 ($E/O=0.54$ [95%CI 0.48-0.62]). Across unadjusted and PRS-adjusted risks, the under-
313 prediction was stronger for males than females (Figure S4, S5).

314
315 For PRS68 and PRS50, we did not find relative under- or over-prediction by absolute risk
316 quintile when re-calibrating the mean overall absolute risks to $E/O = 1$ (Figure 2, Appendix
317 S1).

318 **Discriminative ability of absolute melanoma risks with and without PRS**

319 Figure 3 shows the discriminative ability of the unadjusted (age-and-sex-specific) and PRS-
320 adjusted absolute risks, including sensitivity, specificity, and the AUC. In UKB, PRS-
321 adjusted absolute risks improved discrimination (AUC difference between PRS-
322 adjusted and unadjusted risks for each of the three PRS: ΔAUC 0.07-0.10, $p < 10^{-10}$). In the
323 MCCS, the incremental improvement from PRS-adjustment was smaller (ΔAUC 0.03-0.07,
324 $p < 0.0001$ for PRS68 and PRS50, $p = 0.07$ for PRS45); however, the AUC for the unadjusted
325 absolute risks was higher in the MCCS than UKB, resulting in similar AUCs for the PRS-
326 adjusted absolute risks (0.69 for PRS68 for both cohorts). In both UKB and the MCCS, the
327 PRS68- and PRS50-adjusted absolute risks had higher discriminative ability than the PRS45-
328 adjusted absolute risks ($p < 0.0001$).

329
330
331 At a threshold of the top predicted risk decile, the PRS68- and PRS50-adjusted absolute risks
332 had >90% specificity and >24% sensitivity for predicting melanoma incidence in both UKB
333 and the MCCS (Table S6, Appendix S1). At a threshold of the top predicted risk quartile, the
334 PRS68- and PRS50-adjusted absolute risks had >75% specificity and >46% sensitivity for
335 predicting melanoma incidence in both UKB and the MCCS.

336

337 **Explained variation on the liability scale**

338 PRS68 explained ~5% variation on the liability scale, with similar estimates for PRS50
339 (~4.5%) and a slight decrease for PRS45 (2.1-3.7%, see Table S7 and Appendix S1). PRS-
340 adjusted absolute risks explained more variation than the unadjusted risks (e.g. PRS50-
341 adjusted absolute risk: 6.2% (95%CI 5.3-7.2%) in UKB and 7.0% (4.8-9.6%) in the MCCS;
342 unadjusted risk: 1.4% (1.0-2.0%) and 3.0% (1.6-4.9%), respectively).

343

344 **Estimated 10-year absolute risks by PRS quintile and age**

345 Figure 4 shows the estimated 10-year absolute risks of melanoma by sex, age and
346 top/bottom PRS50 quintile. The population-average 10-year risk of invasive melanoma at
347 age 50 years was estimated to be reached 10-20 years earlier for those in the top PRS50
348 quintile, and 10-30 years later for those in the bottom PRS50 quintile.

349

350 **Sensitivity analyses**

351 The results presented above were robust to the factors assessed in multiple sensitivity
352 analyses (see Appendix S1).

353

354 **Discussion**

355 We evaluated performance of three melanoma PRS using data from large-scale UK and
356 Australian cohorts. We found that PRS-adjustment of age-and-sex-specific absolute risks
357 determined from population data improved discriminative ability of the risk prediction. The
358 PRS68- and PRS50-adjusted sex-and-age-specific absolute risks in our study had AUCs of
359 0.68-0.69, suggesting moderate discriminative ability. Whether this discriminative accuracy
360 would be sufficient for specific risk-stratified melanoma prevention or early detection
361 approaches will depend on the benefits, harms, and costs of the specific interventions. We
362 have provided data on sensitivity and specificity of the PRS-adjusted absolute risks at
363 different risk thresholds to enable health-economic evaluations of risk-stratified strategies.
364 Our analysis of estimated absolute risks by age and top/bottom PRS50 quintile suggests that
365 risk thresholds corresponding to the population-average risks at age 50 years might be
366 reached 10-20 years earlier by individuals with the 20% highest PRS, and 10-20 years later
367 by individuals with the 20% lowest PRS. The prediction could likely be improved by
368 incorporating information from variants that do not meet genome-wide significance
369 thresholds in association studies.²⁴ Due to extensive, population-dependent correlations
370 between genetic variants (linkage disequilibrium), optimising the set of variants used in a
371 PRS and the weights allocated to them is an active area of research. The choice of method
372 can be disease-dependent (a reflection of different genetic architectures), so further work is
373 needed to carry out a comprehensive comparison of different methods for melanoma.

374

375 Calibration of absolute melanoma risks was not straightforward, with age-and-sex-specific
376 risks derived from population data (without consideration of PRS) under-predicting
377 melanoma risks for participants of both cohorts. Adjustment of age-and-sex-specific risks by
378 PRS reduced the under-prediction, but further work will be required to build well-calibrated
379 risk prediction models. Potential reasons for the under-prediction include that our analyses
380 were restricted to European-ancestry participants only, which could lead to higher
381 melanoma incidence compared to the whole UK and Victorian populations. Increases in
382 melanoma risks determined from population-wide data might also be due to changes in risk
383 exposure and/or different diagnostic practices. Finally, the higher incidence of melanoma in

384 UKB and MCCS participants than in the general population could reflect some overdiagnosis
385 in more health-seeking participants, as participants from both cohorts are more likely to
386 have greater health literacy than the general population.²⁵ Melanoma incidence in UKB was
387 higher for participants with higher socioeconomic status, so over-representation of
388 participants with higher socioeconomic status in the cohorts²⁵ could also contribute to the
389 higher incidence than in the general population.

390
391 Comparing the three PRS, we found that PRS68 performed slightly better than PRS50 in
392 UKB, as expected given that UKB data contributed to the meta-analysis from which PRS68
393 was derived. We did not see a difference in performance between PRS68 and PRS50 in the
394 MCCS. This highlights the importance for independence between PRS construction and
395 validation data.²⁴ To our knowledge, this is the first evaluation of the risk prediction
396 performance for the PRS68 and PRS50 melanoma risk scores. A previous evaluation¹² of the
397 PRS45 score in UK and Australia case-control studies, and a Southern-European study²⁶ of a
398 204-variant score including non-genome-wide-significant variants based on pre-2020 data,
399 found similar or slightly weaker associations of PRS with melanoma incidence. Other
400 studies^{27,28} evaluating melanoma PRS used smaller sets of variants (16-24) known prior to
401 2020. A 55-variant PRS based on the 2020 meta-analysis PRS has also been associated with
402 melanoma risk (hazard ratio 1.46 per sd) in Australians aged 70+ years,¹⁵ but no assessment
403 of calibration was performed.

404
405 Considering any melanoma risk models, a recent systematic review identified 46 different
406 models from 40 publications, commonly using nevi (78% of models), but with little
407 consistency regarding other risk factors included.²⁹ Notably, only five publications assessed
408 AUC using independent validation data. For four risk models, the 95%CI for the external
409 validation AUC in the original publication or a later study overlapped the AUC estimates of
410 0.68-0.69 for the PRS50-adjusted absolute risks reported here (see Appendix S1). Only one
411 risk model incorporating clinically-assessed factors had a higher validation AUC (0.73, 95%CI
412 0.70-0.75, see Appendix S1).³⁰ Thus, the discrimination of the models based on genomic risk
413 factors, age and sex examined here is similar to externally validated phenotype-based risk
414 models. Incorporating genomic risk information into the best-available phenotype-based
415 risk models might further improve risk prediction in the future.¹²

416
417 The PRS68 and PRS50 scores were based on a larger meta-analysis than the older studies
418 underlying PRS45, and performed better than the PRS45, especially in the more ethnically
419 diverse MCCS. The 2020 meta-analysis had improved representation of southern European
420 populations compared to previous large-scale melanoma GWAS, which could improve
421 melanoma risk predictions for countries with more diverse European ancestry such as
422 Australia. Previous research^{31,32} demonstrates the importance of increasing diversity of
423 participants in genetic association studies that are used to construct PRS, and failure to do
424 so could lead to widening health disparities when PRS are used in practice.

425
426 Our work has several limitations. We only included participants of European ancestry, so
427 results cannot be extrapolated to different populations. In the MCCS data, self-reported
428 ethnicity could also reflect differences in migration status (e.g. those selecting Australian
429 ethnicity were perhaps more likely to be born in Australia and thus have higher underlying

430 risk due to higher ambient ultraviolet radiation). The results from the MCCS had more
431 uncertainty due to a smaller sample size.

432
433 Our work also has multiple strengths. We used data from two large prospective cohorts,
434 ensuring the evaluations of melanoma PRS were valid in different settings. Moreover, we
435 calculated absolute risks that include competing mortality risks, which is important to not
436 over-estimate melanoma risks for older age groups.³³ The results were reported according
437 to best-practice for PRS reporting³⁴ and details for the PRS have been deposited in the
438 Polygenic Score Catalog³⁵ to enable re-use. We also conducted thorough sensitivity analyses
439 to verify the robustness of our results.

440
441 In conclusion, newer melanoma PRS derived from a larger meta-analysis have better risk
442 prediction performance than earlier PRS, and could be used to tailor melanoma prevention
443 and early detection strategies to different risk levels. Re-calibration of 10-year absolute risks
444 may be necessary when applied to specific populations. Ensuring risk prediction models are
445 well-calibrated and can be applied to diverse populations remain key aspects for future
446 research.

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455 participate in the UK Biobank and Melbourne Collaborative Cohort Study, respectively. We
456 would like to thank the research participants and employees of 23andMe, Inc. for making
457 this work possible (a detailed list is included in the appendix).

458
459

460 **Contributions**

461 Conception and study design: AEC, RJM, JS. UKB data extraction and QC: MMI, DTB. MCCS
462 data collection, data extraction, genotyping, QC and imputation: RJM, RLM, GGG, XW, TND,
463 MCS. MCCS additional QC: JS, JYL. GWAS meta-analysis without UKB data: MHL. Funding
464 acquisition: GJM, AEC, RJM. PRS analyses: JS, RJM, AEC. Interpretation of results: all authors.
465 Writing - original draft: JS, AEC. Writing - editing: all authors.

466
467

468 **Data availability**

469 Bona fide researchers can apply to access the UK Biobank data for health-related research
470 that is in the public interest as described on the UK Biobank website
471 (<https://www.ukbiobank.ac.uk/enable-your-research>, accessed 26/04/2021). The MCCS
472 data can be made available on request to pedigree@cancervic.org.au.

473
474

475 **Supporting Information**

476 Appendix S1: Supplementary Methods and Results
477 Supplementary Figures S1-6
478 Supplementary Tables S1-7
479 PRS Reporting Checklist (based on GRIPS)

480 **Figure Legends**

481

482 **Figure 1. Calibration of PRS-adjusted and unadjusted (based on age and sex only) absolute**
483 **melanoma risks in UK Biobank and the MCCS, by risk quintile.**

484 (a) Calibration of absolute melanoma risks in UKB based on population-wide data for
485 2011-2015.

486 (b) Calibration of absolute melanoma risks in the MCCS based on population-wide data
487 for 2009-2013.

488 Bars show 95% confidence intervals.

489

490 **Figure 2. Calibration of absolute melanoma risks after re-scaling to $E/O=1$ in (a) UKB and**
491 **(b) the MCCS, by risk quintile.**

492 Bars show 95% confidence intervals.

493

494 **Figure 3. PRS-adjustment improves discriminative ability of absolute melanoma risks in (a)**
495 **UKB and (b) the MCCS.** Shaded areas indicate 95% confidence intervals. Absolute risks were
496 calculated based on population-level cancer registry data, age and sex, with and without
497 PRS adjustment. The straight diagonal line (grey) represents the line of no discrimination.
498 AUC: area under the receiver operating characteristic curve (with 95% confidence intervals).
499 ns: $p>0.05$; ** $p<0.001$; *** $p<0.0001$; **** $p<10^{-10}$.

500

501 **Figure 4. Estimated 10-year absolute melanoma risks and 95% CIs by top or bottom PRS50**
502 **quintile and age based on data from (a) UKB and population-wide data for England/Wales,**
503 **and (b) the MCCS and population-wide data for Victoria, Australia.**

504 (a) The estimated population-average 10-year risk of invasive melanoma for 50-year old
505 males in England/Wales was 0.28%, which is estimated to be reached 15 years earlier and
506 15 years later for males in the top and bottom PRS50 quintiles, respectively. Similarly, the
507 estimated population-average 10-year risk of melanoma for 50-year old females in
508 England/Wales was 0.31%, estimated to be reached more than 20 years earlier for those in
509 the top PRS50 quintile, and not before age 80 for those in the bottom quintile. The results
510 were similar for Scotland.

511 (b) For Victoria, the estimated population-average 10-year risk for 50-year old males was
512 0.67%, estimated to be reached about 10 years earlier and 10 years later for those in the
513 top and bottom quintiles, respectively. For females, the population-average 10-year risk at
514 age 50 was 0.52%, estimated to be reached 10 years earlier and not before age 80 for those
515 in the top and bottom PRS50 quintiles, respectively.

Tables

Table 1. Characteristics of UK Biobank (UKB) and Melbourne Collaborative Cohort Study (MCCS) participants. *Participants with non-European ancestry were excluded, so participants who self-reported Australian/New Zealand ethnicity in the MCCS descended from European migrants.

	UKB			MCCS All participants with follow-up		
	Cohort (n=395,647)	Cases (n=1,651)		2 visit (n=20,556)	Subcohort (n=4,528)	Cases (n=303)
Incident melanoma cases: n (%)	1,651 (0.4)	1,651 (100.0)	Incident melanoma cases: n (%)	304 (1.5)	66 (1.5)	303 (100.0)
Death before melanoma: n (%)	11,832 (3.0)	0 (0.0)	Death before melanoma: n (%)	2,585 (12.6)	343 (7.6)	0 (0.0)
Sex: Male (%)	183,417 (46.4)	815 (49.4)	Sex: Male (%)	8,047 (39.2)	1,737 (38.4)	178 (58.7)
Age: mean (SD)	56.91 (7.95)	58.95 (7.49)	Age: mean (SD)	63.5 (7.2)	63.4 (7.2)	65.0 (6.9)
Follow-up time: mean (SD)	7.16 (0.79)	7.26 (0.81)	Follow-up time: mean (SD)	9.7 (1.3)	9.7 (1.2)	9.6 (1.5)
Ethnicity: n (%)			Ethnicity: n (%)			
<i>British</i>	378,305 (95.6)	1,588 (96.2)	<i>Australian/New Zealand*</i>	15387 (74.9)	3,400 (75.1)	277 (91.4)
<i>Irish</i>	11,375 (2.9)	45 (2.7)	<i>British/Irish</i>	1459 (7.1)	316 (7.0)	16 (5.3)
<i>White / Any other white background</i>	5,965 (1.5)	18 (1.1)	<i>Greek/Italian/ Maltese</i>	3705 (18.0)	812 (17.9)	10 (3.3)
Ease of tanning: n (%)			Ease of tanning: n (%)			
<i>Never tan, only burn</i>	69,500 (17.6)	343 (20.8)	<i>Burn, rarely tan</i>	4,940 (24.0)	1,095 (24.2)	85 (28.1)
<i>Mildly or occasionally tanned</i>	83,561 (21.1)	400 (24.3)	<i>Burn, then tan</i>	10,111 (49.2)	2,230 (49.2)	163 (53.8)
<i>Moderately tanned</i>	156,426 (39.6)	641 (38.9)	<i>Tan, rarely burn</i>	5,505 (26.8)	1,203 (26.6)	55 (18.2)
<i>Very tanned</i>	78,190 (19.8)	244 (14.8)				
<i>Missing</i>	7,609 (2.0)	20 (1.2)				

SD: standard deviation

Table 2. Associations between melanoma PRS and with melanoma incidence, after adjustment for traditional risk factors and genetic principal components (all p-values are significant after Bonferroni multiple-testing correction). Associations per PRS quintile and decile see Table S4 and Figure S2.

Covariates	PRS	UKB		MCCS	
		SHR ¹ [95%CI ²]	p-value	SHR [95%CI]	p-value
Age, sex	PRS68	1.78 [1.70-1.86]	<1x10 ⁻¹⁵	1.65 [1.49-1.83]	<1x10 ⁻¹⁵
	PRS50	1.72 [1.65-1.80]	<1x10 ⁻¹⁵	1.64 [1.48-1.82]	<1x10 ⁻¹⁵
	PRS45	1.61 [1.54-1.68]	<1x10 ⁻¹⁵	1.42 [1.28-1.58]	5.7x10 ⁻¹¹
Age, sex, ethnicity	PRS68	1.78 [1.71-1.87]	<1x10 ⁻¹⁵	1.51 [1.35-1.68]	9.5x10 ⁻¹⁴
	PRS50	1.72 [1.65-1.80]	<1x10 ⁻¹⁵	1.50 [1.35-1.67]	1.5x10 ⁻¹³
	PRS45	1.61 [1.54-1.68]	<1x10 ⁻¹⁵	1.29 [1.16-1.45]	7.7x10 ⁻⁶
Age, sex, ethnicity, ease of tanning, 20 genetic PCs	PRS68	1.80 [1.71-1.88]	<1x10 ⁻¹⁵	1.47 [1.31-1.67]	3.6x10 ⁻¹⁰
	PRS50	1.73 [1.65-1.81]	<1x10 ⁻¹⁵	1.47 [1.30-1.66]	5.4x10 ⁻¹⁰
	PRS45	1.61 [1.53-1.69]	<1x10 ⁻¹⁵	1.25 [1.10-1.42]	6.2x10 ⁻⁴
Age, sex, ethnicity, ease of tanning, 20 genetic PCs, education, Townsend deprivation index	PRS68	1.80 [1.71-1.89]	<1x10 ⁻¹⁵	-	-
	PRS50	1.73 [1.65-1.81]	<1x10 ⁻¹⁵	-	-
	PRS45	1.61 [1.54-1.69]	<1x10 ⁻¹⁵	-	-
Age, sex, ethnicity, ease of tanning, 20 genetic PCs, skin colour, hair colour	PRS68	1.74 [1.65-1.83]	<1x10 ⁻¹⁵	-	-
	PRS50	1.67 [1.59-1.76]	<1x10 ⁻¹⁵	-	-
	PRS45	1.55 [1.47-1.63]	<1x10 ⁻¹⁵	-	-
Age, sex, ethnicity, ease of tanning, 20 genetic PCs, education, Townsend deprivation index, skin colour, hair colour	PRS68	1.74 [1.66-1.83]	<1x10 ⁻¹⁵	-	-
	PRS50	1.67 [1.60-1.76]	<1x10 ⁻¹⁵	-	-
	PRS45	1.55 [1.47-1.63]	<1x10 ⁻¹⁵	-	-

¹ SHR: subhazard ratio per 1 standard deviation (sd) of PRS on log-scale, with death as competing risk (for the MCCS, sd was based on subcohort participants)

² 95%CI: 95% confidence interval

References

- 1 Chatterjee N, Shi J, Garcia-Closas M. Developing and evaluating polygenic risk prediction models for stratified disease prevention. *Nat Rev Genet* 2016; **17**: 392-406.
- 2 Kerr A, Broer T, Ross E *et al*. Polygenic risk-stratified screening for cancer: Responsibilization in public health genomics. *Soc Stud Sci* 2019; **49**: 605-26.
- 3 Pashayan N, Guo Q, Pharoah PD. Personalized screening for cancers: should we consider polygenic profiling? *Per Med* 2013; **10**: 511-3.
- 4 Khera AV, Chaffin M, Aragam KG *et al*. Genome-wide polygenic scores for common diseases identify individuals with risk equivalent to monogenic mutations. *Nat Genet* 2018; **50**: 1219-24.
- 5 Ferlay JL, M.; Ervik, M.; Lam, F.; Colombet, M.; Mery, L.; Piñeros, M.; Znaor, A.; Soerjomataram, I.; Bray, F. Global Cancer Observatory: Cancer Tomorrow. In, Vol. 2021. 2020.
- 6 Arnold M, de Vries E, Whiteman DC *et al*. Global burden of cutaneous melanoma attributable to ultraviolet radiation in 2012. *Int J Cancer* 2018; **143**: 1305-14.
- 7 Olsen CM, Carroll HJ, Whiteman DC. Estimating the attributable fraction for cancer: A meta-analysis of nevi and melanoma. *Cancer Prev Res (Phila)* 2010; **3**: 233-45.
- 8 Lu Y, Ek WE, Whiteman D *et al*. Most common 'sporadic' cancers have a significant germline genetic component. *Hum Mol Genet* 2014; **23**: 6112-8.
- 9 Law MH, Aoude LG, Duffy DL *et al*. Multiplex melanoma families are enriched for polygenic risk. *Hum Mol Genet* 2020; **29**: 2976-85.
- 10 Landi MT, Bishop DT, MacGregor S *et al*. Genome-wide association meta-analyses combining multiple risk phenotypes provide insights into the genetic architecture of cutaneous melanoma susceptibility. *Nat Genet* 2020; **52**: 494-504.
- 11 Armstrong BK, Cust AE. Sun exposure and skin cancer, and the puzzle of cutaneous melanoma: A perspective on Fears *et al*. Mathematical models of age and ultraviolet effects on the incidence of skin cancer among whites in the United States. *American Journal of Epidemiology* 1977; **105**: 420-427. *Cancer Epidemiol* 2017; **48**: 147-56.
- 12 Cust AE, Drummond M, Kanetsky PA *et al*. Assessing the Incremental Contribution of Common Genomic Variants to Melanoma Risk Prediction in Two Population-Based Studies. *J Invest Dermatol* 2018; **138**: 2617-24.
- 13 Lewis CM, Vassos E. Polygenic risk scores: from research tools to clinical instruments. *Genome Med* 2020; **12**: 44.
- 14 Moorthie S, Hall A, Janus J *et al*. Polygenic scores and clinical utility. In: PHG Foundation, Cambridge. 2021.
- 15 Bakshi A, Yan M, Riaz M *et al*. Genomic Risk Score for Melanoma in a Prospective Study of Older Individuals. *J Natl Cancer Inst* 2021.
- 16 Lautenbach DM, Christensen KD, Sparks JA *et al*. Communicating genetic risk information for common disorders in the era of genomic medicine. *Annu Rev Genomics Hum Genet* 2013; **14**: 491-513.
- 17 Bycroft C, Freeman C, Petkova D *et al*. The UK Biobank resource with deep phenotyping and genomic data. *Nature* 2018; **562**: 203-9.
- 18 Milne RL, Fletcher AS, MacInnis RJ *et al*. Cohort Profile: The Melbourne Collaborative Cohort Study (Health 2020). *Int J Epidemiol* 2017; **46**: 1757-i.
- 19 Townsend P, Phillimore P, Beattie A. Health and Deprivation: Inequality and the North. *Routledge, London* 1988.

- 20 Mealiffe ME, Stokowski RP, Rhees BK *et al.* Assessment of clinical validity of a breast cancer risk model combining genetic and clinical information. *J Natl Cancer Inst* 2010; **102**: 1618-27.
- 21 Fay MP, Pfeiffer R, Cronin KA *et al.* Age-conditional probabilities of developing cancer. *Stat Med* 2003; **22**: 1837-48.
- 22 Li SX, Milne RL, Nguyen-Dumont T *et al.* Prospective Evaluation of the Addition of Polygenic Risk Scores to Breast Cancer Risk Models. *JNCI Cancer Spectrum* 2021.
- 23 Lee SH, Goddard ME, Wray NR *et al.* A Better Coefficient of Determination for Genetic Profile Analysis. *Genetic Epidemiology* 2012; **36**: 214-24.
- 24 Choi SW, Mak TS, O'Reilly PF. Tutorial: a guide to performing polygenic risk score analyses. *Nat Protoc* 2020; **15**: 2759-72.
- 25 Fry A, Littlejohns TJ, Sudlow C *et al.* Comparison of Sociodemographic and Health-Related Characteristics of UK Biobank Participants With Those of the General Population. *Am J Epidemiol* 2017; **186**: 1026-34.
- 26 Gu F, Chen TH, Pfeiffer RM *et al.* Combining common genetic variants and non-genetic risk factors to predict risk of cutaneous melanoma. *Hum Mol Genet* 2018; **27**: 4145-56.
- 27 Graff RE, Cavazos TB, Thai KK *et al.* Cross-cancer evaluation of polygenic risk scores for 16 cancer types in two large cohorts. *Nat Commun* 2021; **12**: 970.
- 28 Fritsche LG, Gruber SB, Wu Z *et al.* Association of Polygenic Risk Scores for Multiple Cancers in a Phenome-wide Study: Results from The Michigan Genomics Initiative. *Am J Hum Genet* 2018; **102**: 1048-61.
- 29 Kaiser I, Pfahlberg AB, Uter W *et al.* Risk Prediction Models for Melanoma: A Systematic Review on the Heterogeneity in Model Development and Validation. *International Journal of Environmental Research and Public Health* 2020; **17**.
- 30 Vuong K, Armstrong BK, Drummond M *et al.* Development and external validation study of a melanoma risk prediction model incorporating clinically assessed naevi and solar lentigines. *Br J Dermatol* 2020; **182**: 1262-8.
- 31 Martin AR, Kanai M, Kamatani Y *et al.* Clinical use of current polygenic risk scores may exacerbate health disparities. *Nat Genet* 2019; **51**: 584-91.
- 32 Lewis ACF, Green RC. Polygenic risk scores in the clinic: new perspectives needed on familiar ethical issues. *Genome Med* 2021; **13**: 14.
- 33 Bach AC, Lo KS, Pathirana T *et al.* Is the risk of cancer in Australia overstated? The importance of competing mortality for estimating lifetime risk. *Med J Aust* 2020; **212**: 17-22.
- 34 Wand H, Lambert SA, Tamburro C *et al.* Improving reporting standards for polygenic scores in risk prediction studies. *Nature* 2021; **591**: 211-9.
- 35 Lambert SA, Gil L, Jupp S *et al.* The Polygenic Score Catalog as an open database for reproducibility and systematic evaluation. *Nat Genet* 2021.