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1	Observational evidence of increasing global radiative forcing
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18	
19	Key Points
20	• Observed instantaneous radiative forcing has increased, strengthening the top-of-
21	atmosphere radiative imbalance.
22	• Due to cancellations in longwave and shortwave radiation, the sum of rapid adjustments
23	and radiative feedbacks exhibit an insignificant trend.
24	• Observed increases in instantaneous radiative forcing are direct evidence of the
25	anthropogenic effects on the Earth's radiative energy budget.
26 27 28 29	
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32	

33 Abstract

34

35 Changes in atmospheric composition, such as increasing greenhouse gases, cause an initial 36 radiative imbalance to the climate system, quantified as the instantaneous radiative forcing. This 37 fundamental metric has not been directly observed globally and previous estimates have come 38 from models. In part, this is because current space-based instruments cannot distinguish the 39 instantaneous radiative forcing from the climate's radiative response. We apply radiative kernels 40 to satellite observations to disentangle these components and find all-sky instantaneous radiative forcing has increased 0.53±0.11 W/m<sup>2</sup> from 2003 through 2018, accounting for positive trends in 41 42 the total planetary radiative imbalance. This increase has been due to a combination of rising 43 concentrations of well-mixed greenhouse gases and recent reductions in aerosol emissions. These 44 results highlight distinct fingerprints of anthropogenic activity in Earth's changing energy 45 budget, which we find observations can detect within 4 years.

46

### 47 Plain Language Summary

48 Climate change is a response to energy imbalances in the climate system. For example, rising 49 greenhouse gases directly cause an initial imbalance, the radiative forcing, in the planetary 50 radiation budget, and surface temperatures increase in response as the climate attempts to restore 51 balance. The radiative forcing and subsequent radiative feedbacks dictate the amount of 52 warming. While there are well-established observational records of greenhouse gas 53 concentrations and surface temperatures, there is not yet a global measure of the radiative 54 forcing, in part because current satellite observations of Earth's radiation only measure the sum 55 total of radiation changes that occur. We use the radiative kernel technique to isolate radiative

forcing from total radiative changes and find it has increased from 2003 through 2018, accounting for nearly all of the long-term growth in the total top-of-atmosphere radiation imbalance during this period. We confirm that rising greenhouse gas concentrations account for most of the increases in the radiative forcing, along with reductions in reflective aerosols. This serves as direct evidence that anthropogenic activity has affected Earth's energy budget in the recent past.

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### 1. Introduction

The Instantaneous Radiative forcing (IRF) is the initial imbalance of the Earth's top-of-65 66 the-atmosphere (TOA) radiative energy budget directly caused by a change in atmospheric 67 composition, such as increasing greenhouse gases (GHGs), or perturbed surface properties, like 68 from land use change. All anthropogenic climate changes are a response to the IRF, including 69 surface temperature change and associated radiative feedbacks (Sherwood et al. 2015). Despite a 70 sound basis in physics and radiative transfer theory, the IRF is hard to directly diagnose from 71 observations. Multiple remote sensing and in-situ instruments observe net radiative fluxes, but 72 these measurements convolve the IRF with radiative responses to the changing atmospheric 73 state. Some studies have diagnosed a more broadly defined "greenhouse effect" by evaluating 74 observations of clear-sky longwave radiation at the surface (Philipona et al. 2004) and TOA 75 (Raghuraman et al. 2019), but this analysis does not separate the IRF from water vapor feedback 76 processes.

Harries et al. (2001) compared outgoing longwave radiation at the TOA from two
satellite instruments launched decades apart, attributing emission differences at relevant spectral
bands to rising greenhouse gas (GHG) concentrations. However, instrumental uncertainty
between the two platforms complicates interpretation (Jiang et al. 2011). Feldman et al. (2015,

81 2018) used ground observations from the US Department of Energy Atmospheric Radiation 82 Measurement (ARM) program to provide the most observationally-oriented assessment to date 83 of GHG surface radiative forcing, which is proportional to the TOA IRF. However, their 84 analysis was limited to longwave (LW) forcing from CO<sub>2</sub> and CH<sub>4</sub> and was only conducted for 85 two locations. The total IRF has not been directly diagnosed globally from observations. 86 Well understood radiative transfer theory tightly constraints the GHG component of the IRF. 87 Line-by-line radiative transfer models diagnose it within 1% agreement (Collins et al. 2006; 88 Mlynczak et al. 2016; Pincus et al. 2020). However, these highly accurate calculations are 89 computationally expensive, so analysis is often limited to a few idealized atmospheric profiles. 90 Quantifying the IRF globally and over time relies on more efficient but less accurate 91 parameterized radiative transfer models (Soden et al. 2018), which introduces model bias when 92 applied to observations. Diagnosing the IRF from aerosols with these models suffers from the 93 same pitfalls, plus additional uncertainty associated with aerosol optical properties that are not well-observed (Randles et al. 2013; Stier et al. 2013). While there have been recent efforts to 94 95 constrain aerosol IRF with observations (Bellouin et al. 2020; Watson-Parris et al. 2020), results 96 are usually not temporally resolved.

97 Here we circumvent these limitations by applying radiative kernels (Soden et al. 2008) to 98 isolate the IRF from radiative feedbacks and rapid adjustments over time. We demonstrate that 99 the IRF has increased with rising GHG concentrations, accounting for recent, positive trends in 100 the total TOA radiative imbalance. More specifically, we consider this IRF to be largely a 101 consequence of concentration changes after anthropogenic emissions are moderated by natural 102 carbon cycle responses (Friedlingstein et al. 2019).

104 **2. Methods** 

105 106 Variations in the total, all-sky radiative energy balance at the TOA, dR, constrain global 107 surface temperature change and consists of the all-sky instantaneous radiative forcing (IRF) and 108 radiative responses to the IRF: 109  $dR = IRF + dR_{\lambda}$ 110 (1), 111 112 113 where  $dR_{\lambda}$  is net radiative changes caused by surface temperature-mediated radiative feedbacks 114 and rapid adjustments from, to first order, temperature (T), water vapor (a), surface albedo ( $\alpha$ ) 115 and cloud (*C*) changes (Vial et al. 2013; Sherwood et al. 2015): 116  $dR_{\lambda} = dR_T + dR_a + dR_{\alpha} + dR_C$ 117 (2). 118 119 For simplicity, we will not decompose these terms further into feedbacks and rapid adjustments 120 since it has no bearing on diagnosing the IRF. We simply refer to these radiative anomalies as 121 radiative responses. We note that  $dR_{\lambda}$  includes both anthropogenic responses and natural 122 variability (e.g. Trenberth et al. 2015). 123 The Clouds and Earth's Radiant Energy System (CERES) has provided global TOA

energy balance observations since 2000. Here, we diagnose *dR* using radiative flux anomalies from the CERES Energy Balance and Filled (EBAF) Ed. 4.1 product (Loeb et al. 2018a; Loeb et al. 2019). While no observational product measures the radiative response terms in isolation, they can be diagnosed using radiative kernels combined with observations of the relevant state 128 variable, *x* (B. Zhang et al. 2019; Bony et al. 2020). An individual, non-cloud radiative response, 129  $dR_x$ , in linear form is:

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131 
$$dR_x = \frac{\partial R}{\partial x} dx = K_x dx, \ x = T, q, \alpha \quad (3),$$

132

133 where  $K_x$  is a radiative kernel representing direct radiative changes from small, standard 134 perturbations in state variable *x* and *dx* is the actual temperature (*T*), water vapor (*q*) or surface 135 albedo ( $\alpha$ ) climate response. Under clear-sky (CS) conditions:

$$dR^{CS} = IRF^{CS} + dR_{\lambda}^{CS} \qquad (4),$$

- 137
- 138 where:

139 
$$dR_{\lambda}^{CS} = dR_T^{CS} + dR_q^{CS} + dR_{\alpha}^{CS} \qquad (5).$$

140

To diagnose  $dR_x$  or  $dR^{CS}_x$  we use observational-based radiative kernels developed from 141 142 the CloudSat Fluxes and Heating Rates product 2B-FLXHR-LIDAR (Kramer et al. 2019). 143 Unlike GCM-derived radiative kernels, these kernels are free from model bias in the base state, 144 and thus ideal for diagnosing observed radiation changes. Calculating  $K_x$  requires using a 145 radiative transfer model to convert base state perturbations to radiative sensitivities. Therefore, 146 using radiative kernels introduces some radiative-transfer model dependency. We apply the 147 radiative kernels to deseasonalized anomalies of temperature and specific humidity profiles from 148 version 6 Level 3 AIRS retrievals (Aumann et al. 2003) to estimate  $dR_T$  and  $dR_q$  and to surface albedo anomalies from CERES EBAF surface fluxes (Kato et al. 2018) to estimate  $dR_{\alpha}$ . Due to 149 150 computational expense, radiative kernels, including those used here, are often derived from one

151 year of data. However radiative kernel inter-annual variability is small (Pendergrass et al. 2018;

152 Thorsen et al. 2018), therefore applying radiative kernels to the entire observational record is153 justified.

In the traditional radiative kernel technique used here, the cloud radiative response  $(dR_C)$ is calculated as the change in cloud radiative effects (CRE) corrected for cloud masking (Soden et al, 2008; Kramer et al. 2019):

157

158 
$$dR_{c} = dCRE - (dR_{T} - dR_{T}^{CS}) - (dR_{q} - dR_{q}^{CS}) - (dR_{\alpha} - dR_{\alpha}^{CS}) - (IRF - IRF^{CS})$$
(6),

159

160 where CRE is the difference between all-sky and clear-sky radiative fluxes. The cloud masking 161 correction is necessary because CRE includes differences between all-sky and clear-sky non-162 cloud radiative changes, which are not actual cloud radiative responses (Soden et al. 2004). Here 163 dCRE is estimated using the TOA CERES EBAF radiative fluxes. The  $dR_x$  terms are diagnosed 164 using all-sky and clear-sky radiative kernels as described above.

165 The ultimate goal of this study is to derive the IRF from these radiative kernel 166 calculations. Under clear-sky conditions, we simply diagnose IRF<sup>CS</sup> by rearranging Equation 3, 167 whereby:

168 
$$IRF^{cs} = dR^{cs} - dR^{cs}_{\lambda} = dR^{cs} - \left(dR^{cs}_{T} + dR^{cs}_{q} + dR^{cs}_{\alpha}\right)$$
(7),

169

For all-sky conditions, an analogous calculation would require  $dR_C$  to be removed from dR, but since estimating  $dR_C$  as in equation 6 requires the IRF to be known, this differencing technique is not possible. Following common practice (Soden et al. 2008; Vial et al. 2013), we estimate the all-sky IRF as:

$$IRF = \frac{IRF^{CS}}{Cl} \qquad (8),$$

176 where Cl is a constant that accounts for cloud masking of the IRF. For the longwave (LW) Cl, 177 we use a constant of 1.24, derived by dividing clear-sky and all-sky double-call radiative transfer 178 calculations of CO<sub>2</sub> IRF from models (Smith et al. 2018). The cloud mask for the shortwave 179 (SW) is derived from direct output of aerosol IRF from Modern-Era Retrospective Analysis for 180 Research and Applications, Version 2 (MERRA-2) reanalysis (Gelaro et al. 2017). The global-181 mean value is 2.43, in line with a range of observational-based cloud masking estimates by 182 Bellouin et al. (2020). Only the MERRA-2 SW Cl is available over time, but it has an 183 insignificant long-term trend. Consequently, SW IRF has nearly identical trends when computed 184 with a time resolved versus constant SW Cl. 185 This conversion to all-sky conditions accounts for the presence of clouds but not cloud 186 changes. Therefore, the IRF in this study does not include aerosol-cloud interactions, such as

187 cloud albedo effects (Boucher et al. 2013). Instead, these terms are included in  $dR_c$ . Therefore, 188 the aerosol component to the kernel-derived estimates of IRF is akin to aerosol direct radiative 189 effects found throughout the literature (e.g. Thorsen et al. 2020).

The AIRS L3 data has the shortest record among satellite observations used in this study, with 2003 being the first complete year of data. Thus, we compute all deseasonalized anomalies from 2003 through 2018 relative to the mean of that time span. While we refer to the resulting calculation as the IRF for brevity, we actually show anomalies of the IRF. For comparison, we also estimate the IRF by applying the CloudSat radiative kernels to MERRA-2 reanalysis over the same period. This reanalysis product assimilates a variety of satellite observations, including observations of aerosol properties.

197 In climate models, idealized simulations and flux diagnostics from double-call radiative 198 transfer calculations can be used to evaluate the accuracy of radiative kernel estimates of  $dR_{\lambda}$  and 199 IRF (e.g. Vial et al. 2013; Smith et al. 2018). Such a comparison is not possible in the observed 200 record or the MERRA-2 reanalysis, however. Since the IRF is derived from differencing the 201 other radiative terms, there will always be near-perfect energy closure, albeit with some error due 202 to cloud masking assumptions, which is typically small (Chung and Soden 2015). Alternatively, 203 we will compare these kernel-derived estimates to various independent measures of the IRF. 204 To verify the aerosol component of the IRF, we compare radiative kernel-derived SW 205 IRF to direct output of the aerosol direct radiative effect from MERRA-2. We also compare SW 206 IRF to trends in aerosol optical depth (AOD) from MERRA-2 and observations from the 207 Moderate Resolution Imaging Spectroradiometer (MODIS) merged Dark Target and Deep Blue 208 product (Sayer et al. 2014).

209 We compare radiative-kernel derived estimates of the LW IRF to offline radiative 210 transfer calculations of GHG IRF. We apply empirical formulas to observed global-mean 211 concentrations of 5 major greenhouse gases (CO<sub>2</sub>, CH<sub>4</sub>, N<sub>2</sub>O, CFC-11 and CFC-12), provided by 212 NOAA Global Monitoring Division (Hoffman et al. 2006; Montzka et al. 2011). Etminan et al. 213 (2016) derive the empirical formulas from polynomial fits to line-by-line radiative forcing 214 calculations. While these formulas were originally developed for net stratospherically adjusted 215 radiative forcing, we use corrections from additional line-by-line calculations (Hodnebrog et al. 216 2013; Etminan et al. 2016) to calculate TOA IRF, decomposed into a LW and SW component. 217 We also estimate GHG IRF using the SOCRATES offline radiative transfer model 218 (Edwards et al. 1996; Manners et al. 2015) with NOAA GHG concentrations and atmospheric 219 profiles from the MERRA-2 reanalysis. Like the other IRF estimates, these calculations are

presented in anomaly space with the seasonal cycle removed. The IRF from CFCs has decreased recently, but this has been compensated for by a near equal increase from other halocarbons not considered in empirical fit and SOCRATES calculations (Myhre et al. 2013). To account for this, we repeat these calculations with no CFC trend. This only modifies total GHG IRF trends by <5%, however, so hereafter we focus on results without this assumption. The SOCRATES IRF calculations are conducted under pristine, clear-sky conditions and converted to all-sky via Equation 8, like the radiative kernel calculations.

227 The various inputs and assumptions detailed above can contribute uncertainty to the 228 estimated radiative changes. In a Supplemental Appendix we provide a comprehensive uncertainty assessment in the IRF trends due to these contributors, including from observed dR, 229 230 radiative kernels, and the cloud masking constant, Cl. We find these uncertainties are smaller 231 than the trend regression uncertainty associated with timeseries variability. Therefore, all trends 232 presented hereafter are provided with 95% confidence intervals (or roughly 2 standard errors 233 around the mean) associated with the least-squares linear regression. This is common practice 234 when diagnosing CERES trends (e.g. Loeb et al. 2018a,b).

The anomalies of dR,  $dR_{\lambda}$  and the IRF are subject to the same sources of uncertainty as longterm trends. Therefore, Figure 1 and 2 below include uncertainty bounds diagnosed as  $2\sigma$  across multiple estimates of the radiative terms using different radiative flux data products from CERES and alternative radiative kernel sets and model estimates of Cl (see Supplemental Appendix).

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3. Results
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Figure 1a shows a timeseries of global-mean total radiative flux anomalies (dR) from CERES satellite observations and its component from radiative responses ( $dR_{\lambda}$ ), estimated by applying

245 th	e CloudSat-based	l radiative kernels to	CERES and AIRS	observations	hereafter
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- 246 CERES/AIRS). Positive anomalies indicate a net increase in downwelling radiation at the TOA
- 247 (planetary warming). The sum of the radiative responses,  $dR_{\lambda}$ , accounts for nearly all of the total
- short-term *dR* variability, as evident by their strong correlation (r=0.88) and small root-mean-
- squared difference of  $0.024\pm0.003$  W/m<sup>2</sup>; ~3.5% of the standard deviation of *dR*. On inter-annual
- timescales, ENSO strongly influences this variability (Trenberth et al. 2014), which lags by ~5
- 251 months (Supplemental Fig. S1; Loeb et al. 2018b). Long-term dR exhibits a positive, linear trend
- 252 (0.038±0.02 W/m<sup>2</sup>/year) significant with 95% confidence, while  $dR_{\lambda}$  exhibits an insignificant
- trend ( $0.002\pm0.02$  W/m<sup>2</sup>/year) an order of magnitude smaller. This arises from cancelation
- between LW and SW  $dR_{\lambda}$ . The LW  $dR_{\lambda}$  has a negative linear trend (-0.042±0.02 W/m<sup>2</sup>/year)
- 255 (Fig. 1b), mainly from global warming-driven  $dR_T$  decreases (-0.041±0.007 W/m<sup>2</sup>/year)
- 256 (Supplemental Fig. S2). The SW  $dR_{\lambda}$  trend (0.044±0.02 W/m<sup>2</sup>/year) is nearly equal and opposite
- of the LW, driven by increases in SW  $dR_{\alpha}$  (0.023±0.09 W/m<sup>2</sup>/year) and SW  $dR_C$  (0.020±0.13
- 258 W/m<sup>2</sup>/year), a predominantly low cloud response (Loeb et al. 2018b). The latter alone accounts
- 259 for most of the SW interannual variability.





263 Figure 1. Global-mean a) net, b) longwave (LW) and c) shortwave (SW) total radiative flux anomalies from 2003 through 2018 as measured by CERES (black) and the contribution to that 264 265 total from the sum of radiative responses (red). Respective trendlines are displayed as dashed 266 lines. Uncertainty of  $+/-2\sigma$  is shown for each timeseries, computed as described in the Methods. 267 Linear trends and 95% confidence intervals are provided in text.



269 LW and SW components (Supplemental Fig. S3). However, there is a positive trend in LW  $dR_{\lambda}$ 

270 and a negative trend in SW  $dR_{\lambda}$ , opposite from the CERES/AIRS response. This occurs due to a

271 considerably different LW and SW  $dR_C$  (Supplemental Fig. S4) compared to satellite

272 observations.

273	Since neither $dR_{\lambda}$ or its uncertainties account for the positive $dR$ trend, it must be
274	explained by the IRF. Figure 2 shows the timeseries of the total, LW and SW IRF under all-sky
275	conditions, estimated from the radiative kernel technique. The total CERES/AIRS IRF exhibits a
276	significant, positive trend (0.033 $\pm$ 0.007 W/m <sup>2</sup> /year), mostly from increasing LW IRF
277	(0.027 $\pm$ 0.006 W/m <sup>2</sup> /year). The SW IRF exhibits a smaller, yet still significant increase
278	(0.006±0.003 W/m <sup>2</sup> /year). The LW IRF trend is opposite in sign from LW $dR$ , since decreasing
279	LW $dR_{\lambda}$ compensates. In the SW, IRF and $dR$ are both increasing, but SW $dR_{\lambda}$ is the dominant
280	contributor while the IRF trend is much smaller.
281	Rising GHG concentrations explain the positive LW IRF trend. Accordingly, it increases
282	at a similar rate to the GHG IRF estimates from the empirical fit (0.021 $\pm$ 0.0002 W/m <sup>2</sup> /year or
283	$0.022\pm0.0002$ W/m <sup>2</sup> /year if ignoring CFCs [see Methods]) and the SOCRATES radiative
284	transfer model ( $0.023\pm0.0003$ W/m <sup>2</sup> /year) (Fig. 2b), despite these calculations neglecting some
285	GHG forcers found in nature, such as ozone. MERRA-2 exhibits a similar LW IRF trend to
286	CERES/AIRS (0.029 $\pm$ 0.003 W/m <sup>2</sup> /year) while direct output of the LW aerosol IRF from
287	MERRA-2 exhibits no trend. This further indicates GHG increases account for roughly all LW
288	IRF increases.







291 *Figure 2.* Global-mean a) total, b) longwave (LW) and c) shortwave (SW) instantaneous

292 radiative forcing (IRF) estimated from the radiative kernel technique for CERES/AIRS (red) and

293 *MERRA-2 (blue). Additional calculations of greenhouse gas-only IRF are also shown using* 

294 empirical formulas (green) and the SOCRATES radiative transfer model (gray). For reference,

295 the trendline for total radiative flux anomalies (Fig 1a) is displayed with the total IRF as a black

296 dashed line. Uncertainty of  $+/-2\sigma$  is shown with shading for each timeseries, computed as

297 described in the Methods. Linear trends and 95% confidence intervals are provided in text and

*in Table 1.* 298

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Increasing GHG concentrations also contribute (0.002\pm0.00 \text{ W/m}^2/\text{year}) to the total
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negligible in the SOCRATES calculations, but the model version used here does not account for
 the SW absorption of CH<sub>4</sub>.

303 The total SW IRF increase is nearly identical in CERES/AIRS and MERRA-2, and to 304 aerosol-only SW IRF trends from MERRA-2 direct output (Supplemental Fig. S5). They also 305 exhibit similar short-term variability. This suggests aerosols explain most of the SW IRF. The 306 long-term radiative heating is consistent with declining anthropogenic aerosol emissions during 307 this period (Q. Zhang et al. 2019). Towards the end of the timeseries, CERES/AIRS SW IRF has 308 more positive anomalies. Locally, the largest differences with MERRA-2 after 2015 are in major 309 absorbing aerosol source regions (Supplemental Fig. S6), suggesting a contribution from 310 different absorbing aerosol properties.

311 Figure 3 shows local linear trends in kernel-derived, total SW IRF from CERES/AIRS 312 and MERRA-2 and direct MERRA-2 output of aerosol-only SW IRF (Figure 3c). The spatial 313 pattern of the SW IRF trend is generally consistent across all three estimates. A notable 314 hemispheric asymmetry is present, with large changes concentrated in the populous Northern 315 Hemisphere. This includes large positive trends over the Eastern United States, Western Europe 316 and Eastern China, where anthropogenic emissions of reflective aerosols have declined because 317 of government actions to combat poor air quality (Kühn et al. 2014; Ridley et al. 2018; Q. Zhang 318 et al. 2019). In contrast, the SW IRF trends are negative over India, where emissions continue to 319 rise (Dey et al. 2012).

There are some magnitude differences in these major source regions, however. For instance, trends are larger in the Eastern US and India in CERES/AIRS than in MERRA-2. This coincides with differences in the MODIS and MERRA-2 AOD trends (Figure 3d,e), which are also larger in CERES/AIRS. Over Saharan Africa, the sign of the SW IRF trend differs,





0.09

340

341

342

343

344

0°

30°S

60°S

-0.09

-0.06

-0.03

0.00

 $W/m^2/year$ 

0.03

0.06

349	The strong agreement in MERRA-2 trends from kernel differencing versus direct SW
350	aerosol IRF output (Fig 3b,c) highlights the dominant role of aerosols in the total SW IRF trends.
351	It also confirms the accuracy of the radiative kernel technique. The kernel differencing method
352	results in artifacts in the polar regions, however, where large local trends are a consequence of
353	underestimating the SW $dR_{\alpha}$ removed from $dR$ (Supplemental Fig. S8) and not from actual
354	forcing. One possible explanation is surface albedo radiative kernels fail to capture important
355	ice-albedo feedback non-linearities (Block et al. 2013). Nevertheless, the polar region errors
356	have negligible effect on global-mean SW IRF trends.
357	Some inter- and intra-annual variability (hereafter short-term variability) in SW IRF is
358	expected, given natural variations in aerosol concentrations. Consequently, the detrended
359	aerosol-only ( $\sigma$ =0.088 W/m <sup>2</sup> ) and kernel-derived ( $\sigma$ =0.097 W/m <sup>2</sup> ) SW IRF in MERRA-2
360	exhibit similar variability and are highly correlated (r=0.78). The source of the notable short-
361	term variability in LW IRF (Fig. 2b) is less apparent, however, since greenhouse gas
362	concentrations increase relatively steadily on these timescales, as evident in the empirical fit
363	estimate of GHG IRF, which increases almost perfectly linearly.
364	While radiative kernel error may play some role, the LW IRF from CERES/AIRS
365	exhibits considerably more short-term variability ( $\sigma$ =0.24) than MERRA-2 ( $\sigma$ =0.16), despite
366	using the same CloudSat-derived radiative kernels in both estimates. This highlights short-term
367	inconsistencies between the radiative fluxes observed by CERES ( $dR^{cs}$ ) and the AIRS retrievals
368	used to diagnose LW $dR_{\lambda}^{cs}$ . For instance, the difference between CERES/AIRS and MERRA-2
369	$dR_{\lambda}^{cs}$ exhibits considerably more short-term variability than the difference between dR <sup>cs</sup> . This is
370	mostly due to different variability in $dR_T^{cs}$ (Supplemental Fig. S9), and more specifically due to

371 different temperature anomalies at the surface and in the boundary layer between AIRS and 372 MERRA-2 (Supplemental Fig. S10). Since AIRS temperature anomalies are more variable, so is 373 the  $dR_T^{cs}$  estimate. And since this variability is not also observed radiatively by CERES, it is not 374 evident in  $dR^{cs}$ . This ultimately translates to a more variable LW IRF when using the kernel 375 differencing technique. This also explains why LW IRF spatial patterns are noisier for 376 CERES/AIRS than for MERRA-2 (Supplemental Fig. S11). Cloud contamination likely 377 contributes to the AIRS temperature variability, as found previously (Hearty et al. 2014). This is 378 evident at the surface, for example, where the largest differences between AIRS and MERRA-2 379 temperature anomalies tend to occur where clouds are common (Supplemental Fig. S9), 380 especially over land. While global-mean surface temperature anomalies from AIRS closely agree 381 with other, independent datasets (Susskind et al. 2019), it is possible the temperature biases that 382 do exist are magnified in the context of radiative changes. 383 The LW IRF variability may also stem from its sensitivity to the atmospheric base state 384 (Pincus et al. 2015). However, this contribution appears to be small. In the LW GHG IRF 385 estimated from the SOCRATES radiative transfer model, we use daily MERRA-2 temperature, 386 surface albedo and humidity data, thus capturing the GHG IRF sensitivity to the unperturbed, 387 non-cloud base state. Still, the short-term variability from this offline calculation is nearly as 388 small as estimates with the empirical fit, which does not account for base state variability. The 389 LW IRF short-term variability in this comparison (and in the radiative kernel-derived estimates) 390 is not due to variations in the cloud base state since LW cloud masking is always treated as a 391 constant. While clouds may play a greater role in reality, the SW IRF estimated from radiative 392 kernels with constant cloud masking has similar short-term variability to the aerosol-only SW 393 IRF in MERRA-2, which accounts for cloud masking temporal variations. This suggests cloud

- 394 variability may not be important in the global-mean. Lastly, some LW IRF variability in
- 395 MERRA-2 (and in CERES/AIRS) may be due to spatial variability in the GHG concentrations
- 396 (Myhre et al. 2013), which is not present in the empirical fit or the SOCRATES estimates.
- 397

	LW	SW	Net
CERES/AIRS	0.027±0.006	0.006±0.003	0.033±0.007
MERRA-2	0.029±0.003	0.006±0.003	0.035±0.004
Aerosol-Only MERRA-2	-4.2E-4±1.5E-4	0.006±0.003	0.006±0.003

*Table 1.* Global-mean linear trends ( $W/m^2/vear$ ) and 95% confidence bounds in

instantaneous radiative forcing estimated using the radiative kernel differencing

technique (first two rows) and MERRA-2 flux diagnostics (third row).

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- 403 404
  - 4. Conclusions

405

406 We have diagnosed the global instantaneous radiative forcing (IRF) directly from 407 observations using radiative kernels. Table 1 summarizes linear trends. We find that from 2003 408 through 2018, the observed IRF has increased  $0.53\pm0.11$  W/m<sup>2</sup>, almost entirely accounting for 409 the positive trend in CERES Top-of-Atmosphere (TOA) radiative flux anomalies (dR). The 410 intrinsic LW and SW climate radiative responses largely cancel out. This IRF increase mostly 411 occurs in the LW ( $0.43\pm0.1$  W/m<sup>2</sup>), driven by rising greenhouse gas concentrations. This serves 412 as direct observational evidence that anthropogenic activity is impacting the Earth's energy balance. The SW IRF has also increased  $(0.1\pm0.05 \text{ W/m}^2)$ . In part, this is a reflection of 413 414 government-mandated aerosol emission reductions throughout major source regions, which may 415 have a greater direct impact than inferred by the SW IRF, which does not include aerosol cloud-416 albedo effects in this analysis.

417 Diagnosing the observed IRF is important for our fundamental understanding of Earth's 418 response to climate change and a valuable piece of information for policy decisions. 419 Conceivably, observed IRF could be used as a top-down approach for monitoring the climate 420 response to mitigation efforts. By applying published metrics of instrumental uncertainty in 421 AIRS (Tobin et al. 2006; Hearty et al. 2014) and CERES (Loeb et al. 2018), along with the 422 kernel-derived IRF variance and trend, we apply formulas by Leroy et al. (2008) to determine the 423 minimum length of the observational record necessary to detect a climate change signal. These 424 formulas account for trend uncertainty due to natural variability and instrumental uncertainty. 425 Using this approach, we find total IRF trends are detectable, given these sources of uncertainty, 426 within 3.8 years using the satellite data presented in this study. Therefore, the methods 427 introduced here could be useful for near-real time monitoring, especially since the time to 428 detection shortens with the lengthening of the observational record. 429 430 Acknowledgements: We thank the Editor, reviewers and Graeme Stephens for valuable 431 feedback on this work. RJK is supported by an appointment to the NASA Postdoctoral Program 432 administered by Universities Space Research Association. HH and BJS are supported by NASA 433 award 80NSSC18K1032. LO gratefully acknowledges support from NASA's 434 CloudSat/CALIPSO Science Team and MEaSUREs programs. GM, PMF and CJS were

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- 440 **Competing Interests:** Authors have no competing interests.
- 441
- 442 Data and Materials Availability: The CERES radiative flux observations are available at
- 443 <u>https://ceres.larc.nasa.gov/data/</u>. The AIRS temperature and water vapor observations and the
- 444 MERRA-2 reanalysis data are available at <u>https://disc.gsfc.nasa.gov/</u>. The CloudSat/CALIPSO
- 445 radiative kernels used in this study and related code for applying them are available at
- 446 <u>https://climate.rsmas.miami.edu/data/radiative-kernels/</u>.
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- 669
- 670 671
- 672 Figure 2. Global-mean a) net, b) longwave (LW) and c) shortwave (SW) total radiative flux
- anomalies from 2003 through 2018 as measured by CERES (black) and the contribution to that
- 674 total from the sum of radiative responses (red). Respective trendlines are displayed as dashed
- 675 lines. Uncertainty of  $\pm -2\sigma$  is shown for each timeseries, computed as described in the Methods.
- 676 Linear trends and 95% confidence intervals are provided in text.
- 677
- 678 **Figure 2.** Global-mean a) total, b) longwave (LW) and c) shortwave (SW) instantaneous
- 679 radiative forcing (IRF) estimated from the radiative kernel technique for CERES/AIRS (red) and
- 680 MERRA-2 (blue). Additional calculations of greenhouse gas-only IRF are also shown using
- 681 empirical formulas (green) and the SOCRATES radiative transfer model (gray). For reference,
- the trendline for total radiative flux anomalies (Fig 1a) is displayed with the total IRF as a black
- dashed line. Uncertainty of  $\pm -2\sigma$  is shown with shading for each timeseries, computed as
- described in the Methods. Linear trends and 95% confidence intervals are provided in text and
- 685 in Table 1.
- 686

687	Figure 3. Local linear trends from 2003 through 2018 in all-sky shortwave instantaneous
688	radiative forcing (SW IRF) diagnosed in a) CERES/AIRS observations and b) MERRA-2
689	reanalysis using the radiative kernel differencing technique and c) from direct output of
690	MERRA-2 aerosol IRF. Also, local linear trends over the same time period are shown for aerosol
691	optical depth (AOD) from d) MODIS and e) MERRA-2.
692	

- **Table 1.** Global-mean linear trends (W/m<sup>2</sup>/year) and 95% confidence bounds in instantaneous
- radiative forcing estimated using the radiative kernel differencing technique (first two rows) and
- MERRA-2 flux diagnostics (third row).

723	Supplementary Materials for
724	Observational evidence of increasing global radiative forcing
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729 730 731 732 733 734 735 736 737 738 739	<ul> <li><sup>1</sup>Earth Sciences Division, NASA Goddard Space Flight Center, Greenbelt, MD 20771, USA</li> <li><sup>2</sup>Universities Space Research Association, 7178 Columbia Gateway Drive, Columbia, MD 21046, USA</li> <li><sup>3</sup>Rosenstiel School of Marine and Atmospheric Science, University of Miami, 4600</li> <li>Rickenbacker Causeway, Miami, FL 33149, USA</li> <li><sup>4</sup>CICERO Center for International Climate and Environmental Research in Oslo, 0318, Oslo, Norway</li> <li><sup>5</sup> School of Earth and Environment, University of Leeds, Leeds, LS2 9JT, UK</li> <li><sup>6</sup>International Institute for Applied Systems Analysis (IIASA), Laxenburg, 2351, Austria</li> </ul>
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741	
742	Supplemental Appendix:
743	SA1. Uncertainty Quantification
744	Following common practice among previous CERES-focused literature (e.g. Loeb et al.
745	2018a,b), the trend uncertainty quoted throughout the main text is a measure of the linear
746	regression uncertainty, which is largely driven by the internal variability of the timeseries being
747	analyzed. The 95% confidence intervals are given. It is worthwhile to also evaluate uncertainty
748	due to the various assumptions and diagnostic tools that contribute to the estimate of the IRF. As
749	illustrated by equations 7 and 8 in the main text, all-sky IRF is estimated by subtracting
750	radiative-kernel derived, clear-sky radiative responses from the overall clear-sky TOA radiative
751	imbalance (dR <sup>CS</sup> ). This difference, an estimate of the clear-sky IRF (IRF <sup>CS</sup> ) is then divided by a
752	cloud masking constant, Cl, to convert IRF <sup>CS</sup> into an all-sky IRF. In this supplementary section,

we diagnose uncertainty in IRF trends associated with observations of dR<sup>CS</sup>, radiative kernels, 753 754 and Cl. We do so by repeating calculations of the IRF, each time substituting in different values 755 for these terms from different sources as explained below, while keeping all other terms 756 unchanged from the method and data described in the main text. Since the standard trend 757 uncertainty is dependent on these additional sources of uncertainty, it is not practical to combine 758 them to quantify a total, comprehensive measure of uncertainty. We therefore discuss these 759 sources individually, compare their relative magnitude, and summarize the uncertainty budget in 760 Table SA4. We focus on the observational estimate of the IRF.

761 SA1.1 Uncertainty in  $dR^{CS}$ 

762 In this work, observed, total TOA radiative anomalies are diagnosed using radiative flux 763 data from CERES EBAF 4.1. This is identical to CERES EBAF4.0 (Loeb et al. 2018a), except it 764 includes an additional clear-sky radiative flux dataset. While the traditional clear-sky products 765 are comprised only of pixels designated as cloud-free, the new product uses an adjustment factor 766 to mimic a total absence of clouds for all regions, similar to how clear-sky is defined in model 767 simulations (Loeb et al. 2019). While CERES has well documented uncertainty in the magnitude 768 of the TOA radiative flux measurements, our work to estimate the IRF is conducted in anomaly 769 space, where uncertainty in absolute fluxes is irrelevant. Instead, it is the uncertainty due to the 770 stability (or lack thereof) of the observing platform that is important. The presence of spurious 771 trends is frequently assessed by comparing EBAF products to the CERES SSF1deg product, 772 which is considered to be extremely stable (e.g. Loeb et al. 2018a,b). To determine associated uncertainty from stability in the observed timeseries of dR<sup>CS</sup>, we recompute the IRF using four 773 774 sources of clear-sky radiative fluxes: CERES EBAF 4.1 assuming clear-skies over the total 775 region, CERES EBAF 4.1 assuming clear-skies over cloud-free regions only (the traditional

- method), CERES SSF1deg from Terra and CERES SSF1deg from Aqua. All other components
- of the IRF calculation are consistent across the four estimates. Linear trends of the global-mean

CERES clear-sky Source	Net	LW	SW
EBAF 4.1 – clear-sky (for total	0.033	0.027	0.0061
region)			
EBAF 4.1 – clear-sky (for cloud-free	0.026	0.019	0.0064
areas of region)			
SSF1deg Terra	0.027	0.026	0.0015
SSF1deg Aqua	0.024	0.025	-0.0004
Standard Deviation	0.0041	0.0035	0.0034

778 IRF are summarized in Table SA1.

779 *Table SA1.* Linear trends from 2003 through 2018 in global mean net, longwave (LW) and

shortwave (SW) all-sky Instantaneous radiative forcing, estimated with differences observational

781 sources of clear-sky radiative fluxes used to diagnose clear-sky TOA radiative flux anomalies.

782 Units are  $W/m^2/yr$ .

783

Across the four estimates, we find a standard deviation of  $\sigma = 0.0041$  W/m<sup>2</sup>/yr for the Net 784 785 IRF,  $\sigma = 0.0035$  W/m<sup>2</sup>/yr in the LW and  $\sigma = 0.0034$  W/m<sup>2</sup>/yr in the SW. We consider this to be 786 an upper bound on uncertainty associated with the stability of the CERES observations, since 787 stability is not the only source of differences between these datasets. Ultimately, we use the new 788 EBAF 4.1 clear-sky fluxes, representing cloud absence over all regions, in the main analysis 789 since it is more consistent with the way clear-sky is defined in radiative kernels, the additional 790 offline radiative transfer calculations of LW GHG IRF, MERRA-2, and in climate models. 791 SA1.2 Radiative kernel uncertainty 792 Radiative kernels based on CloudSat/CALIPSO observations are used in this study to 793 quantify radiative responses to changes in temperature, water vapor and surface albedo. 794 Radiative kernels are constant in time (beyond a seasonal cycle) and therefore do not contribute 795 to any spurious trends in the diagnosis of the IRF. However, there is uncertainty in the magnitude 796 of the radiative kernels which can contribute to uncertainty in the anomalies and trend of the

797 IRF. To quantify this, we estimate the IRF using four different sets of radiative kernels: those

based on CloudSat/CALIPSO discussed in the main text, as well as radiative kernels derived

from the GFDL (Soden et al. 2008), ECHAM6 (Block and Mauritsen 2013) and HadGEM3

800 (Smith et al. 2020) climate models. All other components of the calculation are consistent across

801 the four estimates. Linear trends of global-mean IRF are summarized in Table SA2.

802

Radiative Kernel	Net	LW	SW
CloudSat/CALIPSO	0.0333	0.0272	0.0061
GFDL	0.0313	0.0286	0.0027
ECHAM6	0.0320	0.0297	0.0023
LL ACEM2	0.0222	0.0262	0.0060
HadGEMIS	0.0323	0.0265	0.0000
Standard Deviation	0.0008	0.0015	0.0020

803 **Table SA2.** Linear trends from 2003 through 2018 in global mean net, longwave (LW) and

shortwave (SW) all-sky Instantaneous radiative forcing, estimated using different sets of

805 radiative kernels. Units are  $W/m^2/yr$ .

806 We find a standard deviation in trend across the four estimates of  $\sigma = 0.0008 \text{ W/m}^2/\text{yr}$  for

807 Net IRF,  $\sigma = 0.0015 \text{ W/m}^2/\text{yr}$  in the LW and  $\sigma = 0.0020 \text{ W/m}^2/\text{yr}$  in the SW.

808 SA1.3 Uncertainty in the cloud masking term.

809 The cloud masking constant, Cl, used to estimate all-sky IRF accounts for the effect of

810 the presence of clouds on the magnitude of the IRF, relative to clear-sky conditions. This

811 quantity is not directly observable and typically requires separate radiative transfer calculations

- 812 to diagnose. Therefore, like radiative kernels, it contains uncertainty due to radiative transfer
- 813 error and due to biases in the cloud climatology used in those calculations.

A lack of data prohibits accurately computing the cloud masking directly from

815 observations. All- and clear-sky double-call calculations of the IRF from model simulations offer

816 the best alternative. However, as discussed by Soden et al. (2018), these diagnostics are rarely

817 conducted with model simulations. To the best of our knowledge, none are available for realistic,

818 historical forcing scenarios.

819 With these limitations, we assume the LW cloud masking is equivalent to the masking of 820 IRF from CO<sub>2</sub> perturbations in this study, which is the dominant GHG driver over the observed 821 period being evaluated. The Coupled Model Intercomparison Project phase 5 (CMIP5, Taylor et 822 al. 2013) includes the necessary double-call calculations from four models to diagnose CO<sub>2</sub> 823 cloud masking, using prescribed sea surface temperature, atmosphere-only simulations where 824 CO<sub>2</sub> concentrations are quadrupled. To diagnose uncertainty in the IRF trends due to Cl, we 825 recompute observed all-sky LW IRF by applying Cl estimated from these four models to the 826 observed clear-sky IRF.

For the SW, there are analogous simulations in CMIP5 (and CMIP6) for aerosol forcing scenarios, but there are no double-call calculations available to diagnose Cl. Instead, we use clear-sky and all-sky Direct Radiative Forcing (DRF) in 15 models included in the AeroCOM Phase II project (Myhre et al. 2013). Although DRF only includes anthropogenic aerosols, the model-mean Cl from these simulations is 2.70, close to the value from MERRA-2 for SW IRF used in the main text (2.43).

To determine associated uncertainty in IRF trends, we recompute IRF with each LW and SW value of Cl from the model simulations discussed above. Results are summarized in Table SA3.

Model	LW IRF	Model	SW IRF
CanAM4	0.0281	BCC	0.00354
HadGEM2-A	0.0272	CAM4-Oslo	0.00631
INMCM4	0.0253	GEOS_CHEM	0.00627
IPSL-CM5A-LR	0.0243	GISS_MATRIX	0.01081
		GISS-ModelE	0.01024
		GMI	0.00841
		GOCART	0.00913
		HadGEM2	0.00634
		IMPACT-Umich	0.00306
		INCA	0.00726
		ECHAM5-HAM	0.00502
		NCAR-CAM3.5	0.00557
		OsloCTM2	0.00362
		SPRINTARS	0.00290
		TM5	0.00923
Standard Deviation	0.00150	Standard Deviation	0.00253

839 *Table SA3.* Linear trends from 2003 through 2018 in global mean net, longwave (LW) and

shortwave (SW) all-sky Instantaneous radiative forcing, estimated using different cloud masking
constants derived from 4 CMIP5 models for the LW and 14 AeroCOM models for the SW that

842 provided the radiative flux diagnostics necessary for this calculation.

843 The standard deviation across the 4 estimates of LW IRF is  $\sigma = 0.00150$  W/m<sup>2</sup>/yr and  $\sigma =$ 

844 0.00253 W/m<sup>2</sup>/yr across the 14 estimates of SW IRF. Since different models are used for the

LW and SW component, we estimate the standard deviation of the net IRF ( $\sigma = 0.00294$ 

 $W/m^2/yr$ ), by summing every possible pair of LW and SW IRF trends listed in Table SA3.

837

849 SA1.4 Summary

Table SA4 summarizes the results above and additionally shows the trends and 95% confidence intervals for global-mean IRF as outlined in the main text. The 95% confidence intervals represent roughly  $\pm 2$  standard errors around the mean. To make the additional measures of uncertainty comparable, the values shown in Table SA4 are doubled from the standard deviations outlined in Tables SA1-3 and are divided by the square root of the number of samples that contributed to each uncertainty calculation (to represent of  $\pm 2$  standard errors around the mean)

IRF	Trend	95% Confidence	dR <sup>CS</sup> uncertainty	Radiative kernel	Cloud Mask
		Interval (±)		uncertainty	Uncertainty
Net	0.033	0.007	0.004	0.001	0.0015
LW	0.027	0.006	0.0035	0.0015	0.0015
SW	0.006	0.003	0.0035	0.002	0.0018

Table SA4. Linear trends and 95% confidence intervals (± value) for observed, global-mean net,
 longwave (LW) and shortwave (SW) Instantaneous Radiative Forcing diagnosed using the

859 *methods and data described in the main text as well as uncertainty (±2 standard errors) from* 

860 clear-sky TOA radiative anomalies ( $dR^{CS}$ ), radiative kernels and the cloud masking constant.

861

All uncertainties are an order of magnitude smaller than the Net and LW IRF trend and of similar magnitude to the trend in the SW. The IRF trends never cross the zero W/m<sup>2</sup>/yr threshold given the sources of uncertainty presented. Therefore, the trends are significantly positive. The largest source of uncertainty is in the linear regression itself, represented by the 95% confidence intervals, followed by uncertainty in dR<sup>CS</sup>. For the Net IRF, this is arguably to be expected, since the trend from the radiative kernel-derived radiative responses is insignificant.





874 Figure S3. Global-mean total radiative flux anomalies (black) as measured by CERES and the

875 contribution from radiative feedback processes (red). Both quantities are smoothed with a 12-

876 month moving average. The Niño3.4 Index (NOAA/NCEP CPC) is overlaid (blue dashed).



**Figure S2.** The total a) longwave (LW) and b) shortwave (SW) radiative response and its

883 decomposition into individual radiative responses in CERES/AIRS observations.



Figure S3. Global-mean, MERRA-2 a) net, b) longwave (LW) and c) shortwave (SW) total
radiative flux anomalies (black) from 2003 through 2018 and the contribution to that total from
the sum of radiative responses (red). Respective trendlines are displayed as dashed lines.



**Figure S4.** Same as Figure S2 but for MERRA-2.



Figure S5. Global-mean a) total shortwave instantaneous radiative forcing (SW IRF) from
MERRA-2 derived from the kernel differencing technique and b) aerosol-only SW IRF from

897 direct output of MERRA-2 radiative flux diagnostics.



**Figure S6.** Average difference (CERES/AIRS minus MERRA-2) in SW IRF from 2016 through

902 2018.





**Figure S7.** Local linear trends from 2003 through 2018 in dust aerosol optical depth from

- 909 MERRA-2 reanalysis.



Figure S8. Local linear trends from 2003 through 2018 in clear-sky surface albedo radiative
response, used in the kernel differencing method to derive shortwave instantaneous radiative
forcing (SW IRF).







929 Figure S10. Correlation of the global-mean differences in the temperature feedback between
930 CERES/AIRS and MERRA-2 with differences in temperature anomalies at each vertical level
931 and the surface.



**Figure S11.** Local linear trends from 2003 through 2018 in all-sky longwave instantaneous radiative forcing (LW IRF) diagnosed in a) CERES/AIRS observations and b) MERRA-2 reanalysis.