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1 **Individual variation in early-life telomere length and survival in a wild mammal**

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13

14 **Abstract**

15 Individual variation in survival probability due to differential responses to early-life environmental
16 conditions is important in the evolution of life-histories and senescence. A biomarker allowing
17 quantification of such individual variation, and which links early-life environmental conditions with
18 survival by providing a measure of conditions experienced, is telomere length. Here, we examined
19 telomere dynamics among 24 cohorts of European badgers (*Meles meles*). We found a complex cross-
20 sectional relationship between telomere length and age, with no apparent loss over the first 29
21 months, but with both decreases and increases in telomere length at older ages. Overall, we found low
22 within-individual consistency in telomere length across individual lifetimes. Importantly, we also
23 observed increases in telomere length within individuals, which could not be explained by
24 measurement error alone. We found no significant sex differences in telomere length, and provide
25 evidence that early-life telomere length predicts lifespan. However, while early-life telomere length

26 predicted survival to adulthood (≥ 1 year old), early-life telomere length did not predict adult survival
27 probability. Furthermore, adult telomere length did not predict survival to the subsequent year. These
28 results show that the relationship between early-life telomere length and lifespan was driven by
29 conditions in early-life, where early-life telomere length varied strongly among cohorts. Our data
30 provide evidence for associations between early-life telomere length and individual life-history, and
31 highlight the dynamics of telomere length across individual lifetimes due to individuals experiencing
32 different early-life environments.

33

34 **Keywords:** telomere length, early-life conditions, biomarker, senescence, wild population, mammal

35

36 **1. Introduction**

37 Species from most taxa exhibit a loss of performance in later-life that increases the probability of
38 mortality (Medawar 1952; Williams 1957). This process of senescence is common, but highly variable
39 across taxa (Jones *et al.* 2014) and even within species (Campbell *et al.* 2017; Dugdale *et al.* 2011;
40 Nussey *et al.* 2009). Pioneering laboratory studies using controlled environments have provided
41 important insights into senescence patterns, but cannot explain the remarkable variation in the onset
42 and rate of senescence in wild populations, where selection acts under naturally varying conditions
43 (Partridge & Gems 2007). Hence, studies of wild populations have informed understanding of how
44 early-life environments shape individual senescence patterns (Cooper & Kruuk 2018; Lemaitre *et al.*
45 2015; Nussey *et al.* 2013). This understanding has been further improved by quantification of extrinsic
46 effects through biomarkers that reflect ecological effects that are otherwise difficult to measure
47 (Bebbington *et al.* 2016; Spurgin *et al.* 2017).

48 Telomere length, which reflects the physiological consequences of within-individual
49 experiences and facilitates between-individual comparisons, is a biomarker of senescence (Monaghan
50 & Haussmann 2006). Telomeres are non-coding hexameric repeats (5'-TTAGGG-3') that, with

51 associated shelterin proteins, prevent end-to-end fusion of linear chromosomes and maintain genomic
52 integrity (Blackburn 2000; de Lange 2004). Telomeres shorten with age due to incomplete DNA-
53 replication at the 3'-end of the DNA-strand (Olovnikov 1973). This occurs more rapidly in early-life due
54 to higher levels of cellular division during growth (Frenck *et al.* 1998; Hall *et al.* 2004), or in response
55 to metabolically demanding activities (e.g. reproduction; Heidinger *et al.* 2012; coping with
56 stress/disease; Epel *et al.* 2004; Willeit *et al.* 2010). The amount of telomeric DNA lost in each cell
57 division depends on cellular conditions (Monaghan & Ozanne 2018) and oxidative stress (Reichert &
58 Stier 2017; von Zglinicki 2002; but see Boonekamp 2017). Telomeres can, however, be replenished by
59 telomerase, the telomere-elongating enzyme (Blackburn *et al.* 1989). Telomerase is transcriptionally
60 repressed later in development (Blackburn *et al.* 1989), but alternative pathways for telomere
61 lengthening do exist (Cesare & Reddel 2010; Mendez-Bermudez *et al.* 2012). Telomere shortening
62 occurs until cells enter a state of arrest, inducing replicative senescence, where the accumulation of
63 senescent cells, due to progressive loss of regenerative capacity (Campisi & di Fagagna 2007), can
64 impair tissue functioning (Armanios & Blackburn 2012; Campisi 2005).

65 Variation in the rate of telomere shortening occurs among organisms (Monaghan 2010). For
66 example, mean human leukocyte telomere length shows a biphasic decline with age, with rapid
67 shortening in early-life followed by slower attrition in adulthood (Aubert & Lansdorp 2008).
68 Correlations among within-individual telomere measurements in humans were high (0.82 – 0.93;
69 Benetos *et al.* 2013), which corroborates the high individual repeatability (i.e. 81 – 83%) in telomere
70 length in wild populations using TRF (telomere restriction fragment) methods (Bauch *et al.* 2013;
71 Boonekamp *et al.* 2014). However, longitudinal studies in wild populations using a qPCR (quantitative-
72 PCR) approach across individual lifetimes reported much lower (i.e. 7 – 13%) individual repeatability in
73 telomere length (Fairlie *et al.* 2016; Spurgin *et al.* 2017), indicating that telomeres are highly dynamic
74 over individual lifetimes. Indeed, telomere length can both decrease and increase with age (Bateson
75 & Nettle 2016), which has been attributed to measurement error (Steenstrup *et al.* 2013) but cannot

76 be explained by measurement error alone (Spurgin *et al.* 2017). Telomere length can therefore exhibit
77 complex relationships with age, explained by within-individual changes, and provide a measure of
78 conditions experienced that links to individual life-history.

79 Telomere length has been linked positively to survival to adulthood and/or annual adult
80 survival probability in both captive (Heidinger *et al.* 2012) and wild populations (Asghar *et al.* 2015b;
81 Barrett *et al.* 2013; Cram *et al.* 2017; Fairlie *et al.* 2016; Hausmann *et al.* 2005). Even though other
82 studies have tested for, but not found such associations (Beaulieu *et al.* 2011; Sudyka *et al.* 2014), a
83 meta-analysis in non-human vertebrates reported an overall association between short telomeres and
84 higher mortality risk (Wilbourn *et al.* 2018). While this provides evidence for a link between telomere
85 length and life-history, whether telomere length plays a direct causal role in senescence, because
86 telomeres are integral to organismal function, or acts as a non-causal biomarker of somatic integrity
87 remains currently unclear (Simons 2015; Young 2018).

88 Compelling evidence exists that early-life conditions such as maternal effects, developmental
89 stress and competition for resources (e.g. Asghar *et al.* 2015a; Hausmann *et al.* 2012; Cram *et al.* 2017)
90 can be particularly influential in shaping telomere length. The greater strength of early-life than late-
91 life effects could be due to stronger forces of selection, since natural selection acts on the proportion
92 of a cohort that is alive, which is greatest in early-life (Hamilton 1966). However, greater selection in
93 early-life is affected by a trade-off between parental and offspring survival (Lee 2008; Lee 2003),
94 causing the evolutionary paradigm around early-life telomere length to remain relatively poorly
95 understood (Vedder *et al.* 2017). Nevertheless, early-life telomere length might be an important
96 predictor of life-histories (Monaghan 2010; Wilbourn *et al.* 2018; Young 2018). While studies into the
97 effects of the environment on telomeres are emerging in wild mammals (Cram *et al.* 2017; Izzo *et al.*
98 2011; Lewin *et al.* 2015), longitudinal studies in wild mammals remain relatively rare (Beirne *et al.*
99 2014; Fairlie *et al.* 2016). Gaining a better understanding of telomere dynamics, its relationship with
100 survival, and early-life effects requires more comprehensive longitudinal studies in wild populations.

101 The European badger (*Meles meles*; henceforth ‘badger’) provides an informative mammalian
102 model species for studying the effects of early-life conditions on telomere length and senescence
103 patterns. We benefit here from a long-term study of badgers at Wytham Woods (Oxford, UK;
104 Macdonald *et al.* 2015); an almost closed population (see Macdonald *et al.* 2008) with a high and
105 relatively consistent annual recapture rate of 84% (SE = 1.3%; Macdonald *et al.* 2009) over 1726 life-
106 histories monitored seasonally since 1987. In this population, badgers live in polygynandrous social
107 groups (mean group size: 11.3, range: 2 – 29; da Silva *et al.* 1994; Macdonald *et al.* 2015), and show
108 reproductive senescence (Dugdale *et al.* 2011). Badgers have one litter per year (mean litter size $1.4 \pm$
109 0.06 SE; range 1 – 4; Dugdale *et al.* 2007), where cubs emerge from underground dens at 6 – 8 weeks
110 of age, are weaned at 12 weeks, and reach independence at 14 – 16 weeks old (Fell *et al.* 2006). Cub
111 survival probability ranges from 61 – 94% (mean \pm SE = $67\% \pm 3\%$; Macdonald *et al.* 2009), and cub
112 cohorts are negatively impacted by early-life exposure to endo-parasitic coccidia infection (Newman
113 *et al.* 2001), oxidative stress (Bilham *et al.* 2018) and unseasonable weather variation (Macdonald *et*
114 *al.* 2010; Noonan *et al.* 2014; Nouvellet *et al.* 2013). We therefore posit that strong selection pressures
115 on badger cubs may be reflected in their telomere length and survival probability.

116 Here, we investigate longitudinal telomere dynamics among 24 cohorts in wild badgers.
117 Relative leukocyte telomere length (RLTL) measurements were used to test: (i) age-related variation in
118 RLTL and the extent to which this was driven by within-individual changes, and both cohort and sex
119 effects; (ii) the repeatability of RLTL and whether within-individual changes in telomere length are
120 attributed to measurement error; and (iii) whether early-life and adult RLTL predict survival and
121 lifespan.

122

123 **2. Methods**

124 2.1 Study system

125 We conducted this study in Wytham Woods, Oxfordshire, UK (51°46'24"N, 1°20'04"W), a 424 ha mixed
126 semi-natural woodland site surrounded by mixed arable and permanent pasture (Macdonald &
127 Newman 2002; Macdonald *et al.* 2004; Savill 2010). The resident high-density badger population
128 (range = 20.5 – 49.5 badgers/km²; Macdonald *et al.* 2015) forms large social groups (Johnson *et al.*
129 2000). Badger social groups have clearly demarcated territories (Buesching *et al.* 2016; Delahay *et al.*
130 2000), although badgers do cross these borders when foraging and meet amicably with neighbouring
131 groups (Ellwood *et al.* 2017; Noonan *et al.* 2015). Mean annual adult survival rates in this population
132 are 0.83 (\pm 0.01 SE, Macdonald *et al.* 2009) with a mean adult lifespan of 4.96 years (\pm 3.21 SD; Bright
133 Ross, J., Pers. Comm.).

134 Trapping has been undertaken three or four times per year since 1987, for two to three
135 consecutive days per social group. Trapped badgers were anaesthetised using an intra-muscular
136 injection of 0.2 ml ketamine hydrochloride per kg body weight (McLaren *et al.* 2005) and identified by
137 a unique tattoo number on the left inguinal region. Capture date, sett, social group (comprising several
138 setts, i.e. burrow systems), sex, age-class (cub <1 year; adult \geq 1 year) and morphometric
139 measurements (i.e. length, weight, tooth wear; da Silva & Macdonald 1989; Macdonald *et al.* 2009)
140 were recorded for each badger. Badger age was defined as the number of days elapsed since the 14th
141 of February in their respective birth year (reflecting the February birth peak; Yamaguchi *et al.* 2006) .
142 Blood was collected by jugular venipuncture into vacutainers with an EDTA anticoagulant, and stored
143 at -20°C immediately. Badgers were released at their setts, after full recovery from anaesthesia.

144

145 2.2 Telomere analyses

146 We selected 1248 blood samples from 612 individuals, representing 308 males and 304 females,
147 comprising individuals varying in lifespan (range: 14 – 233 months; mean \pm SE = 97.2 \pm 1.88 months)
148 and from different cohorts (n = 24). Only badgers for which age could be determined, either trapped
149 as a cub (n = 545) or inferred through low tooth wear, were included (n = 67; males = 26, females = 41;

150 tooth wear 1 indicates a cub and tooth wear 2 indicates a 1-year old adult (da Silva & Macdonald 1989;
151 Macdonald *et al.* 2009), where young individuals also had to have length <685 mm and weight <8 kg).
152 Individuals were either sampled once ($n = 163$) or more ($n = 449$ badgers; 2 – 9 times per individual)
153 for telomere length analyses. Only badgers which were considered dead at the time of analysis were
154 included. All analyses were also run without the 67 individuals for which age was determined through
155 tooth wear, to confirm that inclusion of these samples did not bias the results (see supporting results
156 S1).

157 Genomic DNA was extracted from whole blood using the DNeasy Blood & Tissue kit (Qiagen,
158 Manchester, UK) according to the manufacturer's protocol, with adjustments using 125 μ l of
159 anticoagulated blood and a double elution step (2x 75 μ l AE buffer). DNA integrity was assessed by
160 running a random selection of DNA extracts (ca. 20%) on agarose gels to check for high molecular
161 weight. DNA concentration of all samples was quantified using the Fluostar Optima fluorometer (BMG
162 Labtech, Ortenberg, Germany) and standardized to 20 ng/ μ l, after which samples were stored at -20
163 °C.

164 Relative leukocyte telomere length (RLTL) measurements were made using the monochrome
165 multiplex qPCR method described by Cawthon (2009). This method provides a ratio of the abundance
166 of telomeric sequence to that of the control gene IRBP, the T/S ratio, analysed in the same well which
167 should reduce measurement error by excluding pipetting errors and well effects. DNA samples were
168 assayed using SYBR® Select Master Mix (Applied Biosystems, Warrington, UK) with telomere primers
169 telg (5'-ACA-CTA-AGG-TTT-GGG-TTT-GGG-TTT-GGG-TTA-GTG-T-3') and telc (5'-TGT-TAG-
170 GTA-TCC-CTA-TCC-CTA-TCC-CTA-TCC-CTA-TCC-CTA-ACA-3') at a concentration of 900 nM. A GC-clamp
171 was added to the control gene (inter-photoreceptor retinoid-binding protein; IRBP) primers to allow
172 for sufficiently different melt temperatures between the control gene and telomeric sequences, using
173 GC-clamped IRBP primers IRBP-F (5'-CGG-CGG-CGG-GCG-GCG-CGG-GCT-GGG-CGG-GCC-ACA-TTT-CTG-
174 GTA-TCC-CCT-3') and IRBP-R (5'-GCC-CGG-CCC-GCC-GCG-CCC-GTC-CCG-CCG-GGG-CGG-TCG-TAG-ATG-

175 GTA-TC-3') at a concentration of 900 nM. Subsequent melt-curve analysis confirmed differential melt-
 176 curves and lack of primer-dimer formation. Semi-skirted 96-well polypropylene qPCR plates were
 177 loaded manually with initial reaction volumes of 20 µl. Each well contained 10 µl of SYBR® Select
 178 Master Mix (Applied Biosystems, Warrington, UK), 4.9 µl of nuclease free water, 0.9 µM of both the
 179 forward and reverse primers (900 nM) and 1.5 µl of 20 ng/µl DNA sample (which was replaced with
 180 1.5 µl of nuclease free water in controls) and sealed with PCR-plate film adhesive. Cycling conditions
 181 in the Quantstudio 12K flex real-time PCR system (Applied Biosystems, Warrington, UK) were: 50°C for
 182 2 min and 95°C for 2 min, followed by 2 cycles at 94°C for 15 sec and 49°C for 15 sec, then 40 cycles at
 183 94°C for 15 sec, at 60°C for 10 sec, at 74°C for 15 sec, at 84°C for 10 sec and 86°C for 15 sec. A serially
 184 diluted (4x from 80 to 0.3125 ng/µl) 'reference' sample was included on each qPCR plate to produce a
 185 standard curve to calculate plate efficiencies, where the 20 ng/µl dilution was used as a calibrator. The
 186 reference sample was collected from a badger in 2005 and was subject to the same capture methods
 187 and long-term storage as the other samples that we analysed.

188 Samples were randomly allocated to qPCR plates and run in duplicate in adjacent wells, after
 189 which amplicon lengths and telomeric sequences were confirmed on the Agilent TapeStation 4200 and
 190 3730 DNA Analyzer (Applied Biosystems, Warrington, UK) with the Big Dye 3.1 cycle sequencing kit
 191 (Applied Biosystems, Warrington, UK). Cq-values on the 34 qPCR plates declined in a log-linear fashion
 192 ($r^2 > 0.99$). Using LinRegPCR 2017.1 (Ruijter *et al.* 2009) we corrected for baseline fluorescence,
 193 determined the windows of linearity for the amplification curves (0.432 for IRBP and 0.694 for
 194 telomeres) and calculated efficiencies and Cq-values for each well. Reaction efficiencies were (mean ±
 195 SE) 1.793 ± 0.004 for IRBP and 1.909 ± 0.004 for telomeres, and we calculated RLTL according to Pfaffl
 196 (2001):

$$197 \quad RLTL = \frac{(E_{tel}^{(Cq_{tel(calibrator)} - Cq_{tel(sample)})})}{(E_{IRBP}^{(Cq_{IRBP(calibrator)} - Cq_{IRBP(sample)})})}$$

198 where E_{tel} and E_{IRBP} represent the mean well efficiencies for each of the amplicons, $Cq_{tel(calibrator)}$ and
199 $Cq_{IRBP(calibrator)}$ are the mean Cq-values for the calibrators (20 ng/μl) for each amplicon and $Cq_{tel(sample)}$ and
200 $Cq_{IRBP(sample)}$ are the mean Cq-values for both amplicons in each sample.

201 Inter-plate repeatability (intraclass correlation coefficient), calculated with rptR 0.9.2 (Stoffel
202 *et al.* 2017), was calculated with the reference sample by comparing variance among duplicates of the
203 reference sample within a plate, to variance of the reference sample among plates and estimated at
204 0.82 (95% CI = 0.76 – 0.87; $n = 142$ samples; 34 plates). Intra-plate repeatability was calculated with
205 duplicates of the same sample on the same plate, while controlling for plate effects, and estimated at
206 0.90 (95% CI = 0.86 – 0.93; $n = 1248$ samples; 34 plates) for IRBP, 0.84 (95% CI = 0.79 – 0.90; $n = 1248$
207 samples; 34 plates) for telomere Cq-values and 0.87 (95% CI = 0.82 – 0.91; $n = 1248$ samples; 34 plates)
208 for RLTL measurements (for further details on quality control see supporting methods).

209

210 2.3 Statistical analyses

211 Statistical analyses were conducted in R 3.3.1 (R Development Core Team 2019), with RLTL
212 measurements square-root transformed to meet the assumptions of Gaussian error distributions in
213 models with RLTL as the response variable.

214

215 2.3.1 Age, sex and cohort effects on telomere length

216 We assessed the relationship between RLTL and age (months), and the interaction with cohort,
217 following Fairlie *et al.* (2016) and Spurgin *et al.* (2017). We tested a variety of age functions in General
218 Linear Mixed Models (GLMMs; Bates *et al.* 2015) that included individual ID, plate ID and year as
219 random effects, and sex, sample storage time (months), and in some models cohort, as fixed effects.
220 We checked for collinearity and found that sample storage time and cohort were collinear (VIF>3),
221 since sample storage time is similar within cohorts. We therefore first determined that sample storage
222 time was not associated with telomere length ($\beta = -0.006 \pm 0.010$ SE, $X^2 = 0.383$, d.f. = 1, $P = 0.536$) and

223 then excluded it from subsequent models. We considered a null model (without the age terms),
224 polynomial age terms (linear, quadratic, cubic), a full-factorial age term and a variety of threshold
225 functions. Visual inspection of the data indicated inflection points, with further specification of
226 inflection points through comparison of AIC values, at 29, 65 and 112 months of age. These threshold
227 models (with either a single, double or triple threshold) were compared to all other models. We ran
228 additional models to test whether adding a cohort fixed effect and an interaction between age and
229 cohort improved the model, using AIC values. We did not fully apply model selection or averaging, as
230 we aimed to compare a set of specifically defined models, where the model with the lowest AIC fits
231 these data best, but we considered all plausible models with $\Delta\text{AIC} < 7$.

232 We then tested age-specific sex differences in telomere length through an interaction between
233 age and sex in the best fitting age model and all non-significant interactions were dropped. In the same
234 model we included age at last capture (α_i), as a measure of lifespan (van de Pol & Verhulst 2006), to
235 test if selective disappearance of individuals contributed to the age pattern observed. We also
236 compared, in the same model, within-individual (β_w) to between-individual (β_B) slopes, where the
237 difference between these slopes is exactly the effect of selective disappearance (van de Pol & Verhulst
238 2006). In a separate model we tested the significance of the between-individual component by
239 replacing age parameters by within-group deviation scores (age - α_i).

240

241 2.3.2 Individual repeatability and telomere elongation

242 Individual repeatability (across multiple samples from the same individual) was calculated by dividing
243 the variance explained by individual identity by total phenotypic variance, in a Gaussian-distributed
244 model (identity link function), across all samples ($n = 1248$) and only for adult samples ($n = 779$). These
245 models included RLTL as the response variable and the best fitting age variable and cohort as fixed
246 effects, with individual ID and qPCR-plate as random effects. The variance explained by qPCR-plate was
247 then excluded from the total phenotypic variance as it is a source of experimental measurement error

248 and therefore not biologically relevant phenotypic variance; thus, it could lead to underestimation of
249 repeatability (Dochtermann *et al.* 2015). Additionally, we determined the correlation between within-
250 individual telomere measurements, using the marginal R^2 (Nakagawa & Schielzeth 2013), in a Gaussian-
251 distributed model (identity link function) with RLTL as the response variable, RLTL at $t+1$, cohort and
252 age (months) as fixed effects and individual ID as a random effect.

253 We examined increases in RLTL with age by estimating differences in telomere lengths among
254 technical replicates, i.e. duplicates next to each other within a qPCR-plate, and among within-individual
255 samples, i.e. difference in RLTL between within-individual samples. We used MCMCglmm (Hadfield
256 2010) with an inverse Wishart prior ($\nu = 1$, $\nu = 0.002$), 600,000 iterations, a thinning of 300 and burn-
257 in period of 15,000 iterations, to test whether within-individual changes in RLTL were greater than
258 measurement error. We randomly selected two samples per individual, and built a model with
259 telomere length as the response variable and individual ID and qPCR-plate as random effects ($n = 898$
260 samples; 449 individuals). We then randomly selected one set of duplicates per individual, and
261 constructed a model with telomere length for each of the technical replicates as the response variable
262 and individual ID as a random effect ($n = 898$ samples; 449 individuals). We compared the explained
263 variance by the random effect for individual ID between these two models and whether the 95%
264 credible intervals overlapped. Additionally, we separated the dataset into groups that either increased
265 or decreased in RLTL and ran these models again for these groups separately. We also tested if the
266 residual error variance ($\bar{\sigma}_{\epsilon}^2$) was smaller than the error variance in RLTL, when RLTL can increase or
267 decrease ($\sigma_{\epsilon}^{\prime 2}$), following Simons *et al.* (2014), which would reject the hypothesis that RLTL shows no
268 elongation.

269

270 2.3.3 Telomere length, survival and lifespan

271 We used GLMMs to test the relationship between early-life RLTL (<1 year old) and lifespan ($n = 435$).

272 In the following models, we conducted model averaging, using an information theoretic approach to

273 select plausible models and estimate the relative importance of fixed effects for models with $\Delta\text{AIC} < 7$
274 with the “natural average method” (Burnham *et al.* 2011). All four models included sex as a fixed factor,
275 and plate and natal social group as random effects. Early-life RLTL did not vary with age ($n = 435$, β
276 $= -0.002 \pm 0.006$ SE, $\chi^2 = 0.160$, d.f. = 1, $P = 0.690$); therefore, age was not included in GLMMs with
277 early-life RLTL as a fixed effect. Firstly, early-life RLTL as a predictor of lifespan was modelled with
278 lifespan as the response variable ($n = 435$), including early-life RLTL and cohort as additional fixed
279 effects in a Poisson-distributed model (log link function). We also controlled for overdispersion by
280 including observation (for each unique measure) as a random effect (Harrison 2014). Lifespan was
281 determined as the age at last capture. To ensure the different survival probabilities for cubs and adults
282 did not alter the results we also ran a model (see Table S1) with lifespan calculated in months as the
283 difference between the date of birth and last capture, with 24 months added when last captured as
284 adults, due to a 95% recapture interval of 2 years (Dugdale *et al.* 2007), and 12 months as cub due to
285 their different survival rates (Macdonald *et al.* 2009). Secondly, we modelled survival to adulthood (≥ 1
286 year old) using a binary term in a binomial (logit link function) mixed-effects model with early-life RLTL
287 ($n = 435$) and cohort as additional fixed effects. Thirdly, we used a Cox mixed-effects model to test
288 whether early-life RLTL predicts annual adult survival probability over the lifetime of individuals that
289 survived their first year. The model included early-life RLTL ($n = 336$) as an additional fixed effect, and
290 cohort as an additional random effect. Finally, we tested the relationship between adult RLTL ($n = 779$)
291 and survival to the subsequent year, in a binomially-distributed model (logit link function) with RLTL
292 interacting with age (based on the best fitting model) as an additional fixed effect and individual ID
293 (correcting for multiple measures per individual), cohort, current social group and year as additional
294 random effects.

295

296 **3. Results**

297 **3.1 Age, sex and cohort effects on telomere length**

298 Across all samples, after no change up to and including 29 months of age, RLTL increased up to and
299 including 65 months, followed by a decline up to and including 112 months, with a second increase in
300 RLTL in older age (Table 1; Figure 1). Two models had $\Delta\text{AIC} < 7$, with the top model including all
301 thresholds, and the second-best model with thresholds at 65 and 112 months, where both models
302 included a fixed factor for cohort (Table S2 and Figure S1). Males and females had similar telomere
303 lengths (Table 1) and there was no evidence for different age patterns by sex. Cohorts from earlier
304 years (1987 – 1992) had lower and more variable early-life RLTL measurements than those from
305 subsequent years (Figure 2a). We thus repeated these analyses where these cohorts were omitted,
306 which showed that these cohorts did not alter the results (see supporting results S2).

307 Selective disappearance of individuals was accounted for by including age at last capture (β_s)
308 in the best fitting age model, which was borderline significant (Table 1). However, there was a
309 between-individual effect (β_B) and a within-individual effect (β_w) for individuals aged 29 months or
310 older, where the difference between these slopes is due to selective disappearance of individuals with
311 shorter telomeres (Table 1). Consequently, selective disappearance of individuals with shorter
312 telomeres did contribute to the age pattern observed.

313

314 3.2 Individual repeatability and telomere elongation

315 Individual repeatability was 0.017 (95% CI = 0.001 – 0.098) including cub and adult RLTL estimates, and
316 0.026 (95% CI = 0.001 – 0.143) using only RLTL measurements from adulthood. These repeatabilities
317 changed to 0.022 (95% CI = 0.001 – 0.103) and 0.039 (95% CI = 0.001 – 0.154), respectively, when plate
318 variance (measurement error) was removed from the phenotypic variances, so 2.2% of the variance in
319 RLTL was explained by within-individual consistency among samples. There was no significant
320 correlation between RLTL measured at different time points in the same individual (marginal $R^2 =$
321 0.067; $X^2 = 0.92$, $P = 0.336$; Figure 2b).

322 Increases (in the range of 0.004 – 5.829% per month) in RLTL were identified in 61.2% of within-
323 individual changes (Figure 2c) for individuals with ≥ 2 samples ($n = 449$). When accounting for plate
324 effects using MCMCglmm, the random effect estimate for individual ID with technical replicates was
325 0.0331 (95% CI = 0.0290 – 0.0376), whereas for within-individual samples the random effect estimate
326 was 0.0014 (95% CI = 0.0003 – 0.0044; Figure 2d). For the group that exhibited increases in RLTL the
327 random effect estimate for individual ID with technical replicates was 0.0345 (95% CI = 0.0289 –
328 0.0424), whereas for within-individual samples this estimate was 0.0016 (95% CI = 0.0003 – 0.0058).
329 The random effect estimate for technical replicates in the group that exhibited decreases in RLTL was
330 0.0359 (95% CI = 0.0310 – 0.0452) and for within-individual samples this estimate was 0.0006 (95% CI
331 = 0.0003 – 0.0045), where none of the 95% credible intervals from the technical replicates and within-
332 individual samples overlapped. Additionally, residual variance among samples was smaller ($\bar{\sigma}_{\epsilon}^2 = 0.041$)
333 than the overall change in RLTL ($\sigma_{\epsilon}^{\prime 2} = 0.922$; $F_{31,40} = 22.48$, $P < 0.001$). These within-individual increases
334 in RLTL were therefore not solely due to measurement error.

335

336 3.3 Telomere length, survival and lifespan

337 Early-life RLTL (<1 year old) was positively associated with lifespan (Figure 3 and 4a; Table S3 and S4),
338 where individuals with longer telomeres in early-life had longer lifespans, such that an increase of 1
339 T/S ratio was associated with 13.3% greater longevity. However, this association was underpinned by
340 survival benefits in early-life and not in adulthood as early-life RLTL only predicted survival to
341 adulthood (Figure 5 and 4b; Table S4 and S5). In contrast, early-life RLTL showed no relationship with
342 annual adult survival probability (Table S4) and adult RLTL showed no association with survival to the
343 subsequent year (Figure 4c; Table S4 and S6), but all models indicated an effect of cohort on survival
344 and lifespan (Figure S2; Table S4).

345

346 4. Discussion

347 We found complex telomere dynamics with no apparent change (≤ 29 months of age), decreases (i.e.
348 between 65 and 112 months) and increases in RLTL with age (> 29 and ≤ 65 , and > 112 months). This
349 pattern was mainly due to within-individual changes. However, selective disappearance of individuals
350 with shorter telomeres contributed to the age pattern observed when age at last capture was included
351 (as a measure of selective disappearance) and within- and between-individual slopes were compared.
352 While the lack of change in RLTL in early-life contrasts with previous studies that have reported rapid
353 declines in RLTL with age in early-life (Aubert & Lansdorp 2008; Baerlocher *et al.* 2003), we are unable
354 to sample individuals until at least 3 months of age, due to welfare legislation (Protection of Badgers
355 Act, 1992), and therefore we may miss the period where the greatest changes in RLTL occur. The
356 combination of selective mortality and within-individual changes in RLTL was also reported in wild Soay
357 sheep (*Ovis aries*; Fairlie *et al.* 2016), providing evidence for complex relationships between telomere
358 length and age.

359 Male and female badgers had similar telomere lengths across all ages, corroborating recent
360 findings in wild meerkats (*Suricata suricatta*) and European badgers in Woodchester (Beirne *et al.*
361 2014; Cram *et al.* 2017), but contrasting with age-specific sex differences in telomere length in Soay
362 sheep (*Ovis aries*; Watson *et al.* 2017). The lack of age-specific sex differences in badgers and meerkats
363 could be due to males and females having similar lifespans, whereas in Soay sheep females live much
364 longer than males (Cram *et al.* 2017; Fairlie *et al.* 2016; Macdonald & Newman 2002).

365 Individual repeatability in RLTL was only 2.2% throughout an individual's lifespan. The point
366 estimate was higher (3.9%) when only including RLTL measurements in adulthood, but the 95%
367 confidence intervals overlapped greatly, and within-individual RLTL measurements were not
368 correlated. Within-individual RLTL correlations in humans were high (0.82 – 0.93; Benetos *et al.* 2013)
369 and individual repeatability in RLTL in avian TRF studies was also high (81% – 83%; Bauch *et al.* 2013;
370 Boonekamp *et al.* 2014). In contrast, lifelong qPCR studies in wild populations provide substantially
371 lower repeatability estimates (7%, Spurgin *et al.* 2017; 13%, Fairlie *et al.* 2016). The individual

372 repeatability estimate in RLTL in our system is in the lower spectrum of qPCR-studies. Such a low
373 individual repeatability indicates that the within-individual slopes in RLTL across ages are different.
374 RLTL is therefore highly variable within individuals across their lifetimes, where positive within-
375 individual changes indicate some active process in increasing telomere length.

376 Telomere elongation, particularly in qPCR-based studies, is often attributed to measurement
377 error (Steenstrup *et al.* 2013; Verhulst *et al.* 2015). It is, however, becoming more apparent in wild
378 population studies that telomeres do elongate (Fairlie *et al.* 2016; Hoelzl *et al.* 2016a; Hoelzl *et al.*
379 2016b; Kotrschal *et al.* 2007; Spurgin *et al.* 2017). Our study supports this, using monochrome
380 multiplex qPCR that, in principle, reduces measurement error due to reactions occurring in the same
381 well. Additionally, we found that residual variance among samples was smaller than the overall change
382 in RLTL, and variance among technical replicates was smaller than among-sample variation, indicating
383 that increases in mean telomere length with age were not due to measurement error alone.

384 Aside from actual telomere elongation, however, we acknowledge the potential for competing
385 mechanisms that could alter mean RLTL, notably changes in leukocyte cell composition with age
386 (Kimura *et al.* 2010; Linton & Dorshkind 2004; Pawelec *et al.* 2010; Weng 2012). Mammalian leukocytes
387 are nucleated and different leukocyte cell types have different telomere lengths due to their respective
388 functional capacities to proliferate and express telomerase (Aubert & Lansdorp 2008; Weng 2001), and
389 these vary in ratio over time with health/immune status (see Davis *et al.* 2008). For instance, an innate
390 immune response can cause a granulocyte-biased leukocyte ratio, where in humans and baboons the
391 granulocytes have longer telomeres than lymphocytes (Baerlocher *et al.* 2007; Kimura *et al.* 2010).
392 While a previous study of RLTL in wild Soay sheep did not find changes in leukocyte cell composition
393 with age (Watson *et al.* 2017), leukocyte cell composition in badgers does vary between similar aged
394 cubs and across an individual's lifespan due to changes in immune system activation (Montes 2007). A
395 greater metabolic rate while clearing infection could also modify leukocyte cell composition and
396 potentially affect mean RLTL directly. For instance, badger cubs are typically infected with coccidia

397 (Newman *et al.* 2001), causing a strong innate immune response and oxidative stress (Bilham *et al.*
398 2018; Bilham *et al.* 2013). A change in an individual's immunological status, along with age, may
399 therefore alter individual leukocyte cell composition and might contribute to RLTL elongation in this
400 study.

401 Our study shows a positive relationship between early-life RLTL and lifespan, driven by survival
402 benefits of long telomeres in early-life, rather than in adulthood. This is congruent with previous
403 studies reporting that early-life RLTL predicts lifespan more strongly than RLTL in adulthood (Fairlie *et*
404 *al.* 2016; Heidinger *et al.* 2012) and where early-life RLTL predicts survival to adulthood in non-human
405 mammals (Cram *et al.* 2017; Fairlie *et al.* 2016). Early-life RLTL in badgers does predict survival to
406 adulthood, but not adult survival probability. Cubs have higher mortality rates than adults (Macdonald
407 *et al.* 2009), which could drive this association between early-life RLTL and lifespan. In contrast, adult
408 RLTL in badgers did not predict survival to the following year, whereas other studies found that adult
409 RLTL does predict survival to the next year (e.g. Barrett *et al.* 2013). The lack of such an association in
410 our study system could be due to, for example, most of our RLTL measurements in later adulthood (≥ 2
411 years) being from long-lived individuals, indicating a sampling bias with fewer samples in later
412 adulthood from individuals with shorter lifespans. The interplay between adult RLTL and the adult
413 environment, or in combination with the early-life environment, also requires understanding to
414 explain the link between adult RLTL and adult survival to the next year. Even though early-life RLTL
415 predicts survival probability in badgers, it remains currently unclear how RLTL and life-history are
416 linked (Simons 2015; Young 2018). A direct link might exist through delayed cellular senescence when
417 telomeres are longer (von Zglinicki *et al.* 2001). However, an indirect link exists when telomeres
418 function as a biomarker of somatic redundancy and reflect the accumulated damage to other biological
419 structures that have deleterious effects on fitness (Boonekamp *et al.* 2013; Young 2018).

420 The early-life environment clearly exerted a strong effect on early-life RLTL, apparent from the
421 pronounced variation in early-life RLTL we noted among cohorts, which corroborates the variation in

422 survival rate and lifespan among cohorts in our study system (Macdonald & Newman 2002; Macdonald
423 *et al.* 2010). Badgers in our study are exposed to variable environmental conditions and have a limited
424 tolerance for, for example, cohort-specific weather conditions (i.e. higher cub recruitment and survival
425 probability with intermediate levels of rainfall and restricted deviation from the mean temperature;
426 Nouvellet *et al.* 2013; Macdonald *et al.* 2010) and exposure to diseases (i.e. lower cub survival
427 probability with higher intensities of coccidia; Newman *et al.* 2001). These variable environmental
428 conditions may be reflected in the variation in early-life telomere length seen in our study system.
429 Similarly, previous studies in birds have shown that higher levels of early-life competition can
430 accelerate telomere shortening (Boonekamp *et al.* 2014; Nettle *et al.* 2015), although studies that do
431 not find stressors affecting early-life telomere length do exist (reviewed in Vedder *et al.* 2017). In
432 mammals, studies on social and ecological effects on telomere dynamics are emerging (Cram *et al.*
433 2017; Izzo *et al.* 2011; Lewin *et al.* 2015; Watson *et al.* 2017; Wilbourn *et al.* 2017), showing that, for
434 example, socially dominant spotted hyaenas (*Crocuta crocuta*) have longer telomeres (Lewin *et al.*
435 2015) and that meerkat pups experiencing more intense early-life competition have shorter telomeres
436 (Cram *et al.* 2017).

437 As well as environmental effects, variation in early-life RLTL can also be caused by additive
438 genetic effects (Dugdale & Richardson 2018). In wild populations, using a quantitative genetic ‘animal
439 model’, no heritability of telomere length was found in white-throated dippers (*Cinclus cinclus*; Becker
440 *et al.* 2015), and high heritability (0.35 – 0.48) was found in the great reed warbler (*Acrocephalus*
441 *arundinaceus*; Asghar *et al.* 2015). Even though we currently have no heritability estimates from wild
442 mammals, the likelihood for additive genetic effects in our study system to contribute to early-life RLTL
443 is small given that individual repeatability, which sets the upper limit for heritability (unless indirect
444 genetic effects occur), in RLTL is low. This indicates that the individual variation in RLTL in our study
445 system is likely driven by early-life environmental conditions.

446 Our findings demonstrate that telomeres reflect the effects of early-life conditions on
447 individual life-history, and elaborate on the dynamic way that telomeres function as a biomarker of
448 senescence in a wild mammal, where within-individual telomere length is highly variable. Further work
449 on how specific early-life environment conditions impact telomere lengths in wild mammals and
450 quantifying the relative contribution of environmental effects (e.g. cohort, year and social group) on
451 telomere length will provide insight into the evolution of senescence.

452

453 **Ethics**

454 All work was approved by the University of Oxford's Animal Welfare and Ethical Review Board, ratified
455 by the University of Leeds, and carried out under Natural England Licenses, currently 2017-27589-SCI-
456 SCI and Home Office Licence (Animals, Scientific Procedures, Act, 1986) PPL: 30/3379.

457

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469

470 **Authors' contributions**

471 The study was conceived by S.H.J.v.L., A.B. and H.L.D., and developed by C.N., C.D.B. and D.W.M.;
472 Samples were collected by S.H.J.v.L., C.N., C.D.B., D.W.M. and H.L.D.; S.H.J.v.L. conducted laboratory
473 work and statistical analyses with input from H.L.D.; the paper was written by S.H.J.v.L. and H.L.D. and
474 all authors critiqued the output for important intellectual content. All authors gave final approval for
475 publication.

476

477 **Data Accessibility**

478 Data available from the Dryad Digital Repository: <https://doi.org/10.5061/dryad.64hm348>.

479

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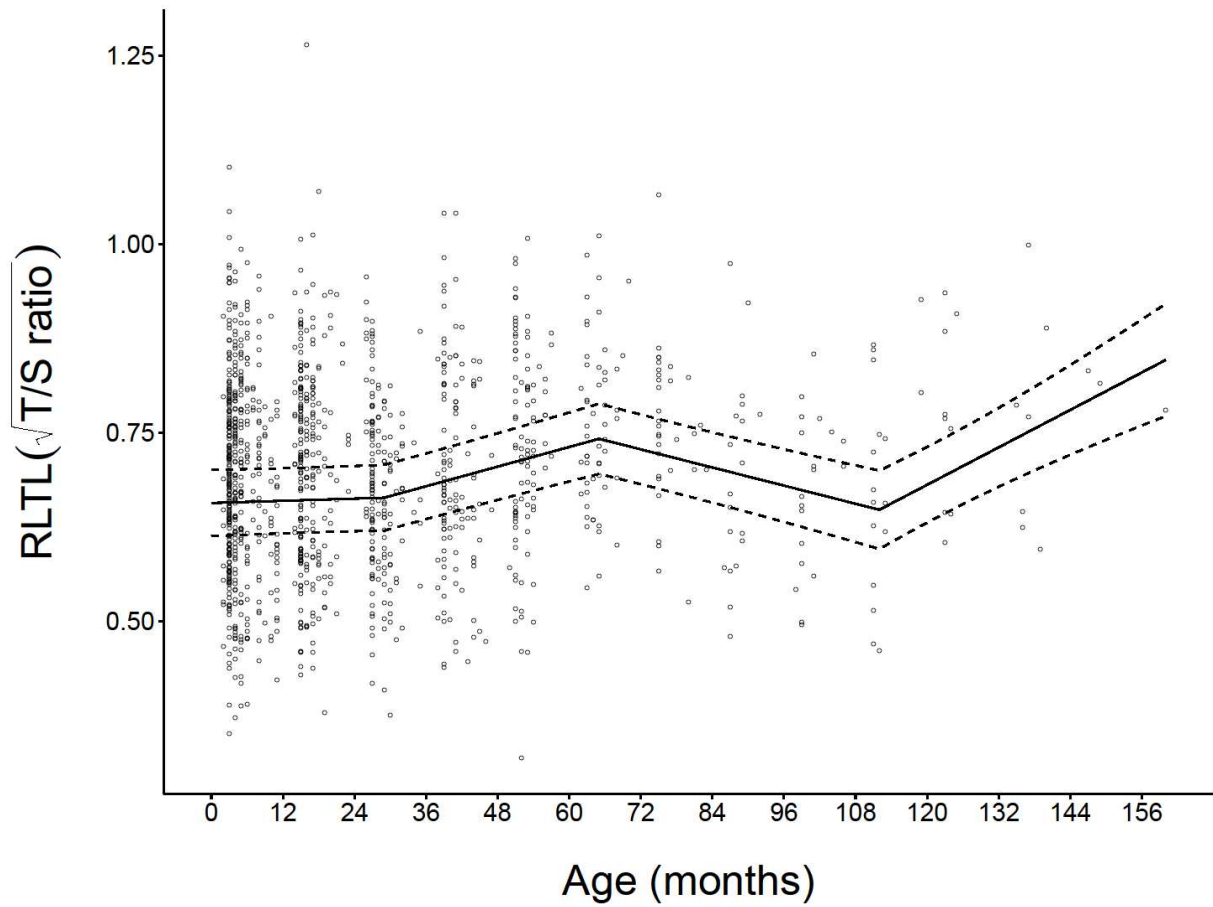
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801 **Figures & tables**

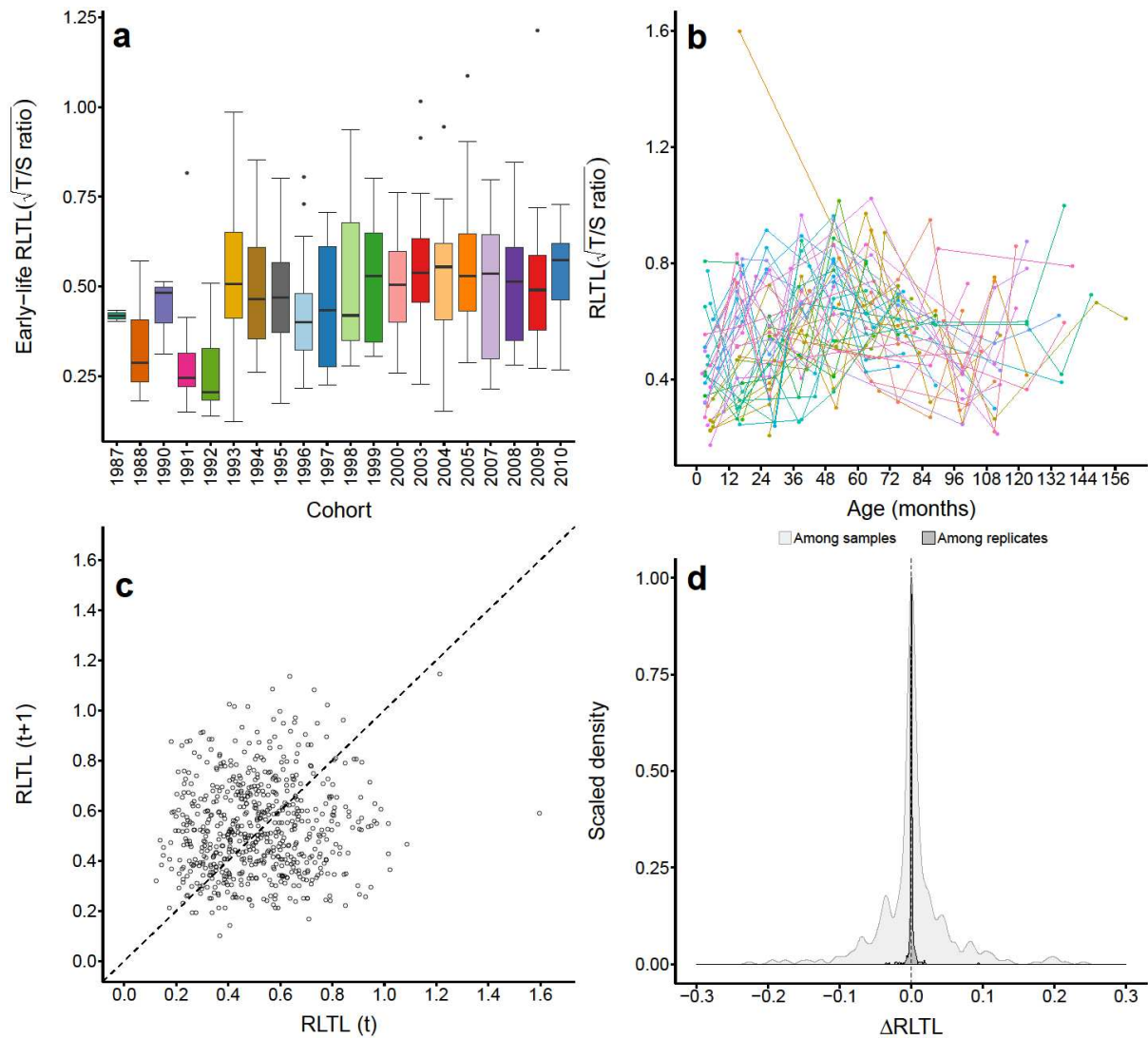
802 **Table 1:** Parameter estimates from the models that best explained the relationship between telomere
 803 length and age, when accounting for selective disappearance ($n = 1248$ samples; 612 individuals). β_w
 804 = within-individual slope, β_s = selective disappearance according to age at last capture, β_B = between-
 805 individual slope, α_i = between-individual component, S.E. = standard error, d.f. = degrees of freedom.
 806 P -values from log-likelihood ratio tests, where significant parameters are in bold.

Parameters	β	S.E.	d.f.	P-value	$\beta_B (\beta_s + \beta_w)$
Model 1 [†] :					
Intercept	0.6259	0.0527			
Age (≤ 29 months) (β_w)	0.000029	0.00054	1	0.958	0.000199
(>29, ≤ 65 months) (β_w)	0.002130	0.00051	1	<0.001	0.002301
(>65, ≤ 112 months) (β_w)	-0.00210	0.00063	1	<0.001	-0.001924
(> 112 months) (β_w)	0.004008	0.00143	1	0.005	0.004179
Sex (male)	0.008045	0.00687	1	0.242	
Cohort[§]			23	<0.001	
Lifespan (β_s)	0.000171	0.000093	1	0.068	
Model 2 [†] :					
Intercept	0.6259	0.0527			
Age (≤ 29 months) (β_w)	0.000029	0.00054	1	0.958	
(>29, ≤ 65 months) (β_w)	0.002130	0.00051	1	<0.001	
(>65, ≤ 112 months) (β_w)	-0.00210	0.00063	1	<0.001	
(> 112 months) (β_w)	0.004008	0.00143	1	0.005	
Sex (male)	0.008045	0.00687	1	0.242	
Cohort[§]			23	<0.001	
$\alpha_i (\beta_B)$	0.004242	0.00138	1	0.004	

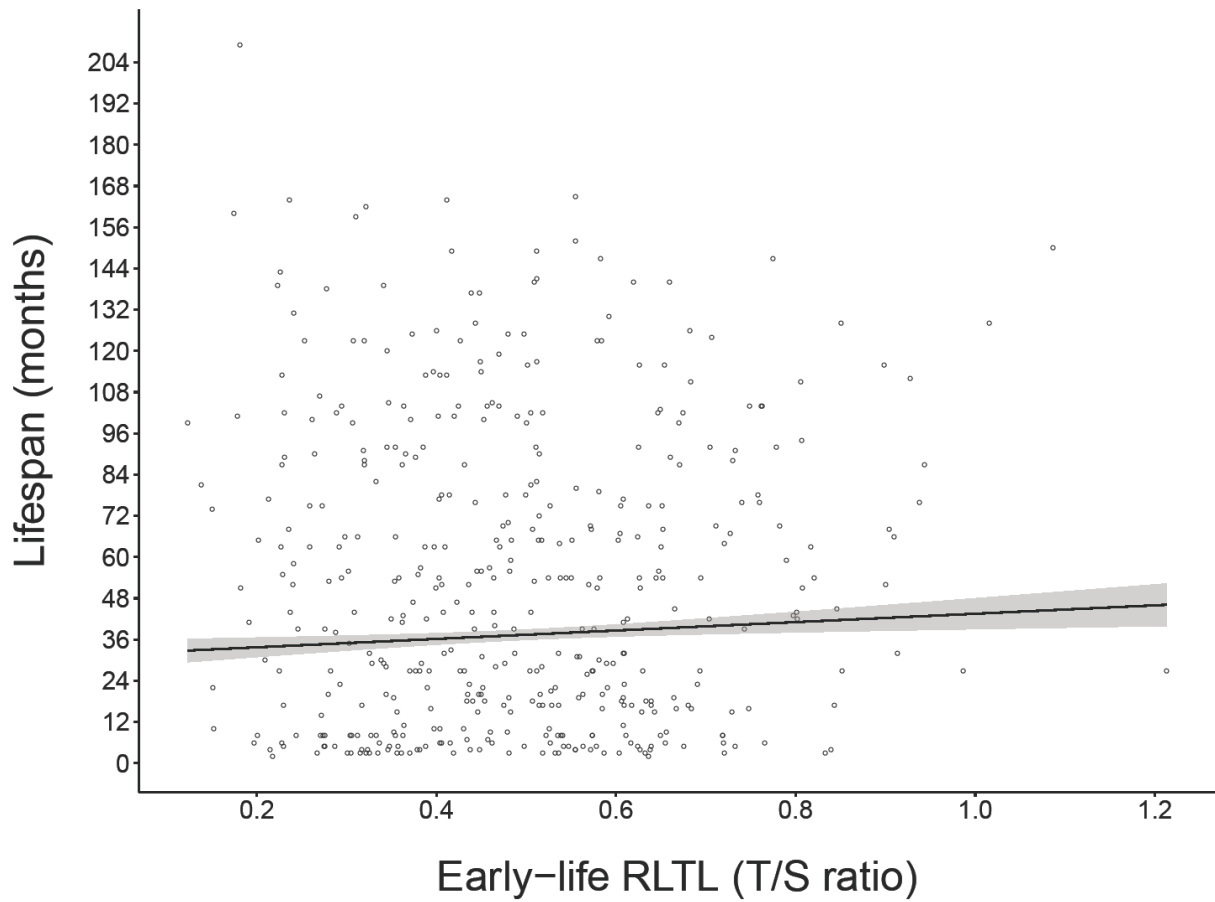
807 Random effect estimates (variance): [†]Individual ID (4.851×10^{-5}), Plate (1.067×10^{-3}), Social group (6.062×10^{-5}), Year
 808 (3.731×10^{-3}), Residual (1.295×10^{-2}); [§]Estimates \pm S.E. for 24 cohorts are in the supporting information (Figure S1).



809
 810 **Figure 1:** Age-related variation in relative leukocyte telomere length (RLTL), with inflection points at
 811 29, 65 and 112 months of age. Raw data points ($n = 1,248$) are shown with fitted lines representing the
 812 model prediction for RLTL (T/S ratio) with 95% confidence intervals.
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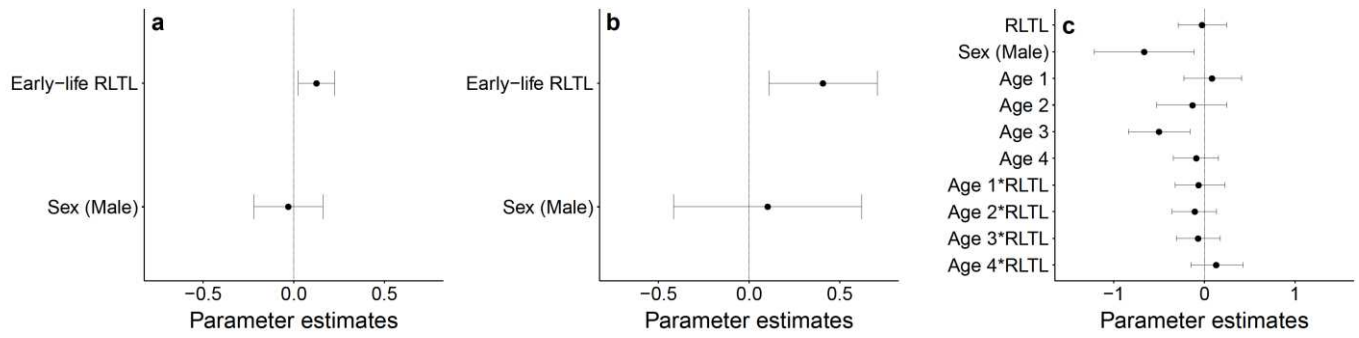


814
 815 **Figure 2:** Telomere dynamics in European badgers. a) Variation in early-life relative leukocyte telomere
 816 length (RLTL) among cohorts. b) Longitudinal telomere dynamics for 41 individuals that were measured
 817 at least four times. c) Within-individual variation in RLTL over consecutive time points (t and $t+1$).
 818 Dashed line represents parity, thus data points above and below this line represent increases and
 819 decreases in telomere length, respectively. d) Scaled density plots of changes in RLTL among technical
 820 replicates (dark grey) and among individual samples (light grey) with a dotted line representing no
 821 change. Areas left of the dotted line represent decreases in RLTL, while to the right represent increases.



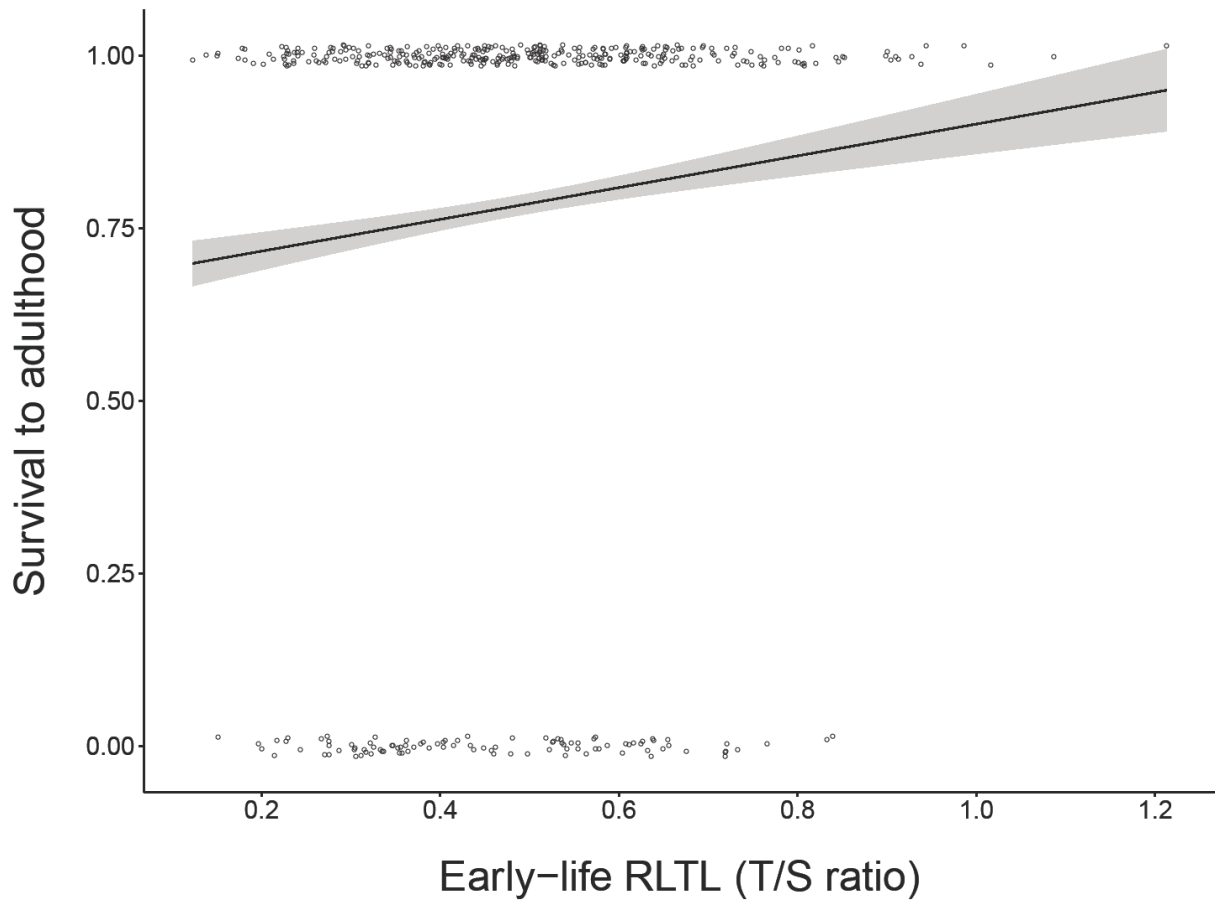
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Figure 3: Early-life (<1 year old) relative leukocyte telomere length (RLTL) predicts lifespan. Raw data ($n = 435$) are shown as open circles, the regression from the GLMM as a black line, and the 95% confidence interval as the shaded area.



827
828

829 **Figure 4:** Parameter estimates and 95% confidence intervals of fixed effects from models investigating
 830 the effect of: a) Early-life RLTL (relative leukocyte telomere length) on lifespan; b) Early-life RLTL on
 831 survival to adulthood; and, c) Adult RLTL on survival to the next year. Age parameters in plot c) refer
 832 to threshold model where Age 1 ≤ 29 months old, Age 2 > 29 and ≤ 65 months old, Age 3 > 65 and ≤ 112
 833 months old and Age 4 > 112 months old. Scale differs in plot c). For cohort effects see Figure S2. *
 834 represents an interaction.



835
836 **Figure 5:** Early-life (<1 year old) relative leukocyte telomere length (RLTL) predicts survival to adulthood
837 (>1 year old). The regression line from a binomial GLMM is shown, with associated 95% confidence
838 interval as a shaded area, and raw jittered data as open circles ($n = 435$).