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1 **Modelling long-distance route choice using mobile phone call detail**
2 **record data: A case study of Senegal**

3 The growing mobile phone penetration rates (world-wide) have led to the emergence of
4 large scale call detail records (CDRs) that could serve as a low-cost data source for travel
5 behaviour modelling compared to the commonly used data sources such as stated
6 preference (SP) data. However, to the best of our knowledge, there is no previous study
7 specifically evaluating the potential of CDR data in the context of route choice behaviour
8 modelling. Being event-driven, the data is discontinuous and only able to yield partial
9 trajectories, thus presenting serious challenges for route identification, especially in
10 highly overlapping networks. This paper proposes techniques for inferring the users'
11 chosen routes or subsets of their likely routes from partial CDR trajectories, as well as
12 data fusion with external sources of information such as route costs and travel times, and
13 then adapts the broad choice framework to the current modelling scenario where the road
14 network is highly overlapped. The model results show that CDR data can capture the
15 expected sensitivities towards route attributes and the behaviour towards overlapping
16 routes. The value of travel time estimates derived from the models are found to be
17 realistic for the study area (Senegal). These findings are timely for developing countries
18 with low budgets for transport studies.

19 **Keywords:** Route choice behaviour, Broad choice, Mobile phone data, Call detail
20 records, Value of travel time

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1 **1. Introduction**

2 The modelling of route choice behaviour for long journeys, inter-city and inter-
3 regional trips is an important component of large-scale transport planning models. The
4 emergence of technologies that enable passive collection of mobility trajectories has led
5 to a step change in route choice modelling. Most of these studies are based on Global
6 Positioning System (GPS) data, primarily from navigational devices (Li et al., 2018,
7 Hess et al., 2015, Broach et al., 2012, Bierlaire and Frejinger, 2008), and more recently
8 from smartphones (Bierlaire et al., 2010, Papinski et al., 2009). The navigational
9 devices are still not used widely in all parts of the world (in countries of the Global
10 South in particular). Although smartphone ownership is on the rise throughout the
11 world, the quality of data from these devices strongly rely on the availability of the
12 Assisted-GPS (A-GPS)¹ feature, internet connectivity and data storage capacity, which
13 leads to small sample sizes as seen in most related studies (e.g. Nitsche et al., 2014,
14 Bierlaire et al., 2013, Nitsche et al., 2012, Bierlaire et al., 2010). Thus both sources pose
15 a substantial risk of sampling bias.

16 This problem can be overcome by taking advantage of the large scale
17 anonymous datasets that are already being passively collected by operators for different
18 purposes, and applying these to transport studies. Such datasets have already yielded
19 promising results various mobility studies. Examples include; social media data
20 (Hawelka et al., 2014, Hasan et al., 2013), smart card data (Chakirov and Erath, 2012,
21 Agard et al., 2006), and network-generated mobile phone data such as Call Detail

1 Records (CDRs)² and Global System for Mobile communications (GSM)³ data (Çolak
2 et al., 2015, Jiang et al., 2013, Schlaich et al., 2010). However among these, network-
3 generated mobile phone data is particularly a promising source due to the high mobile
4 phone penetration rate world-wide (GSM Association, 2017).

5 A review of literature shows that there have been a few route identification
6 studies using network-generated mobile phone data (e.g. Nie et al., 2015, Leontiadis et
7 al., 2014, Hoteit et al., 2014, Schlaich et al., 2010), however, most of these studies end
8 on route identification and do not attempt to investigate the factors affecting route
9 choice behaviour. At the moment, only Schlaich (2010) combines GSM trajectories
10 with traffic state information to analyse the influence of variable message signs (VMS)
11 and other factors on route choice. However, the success of Schlaich's study could in
12 part be attributed to the use of GSM data, which is network-driven and semi-continuous
13 in nature as opposed to CDR data, which is event-driven and discontinuous, thus
14 presenting serious trajectory identification challenges. Despite these challenges, the data
15 is more readily available as it is stored longer for billing purposes as opposed to GSM
16 data, which is discarded after each location area update operation to save computer
17 memory. Thus, methods developed to harness the potential of such a widely available
18 data source can have wide benefits in both the developed and developing worlds. This
19 motivates this research where we investigate the feasibility of CDR data for modelling
20 route choice behaviour.

² CDR data reports the time stamped locations of communication events (i.e. voice calls, text messages, and data calls) as well as the details of the request (i.e. the duration and direction).

³ GSM data reports the IDs of all the GSM cells traversed by an active mobile phone at regular time intervals (irrespective of the calling or texting patterns of the users).

1 However, it is important to underscore the practical challenges that stem from
2 the use of CDR data, and how this impacts our work. Since the data is event-driven,
3 there is an increased risk of not capturing any trajectories, especially for very close O-D
4 pairs with short travel times. For this reason, we limit our scope to long-distance inter-
5 regional route choice as there is an increased possibility that a user will use his/her
6 phone at different points during the trip, thus enabling the capture of his/her partial
7 trajectory. The fact that only partial trajectories are observed means that in some cases,
8 it is not possible to precisely infer the chosen routes. In such cases, route choice is
9 observed at a broad sub-group level (e.g. northern, southern etc.), where each sub-group
10 comprises of a small set of possible routes. This prompts us to adapt the ‘Broad Choice
11 Modelling Framework’, developed in the context of vehicle type choice (Wong, 2015)
12 to route choice modelling using noisy CDR data. To the best of our knowledge, the
13 Broad Choice Modelling Framework is the only model structure that can deal with the
14 case when the choice data (dependent variable) has multiple levels of aggregation (i.e.
15 unique choice for some observations while broad sub-group for the other ones) as in this
16 case. We note that this may be problematic in dense inter-urban networks, where it
17 would be difficult to identify a small enough subset of possible routes using a few CDR
18 locations. However, with the increasing trend of mobile internet usage (Gerpott and
19 Thomas, 2014), the frequency of CDR locations is likely to improve significantly in the
20 near future, which adds further to the timeliness of the present paper.

21 Although CDR data is only able to capture the partial trajectories of frequent
22 phone users, the samples are usually large, thus increasing the possibility that they are
23 representative enough to capture rational route choice behaviour. The need to
24 investigate this assertion forms the basis of our validation exercise where we obtain
25 stable results. The developed models are used to estimate the value of travel time (VTT)

1 for Senegal (the study area), yielding reasonable estimates. The study is timely in the
2 sense that it extends the application of CDR data beyond travel pattern visualisation to
3 econometric modelling of travel behaviour. This could motivate reliable and low cost
4 policy formulation in both developed and developing countries.

5 The rest of the paper is arranged as follows; section 2 presents a review of
6 relevant literature, section 3 presents the data description, section 4 presents the data
7 processing conducted, section 5 presents the modelling framework, section 6 discusses
8 the model results, while section 7 presents the summary and conclusions.

9 **2. Literature review**

10 This section briefly reviews literature on the applications of mobile phone data in
11 transport studies, as well as different models of route choice.

12 **2.1. Previous applications of mobile phone data to transportation studies**

13 The last few decades have seen significant research effort aimed at investigating the
14 potential application of mobile phone data to various aspects of transportation studies,
15 so far yielding promising results. To date, the data has been widely applied in human
16 mobility modelling to understand individual's travel patterns (e.g. Ahas et al. 2010,
17 Ahas et al. 2015, Deville et al., 2016, Diao et al., 2016, Xu et al., 2015, Calabrese et al.,
18 2013, Isaacman et al., 2012, Shaw et al. 2016, Song et al., 2010, Yuan et al. 2012, Zhao
19 et al. 2013) and correlating it with phone usage patterns (e.g. Yuan et al. 2012), tourism
20 surveys to understand the attractiveness of different tourist sites (Ahas et al., 2008,
21 Ahas et al., 2007), urban area (land use) classification to facilitate urban planning (Pei et
22 al., 2014, Yuan and Raubal, 2012), analysing commute (Kung et al., 2014, Ahas et al.,
23 2010) and migration patterns (e.g. Wang et al. 2019), estimating trip making rates (e.g.
24 Çolak et al., 2015), developing origin-destination matrices (e.g. Çolak et al., 2015, Iqbal

1 et al., 2014, Calabrese et al., 2011, White and Wells, 2002), detecting travel modes to
2 estimate mode shares (e.g. Qu et al., 2015, Wang et al., 2010, Reddy et al., 2008), and
3 estimating traffic parameters to assess the traffic conditions of key roads (e.g. Bolla et
4 al., 2000). However, we place focus on studies related to route identification. A few of
5 these studies have used GSM data, which reports the complete mobile phone location
6 area sequences of each user, thus enabling the easy identification of routes through
7 sequence matching (e.g. Ma et al. 2013, Tettamanti et al., 2012, Schlaich et al., 2010)
8 and probabilistic methods such locality-sensitive hashing and graph clustering (e.g.
9 Görnerup, 2012). Furthermore, a growing number of studies have focussed on analysing
10 the potential of CDR data, which is more widely available compared to GSM data and
11 yet challenging to use as mentioned in the introduction section. For example, Doyle et
12 al. (2011) use the virtual cell paths technique to extract user trajectories from CDR data,
13 and generate the kernel density paths for different routes to validate their findings.
14 Saravanan et al. (2011) analyse the spatial and temporal information of CDR events
15 over a long period of time to establish the daily routines and routes of the users. Hoteit
16 et al. (2014) join subsequent triangulated CDR locations using linear, cubic, and nearest
17 - neighbour interpolation to model the potential trajectories. Leontiadis et al. (2014)
18 calculate the weights for each road segment within the cell areas linked to a user's
19 communication events and determine the shortest weighted path for a given OD pair.
20 Nie et al. (2015) mark each route with a subset of k optimal cell handover sequences
21 extracted from the full set of possible handover sequences and match these with those
22 observed in the cell phone hand over data (similar to CDR data) based on the degree of
23 similarity. In this study, we follow a slightly similar approach, however, instead of
24 using a similarity index, we pursue the idea of unique and shared location area
25 sequences in the context of broad choice modelling as explained later.

1 **2.2. Existing route choice models**

2 The vast majority of route choice models belong to the family of discrete choice models
3 (see Ben-Akiva and Lerman, 1985 for details), with the multinomial logit (MNL) model
4 (McFadden, 1974) being the most widely used. However, the MNL model is affected by
5 the irrelevance of independent alternatives (IIA) property, which can be problematic for
6 highly overlapping routes (Ramming, 2002). This has motivated the development of more
7 advanced route choice models to address this challenge. Examples include the nested
8 recursive logit model (Mai et al., 2015), the c-logit model (Cascetta et al., 1996), the path
9 size logit model (Ben-Akiva and Ramming, 1998), the link nested (cross-nested) logit
10 model (Vovsha and Bekhor, 1998), and the multinomial probit model with logit kernel
11 (Daganzo et al., 1977). Details of how some of these models overcome the overlapping
12 route problem are discussed in section 5.3 of this paper.

13 An important point to note is that the complexities of route choice modelling go
14 beyond the overlapping route problem. Choice set generation is a key challenge,
15 especially in highly overlapping dense urban networks, where several alternative routes
16 can be possible, and yet individuals do not consider all the alternatives while making
17 choices (Prato, 2009). Several choice set generation methods have been proposed in the
18 literature including, the k-shortest path algorithms (e.g. Shier, 1979, Bellman and Kalaba,
19 1960), the labelling approach (Ben-Akiva et al., 1984), link elimination approaches (e.g.
20 Azevedo et al., 1993, Bellman and Kalaba, 1960), link penalty approaches (e.g. Roupail
21 et al., 1995, De La Barra et al., 1993), simulation approaches (e.g. Sheffi and Powell,
22 1982), doubly stochastic generation functions (e.g. Nielsen, 2000), constrained
23 enumeration methods (e.g. Prato and Bekhor, 2006), and probabilistic methods (e.g.
24 Cascetta and Papola, 2001, Manski, 1977). However, since the focus of this paper is long

1 distance trips, where the alternatives are usually few in number, choice set determination
2 is more straightforward as discussed later in section 4.2 of this paper.

3 As mentioned earlier, the discontinuous nature of CDR data presents serious route
4 identification challenges, resulting in route choice observations at different levels of
5 aggregation (i.e. unique route choices for some travellers while broad sub-groups of the
6 possible routes for the rest). Although there is no specific route choice model dealing
7 with such a scenario, recent advances in other fields of choice modelling have addressed
8 this challenge using the broad choice framework (proposed by Wong, 2015 and later
9 used by Habibi et al. 2017, Lloro and Brownstone 2018, Wong et al. 2018, Yip et
10 al.2018 in the context of vehicle choice and by Tran et al. 2019 in the context of hotel
11 choice), which can be adapted to deal with the current route choice modelling scenario,
12 discussed later in this paper. It may be noted that according to the choice modelling
13 literature, this framework is the only one that can deal with the case where the choices
14 are uniquely observed only for some respondents and for the rest, only the sub-group of
15 choices are observed.

16

17 **3. Data**

18 This study uses CDR data collected from Senegal as part of the Orange Data for
19 Development (D4D) challenge (de Montjoye et al., 2014).

20 **3.1. Study area**

21 Senegal is located in West Africa with a population of approximately 13.5 million
22 according to the 2013 population census (ANSD, 2016). Road transport accounts for
23 over 99% of all passenger travel (World Bank, 2004). The only long-distance train
24 service (the Dakar-Niger line) was discontinued in May 2010 (Imedia and Calao

1 Production, 2013). The country has a sparse national road network (see Figure 1), and
2 for some O-D pairs, there is only one feasible alternative, making them unsuitable for
3 route choice modelling. This study ignores such O-D pairs and only considers those
4 where alternative routes exist. In total, twelve distant O-D pairs are considered, and
5 these are; Dakar-Bakel, Dakar-Matam, Thies-Bakel, Thies-Matam, Diourbel-Bakel,
6 Diourbel-Matam, and the corresponding O-D pairs for the reverse directions as shown
7 in Figure 1. The long travel times between these regions increase the possibility of
8 capturing the users' partial trajectories as explained earlier.

9 Figure 1. Study area

10

11 **3.2. CDR data**

12 The CDR data was collected between January and December 2013 and aggregated to
13 the arrondissement (district) level by the data provider. The geographical location of the
14 arrondissements is presented in Figure 1 where we also show the tower locations for
15 illustration purposes.

16 The original CDR data comprised of 9 million unique users (67% of the study
17 area population). This was pre-processed to retain frequent phone users (i.e. those with
18 interactions on 75% of the days in a year) and randomly split into smaller monthly
19 rolling sub-samples made available for research (see de Montjoye et al., 2014 for
20 details). The user IDs in each sub-sample are anonymised to prevent possible re-
21 identification across the different months.

22 The data for each month comprises of about 150,000 users. Assuming that the
23 most commonly observed arrondissement for each user during the month is their home
24 district, the monthly population sampling rate ranged from 2.4% in Dakar (the capital)

1 to 0.4% in the rural regions. On average, these users together generated over 40 million
2 records per month (see excerpt of the CDR data in Table 1a). The data is reduced to
3 remove duplicate records resulting in the processed arrondissement visitation data
4 presented in Table 1b.

5 Table 1a. Excerpt of the raw CDR data

6 Table 1b. Excerpt of the processed arrondissement visitation data

7

8 The overall level of user mobility is illustrated in Figure 2. As shown, most
9 users visited less than three unique arrondissements per month. The low levels of inter-
10 arrondissement mobility led to the capture of few trajectories as reflected in the final
11 sample size (see Section 4.2).

12 Figure 2. Average monthly arrondissement observation frequency distribution

13

14 Although it is difficult to detect false tower jump movements in aggregate CDR
15 data (Çolak et al., 2015, Iqbal et al., 2014), this is not a big factor for distant O-D pairs
16 where the origin and destination arrondissements have been grouped into regions.

17 **4. Data preparation for analysis**

18 This section describes the analysis carried out on the processed arrondissement
19 visitation data in Table 1b to identify the routes followed, as well as the processes of
20 estimating the route attributes.

21 **4.1. Route identification**

22 The route identification process is summarised in Figure 3. This is divided into two
23 main stages as described in the subsequent sections.

1 Figure 3. Summary of the route identification process

2

3 4.1.1. Generation of unique and shared CDR arrondissement sequences

4 Route arrondissement sequences (which are extracted from maps) show the order of all
5 the arrondissements traversed by a particular route between a given O-D pair. On the
6 other hand, CDR arrondissement sequences (which are extracted from the CDR data)
7 show the order of the arrondissements in which a user used his/her phone during the
8 trip, and are subsets of the route arrondissement sequences.

9 For any given trip along a particular route, several possible CDR arrondissement
10 sequences can be observed depending on the number and the location of the CDR
11 events. These can be obtained by generating permutations of different sizes based on the
12 route arrondissement sequence (in which order matters and no repetitions are allowed).
13 However, since most of the O-D pairs have overlapping routes, some of the CDR
14 arrondissement sequences can be linked to more than one route if all the intermediate
15 CDR events occurred along the shared sections. In this case, it would only be possible
16 to observe the subset of routes that were potentially followed (see illustration in Figure
17 4 where the blue areas indicate the shared arrondissements, while the grey and green
18 areas indicate the unique arrondissements for the Northern 1 and 2 routes respectively).

19 Figure 4. Arrondissement paths (Dakar-Bakel O-D pair as an example)

20

21 The generated CDR arrondissement sequences linked to each route were cross-
22 referenced to identify the permutations linked to unique routes (i.e. unique CDR
23 arrondissement sequences) and those shared across multiple routes (i.e. shared CDR
24 arrondissement sequences). The outcome of this analysis was a list of all the possible

1 CDR arrondissement sequences labelled with the associated routes or sub-groups of the
2 possible routes (see example in Table 2).

3 4.1.2. Extraction the O-D pair trip trajectories from CDR data

4 The processed arrondissement visitation data for each user (see excerpt in Table 1b) was
5 analysed to extract sub-sequences linked with possible trips between the regions of
6 interest following the criteria below;

- 7 • The first and the last arrondissements in the sub-sequence must be located
8 within different regions of interest, and the user must not be observed in an
9 upstream or a downstream region of interest within the same day for origins and
10 destinations respectively;
- 11 • The dwell time in the origin and the destination regions of interest must be
12 longer than that required to directly traverse each of them to increase the
13 possibility that these are the trip start and end locations. A user needs to use
14 his/her phone at least twice in each of these regions to calculate the dwell time;
- 15 • The intermediate arrondissements in the sub-sequence must all be associated
16 with one of the defined arrondissement/corridors paths (see Figure 4) to ensure
17 only direct trips are retained; and
- 18 • The timestamp difference between the origin and the destination must not
19 exceed 24 hours, which is used as an upper limit to distinguish between users
20 with direct trips but delay to use their phones on arrival, and those with
21 intermediate destinations, thereby arriving late.

22 The extracted sub-sequences meeting all the above criteria were either assigned
23 to unique routes or sub-groups of the possible routes by cross-referencing with the

1 labelled lists generated in Section 4.1.1. Table 2 presents an excerpt of the route
2 assignment data.

3 Table 2. Excerpt of the route assignment data
4

5 In this data, 70% comprised of unique assignments, while 30% comprised of
6 broad assignments. Since this is a scenario where for some users and/or trips, we know
7 the chosen route at a more disaggregate level than for others, we use the Broad Choice
8 Modelling Framework (Wong, 2015) to analyse route choice behaviour.

9 **4.2. Estimation of route attributes**

10 For model estimation, it is critical to determine the choice set and the attributes of the
11 alternatives. We assumed that the choice set is comprised of the routes that have ever
12 been chosen by the different users. Routes not chosen by any user for the whole year
13 were excluded from the choice set. On average, each origin-destination pair had four
14 alternatives. Given that the total dataset is comprised of 9,453 records from 6,497 users,
15 this is not a very restrictive assumption. Given the choice sets, we reviewed previous
16 studies to identify the attributes typically used in route choice models and their
17 availability status for Senegal as summarised in Table 3. Although data on six
18 explanatory variables was available, the final model specification contains three
19 explanatory variables only as the inclusion of the other variables led to correlation
20 problems and/or illogical model results. A detailed explanation of the variable
21 specification process is presented in Section 6.1. The subsequent sections summarise the
22 processes of estimating some of these attributes.

23 Table 3. Attributes typically used in route choice models
24

1 4.2.1. Link length and surface attributes

2 Data on the link lengths and surface attributes (i.e. paved or unpaved) was derived from
3 the Senegal roads GIS layer (ArcGIS, 2013). This was updated to reflect the situation in
4 2013 relying on road condition reports sourced from government and other relevant
5 websites (Ageroute Senegal, 2017, ANSD, 2017, Logistics Cluster, 2013).

6 4.2.2. Travel time

7 Travel time cannot be reliably estimated from the CDR data as users do not necessarily
8 use their phones at the moment of departure or arrival. The typical travel times for most
9 links in 2013 were obtained from the website of Logistics Capacity Assessment
10 (Logistics Cluster, 2013). For links not covered by this website, we relied on Google
11 Maps (Google Maps, 2017a).

12 4.2.3. Travel cost

13 Travel cost was estimated in terms of the vehicle operating costs (VOCs) per user (i.e.
14 fuel and non-fuel costs). After a review of several VOC estimation techniques, we
15 settled for the HDM-III model (Watanatada et al., 1987) due to its applicability to
16 developing countries and input data availability. The HDM-III model is an earlier
17 version of the more advanced HDM-4 (Kerali, 2000), which we could not use due to
18 input data constraints.

19 The HDM-III model relies on vehicle calibration data (where we used default
20 values) and other basic input data (see Table 4) to estimate the VOCs for each vehicle
21 type. The model works by defining link-specific relationships between the International
22 Roughness Index - IRI (Sayers et al., 1986) and speed, and using the IRI values at the
23 respective average link speeds (derived from Sections 4.2.1 and 4.2.2) to estimate the
24 link-specific VOCs, which are summed to estimate the route VOCs.

1 Table 4. HDM-III basic input data

2

3 The estimated route VOCs need to be converted to person costs. Given the
4 anonymous nature of CDR data, we use information on the typical occupancy rates and
5 mode shares (see Table 5) to estimate the weighted average VOCs per user for each
6 route.

7 In this research, we have used the same average travel cost for all the users along
8 a particular route between a given O-D pair. An improved approach would have been to
9 estimate user-specific travel costs based on the corresponding travel speeds, however,
10 this was not possible due to difficulties in obtaining the user-specific travel times as
11 discussed in the previous section.

12 Table 5. Typical occupancy rates and mode shares in Senegal (World Bank, 2004)Other
13 attributes

14 The other attributes considered were the scenic characteristics and the urban
15 developments along each route. Scenic variables were estimated in terms of route
16 lengths traversed through nature reserves (Google Maps, 2017b), while urban
17 developments were reflected as the number of towns along each route as shown in
18 Figure 1 (Worldatlas, 2017).

19 **5. Modelling framework**

20 We use discrete choice models in this study since route choices are discrete and
21 mutually exclusive. To develop these models, we apply the random utility theory
22 (Marschak, 1960), a well-established approach for estimating discrete choice models.

23 **5.1. Basic model**

24 Suppose U_{nr} is the utility of choosing route r by individual n . This can be expressed as;

$$U_{nr} = V_{nr} + \varepsilon_{nr} \quad (1)$$

Where V_{nr} and ε_{nr} are the systematic and the random parts utility respectively. The systematic utility is a function of the observed route attributes, and may be expressed as $V_{nr} = \beta' X_{nr}$, where X_{nr} is a vector of the attributes of route r for individual n and β is a vector the model parameters. We assume that the random term ε_{nr} is independently and identically distributed across the alternatives following a type I extreme value distribution, and use the Multinomial Logit (MNL) model to estimate the route choice probabilities as follows (see McFadden, 1974 for details);

$$P_n(r) = \frac{\exp(V_{nr})}{\sum_{r^* \in C_n} \exp(V_{nr^*})} \quad (2)$$

Where, $P_n(r)$ is the probability of individual n choosing route r , and C_n is the choice set. Given the route choice probabilities, the model parameters can be estimated by maximising the log-likelihood function below;

$$LL = \sum_n \sum_r [Z_{nr} \cdot \ln(P_n(r))] \quad (3)$$

Where Z_{nr} is a dummy variable, which is equal to 1 if and only if user n chooses route r .

5.2. Accounting for broad choices

The log-likelihood function in Equation (3) assumes that all the route choices are uniquely observed, and is inadequate for the current scenario, where we also have broad sub-group choices. Therefore, we use the broad choice modelling framework proposed by Wong (2015) to account for this situation.

1 In the broad choice framework, the choice probabilities of the broad sub-groups
 2 are expressed as a sum of the choice probabilities of the member alternatives. For
 3 example, the choice probability of the ‘Northern’ broad sub-group is the sum of the
 4 ‘Northern 1’ and the ‘Northern 2’ route choice probabilities (see Figure 4 and Table 2).
 5 The joint probabilities of the broad sub-groups capture the aggregate shares at the
 6 unique route choice level using the relative probabilities of the constituent routes.

7 The goal of model estimation is to maximise the probabilities of both the
 8 observed routes and the broad sub-groups for users with unique and broad choices
 9 respectively. The log-likelihood function is specified as follows (Wong, 2015);

$$10 \quad LL = \sum_n \sum_b \left[Z_{nb} \cdot \ln \left(\sum_{r \in S_b} P_n(r) \right) \right] \quad (4)$$

11 Where S_b is a set comprising of the routes in broad category b . For uniquely
 12 assigned trips, set S_b comprises of only one alternative. Z_{nb} is a dummy variable,
 13 which is equal to 1 if and only if user n is associated with category b .

14 **5.3. Accounting for overlap**

15 A major weakness of the MNL model (Equation 2) is the IIA property, which could
 16 lead to illogical route choice probabilities for highly overlapping routes as is the case in
 17 this study. This is illustrated using the overlapping route problem (Ramming, 2002,
 18 Cascetta et al., 1996) in Figure 5.

19 Figure 5. Overlapping route problem (Cascetta et al., 1996, Ramming, 2002)

20
 21 Here, all three routes have the same total length L , however, routes 1 and 2
 22 follow the same alignment for length l followed by distinct sections, each of length $L -$
 23 l . The MNL model predicts equal shares for each of the routes irrespective of the

1 overlap length. However, as the overlap increases (as l tends to L), it becomes difficult
 2 to distinguish between routes 1 and 2, and it is expected that route 3 will take a share of
 3 50%, while routes 1 and 2 will each take a share of 25%.

4 Various models accounting for overlap were presented in Section 2.2, however,
 5 this study only uses the c-logit (Ramming, 2002, Cascetta et al., 1996) and the path size
 6 logit (Ben-Akiva and Ramming, 1998) models for illustration purposes as the modelling
 7 scenario did not require complex formulations. In general, these models are modifications
 8 of the MNL model, where the systematic route utilities are adjusted using certain
 9 correction factors as follows;

$$10 \quad P_n(r) = \frac{\exp(V_{nr} + \tau_{nr})}{\sum_{r^* \in C_n} \exp(V_{nr^*} + \tau_{nr^*})} \quad (5)$$

11 Where τ_{nr} is the systematic utility correction factor for route r .

12 5.3.1. C-logit model

13 For the c-logit model, the correction factor τ_{nr} is a commonality factor (Cascetta et al.,
 14 1996). Different possible specifications have been proposed, however, this study uses
 15 the following common specification (Ramming, 2002, Cascetta et al., 1996);

$$16 \quad \tau_{nr} = \beta_{CF} \ln \left[\sum_{r^* \in C_n} \left(\frac{L_{rr^*}}{\sqrt{L_r L_{r^*}}} \right)^{\gamma_{CF}} \right] \quad (6)$$

17 Where L_{rr^*} is the overlap length between routes r and r^* , L_r and L_{r^*} are the total
 18 lengths of routes r and r^* respectively, β_{CF} and γ_{CF} are the unknown parameters to be
 19 estimated. From Equation 6, the ratio in the brackets is proportional to the degree of
 20 overlap, while the corresponding logarithm has an inverse negative relationship. Thus

1 β_{CF} and γ_{CF} are expected to have negative and positive signs respectively to allow for
2 the positive adjustment of route utility with decreasing overlap (Ramming, 2002).

3 5.3.2. Path size logit model

4 For the path size logit model, the correction factor τ_{nr} is a path size term, which is
5 computed as the weighted average of the constituent link sizes. The specification
6 adopted for this study is as follows (Ben-Akiva and Ramming, 1998);

$$7 \quad \tau_{nr} = \beta_{ps} \ln \left[\sum_{a \in \Gamma_r} \left(\frac{l_a}{L_r} * \frac{1}{N_{ar}} \right) \right] \quad (7)$$

8 Where $1/N_{ar}$ is the inverse of the number of routes sharing the link a (the link
9 size), and l_a/L_r is a weight representing the proportion contributed by link a to the
10 overall route size. From Equation 7, it is observed that route size is inversely
11 proportional to the degree of overlap, while the corresponding logarithm has a negative
12 proportional relationship. Thus, the path size parameter is expected to be positive to
13 allow for negative adjustment of route utility with increasing overlap.

14 The models accounting for overlap are generally expected to have better fit than
15 the MNL model. Model fit is evaluated using the adjusted-rho square and the likelihood
16 ratio tests (see Ben-Akiva and Lerman, 1985 for details).

17 **6. Model results**

18 This section presents the modelling results. We start by discussing the variable
19 specification, followed by the model estimation and validation results.

20 **6.1. Variable specification**

21 The attributes available for possible inclusion in the model are summarised in Table 3.

1 However, these could not all be specified together or in the same way for various
2 reasons as explained below.

3 For travel time and cost, after initial tests using the linear specification, we used
4 the log-transforms of the variables to allow for utility damping with respect to
5 increasing time and cost (see Daly, 2010 for details). Various interactions of these
6 variables with others (such as surface type) were tested, however, this led to correlation
7 problems, and hence generic variables were specified.

8 The urban developments along the alternative routes were incorporated in terms
9 of the average distance between towns rather than the number of towns to avoid
10 situations where longer routes also have more towns. Again, this was specified using the
11 log-transform of the variable for similar reasons as the travel time and cost. This being a
12 largely rural road network with no traffic signals, the average distance between towns is
13 the only variable we could use to capture traffic flow interruptions.

14 An attempt was made to incorporate scenic beauty into the model using either
15 the length or the proportion of route length traversed through nature reserves, however,
16 we could not achieve intuitive model results potentially due to our lack of detailed
17 knowledge about the characteristics of these reserves, and the security levels of the
18 corresponding routes. The final systematic utility specification is as follows;

$$19 \quad V_{nr} = \beta_{1-\text{cost}} \ln(C_{nr}) + \beta_{1-\text{time}} \ln(T_{nr}) + \beta_{1-\text{dtown}} \ln(Dt_{nr}) \quad (8)$$

20 Where C_{nr} , T_{nr} , and Dt_{nr} give the travel cost, the travel time, and the average
21 distance between towns respectively of route r for individual n , and the βs are the
22 corresponding model parameters to be estimated.

1 **6.2. Estimation results**

2 We present the estimation results of the MNL model, the c-logit model, and the path
3 size logit based on the full sample in Table 6 for comparison purposes.

4 Table 6. Estimation results

5 6.2.1. Statistical performance of the models

6 The estimated parameters in each of the three models are statically significant at the
7 95% level of confidence. The only exceptions are the parameters associated with the
8 average distance between towns and the beta coefficient of the commonality term in the
9 c-logit model, however, these were retained in the model as they are important.

10 To evaluate the collective statistical significance of the models, we used the
11 likelihood ratio test (see Ben-Akiva and Lerman, 1985 for details), whose values are
12 reported in the last two rows of Table 6. From these, it is noted that the p-values for all
13 the models are less than 0.01. Therefore, the hypothesis that the respective model
14 parameters are collectively equal to zero is rejected at the 99% level of confidence.

15 6.2.2. Route variables

16 The parameter signs for the travel cost variable are consistent with a priori expectations
17 in each of the three models. In general, an increase in the cost of an alternative is
18 expected to have a negative impact on its utility, hence the negative parameter sign. The
19 same explanation holds for the travel time parameters as individuals generally prefer
20 shorter travel times.

21 On the other hand, the average distance between towns has a positive parameter
22 sign. As earlier mentioned, this variable gives an indication of the amount of
23 uninterrupted flow. Although no traffic congestion problems have been reported in
24 these towns, traffic generally slows down due to speed control measures leading to

1 delays. An increase in the average distance between towns therefore indicates more
2 uninterrupted flow, hence the positive parameter sign.

3 6.2.3. Overlap correction parameters

4 For the c-logit model, it is observed that the beta parameter of the commonality term is
5 negative while the gamma parameter is positive. Similarly, the path size parameter in
6 the path size logit model has a positive parameter sign. These results are in line with
7 behavioural expectations as discussed earlier under Equations 6 and 7, an indication that
8 CDR data is able to capture the behaviour towards overlapping routes.

9 6.2.4. Model comparison

10 A comparison of the adjusted rho-square values in Table 6 shows that the models
11 accounting for overlap (i.e. the c-logit and the path size logit models) perform better
12 than the MNL model. This is as expected given that the national road network of
13 Senegal is highly overlapping (see Figure 1 and discussion under Figure 5). It is also
14 worth noting that the path size logit model outperforms the c-logit model because the
15 behavioural underpinning of the systematic utility adjustment process in the path size
16 logit model is stronger than that in the c-logit model (Ramming, 2002).

17 The statistical significance of the improvements associated with accounting for
18 overlap are evaluated using the likelihood ratios of the c-logit and the path size logit
19 models with respect to the MNL model (see Table 7).

20 Table 7. Statistical comparison of the models

21

22 From Table 7, it is noted that the p-values for the c-logit and the path size logit
23 models are all less than 0.01, an indication that accounting for overlap has a statistically
24 significant effect (at the 99% confidence level) beyond the improvements contributed

1 by the additional degrees of freedom resulting from the extra parameters (see Ben-
2 Akiva and Lerman, 1985 for details).

3 6.2.5. Policy insights

4 This section highlights the policy implications of the reported results in terms of the
5 value of travel time (VTT). This metric quantifies the benefits derived from reduced
6 travel time in monetary terms, and is useful in transportation cost-benefit analysis
7 (Mackie et al., 2001). The value of travel time is calculated by taking the ratio of the
8 partial derivatives of the systematic utility function (V) with respect to the travel time
9 (T_{nr}) and cost (C_{nr}) as follows;

$$10 \quad \text{VTT} = \frac{\partial V_{nr} / \partial T_{nr}}{\partial V_{nr} / \partial C_{nr}} = \frac{\beta_{1-\text{time}} C_{nr}}{\beta_{1-\text{cost}} T_{nr}} \quad (9)$$

11 Figure 6 shows the variations of VTT with respect to the cost per hour of the
12 alternatives for the range in the estimation data, and it is observed that the values
13 increase as the alternatives become more expensive, which is expected. It is noted that
14 the path size logit model gives the lowest values, which are deemed to be more accurate
15 due to the superior performance (see Tables 6 and 7).

16 Figure 6. Variations in VTT across the alternatives and models

17

18 However, to assess how realistic these estimates are, we computed the average
19 values for each model using the estimation data, and compared these values with those
20 derived from other studies and other relevant statistics as summarised in Table 8.

21 Table 8. Comparison of the VTT estimates with other sources

22

1 In the Africa-wide meta-analysis by Teye et al. (2017), VTT is estimated as a
2 function of the GDP per capita. However, the reported mean value (4.3213 USD/hr)
3 seems high when compared to the toll being charged on the new Dakar–Diamniadio toll
4 highway for a time saving of one hour (2.3411 USD/hr), a value that was highly
5 criticised by the Senegalese media as being extremely high (Gainer and Chan, 2016).

6 Although the median hourly wages do not necessarily translate into the value of
7 travel time, they give a good indication of the range in which these values should fall,
8 and as observed in Table 8, the average VTT estimate for the path size logit model is
9 very close to the Senegalese median hourly wage. We consider this VTT estimate to be
10 more reasonable for Senegal.

11 **6.3. Validation results**

12 The models based on the full sample provide intuitive results in terms of the parameter
13 signs and the relative model performance. To assess the stability and the predictive
14 performance of the models, the dataset was randomly split into five parts at the
15 individual level. Five rolling subsets, each comprising of 80% of the users were
16 generated for model estimation purposes. For each of these, a complementary subset
17 comprising of 20% of the users was generated for validation purposes. The models were
18 re-estimated on each of the 80% subsets, and the parameter estimates applied to the
19 corresponding 20% hold-out subsets to estimate the predictive measures of fit. Table 9
20 presents the summary outputs from this process.

21 Table 9. Validation results

22
23 The general interpretation of the parameter signs and the relative model
24 performance in each of the 80% estimation subsets remained the same as in the full

1 sample, an indication that the data is representative (detailed results available on
2 request). A comparison of the measures of fit in estimation and validation shows that
3 there is no significant loss in model fit, an indication that the performance of the models
4 during estimation is not due to overfitting, rather it is due to the strong explanatory
5 power of the variables.

6 A comparison of the predictive measures of fit shows that the relative model
7 performance during estimation is mirrored on the holdout samples, with the path size
8 logit model still giving the best model performance due to its behavioural superiority. It
9 would have been interesting to further validate the above results with outputs of route
10 choice models based on traditional data or GSM data, but this was not possible due to
11 lack of data in Senegal.

12 **7. Summary and conclusions**

13 This paper has successfully demonstrated the potential of CDR data to capture rational
14 route choice behaviour for long-distance inter-regional O-D pairs. The broad choice
15 framework was used to leverage the limitations of CDR data where unique route
16 choices could not be observed for some users, and only the broad sub-groups of the
17 possible routes were identifiable. This study is unique in the sense that it adapts the
18 broad choice framework to the context of route choice modelling using noisy CDR data.

19 An examination of the parameter signs shows that CDR data is able to capture
20 the expected sensitivities towards particular route attributes. A review of different
21 models accounting for overlap was conducted, and among these, the c-logit and the path
22 size logit models were considered. A comparison of these models against the
23 multinomial logit model (which does not account for overlap) showed significant
24 improvements in model fit, with the path size logit model giving the best performance.
25 The validation runs based on the 20% holdout samples largely showed the same

1 advantages in prediction, especially for the path size logit model. These results show
2 that CDR is able to capture the expected behaviour towards overlapping routes.

3 This study is timely as it extends the application of CDR data beyond travel
4 pattern analyses to econometric modelling of route choice that can be used for
5 forecasting. The proposed framework can help in the assessment of different policy
6 implications at a low cost compared to traditional approaches, which involve expensive
7 data collection. For example, the models developed in this study can be used to reliably
8 estimate the value of travel time (VTT) as we have demonstrated. This study is thus
9 beneficial to developing countries where budget constraints on transport studies are
10 common and traditional data for transport studies is scarce.

11 We conclude that the study findings serve as a proof-of-concept that CDR data
12 can be successfully used to model route choice behaviour for long-distance inter-
13 regional trips, where there is a strong possibility that a user will use his/her phone
14 during the trip, thereby enabling the capture of their partial trajectories needed for route
15 identification. It may be noted that with the increasing trend of mobile internet usage
16 (Gerpott and Thomas, 2014), the temporal resolution of CDR locations is likely to
17 improve significantly in the near future, and this could make CDR data suitable for
18 evaluating route choice behaviour for short trips.

19 A comparison of the study findings with those based on traditional data from
20 Senegal would have been insightful, however, this was not possible due to data
21 unavailability. Investigating the performance of the proposed approach in urban or
22 intra-city scenarios would be an interesting direction for future research.

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