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Clinimetrics of Ultrasound Pathologies in Osteoarthritis: Systematic Literature Review and Meta-analysis

3 Introduction

| 4 | Osteoarthritis (OA) is the ubiquitous joint disease, predisposing to severe disability and |
|----|--|
| 5 | economic burden on the community [1], with its prevalence surging world-wide due to an |
| 6 | increase in ageing population [2]. Pathophysiology of OA is complex and involves multiple |
| 7 | tissue pathologies; there is currently no consensus on which manifestations should be measured |
| 8 | in OA clinical studies. In attempting to objectively evaluate OA structural components, X-ray |
| 9 | and MRI have been commonly employed as they visualize constructs related to cartilage. |
| 10 | Ultrasound has been less well studied, but does provide certain advantages such as real-time |
| 11 | assessment of multiple joints, sensitive visualisation of synovitis without the need for contrast |
| 12 | agents [3-5], its detection of pathologies such as meniscus extrusion [6-9], osteophytes [10-12], |
| 13 | degeneration of femoral trochlear cartilage [13-16], and effusions (which might be missed on |
| 14 | clinical examination or plain radiography) [5, 17-19]. As a result of these attributes, and likely |
| 15 | because of widespread uptake in the rheumatology community, ultrasound has increasingly been |
| 16 | applied as an outcome tool in OA clinical studies over the last decade. |
| 17 | Since Keen et al. reported its clinimetrics, mainly with a focus on validity, in a |
| 18 | systematic review in 2009, based on PubMed and Medline database searches [20], many |
| 19 | ultrasound OA studies have been published according to recent narrative reviews [21, 22], with |
| 20 | most papers having sound methodology, utilizing more advanced technology such as high- |

| 21 | frequency probes, and use of definitions and techniques from Outcome Measures in |
|----|--|
| 22 | Rheumatology (OMERACT) [23] and European League Against Rheumatism (EULAR) |
| 23 | Ultrasound Working Groups [24]. The increase in knowledge base in this area, therefore, |
| 24 | warrants an update of the previous review in terms of clinimetrics (clinical measurement) such as |
| 25 | reliability, validity, responsiveness [25]. Moreover, there is no published meta-analysis on these |
| 26 | clinimetric of commonly assessed ultrasound pathologies in OA. |
| 27 | Therefore, the purposes of this study were: (1) to systematically review the performance |
| 28 | metrics of ultrasound as applied to the detection of commonly assessed pathologies in people |
| 29 | with OA with a focus on knee, hand and hip jointsand (2) to conduct a meta-analysis of each |
| 30 | clinimetric property for the ultrasound findings if feasible. |
| 31 | |
| 32 | Methodology |
| 33 | Selection criteria |
| 34 | Manuscripts were included if 1) they reported clinimetrics of commonly assessed |
| 35 | ultrasound pathologies in knee or hand or hip OA in adults, and 2) separate clinimetrics for OA |
| 36 | were recorded if the sample included different rheumatic diseases. Articles were excluded if 1) |
| 37 | they were not related to the use of B-mode or color/power Dopplerultrasound, 2) they utilized |
| 38 | ultrasound only for injection guidance, 3) they did not provide any ultrasound clinimetrics, or 4) |
| 39 | they were review or editorial articles, non-human or non-English publications. The study |
| 40 | protocol was registered in PROSPERO database with CRD42016039954. |

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42 Information source and selection process

| 43 | One reviewer (WMO) searched MEDLINE via Ovid, EMBASE, and Cochrane Library |
|----|---|
| 44 | databases from their respective inception to September 2016. The search strategy for each |
| 45 | database was developed in consultation with an experienced librarian (supplementary data 1). |
| 46 | The same reviewer implemented the secondary searching in reference lists of included articles, |
| 47 | ultrasound chapters in reference books, and conference abstracts of Osteoarthritis Research |
| 48 | Society International (OARSI), EULAR and American College of Rheumatology (ACR) from |
| 49 | 2014 to 2016. |
| 50 | The retrieved articles were imported into Covidence systematic review software [26], |
| 51 | and two reviewers (WMO and MD) screened the titles and abstracts independently. |
| 52 | Subsequently, the full texts of the selected articles were retrieved and judged against the |
| 53 | inclusion and exclusion criteria. Any disagreement was resolved with a third reviewer (DJH). |
| 54 | When the included studies referred to a previous paper for methodology or reliability, it was |
| 55 | obtained, and appraised if it met the selection criteria. |
| 56 | |
| 57 | Data extraction and quality assessment |
| 58 | According to the OMERACT Instrument Selection Algorithm [27], the same two |
| 59 | reviewers conducted data extraction with a standardized excel template including: 1) |
| 60 | characteristics of studies such as study design, setting, sample size, participants selection and |

| 61 | diagnostic criteria; 2) technical features such as ultrasound mode (i.e. B-mode, Power Doppler), |
|----|---|
| 62 | machine settings, scanning methods, the particular joints and structures scanned; 3) pathological |
| 63 | findings such as ultrasound definitions of pathologies and scoring methods; 4) types of |
| 64 | clinimetrics. |
| 65 | For reliability, imaging and operator characteristics were recorded. Construct validity |
| 66 | was defined if the study correlated ultrasound findings with clinical assessment, plain |
| 67 | radiography or MRI. Criterion/predictive validity was defined when ultrasound findings were |
| 68 | concurrently or predictively compared with the gold standard, i.e. histopathology, arthroscopy. |
| 69 | Discriminative validity was also assessed in two aspects: internal responsiveness (the ability of |
| 70 | ultrasound measure to change over a pre-specified time frame) or external responsiveness (the |
| 71 | extent to which changes in ultrasound measure relate to corresponding changes in a reference |
| 72 | measure of health status) for interventional studies. Feasibility was calculated in scanning time |
| 73 | required for the whole ultrasound examination. One reviewer (WMO) appraised the |
| 74 | methodological quality, using the modified 19-item version (supplementary data 2) derived from |
| 75 | Downs and Black score system [28, 29] for all included papers, and 11-item Quality Appraisal of |
| 76 | Diagnostic Reliability (QAREL) score for reliability papers [30]. |
| 77 | |
| 78 | Pooling Criteria for Meta-analysis |
| 79 | For meta-analysis, data were pooled if the paper reported sufficient data to calculate 1) |

80 kappa or ICC for reliability, 2) Pearson and Spearman correlation coefficients for validity, 3)

| 81 | standardized mean difference for internal responsiveness, 4) correlation coefficient for external |
|-----|--|
| 82 | responsiveness. For validity, all types of regression coefficients (β) were omitted from pooling |
| 83 | due to controversy in combining them [31]. |
| 84 | |
| 85 | Statistical analysis. |
| 86 | Qualitative analysis |
| 87 | Frequencies and percentages were computed for categorical variables of included papers. |
| 88 | Meta-analysis and Meta-regression |
| 89 | Unit of analysis: Each sample of subjects from studies was assumed as one unit of |
| 90 | analysis. When two or more articles documented reliability/correlation coefficients, using the |
| 91 | same sample, the coefficient was included only once as the unit of analysis. When one article |
| 92 | reported more than one reliability/correlation coefficients of the same clinimetric measurement |
| 93 | from the same sample, the mean coefficient was calculated, and then analyzed in the meta- |
| 94 | analysis. If the study comprised independent subgroups, the subgroups were pooled as a separate |
| 95 | unit of analysis [32]. |
| 96 | |
| 97 | Pooling data: Separate meta-analyses were performed for each type of clinimetrics: 1) |
| 98 | kappa or ICC for inter-rater or intra-rater reliability 2) construct validity against healthy control, |
| 99 | pain, functional assessment, conventional X-rays, MRI, or biomarkers, 3) internal or external |
| 100 | responsiveness. These data were pooled, based on each ultrasound pathology (synovitis/effusion/ |

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| 101 | osteophyte/etc.) to be clinically meaningful. For reliability statistics, pooling was stratified for |
|-----|--|
| 102 | each grading method (binary/semi-quantitative/quantitative) of the same ultrasound pathology. |
| 103 | For weighted meta-analysis of kappa estimates, when the standard error (SE) was |
| 104 | unavailable, it was calculated from 95% confidence interval (CI) bounds [33]. If both SEs and |
| 105 | CIs were not reported, the largest observed SE from the included studies was used. For ICC |
| 106 | statistics of reliability and Pearson or Spearman correlation coefficients of validity, effect sizes |
| 107 | were first obtained through the z-transformations, and then the resulting pooled effect sizes were |
| 108 | back-transformed (z to r transformation) to the level of original coefficients for easier |
| 109 | interpretation [34]. For merging odd ratios in validity studies, the log odds ratio and the standard |
| 110 | error of the log odds ratio were determined [35]. The standardized mean difference (SMD), using |
| 111 | Hedges' g due to inclusion of small studies (<30 patients/joints), was calculated for internal |
| 112 | responsiveness [36], and correlation coefficients were pooled for external responsiveness |
| 113 | through the z-transformations [37]. |
| 114 | For assessment of heterogeneity, Cochran Q test was computed [34]. The I^2 was used to |
| 115 | quantify how much of the total variability can be attributed to heterogeneity [38]. To scrutinize |
| 116 | possible publication bias, it was intended to evaluate with funnel plot techniques [39], Begg's |
| 117 | rank test [40] and Egger's regression test [41], as appropriate, given the known limitations of |
| 118 | these methods, if the minimum number of studies could be pooled. All analyses for calculating |
| 119 | the estimates from primary studies, and for pooling data were carried out by using the SPSS, |
| 120 | Excel and Comprehensive Meta-analysis software. |

| 121 | | |
|-----|--|--|
| | | |

| 122 | Results |
|-----|--|
| 123 | Identification of included studies |
| 124 | Our search identified 1246 records (468 Medline, 774 Embase and 4 Cochrane library) |
| 125 | with 120 duplicates. After screening the titles and abstracts, 195 articles remained. Furthermore, |
| 126 | 9 articles were retrieved from the reference lists, totalling 204 articles eligible for full-text |
| 127 | review. Of these, 100 articles were selected as shown in the PRISMA flow diagram (Figure. 1). |
| 128 | |
| 129 | Study characteristics |
| 130 | One hundred articles (listed in supplementary data 3), having a total of 8542 patients |
| 131 | and 32373 OA joints, and published between 1982 and 2016, were included in the systematic |
| 132 | review. The studies' characteristics were summarized in supplementary data 4. Majority of |
| 133 | studies (79%) were documented after 2008. Knee OA was the most widely investigated (n=64), |
| 134 | followed by hand OA (n=28), and hip OA (n=8). |
| 135 | According to Oxford Centre for Evidence-Based Medicine guidelines (www.cebm.net/), |
| 136 | 42 papers utilized a cross-sectional design (42%) and 28 papers applied a cohort design (28%). |
| 137 | The participants were recruited from out-patient rheumatology clinics in 46 papers; the setting |
| 138 | was not mentioned in 23 papers. The selection method was not described in half of the studies, |
| 139 | followed by a consecutive method (n=40), convenience (n=5) and random methods (n=5). ACR |
| 140 | criteria was employed for diagnosis in most of studies (n=81); 14 papers did not disclose |

| 141 | diagnostic criteria. The mean age of included studies ranged from 50.1±9.2 to 71.9±5.9 years; |
|-----|---|
| 142 | female participants varied from 37% to 100%; the mean BMI from 22.2 \pm 2.6 to 33.5 \pm 4.6 kg/m ² . |
| 143 | Eight studies recruited mixed samples with different diseases, but delineated separate |
| 144 | clinimetrics of OA sub-group. |
| 145 | |
| 146 | Ultrasound scanning techniques and definition |
| 147 | For simplicity, the EULAR scanning method [42] and OMERACT definitions [23] were |
| 148 | assumed as the standard criteria to identify respective OA pathologies. Out of 100 papers, power |
| 149 | Doppler was inverstigated in 31 (supplementary data 5). Doppler specifications were detailed in |
| 150 | 19 papers: Doppler frequency was reported in 9 (from 12 MHz to 6.3 MHz); pulse repetition |
| 151 | frequency (PRF) in 10 (from 13.2KHz to 3 Hz); wall filter and gain in 17. One paper examined |
| 152 | contrast ultrasound. |
| 153 | Eighty-eight papers defined ultrasound pathology; 26 papers referred the EULAR |
| 154 | scanning protocol; 59 papers administered their own methods or modification from previous |
| 155 | papers; 13 papers did not delineate the specific scanning method. Thirty-nine studies applied the |
| 156 | OMERACT definitions, which were found to be increasingly used across the years from 1 paper |
| 157 | in 2008, and then 5 papers in 2012 to 10 papers in 2016 (Supplementary data 6). |
| 158 | |
| 159 | Ultrasound lesions and scoring system |

160 Overall, synovial pathologies were more extensively examined, i.e, effusion (52%),

| 161 | synovial hypertrophy (37%), Doppler activity (31%), Baker's cyst (25%), compared to structural |
|-----|---|
| 162 | lesions, i.e, osteophyte (29%), cartilage thinning (28%). A variety of grading systems was |
| 163 | evaluated [binary (n=49,49%), semi-quantitative (n=42, 42%), and quantitative (n=40,40%)]. |
| 164 | |
| 165 | Qualification of ultrasound operator |
| 166 | Only twenty papers declared the number of operator's training years in musculoskeletal |
| 167 | ultrasound, ranging from 3 months to 24 years. The operator/readers were also of diverse |
| 168 | academic backgrounds: rheumatologist (27% of all papers), ultrasonographer (16%), radiologist |
| 169 | (11%), others such as physiatrist, surgeon, fellow-in-training (26%), and no report (20%). |
| 170 | |
| 171 | Methodological quality |
| 172 | The average quality score across the studies assessed with the modified Downs and |
| 173 | Black instrument was 13.01 out of 19 items (taking into account the questions that were not |
| 174 | applicable for certain studies). The chart in supplementary data 7 outlined the proportion of the |
| 175 | 100 studies that met each of the quality assessment items. The papers, in general, had a good |
| 176 | rating (>60%) on the 13 items. However, most papers fell short severely on some items such as |
| 177 | reporting of sample size calculation and sufficient power (10%). |
| 178 | The average QAREL score was 5.93 out of 11 items across all reliability studies (n=43). |
| 179 | Blindness to other raters, own prior findings, clinical information and non-clinical clues were |
| 180 | described in 40% (n=17), 28% (n=12), 56% (n=24) and 5% (n=2), respectively (supplementary |

| 181 | data 8). Randomization of patients/raters was found only in 53% (n=23). As there was no |
|-----|---|
| 182 | definite consensus related to time interval for stability of ultrasound findings between repeated |
| 183 | measurements, only evaluation of stored images was given as yes (n=17), and rating of the |
| 184 | acquired image as unclear (n=26). Overall, the regression plot displayed the significant |
| 185 | improvement of QAREL quality score across the years (β =0.40, P=0.01) (Supplementary data |
| 186 | 9). |
| 187 | |
| 188 | Clinimetric properties |
| 189 | Among the 100 studies, 32 papers were identified for the intra-rater reliability, 25 for |
| 190 | inter-rater reliability, 57 for construct validity, 5 for criterion validity in knee, 10 for clinical |
| 191 | predictive validity, 6 for structural predictive validity, 21 for intrinsic responsiveness, 8 for |
| 192 | extrinsic responsiveness and 7 for feasibility. |
| 193 | |
| 194 | Meta-analysis |
| 195 | The meta-analysis was conducted only for knee OA. Pooling could not be performed for |
| 196 | hand and hip OA due to a paucity of reported clinimetric data for ultrasound, and so descriptive |
| 197 | analysis was presented. Publication bias was not examined due to inadequate numbers of |
| 198 | included papers for a specific OA pathology, which did not allow proper assessment of funnel |
| 199 | plots or more advanced regression-based assessments. |

Knee OA

Reliability:

| 203 | Inter-rater reliability: According to the pooling criteria, stratified kappa meta-analysis |
|-----|---|
| 204 | was conducted across 11 knee studies, including 38 kappa estimates and 556 joints of 506 |
| 205 | patients. ICC estimates was pooled across 7 knee studies with a total of 19 ICC estimates in 340 |
| 206 | joints of 308 participants. Kappa coefficients were interpreted according to Landis and Koch |
| 207 | (0:poor; 0.01-0.20:slight; 0.21-0.40:fair; 0.41-0.60:moderate; 0.61-0.80:substantial; 0.81- |
| 208 | 1.00:almost perfect) [43]. |
| 209 | The pooled kappa of binary score (Table 1) was almost perfect for Baker's cyst |
| 210 | [0.92(0.83-1)], and substantial for effusion [0.75(0.41,1)] (Figure 2), with nearly all pathologies |
| 211 | revealing considerable heterogeneity ($I^2=70$ to 99). For semi-quantitative score, pooled kappa |
| 212 | values were moderate for cartilage thinness [0.44(0.15-0.74)], and substantial for all pathologies, |
| 213 | with high heterogeneity (I^2 = 78-98). For quantitative scores, all pathologies provided almost |
| 214 | perfect reliability for pooled ICC estimate. |
| 215 | Intra-rater reliability: Stratified kappa meta-analysis was peformed from 8 knee |
| 216 | studies, including a total of 23 kappa estimates for 502 joints of 465 patients. For ICC values, |
| 217 | data were pooled from 9 knee studies with a total of 21 ICC estimates for 566 joints of 490 |
| 218 | participants. |
| 219 | The pooled kappa of semi-quantitative score (Table 2) was varied from moderate for |

cartilage thinness [0.55(0.45-0.66)], substantial for synovitis [0.69(0.60-0.78)] and osteophyte

| 221 | [0.74(0.67-0.81)] to almost perfect for meniscal extrusion [0.81(0.66-0.96)], exhibiting low |
|-----|--|
| 222 | heterogeneity ($I^2=7$ to 51). For quantitative scores, reliability was almost perfect in all |
| 223 | pathologies. |
| 224 | |
| 225 | Validity |
| 226 | Meta-analysis was stratified for each comparator such as asymptomatic controls, pain, |
| 227 | function, X-rays, MRI or blood biomarkers or histology or arthroscopy. Correlation coefficients |
| 228 | were interpreted according to the Evans' classification [44], <0.20:very weak; 0.20-0.39:weak; |
| 229 | 0.40-0.59:moderate; 0.60-0.79;strong and >0.80:very strong. |
| 230 | |
| 231 | Construct validity against asymptomatic controls: Six studies, including 643 joints |
| 232 | from 582 participants, provided 23 odd ratios. In symptomatic patients (Table 3), the pooled odd |
| 233 | ratio demonstrated a very strong association with effusion [7.46(2.56,21.70)], and a strong |
| 234 | association with Baker's cyst [3.23(1.57,6.67)] and meniscal extrusion [3.08(1.06,8.92)]. |
| 235 | Heterogeneity was generally moderate ($I^2=41$ to 61). |
| 236 | |
| 237 | Construct validity against pain: Pooling 37 estimates out of 16 studies, including 2577 |
| 238 | joints from 2085 patients, revealed weak correlation with trivial heterogeneity $[I^2=0]$ (Table 4). |
| 239 | |
| 240 | Construct validity against function: Meta-analysis of 15 estimates out of 9 studies, |

| 241 | including 1333 joints and 802 patients, resulted in weak correlation, and mild heterogeneity |
|-----|--|
| 242 | [I ² =20-38] (supplementary data 10). Six studies used WOMAC [45]. |
| 243 | |
| 244 | Construct validity against X-rays: Pooling across a total of 49 estimates from 11 |
| 245 | studies (1956 joints, and 1530 patients) indicated strong correlation with osteophyte |
| 246 | [0.60(0.45,0.71)], moderate correlation with effusion [0.54(0.37,0.68)] and meniscal extrusion |
| 247 | [0.48(0.34,0.60)], and weak association with cartilage thickness [0.35(0.12,0.55)]. Heterogeneity |
| 248 | was moderate [I ² =34-52] (Table 5). Kellgren Lawrence score [46] was applied in 10 studies. |
| 249 | |
| 250 | Construct validity against MRI : Strong correlation (r>0.60) was detected on pooling |
| 251 | 29 estimates across 4 studies examining 306 knee joints in 230 patients, using 0.2T to 1.5 T MRI |
| 252 | with dedicated knee coils (supplementary data 10). |
| 253 | |
| 254 | Construct validity against biomarkers: Twenty-three estimates of serum cartilage |
| 255 | oligomeric matrix protein (COMP) were pooled across 4 studies involving 95 knee joints from |
| 256 | 95 patients, generating weak correlation [r=0.003 to 0.21] with trivial heterogeneity [I ² =0] |
| 257 | (supplementary data 10). |
| 258 | |
| 259 | Criteria validity against histology: Pooling of four estimates from 2 studies, |
| 260 | evaluating histological cartilage thickness in 190 knee joints from 113 patients, produced a |

| 261 | moderate correlation [$r=0.66(-0.05-0.93)$], and considerable heterogeneity [I ² =90] |
|-----|---|
| 262 | (supplementary data 10). |
| 263 | |
| 264 | Criteria validity against arthroscopy: Ultrasound pathologies focused by three |
| 265 | arthroscopic studies, using Noyes' grading scale [47], were not the same among the papers, and |
| 266 | so pooling could not be executed Generally, arthroscopic gradings correlated strongly with |
| 267 | osteophyte [11], moderately with cartilage grading [14] and weakly with subchondral bone [48]. |
| 268 | |
| 269 | Responsiveness |
| 270 | According to Cohen [49], values of 0.0, 0.20, 0.50, and 0.80 or greater represented |
| 271 | trivial, small, moderate, and large responsiveness, respectively. |
| 272 | |
| 273 | Internal responsiveness: Pooling 31 estimates across 10 studies, comprising 480 joints |
| 274 | from 393 patients, produced a moderate effect size for Baker's cyst [0.58(0.40,0.77)], and small |
| 275 | effect size for synovial hypertrophy [0.30(0.05,0.56)], effusion [0.28(0.00,0.56)] and cartilage |
| 276 | thickness [0.20(0.04,0.36)] (Table 6). The interventions included injections of different steroids |
| 277 | (n=6), platelet rich plasma (n=2), glucosamine (n=1), and exercises (n=1). The study duration |
| 278 | ranged from 2 weeks to 6 months. |
| 279 | |

External responsiveness: Pooling 7 estimates across 4 studies with a total of 121 joints

280

| 281 | and 121patients, provided moderate correlation for synovial hypertrophy [0.43(-0.02,0.73)], and |
|-----|--|
| 282 | weak correlation for Baker's cyst [0.35(-0.11,0.69)]. Substantial heterogeneity was detected |
| 283 | $[I^2=68-74]$ (supplementary data 10). The interventions were intra-articular steroid injections |
| 284 | (n=3), and shortwave diathermy (n=1). |
| 285 | (Tables for stratified meta-analysis, and figures for forest plots were also described as |
| 286 | supplementary data 10 and 11). |
| 287 | |
| 288 | Feasibility |
| 289 | Five studies reported the scanning time for complete examination, which varied from 5 |
| 290 | min to 15 min depending on how many pathologies were scanned (supplementary data 10). |
| 291 | |
| 292 | Hand OA |
| 293 | Reliability |
| 294 | There were 4 inter-rater reliability studies for binary scores [50-53], 3 for semi- |
| 295 | quantitative scores [5, 12, 51] and 1 for quantitative scores [54]. The binary scoring system |
| 296 | provided the kappa ranging from slight in cartilage thickness [51] to excellent in synovitis, |
| 297 | effusion and osteophyte [52]. For semi-quantitative score, the kappa values varied from slight in |
| 298 | cartilage thickness [51] to substantial in osteophyte and synovitis [5, 12]. For quantitative score, |
| 299 | ICC was excellent in synovial hypertrophy [54]. |
| 300 | Among intra-reliability studies, 7 studies applied binary scores [5, 10, 12, 50, 51, 55, 56]: |

| 301 | five studies used semi-quantitative scores [5, 12, 51, 57, 58]; one study examined quantitative |
|-----|--|
| 302 | scores [59]. Similar findings of kappa values were reported for different pathologies but with a |
| 303 | higher actual kappa values. |
| 304 | |
| 305 | Validity |
| 306 | Only two studies reported construct validity of ultrasound with pain, disclosing very |
| 307 | weak correlation [57, 60]. Four studies documented ultrasound data for functional correlation |
| 308 | which varied from very weak to weak in most pathologies [55, 57, 60, 61]. Validity of ultrasound |
| 309 | with X-rays was investigated in two studies, providing very weak correlation [56, 60]. However, |
| 310 | ultrasound provided moderate correlation with MRI for osteophyte (r=0.49) and synovitis |
| 311 | (r=0.43) on semi-quantitative scale [62]. |
| 312 | |
| 313 | Responsiveness |
| 314 | Two studies supplied sufficient information to calculate the internal responsiveness. One |
| 315 | study revealed trivial effect size for synovitis and power Doppler outcomes at 12 weeks after |
| 316 | intramuscular methylprednisolone injection [63], and small effect size was detected at 4 weeks |
| 317 | for the same pathologies in another study, using intra-articular injections of hyaluronic acid as an |
| 318 | intervention [64]. |
| 319 | For external responsiveness, one study reported strong correlation of synovial thickening |
| 320 | and power Doppler with VAS pain at 4 weeks [64]. |

and power Doppler with VAS pain at 4 weeks [64].

| 321 | |
|-----|---|
| 322 | Hip OA |
| 323 | Reliability |
| 324 | Inter-rater reliability of binary score ranged from fair in effusion to moderate for |
| 325 | osteophyte in one study [65] while another study recorded excellent reliability for the same |
| 326 | pathologies [66]. |
| 327 | Intra-rater reliability of binary score was moderate in joint effusion and substantial in |
| 328 | osteophyte [65] while the other revealed the excellent kappa [66]. For semi-quantitative scores |
| 329 | by radiologists, excellent kappa was reported for the synovial thickness [67]. |
| 330 | |
| 331 | Validity |
| 332 | Ultrasound synovitis and osteophyte scores demonstrated a strong association with pain |
| 333 | on activity [65]. Weak correlation was documented between effusion and Lequesne index [68], |
| 334 | and between osteophyte and KL grading (r=0.26) [65]. |
| 335 | |
| 336 | Responsiveness |
| 337 | One study applied ultrasound synovial hypertrophy and effusion as outcome measure to |
| 338 | evaluate internal responsiveness, providing moderate effect size (SMD=0.44) at 3 months after |
| 339 | intra-articular injection of 8 mg betamethasone [69]. |
| | |

Discussion

| 342 | Overall, the main findings of our meta-analysis suggest various (weak to very strong) |
|-----|--|
| 343 | construct validity with patients findings and other imaging modalities, depending on pathologies |
| 344 | and comparators, moderate to substantial reliability, strong criterion validity with cartilage |
| 345 | histology, and small to moderate responsiveness to interventions. On qualitative analysis, this |
| 346 | systematic review revealed substantial clinical, technical and methodological heterogeneity of |
| 347 | ultrasound within OA literature, requiring caution in interpreting these meta-analytic results. |
| 348 | However, on quantitative analysis, I ² , which denotes statistical heterogeneity, was only low or |
| 349 | moderate for most of clinimetrics. |
| 350 | Although ultrasound possesses promising potential in OA clinical trials, fewer studies in |
| 351 | hand and hip joints were detected in the literature, compared to the knee. Although |
| 352 | utilization/reporting of OMERACT definitions has gained a significantly positive trend over last |
| 353 | decade, a marked variability of ultrasound scanning characteristics was noted, highlighting the |
| 354 | necessity of following/reporting international consensus protocols in future studies. |
| 355 | In the context of methodological quality, a modified Downs and Black quality |
| 356 | assessment score [28] was administered to identify the potential bias and display the summary of |
| 357 | these bias. All studies, which documented the clinimetric data for each pathology, were pooled |
| 358 | without applying exclusion on the basis of study quality scale because the threshold for |
| 359 | exclusion reduced the precision [70], and was necessarily subjective [71]. According to Detsky et |
| 360 | al, it seemed highly unlikely that these quality scores would generate a linear or monotonically |

| 361 | increasing association with true quality, and no objective reference standard simply existed for |
|-----|---|
| 362 | determining the "true" scientific rigour of a trial [72]. Moreover, due to a limited number of |
| 363 | papers which documented clinimetric data for each ultrasound pathology, the sensitivity analysis, |
| 364 | based on study quality score, could not be examined (i.e. there were some pathologies for each of |
| 365 | which only one paper existed as a unit of analysis.). |
| 366 | In addition, definitions in OA are difficult in terms of what is normal, and what is |
| 367 | defined for OA (radiographic OA or ACR criteria, which means totally different things), making |
| 368 | validity research not easy. |
| 369 | Our meta-analysis results indicated moderate to substantial reliability [minimum |
| 370 | kappa≥0.44(0.15,0.74) and minimum ICC≥0.82(0.73-0.89)] for ultrasound pathologies of knee |
| 371 | OA. Generally, the binary and quantitative scores produced higher reliability statistics than semi- |
| 372 | quantitative score. Some papers calibrated the semi-quantitative scores by utilizing the atlas- |
| 373 | based grading methods [11, 73] while some defined the grading by quantitative cut-offs [6]. The |
| 374 | reliability of Baker's cyst, meniscal extrusion, osteophyte, synovitis and effusion were at least |
| 375 | substantial for the semi-quantitative scores. |
| 376 | The musculoskeletal experience of ultrasound operators ranged from those with short- |
| 377 | course training to very experienced specialist, and so the meta-analysis results represented the |
| 378 | generalizability of reliability statistics across different levels of ultrasound experience. However, |
| 379 | it should be noted that operator-dependent nature of ultrasound measurement and quality of US |
| 380 | machines could largely influence on the performance of the reliability statistics, especially when 19 |

381 smaller joints are addressed.

| 382 | The limited data for criterion validity of OA ultrasound features focused predominantly |
|-----|---|
| 383 | on cartilage histology, with overall strong correlation. Conflicting reports were found for |
| 384 | correlations of synovitis/Doppler signals with synovial vascularity in a mixed sample of |
| 385 | inflammatory arthritis and osteoarthritis [74-77]. Semi-quantitative grading scores currently |
| 386 | applied for OA synovitis were adopted from those validated for inflammatory rheumatoid |
| 387 | arthritis, assuming that synovitis was only quantitatively but not qualitatively different between |
| 388 | the inflammatory arthritis and osteoarthritis [78]. However, replication of these semi-quantitative |
| 389 | scoring systems in osteoarthritis might require consideration due to the low degree of |
| 390 | inflammation, sustained in osteoarthritis compared to rheumatoid arthritis [18], which is likely to |
| 391 | contribute to floor effects, and thereby impairs the capability to detect improvement changes in |
| 392 | interventional studies. |
| 393 | Pooling construct validity of ultrasound findings in case-control studies (OA versus |
| 394 | healthy population) exhibited strong discrimination in some pathologies, suggesting that |
| 395 | ultrasound might be a potential tool for developing ultrasonographic OA propositions, similar to |
| 396 | preliminary OA propositions with MRI [79]. Furthermore, ultrasound demonstrated a strong |
| 397 | correlation with MRI in principal OA features, indicating the promising usefulness of ultrasound |
| 398 | in clinical care where MRI is not readily accessible. |
| 399 | Generally, ultrasound, as expected, had a very weak association with pain, function and |
| | |

400 blood biomarker (COMP). Almost all individual studies incorporated in the meta-analysis

| 401 | consistently denoted weak correlation between ultrasound features and pain (r \leq 0.40). This |
|-----|--|
| 402 | finding may be attributed to a number of reasons such as complex causes of symptoms in OA, |
| 403 | multi-factorial subjective experience of pain (biopsychosocial factors), and that the ultrasound |
| 404 | outcomes used in individual studies might not captured the multi-dimensional nature of pain |
| 405 | (measurement issues) [80]. In contrast, relationship of ultrasound with X rays produced various |
| 406 | values ranging from weak to strong correlation, depending on ultrasound pathologies. |
| 407 | At least small effect size (SMD≥0.2) was documented in most of interventional studies, |
| 408 | and the low I^2 in pooled meta-analysis was detected. Generally, the inflammatory features such |
| 409 | as Baker's cyst, synovial hypertrophy provides greater internal responsiveness, compared to |
| 410 | cartilage changes, perhaps due to short follow-up duration (maximum 24 weeks). However, this |
| 411 | result should be interpreted with caution as the included studies for sensitivity to change were all |
| 412 | small studies with some limitations. Combining external responsiveness of inflammatory |
| 413 | pathologies revealed a moderate correlation with pain while no studies examined external |
| 414 | responsiveness for structural pathologies. |
| 415 | Ultrasound scanning duration largely depended on the number of joints and pathologies |
| 416 | assessed and the scoring systems employed, which were varied across studies. Development of |
| 417 | international consensus guidelines for feasible composite scoring methods is essential, and still |
| 418 | undergoing. |
| 419 | It should be noted that several papers included in the validity assessment of previous |

420 systematic review [20] had to be excluded as our inclusion criteria was focused only on knee,

hand and hip, not other joints such as foot, shoulder, cervical spine, etc and some papers did not
publish the comparator for validity assessment, clinimetric data, etc. However, more than
additional 60 papers were included in this updated review.

Our review had several potential limitations. The first was the considerable clinical and 424methodological heterogeneity of included studies, requiring caution in interpreting the pooled 425results. However, I² was low for validity and responsiveness measures. The second limitation 426 was that we could not rule out some publication bias although a thorough literature search was 427428attempted. The third limitation is the application of SMD for internal responsiveness instead of calculating standardized response mean (SRM), as most interventional studies did not describe 429 standard deviation of mean change [81]. However, in the literature, the best statistics for 430 treatment responsiveness and interpretation is still controversial, and according to mathematical 431432formulae proposed by Norman et al.[36], SRMs tend to be higher than SMDs. The fourth 433 limitation is that we could not appropriately analyze the confounding effects over technology changes over the years because there were numerous confounders such as machine model, probe 434frequency, operator's clinical background, qualification, training period, the severity of the 435sample, the sensitivity of comparator machine models in examining construct validity against X 436 rays and MRI, while a limited number of papers with clinimetric data for each pathology existed, 437causing a lack of power to examine the impact of these confounders on the clinimetrics by 438439regression analysis.

440

To our knowledge, this is the first meta-analytic systematic review comprehensively

| 441 | examining clinimetrics of ultrasound utilized to evaluate common features of OA, covering the |
|-----|--|
| 442 | original OMERACT filter components. Stratified meta-analysis demonstrated moderate to |
| 443 | substantial reliability, various construct validity with several clinical and imaging comparators, |
| 444 | strong criterion validity with cartilage histology and small to moderate responsiveness. Future |
| 445 | studies should improve the conduct and reporting of clinimetric studies especially for the areas |
| 446 | of several poor quality-items. As most of individual studies were of small sample size and just |
| 447 | focused on some individual pathologies, larger studies with comprehensive ultrasound outcomes |
| 448 | in future would provide more clear insight into the clinimetrics of commonly assessed ultrasound |
| 449 | pathologies in OA. |
| 450 | |
| 451 | Contributions: WMO and DJH conceived and designed the study. JML, PGC, HK, SS, JS and |
| 452 | LAD were also involved in the design of the study. WMO, MD and DJH contributed to acquisition |
| 453 | of the main clinimetric data of included papers. WMO had full access to all the data and analysis, |
| 454 | drafted the manuscript and takes responsibility for the integrity of the work from inception to |
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