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**Association between alcohol consumption and abdominal aortic aneurysms: a systematic review and dose-response meta-analysis**

**Running title:** Meta-analysis of aneurysms and alcohol

Shari M Spencer<sup>1,2</sup>, Adam J Trower<sup>1,2</sup>, Xueli Jia<sup>1,3</sup>, D Julian A Scott<sup>1,3</sup>, Darren C Greenwood<sup>1,2</sup>

<sup>1</sup> Leeds Institute of Cardiovascular and Metabolic Medicine, University of Leeds, Leeds, LS2 9JT, UK; <sup>2</sup> Division of Epidemiology and Biostatistics, School of Medicine, University of Leeds, Leeds, LS2 9JT, UK; <sup>3</sup> Department of Vascular Surgery, Leeds General Infirmary, Leeds Teaching Hospitals NHS Trust, Leeds, UK.

**Address for correspondence:**

Dr Darren C Greenwood, PhD  
Leeds Institute of Cardiovascular and Metabolic Medicine  
University of Leeds  
Leeds, LS2 9JT  
United Kingdom  
Tel: +44 113 343 1813  
Email: d.c.greenwood@leeds.ac.uk

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## 1 **Abstract**

2

3 Background: Alcohol is a possible risk factor for Abdominal Aortic Aneurysm (AAA) but  
4 evidence from individual studies is weak and inconsistent. Existing narrative reviews suggest  
5 the possibility of nonlinear associations. We therefore aimed to formally quantify the  
6 association using a systematic literature review followed by dose-response meta-analysis of  
7 prospective studies.

8 Methods: MEDLINE, Embase and Web of Science were systematically searched to January  
9 2017 for relevant prospective studies of alcohol consumption and AAA risk. Summary  
10 estimates of highest versus lowest levels of consumption, linear and nonlinear dose-  
11 response curves were quantified using random-effects models.

12 Results: Eleven relevant cohorts were identified presenting results from 3580 cases amongst  
13 473092 participants. Data were extracted from 10 cohorts for meta-analyses of high versus  
14 low levels of alcohol consumption (risk ratio = 0.93, 95% CI 0.78-1.11,  $p=0.4$ ,  $I^2=47%$ ). The  
15 linear dose-response RR could be derived from 11 cohorts (RR=1.00 per 8g alcohol/day, 95%  
16 CI 0.97 to 1.04,  $p=0.9$ ,  $I^2=73%$ ). Nonlinear dose-response results showed a tick-shaped curve  
17 with lower risks up to 2 units/day but increasing risks beyond that ( $p=0.05$ ). The increase in  
18 risk beyond 2 units/day is stronger in men than in women.

19 *Conclusions:* Whilst the linear dose-response revealed little evidence of an association  
20 between alcohol consumption and AAA risk, a tick-shaped trend in the association was  
21 observed. This nonlinear dose-response revealed reduced risks for alcohol consumption  
22 below 2 units/day masking increasing risks beyond 2 units/day. Randomised controlled trials  
23 are unlikely, so this study may present the strongest level of evidence available.

1

2 *Key Words:* Abdominal aortic aneurysm; Alcohol; Meta-analysis; nonlinear dose-response.

3

4

5

## 1 Introduction

2 Abdominal aortic diameter is an independent predictor of survival<sup>1</sup>. The normal abdominal  
3 aorta is 19 to 22mm. Dilatation that is 25-29 mm is defined as sub-aneurysmal aorta, and  
4  $\geq 30$ mm is defined as aneurysm (AAA). The vast majority of AAAs are asymptomatic until  
5 they expand beyond 55mm, when the risk of rupture increases substantially<sup>2</sup>. This rupture is  
6 associated with sudden onset of severe pain and collapse, and carries an overall mortality of  
7 80%<sup>3</sup>. Prevalence of AAA ranges from 2% to 8% in adults over the age of 65 years<sup>4</sup> and so for  
8 this age group screening is recommended both by the United Kingdom National Health  
9 Service<sup>5</sup> and the United States Preventative Task Force<sup>6</sup>. There are well established risk  
10 factors for AAA development such as increasing age, male sex and smoking status, but the  
11 role of alcohol remains uncertain<sup>4,7-9</sup>.

12

13 Alcohol consumption is common worldwide with an estimated intake of 6.2 litres of pure  
14 alcohol per capita per year worldwide<sup>10</sup> and is a known cardiovascular disease (CVD) risk  
15 factor though it has been suggested that low levels of consumption may be beneficial, with  
16 only higher levels being detrimental<sup>11-16</sup>. Potential mechanisms include up-regulation of  
17 matrix metalloproteinases leading to aneurysm formation with higher intakes<sup>17</sup>, but it is not  
18 known whether the nonlinear associations generally seen in CVD apply to the same extent  
19 with AAA.

20

21 Determining the nature of the association between alcohol and AAA in adult men and  
22 women would be useful because it is a potentially modifiable risk factor. Given that  
23 evidence from individual studies is often weak and has been inconsistent, with only

1 narrative reviews conducted to date, the nature of the association between alcohol and  
2 AAA remains to be quantified. We therefore aimed to quantify the association using a  
3 systematic literature review followed by dose-response meta-analysis of prospective studies  
4 across a broad range of consumption levels, exploring potential sources of heterogeneity. In  
5 particular, we aimed to investigate whether a possible nonlinear trend, similar to that seen  
6 with CVD generally, exists in the relationship between AAAs and alcohol consumption.

7

## 8 **Methods**

### 9 *Search Strategy*

10 A literature search was conducted for all published articles up to January 2017 with the aim  
11 of retrieving all articles that reported associations between alcohol consumption and AAAs.  
12 Three electronic databases were searched: MEDLINE, Embase and Web of Science using  
13 keyword searches and exploded MESH terms based on synonyms for aspects of alcohol  
14 consumption and AAAs. The search was then restricted to studies identified as cohort  
15 studies, case-control studies, cross-sectional studies and randomised controlled trials, with  
16 the aim of identifying all prospective studies for meta-analysis. No date or language  
17 restrictions were imposed. The detailed search strategy used for the MEDLINE database can  
18 be seen in Supplemental Table 2. Reference lists of publications obtained were then hand-  
19 searched for additional relevant articles. The review protocol is published on PROSPERO,  
20 registration number CRD42017055529.

21

### 22 *Study Selection*

1 The titles and abstracts of retrieved publications were screened by two authors (SMS and  
2 DCG) to initially remove articles that were obviously irrelevant such as those pertaining to  
3 surgically associated risks of aneurysm repair and case reports. Of the remaining  
4 publications, the full texts were obtained in order to assess their relevance, again by two  
5 authors (SMS and DCG). Unpublished studies and abstracts were excluded. Only prospective  
6 cohort studies, including historical cohorts and case-control studies nested within a cohort  
7 were eligible for inclusion in the meta-analysis.

8

9 Studies must have been based on the assessment of alcohol consumption prior to the onset  
10 of AAAs to be included. Radiological/clinical diagnosis of an AAA, death due to a ruptured  
11 AAA, ruptured and/or surgical repair of an AAA or autopsy findings of an AAA were the  
12 outcomes included. An estimate of risk ratio (RR) must have been reported with the  
13 corresponding confidence interval (CI).

14

15 Alcohol consumption must have been measured quantitatively. If the number of drinks,  
16 glasses or units was quoted but not its alcohol content, the measurement was assumed to  
17 be a standard alcoholic drink and the associated alcohol content was determined by what is  
18 considered standard in the country/region where the research was conducted<sup>18-20</sup>.

19

20 To conduct the linear dose-response meta-analysis, studies must have reported at least 3  
21 categories of quantified alcohol intake or analysis of alcohol as a continuous variable  
22 together with the associated estimate of RR and CI. If more than one multivariable model  
23 was presented, the model used was the one most consistent with adjustment based on the

1 minimal sufficient adjustment sets from a directed acyclic graph, so that the most  
2 appropriate adjustment for confounding was made. Excluded were studies that reported  
3 episodic drinking patterns only such as binge drinking.

4

#### 5 *Data extraction and quality*

6 The methodological quality of included studies was assessed by two authors (SMS and DCG)  
7 using the Newcastle-Ottawa scale<sup>21</sup>. In the selection category, stars were awarded where  
8 adult participants were sampled from the general population and were representative of  
9 alcohol consumption in that population, if alcohol consumption was assessed by means of  
10 patient records or a structured interview or questionnaire, selection of unexposed  
11 participants from within the same population, and for demonstration that at the beginning  
12 of the study an AAA diagnosis was not present. In the comparability category, stars were  
13 awarded for adjustment for smoking, age and sex. Stars in the outcome category were  
14 awarded for AAA diagnosis made based on death or medical records as opposed to being  
15 self-reported, for follow-up till either the completion of the study, death or AAA diagnosis  
16 with at least 70% follow-up, where participants should have been followed up for at least 15  
17 years.

18

19 The following information was extracted from identified studies: names of authors,  
20 publication year, type of study design, country where the research was conducted, follow-  
21 up duration, sample size, participants' distribution by sex, participants' age range or mean  
22 age, methods of assessment of alcohol intake and outcome assessment, total number of  
23 cases, number of non-cases or person-years and for each exposure category, alcohol intake

1 levels recorded as means, medians, midpoints or ranges for each category or unit of  
2 increment if analysed as continuous, RR estimates and CIs for the categories of alcohol  
3 consumption, covariates included in the multivariable model. Where the distribution of non-  
4 cases or person-years and cases in each category was not reported but the total was  
5 presented, it was estimated based on the definitions of quantiles. Data extraction was  
6 carried out by SMS and DCG, and checked by AJT.

7

### 8 *Data synthesis and analysis*

9 In the analysis of data, three methods were used. The first meta-analysis looked at the  
10 highest versus the lowest level of alcohol consumption. The second was a linear dose-  
11 response analysis and finally any possible nonlinear trend was assessed. A random effects  
12 model was used to derive a summary risk estimate of high versus low consumption with  
13 corresponding 95% CIs using the method described by Der Simonian and Laird<sup>22</sup>. A linear  
14 dose-response trend was obtained for each study using Greenland and Longnecker's  
15 method<sup>23</sup>. The mean or median of the individual alcohol consumption category was used as  
16 the assigned exposure dose, or its midpoint if the mean or median were not given. Where  
17 the category was presented unbounded, and neither the mean nor median was quoted, the  
18 midpoint was calculated by assuming that its width was the same as the adjacent one.  
19 Where the reference category was not the lowest level of intake, estimates were first  
20 recalculated compared to the lowest intake, based on the method of Hamling et al<sup>24</sup>. If a  
21 study presented results for alcohol consumption measured as a continuous exposure then  
22 this was used in preference as that study's estimate in the linear dose-response analysis. If a  
23 difference in mean intakes between AAA cases and non-AAA controls was presented, a

1 linear dose-response was estimated based on the methods of Chêne and Thompson<sup>25</sup>. A  
2 random effects model was then used to pool the dose-response risk estimate from each  
3 study.

4

5 Where separate cohorts were reported in the same publication, the results were included  
6 separately in the pooled analysis. This maintained study independence so that appropriate  
7 heterogeneity estimates would have been obtained. Where two publications contained data  
8 from the same cohort, the one with the most complete adjustment for confounding or most  
9 precise estimate was used, in that order of preference.

10

11 Examination of any possible nonlinearity between alcohol consumption and AAA was done  
12 using restricted cubic splines fitted to each study, using three knots fixed at the 10th, 50th  
13 and 90th percentiles through the distribution of category means, medians or midpoints of  
14 intake across all studies<sup>26</sup>. Studies included in this analysis therefore required more than 3  
15 categories of exposure to have been reported and could not have presented the risk only as  
16 a continuous measure of the exposure. These were then pooled using multivariate meta-  
17 analysis to estimate the linear and nonlinear components of the restricted cubic splines  
18 simultaneously<sup>27,28</sup>.

19

20 Between-study heterogeneity was assessed using the range of individual study estimates  
21 observed<sup>29</sup> and the  $I^2$  statistic for the proportion of total variation explained by between-  
22 study variation<sup>30</sup>. Subgroup analyses were performed for the main linear dose-response  
23 analysis to explore any characteristics of the study quality that could have contributed to

1 the heterogeneity obtained across the various studies. These included geographical  
2 location, follow-up length and adjustment for certain covariates. Additionally, any sex-  
3 specific associations were investigated for analyses of both linear non-linear trends.  
4 Sensitivity analyses were conducted by excluding one study at a time in order to assess the  
5 influence of each study on the overall estimate by observing to what extent the combined  
6 result changed. This also aided in exploring between-study heterogeneity. Potential small  
7 study effects such as publication bias were investigated using funnel plots. The Meta-  
8 analysis of Observational Studies in Epidemiology (MOOSE) guidelines<sup>31</sup> and PRISMA  
9 guidelines<sup>32</sup> were followed when conducting and reporting this review and analysis. All  
10 statistical analyses were performed using Stata version 14.2.

11

## 12 **Results**

13 From the 225 unique articles identified from the literature search, 9 relevant publications  
14 were identified presenting results from 11 separate cohort studies (see Figure 1)<sup>33-41</sup>. These  
15 cohorts included 3580 cases of AAA amongst 473092 participants. Of the 11 included  
16 cohorts, 4 were from Europe and 7 from the USA. Characteristics of these studies are  
17 presented in Supplemental Table 1.

18

19 [Insert Figure 1]

20

21 Ten cohorts identified contributed towards the meta-analysis of highest vs lowest categories  
22 of alcohol intake)<sup>34-41</sup>. The summary risk ratio from this comparison was 0.93 (95% CI 0.78-

1 1.11,  $p = 0.4$ ;  $I^2 = 47\%$ ) (see Figure 2). Estimates for individual studies ranged from around  
2  $RR=0.5$  to  $RR=2.0$ .

3

4 [Insert Figure 2]

5

6 There was no indication of a small study effect such as publication bias, with no evidence of  
7 funnel plot asymmetry ( $p=0.2$ ), though based on only 10 studies.

8

9 Eleven cohorts from all 9 publications could be included in the linear dose-response  
10 analysis<sup>33-41</sup>. The category mean intakes ranged from 0 to approximately 80 mg/day of  
11 alcohol. The summary RR was 1.00 (95% CI 0.97 to 1.04;  $p = 0.9$ ) per 1UK unit (8g) of alcohol  
12 per day (see Figure 3). There was substantial heterogeneity between the studies ( $I^2=73\%$ )  
13 with estimates for individual studies ranging from around  $RR=0.7$  to  $RR=1.7$  per unit/day.

14

15 [Insert Figure 3]

16

17 There was no indication of a small study effect such as publication bias with no evidence of  
18 any funnel plot asymmetry ( $p=0.5$ ).

19

20 When restricted cubic splines were used to investigate the shape of the dose-response  
21 curve there was evidence of nonlinearity ( $p=0.05$ ), with an apparent decreased risk at lower

1 levels of consumption up to about 10-15 mg/day then increasing thereafter, producing a  
2 tick-shaped curve (Figure 4). The nonlinear dose-response curves were estimated on the  
3 basis of six cohorts presenting results with sufficient information in 5  
4 publications<sup>30,34,35,37,38,40</sup>. The data on the very highest intakes though, were based on just  
5 one of the included studies<sup>35</sup>. Tick marks on the horizontal axis indicate the location of  
6 category means, medians or midpoints of studies.

7

8 [Insert Figure 4]

9

10 Risk of bias as assessed by the Newcastle-Ottawa scale was generally low for cohort studies  
11 (Supplemental Table 3). Sensitivity analysis revealed similar estimates when each study was  
12 excluded separately (data not shown. Differential adjustment for anthropometry (e.g. body  
13 mass index) and for socio-economic markers (e.g. education) were significant sources of  
14 heterogeneity, with those ignoring these covariates presenting higher estimates  
15 (Supplemental Table 4). Similarly, studies based in Europe tended to present higher  
16 estimates than those from the US.

17

18 Studies presenting results just for males produced higher overall estimates for the linear  
19 dose-response estimates than studies presenting results just for females, though confidence  
20 intervals were wide (Supplemental Table 4). The nonlinear dose-response curve for males,  
21 however, produced a slightly more pronounced increase in risk at higher levels of alcohol  
22 consumption (Supplemental Figure 1).

1

## 2 **Discussion**

3 This study is the first to quantify the association between alcohol consumption and AAAs in  
4 a meta-analysis and also the first to investigate any nonlinear associations that exist and to  
5 describe the shape of the dose-response curve. The linear dose-response analysis combined  
6 results from 11 large prospective cohort studies, including over 3500 cases identified from  
7 nearly 500000 participants, followed-up for between 5 and 34 years.

8

9 Both the high versus low and the linear dose-response demonstrated no evidence of a  
10 strong association between alcohol consumption and AAA. However, associations may have  
11 been masked by the observed nonlinear trend, where lower levels of alcohol consumption  
12 appeared were associated with lower risk until approximately 15 to 20 g/day, with an  
13 increasing risk thereafter. This produced a tick-shaped curve often observed with alcohol  
14 studies and with turning points similar to those seen in meta-analyses of alcohol with IHD  
15 and stroke<sup>11,42,43</sup>. Results are also consistent with an earlier narrative review<sup>44</sup> and research  
16 into alcohol and aortic diameter<sup>45</sup>. The risk ratio estimates are quite modest, but AAA has a  
17 high prevalence rate, particularly in men. So if the associations are causal, even modest  
18 protection could reduce the number of diagnoses and any subsequent surgery for a large  
19 number of individuals.

20

21 A tick-shaped curve is consistent with people who consume alcohol “in moderation” also  
22 having other aspects of their lifestyle “in moderation” in ways that the individual studies

1 were unable to fully adjust for, or non-consumers being systematically different to the rest  
2 of the cohort they are members of, in a way that puts them at greater risk.

3

4 Meta-analysis of cohort studies is prone to the same potential biases as the contributing  
5 cohorts, so associations cannot be proved to be causal. There may be unknown residual  
6 confounding in some, if not all, of the studies. For example, not all studies adjusted for  
7 smoking, and that that did used self-reported smoking rather than an objective biomarker.

8

9 Between-study heterogeneity was high, despite restriction to prospective studies and  
10 different categories of exposure being combined onto the same scale. Exploration of  
11 potential sources of heterogeneity revealed significant differences between studies that  
12 adjusted for ischemic heart disease (IHD) and those that did not, and between studies that  
13 adjusted for education and those that did not. The studies that did not adjust for IHD  
14 excluded these patients instead, so it is possible that these patients were generally  
15 healthier, potentially eliminating any effect of moderate drinkers appearing to have lower  
16 risk than non-drinkers due to the poorer health status of abstainers. However, the  
17 interpretation of all such exploration of heterogeneity is limited by the relatively small  
18 number of studies in the meta-analyses.

19

20 A further limitation of observational studies is the use of self-reported alcohol consumption.  
21 This exposure estimate is therefore in part subjective and susceptible to differential  
22 measurement misclassification of intake. Additionally, studies in this meta-analysis mostly  
23 used current drinking patterns within the previous year to assess the exposure. If the

1 relevant exposure is longer-term intake, then current intake may not reflect this, but only  
2 one study used an estimate of alcohol intake that was updated with each follow-up visit to  
3 better reflect longer-term use<sup>37</sup>. Use of food frequency questionnaires (FFQs) also requires  
4 averaging intake over the time period in question, such as the past year. This can lead to  
5 difficulties taking episodic drinking into account, and may have included some people with  
6 infrequent heavy drinking as being included in lower categories of intake. It is possible that  
7 episodic drinking may also negate any beneficial effects of lower average consumption<sup>14</sup>.

8

9 The annual growth rate of AAAs less than 55mm is slow<sup>46</sup> but then accelerates beyond that.  
10 There is potential a different strength of association between alcohol consumption and  
11 growth in these two stages or with subsequent rupture. It is also possible that alcohol  
12 consumption may be associated with probability of success or otherwise of endovascular  
13 aneurysm repair. These are possible areas for future research.

14

15 A lack of evidence of increased risk with high levels of alcohol consumption does not imply  
16 evidence of no association. Confidence intervals remain relatively wide, allowing a range of  
17 strength of association. Alcohol consumption is known to have detrimental effects both  
18 acutely and chronically on various aspects of health such as liver cirrhosis, cancers of the  
19 gastrointestinal tract and liver and increased risk of injury, as well as breast cancer even in  
20 moderate amounts<sup>47,48</sup>. Therefore from a public health perspective, our results are  
21 consistent with maintaining current recommendations for alcohol intake<sup>20</sup>.

22

1 The risk for developing an AAA is greater in men than women but the size at which the  
2 female AAA ruptures is potentially smaller than for men. Furthermore, women may benefit  
3 less from intervention than men and may have poorer longer-term survival than men<sup>49-54</sup>.  
4 With these apparent differences between the risk profiles of men and women, it is  
5 therefore relevant that the dose-response curve in males appeared to be steeper than for  
6 women at high levels of consumption.

7

8 Whilst the tick-shaped dose-response curve observed is similar to that seen in other  
9 vascular diseases, there may be differences in pathogenesis. High levels of alcohol have  
10 been shown to upregulate aortic metalloproteases in rats<sup>17</sup>, which has been regarded as a  
11 mechanism in the pathology of AAAs along with inflammatory factors, loss of aortic  
12 elasticity and media thickness<sup>55</sup>. This could therefore account for an increased risk at higher  
13 consumption levels. While it is known that IHD is predominantly an atherosclerotic disease  
14 with the protective effects of alcohol consumption mainly attributed to its antioxidant  
15 properties, reduction in hyperlipidaemia and decreased plaque and thrombus  
16 formation<sup>56,57</sup>, it has been suggested that atheroma formation leads to arterial remodelling  
17 which can subsequently stimulate the biological pathways involved aneurysmal  
18 formation<sup>58,59</sup>. It is therefore possible that factors which reduce the risk of atherosclerosis  
19 can also decrease that of AAAs indirectly, even if not causally.

20

21 Only randomisation of exposure would provide more robust evidence than our review. One  
22 potential route forward may be through a combination of Mendelian randomisation and big  
23 data. Using a genetic variant as a randomly allocated instrument for the environmentally

1 modifiable exposure of alcohol intake<sup>60</sup>, or in some cases even non-genetic instrumental  
2 variables<sup>61</sup>, may provide some more robust insight that the traditional observational cohorts  
3 reviewed to date are unable to provide. Use of biomarkers for alcohol intake would also add  
4 a degree more objectivity to the exposure assessment.

5

6 In conclusion, though results from observational studies should be interpreted cautiously,  
7 this is the first study to quantify the dose-response curve for the association between  
8 alcohol consumption and AAA and find it to be similar to that seen in IHD.

9

10

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2 None.

3

4 **Disclosure**

5 The authors declare that there are no conflicts of interest.

6

7

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## 1 **Figure legends**

- 2 Figure 1: PRISMA Flowchart of article retrieval and screening process.
- 3 Figure 2: Forest Plot of highest versus lowest categories of alcohol consumption and AAA.
- 4 Figure 3: Forest Plot of linear dose-response analysis of alcohol consumption and AAA.
- 5 Figure 4: Nonlinear dose-response curve for alcohol consumption and AAA risk.