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Keywords: cluster randomised controlled trial; screening; prostate cancer; prostate cancer mortality; cause of death; death certification; sensitivity; specificity

Contemporary accuracy of death certificates for coding prostate cancer as a cause of death: Is reliance on death certification good enough? A comparison with blinded review by an independent cause of death evaluation committee

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Background: Accurate cause of death assignment is crucial for prostate cancer epidemiology and trials reporting prostate cancer-specific mortality outcomes.

Methods: We compared death certificate information with independent cause of death evaluation by an expert committee within a prostate cancer trial (2002–2015).

Results: Of 1236 deaths assessed, expert committee evaluation attributed 523 (42%) to prostate cancer, agreeing with death certificate cause of death in 1134 cases (92%, 95% CI: 90%, 93%). The sensitivity of death certificates in identifying prostate cancer deaths as classified by the committee was 91% (95% CI: 89%, 94%); specificity was 92% (95% CI: 90%, 94%). Sensitivity and specificity were lower where death occurred within 1 year of diagnosis, and where there was another primary cancer diagnosis.

Conclusions: UK death certificates accurately identify cause of death in men with prostate cancer, supporting their use in routine statistics. Possible differential misattribution by trial arm supports independent evaluation in randomised trials.

Prostate cancer is the second commonest cause of cancer death in UK men (UK National Screening Committee, 2015). Death certificates are used in routine mortality statistics, large-scale epidemiological studies and randomised controlled trials. However, prostate cancer can be misattributed as the underlying cause of death on death certificates in men diagnosed with prostate cancer (Feuer *et al*, 1999). A review of US medical records (Albertsen *et al*, 2000) suggested that 29% of men with prostate cancer as the

underlying cause of death on death certificates had died of some other condition (Albertsen *et al*, 2000).

The possibility of differential attribution bias in trials, where the primary end point is prostate cancer-specific mortality, is also a concern (Black *et al*, 2002). All-cause mortality is least open to bias (Black *et al*, 2002), but because prostate cancer death is relatively uncommon (UK National Screening Committee, 2015), all-cause mortality is less sensitive to the effects of screening.

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The possibility of attribution bias has led us and others (Miller et al, 2000; de Koning et al, 2003; Miller et al, 2015) to conclude that assignment of the underlying cause of death in prostate cancer trials must be confirmed by an independent expert (CoDE) committee. We have compared the underlying cause of death determined by an independent CoDE committee, with the underlying cause of death listed on official death certificates in UK men with prostate cancer participating in a UK-wide trial (Lane et al, 2014; Turner et al, 2014).

METHODS

Follow-up and identification of a prostate cancer-related event. All 413 000 men enrolled in the Cluster randomised trial of PSA testing for prostate cancer (CAP) trial have been traced and flagged for vital status follow-up at the Health and Social Care Information Centre (Turner *et al*, 2014). Blinded to death certificate and underlying cause of death, detailed information was obtained from the medical records of all men with a potential prostate cancer death (see Supplementary Table 1 for the triggers used to review a potential prostate cancer death), and used to generate a short structured clinical vignette (Williams *et al*, 2015); (see Supplementary Material 1).

Determination of cause of death. Members of an international CoDE committee reviewed the vignettes. They completed a questionnaire that when followed in sequence, and using detailed definitions adapted from the European Randomised Study of Screening for Prostate Cancer trial (de Koning et al, 2003) acted as an algorithm for assigning cause of death into the following categories: definite, probable, possible, unlikely or definitely not prostate cancer, and definite or probable intervention-related mortality (Supplementary Material 2). The committee was divided into three teams, each comprising four consultants from the following specialties: pathology, palliative care, urology and cancer surgery (see Figure 1 and Supplementary Material 1). The death certificate underlying cause of death was accepted where a review was not triggered (i.e., in the absence of a potential prostate cancer death, these were recorded as due to other causes).

Analysis. We examined the agreement between prostate cancer assigned as the underlying cause of death on the death certificate with prostate cancer (definite, probable or intervention-related) assigned as the cause of death after expert review. We calculated 'sensitivity' as the proportion of confirmed prostate cancer deaths as assigned by the expert review process (denominator), which were listed as an underlying cause of death of prostate cancer on the death certificate (numerator) (Box 1). We calculated 'specificity' as proportion of confirmed non-prostate cancer deaths as above. Each of sensitivity and specificity are accompanied by exact binomial (Clopper-Pearson) confidence intervals (Clopper and Pearson, 1934). We investigated whether sensitivity and specificity varied by age-group (splitting age at death into three groups of approximately equal size), the interval between date of diagnosis and date of death, and the presence of another cancer diagnosis. All parameters were calculated with their respective 95% confidence intervals.

Ethics. Ethical approval was provided by Trent Research Ethics Committee (MREC/01/4/025; MREC/03/4/093; 05/MRE04/78) and the Confidentiality Advisory Group (PIAG 4–09 (k)/2003; PIAG 1-05(f)/2006).

RESULTS

Over 50 000 deaths were notified to the CAP investigators by the Health and Social Care Information Centre between 2002 and

2015. Of these, 2069 men had died of a potential prostate cancerrelated death, and the underlying cause of death has been established for 1236 (60%) of these men to date.

Table 1 shows the number of deaths assigned as prostate cancer or other cause on the death certificates compared with the expert review process. Of a total of 1236 potential prostate cancer-related deaths, the independent CoDE committee attributed 523 (42%) to prostate cancer and 713 (58%) to other causes. The corresponding cause of death categories based on death certificates were 535 (43%) and 701 (57%), respectively. The expert committee agreed with death certificate derived underlying cause of death (prostate cancer or other) in 1134 cases (92% agreement, 95% CI: 90, 93%).

Eight per cent of deaths categorised as due to other causes after review of the case vignettes had been assigned as prostate cancer on death certificates (death certificate specificity: 92%; 95% CI: 90%, 94%). On the other hand, 9% of deaths classified as due to prostate cancer by the reviewers were assigned to other causes on the death certificates (death certificate sensitivity: 91%; 95% CI: 89%, 94%).

For men who died within 1 year of their diagnosis of prostate cancer, the death certificates had a sensitivity of 78% (95% CI: 69%, 85%) and specificity of 87% (95% CI: 80%, 93%); for men who died between 1 and 3 years of diagnosis, specificity was 87% (95% CI: 80%, 92%) and sensitivity was 93% (95% CI: 89%, 96%). The presence of another cancer diagnosis, either notified by Health and Social Care Information Centre or present on the death certificate, was also associated with a lower sensitivity (77%; 95% CI: 65%, 86%) and specificity (89%; 95% CI: 85%, 91%). The age at death had little impact on sensitivity or specificity. Three of the 1236 deaths were categorised as intervention-related deaths by the committee.

DISCUSSION

These data suggest that relying on underlying cause of death abstracted from official UK death certificates rather than an independent expert committee would result in some misattribution. Specifically, 9% of deaths assigned as being due to prostate cancer by the CoDE committee were recorded on death certificates as deaths from other causes (false negatives), and 8% of deaths considered on the death certificate to be due to prostate cancer were assigned to other causes (false positives) by the expert review. The impact of age was minimal, suggesting the use of UK death certificates could provide a relatively accurate means of evaluating population trends in prostate cancer mortality. However, where there was a death within 1 year of diagnosis of prostate cancer, both false positives (22%) and false negatives (13%) increased. This could reflect a tendency for competing causes of death to be less frequently considered by doctors completing death certificates when a prostate cancer diagnosis has only recently been made. The presence of another primary cancer, either on the death certificate or diagnosed as alive, also resulted in increased false positives (23%) and false negatives (11%). This could be because the clinical picture is unclear in these cases.

These results are based only on those deaths that were triggered for in-depth review by the expert committee because they were potential prostate cancer deaths. We did not review the other 49 000 deaths where there was no evidence of prostate cancer ever being diagnosed or where there was no evidence of other conditions that could have indicated a potential prostate cancer death, such as bone cancer (conceptualised as a potential misclassified bony metastasis). If all these other 49 000 deaths are

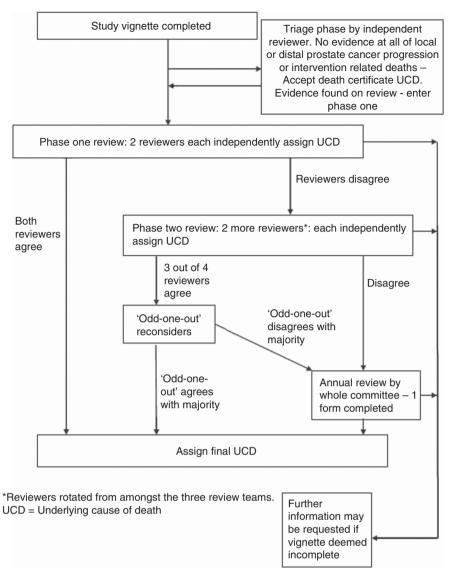


Figure 1. Process for evaluating cause of death.

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Sensitivity = \frac{N \text{ of PCa deaths (agreement between CoDE and death certification)}}{\text{Total } N \text{ of PCa deaths reviewed}} \times 100
\text{Specificity} = \frac{N \text{ of PCa deaths (agreement between CoDE and death certification)}}{\text{Total } N \text{ of deaths other causes (agreement between CoDE and death certification)}}} \times 100
\text{PCa} = \text{prostate cancer}
\text{CoDE} = \text{cause of death evaluation}
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correctly assumed not to be due to prostate cancer, this will have resulted in near perfect specificity for all deaths.

Similar level of agreement between death certification and expert review were observed in the USA (Albertsen *et al*, 2000), Sweden (Godtman *et al*, 2011) and Finland (Makinen *et al*, 2008). Common reasons for misclassification were cardiovascular or cancer comorbidities (Albertsen *et al*, 2000). In a recent study (Miller *et al*, 2015), agreement between death certificate and death review committee was >90%, but death certificated causes of death missed treatment-related deaths and the misattribution was differential by trial arm.

The study's strength was that it was based on large trials, we identified intervention-related deaths, and we successfully masked the trial arm from the expert committee (Williams *et al*, 2015), even though this was reported to be difficult in another trial (Barry *et al*, 2013). Limitations are that the results may not be generalisable beyond the cohorts included in the trials, and the assumption that CoDE results were near perfect in accuracy.

UK death certificates provide a relatively accurate means for evaluating cause of death that would be acceptable for routine public health monitoring and large-scale epidemiological studies.

Table 1. Sensitivity and specificity of prostate cancer as an underlying cause of death on the death certificate vs prostate cancer assigned as the underlying cause of death after expert review of clinical vignettes, stratified by age at death, time between diagnosis and death, and presence or absence of another primary cancer diagnosis (n = 1236)

1236	91 (89, 94); N=523	92 (90, 94); N=713
		72 (70, 7 4), IN = 7 13
287	93 (87, 97); N = 119	92 (86, 95); N=168
365	91 (85, 94); N = 169	92 (88, 96); N = 196
584	91 (87, 94); N = 235	92 (89, 95); N = 349
216ª	100 (54, 100); N=6	99 (97–100); N=210
231	87 (80, 93); N = 117	78 (69, 85); N = 114
390	93 (89, 96); N = 249	87 (80, 92); N = 141
399	92 (87, 96); N = 151	96 (92, 98); N = 248
369	77 (65, 86); N = 65	89 (85, 92); N = 304
867	93 (91, 96); N = 458	94 (91, 96); N = 409
	365 584 216 ^a 231 390 399	365 91 (85, 94); N = 169 584 91 (87, 94); N = 235 216a 100 (54, 100); N = 6 231 87 (80, 93); N = 117 390 93 (89, 96); N = 249 399 92 (87, 96); N = 151 369 77 (65, 86); N = 65 867 93 (91, 96); N = 458

In the context of randomised controlled trials, the potential for even a small amount of misattribution bias that is differential by trial arm means that an independent cause of death evaluation is likely to be necessary to provide unbiased outcome data.

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

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