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# The development and validation of a measure of eating disorder-specific interpersonal problems: The Interpersonal Relationships in Eating Disorders (IR-ED) scale

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**Abstract** 27

> Clinical reports suggest that interpersonal problems are associated with the onset and maintenance of eating pathology, but existing measures of such problems have limited links to eating pathology. Therefore, the aim of this study was to develop an eating-specific measure of interpersonal problems. The new measure, the Interpersonal Relationships in Eating Disorders scale (IR-ED), was administered to a large community sample, a non-clinical replication sample, and a clinical group of eating disorder patients. In Study 1, the psychometric properties of the IR-ED were established, and they were tested using confirmatory analyses in Study 2. Study 3 determined the validity of the test score interpretations in a clinical sample. The final 15-item version of the IR-ED demonstrated three distinct factors with reliability of test scores - Food-Related Isolation; Avoidance of Body Evaluation; and Food-Related Interpersonal Tension. Study 2 demonstrated that the IR-ED comprises a common Interpersonal Problems factor and a specific group factor - Avoidance of Body Evaluation. Study 3 showed that the clinical group had higher IR-ED scores than a non-clinical group. Across the studies, Avoidance of Body Evaluation was the strongest correlate of eating pathology in this group. The IR-ED has strong psychometric properties and its test scores appear to be more valid than those of a generic measure of interpersonal problems. Avoidance of Body Evaluation is the strongest facet of such interpersonal problems, and has meaningful links to models of eating psychopathology.

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words: eating; interpersonal problems; validation; measurement; assessment; psychometrics; eating disorders

# **Public Significance Statement**

Interpersonal problems are commonly reported by individuals with eating disorders, but clinicians have not previously had an evidence-based way of formulating or measuring such problems. This study provides a validated measure of interpersonal problems that clinicians can use to formulate eating problems and plan treatment for eating-disordered patients.

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# The development and validation of a measure of eating disorder-specific interpersonal problems: The Interpersonal Relationships in Eating Disorders (IR-ED) scale

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Interpersonal problems are difficulties in how people relate to, compare themselves to, or interact with others, and have been linked to many psychological difficulties (e.g., Barrett, & Barber, 2007; Eng, & Heimberg, 2006; Grisham, Steketee, & Frost, 2008; Kleiner, & Marshall, 1987; Lazarus, Cheavens, Festa, & Rosenthal, 2014). Clinical accounts and research suggest that interpersonal problems can influence and maintain non-clinical and clinical eating concerns and behaviours (e.g., Abraham, & Beumont, 1982; Broberg, Hjalmers, & Novenen, 2001; Lampard, Byrne, & McLean, 2011; Lieberman, Gauvin, Bukowski, & White, 2001; Murphy, Straebler, Basden, Cooper, & Fairburn, 2012; Rieger et al., 2010; Schmidt, & Treasure, 2006; Steiger et al., 1999; Tanofsky-Kraff, Wilfrey, & Spurrell, 2000). However, it is not clear whether interpersonal problems have any specificity to eating pathology or whether they are relatively generic to all mental disorders. There might be specific interpersonal difficulties linked to eating pathology that are not captured by generic measures of interpersonal problems (e.g., concern about others' evaluation of one's body). Identifying eating-specific interpersonal problems could be important in assessing, formulating, and treating eating disorders. For example, both interpersonal psychotherapy and cognitive behavioural therapy stress the need to address interpersonal issues where they maintain an eating disorder (e.g., Fairburn, Cooper, & Shafran, 2003; Wilfley et al., 2002).

If the link between eating pathology and interpersonal problems were non-specific, one would expect a generic measure of interpersonal problems to account for a substantial amount of variance in eating pathology. There are several such generic measures, such as the Interpersonal Relationship Inventory (Tilden, Nelson, & May, 1990) and the Interpersonal Relationship Scale (Guerney, 1977). However, most have not been considered for their utility when understanding eating disorders. When McEvoy, Burgess, Page, Nathan and Fursland (2013) used one of the most well-established non-specific measures (the Inventory of Interpersonal Problems–32 [IIP-32] - Barkham, Hardy, & Startup, 1996), they found that it has limited utility in detecting interpersonal problems in eating disorder patients. Five out of eight IIP-32 subscales were not related to variance in eating pathology. Raykos, McEvoy and Fursland (2017) have also demonstrated that generic socialising problems (as measured by the IIP-32) do not have a direct relationship with eating disorder symptoms.

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Therefore, it appears that generic interpersonal measures such as the IIP-32 do not adequately address eating-specific interpersonal problems, such as avoidance of others' judgements about one's appearance. Hence, it is important to consider whether a more specific measure of the interpersonal problems faced by people with eating disorders would have greater utility in understanding eating pathology. As no such measure exists, the aims of the first study are to detail the development and initial validation of a measure of interpersonal issues related to eating disorders (IR-ED) and to determine whether it has greater utility than generic measures of interpersonal problems. As detailed in Study 1, the IR-ED items were generated through discussion and revision by the authors, based on substantial experience in working with eating disorders. Inclusion of items was on the basis of clinical relevance, but avoiding redundant items. To demonstrate psychometric and clinical utility, the measure should: have a clear factor structure with adequate internal consistency of the resulting scales; show strong reliability of test-retest scores; be as strongly associated with non-eating pathology (anxiety, depression, social anxiety) as a generic measure of interpersonal problems; and be more strongly associated with eating pathology than a generic measure.

Study 1 aimed to develop the IR-ED measure based on factor analysis with a homogenous non-clinical female sample, and then to determine its initial utility with males and individuals with self-reported eating disorders. Further studies then aimed to replicate the measure's psychometric properties (Study 2) and to validate its scores with a clinical sample

(Study 3). The first hypothesis was that the IR-ED will have a clear and meaningful factor structure, which can be assessed for clinical utility and which can be compared and contrasted with a generic measure of interpersonal problems. The second hypothesis was that the IR-ED's psychometric properties will be replicable. The final hypothesis was that the IR-ED will show clinical utility among patients with eating disorders.

# STUDY 1 – DEVELOPMENT AND INITIAL VALIDATION OF THE MEASURE

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# **Participants**

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Participants were recruited using online survey methods, including university staff and student email lists and advertisements on Facebook and Twitter. They were not compensated in any form. Individuals who reported an eating disorder were not included, in order to ensure that any association with eating disorders in subsequent studies and analyses were not a product of bias introduced by such individuals in the development of the IR-ED. Five hundred and eighty-nine people logged onto the study. Fifty-eight (9.8%) dropped out before completing all measures, leaving 531 completers (393 female, 136 male, 2 no gender specified). A total of 261 completers consented to be contacted for a follow-up, with 142 participants completing the retest stage (54.4%). Exploratory analyses showed that there were no differences in stage 1 measures between those who did or did not agree to or actually undertake the second stage (p < .05 in all cases).

Overall, 31 participants (5.83%; 29 females) reported a current or past eating disorder diagnosis. Most women did not report having an eating disorder diagnosis (n = 364; M age = 33.13, SD = 11.38; M body mass index (BMI) = 25.10, SD = 6.49), with the remainder selfreporting an eating disorder diagnosis (n = 29; M age = 26.34, SD = 8.05; M BMI = 21.15, SD = 1.16). Likewise, most men did not report having an eating disorder diagnosis (n = 134; M age = 37.43, SD = 13.57; M BMI = 26.64, SD = 4.66), with the remainder self-reporting an eating disorder diagnosis (n = 2; M age = 23.00, SD = 7.07; M BMI = 31.50, SD = 9.73). Most of the sample was UK-based (91%), with the next largest contributor being the USA (2.64%).

#### Measures

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Following completion of demographic information (e.g., nationality, age, self-reported weight and height, history of diagnosis of an eating disorder), each participant completed the following six measures within the online survey (completed on Qualtrics).

**Interpersonal Relationships in Eating Disorders scale** (IR-ED). The IR-ED was developed for the purposes of this study. The measure asked participants to rate the extent of various interpersonal issues related to eating pathology, using a five-point Likert scale ranging from 1 ("Not at all") to 5 ("All the time"). A detailed description of the IR-ED's development is detailed in the Procedure.

Inventory of Interpersonal Problems (IIP-32, Barkham et al., 1996). The IIP-32 is a 32-item questionnaire that addresses interpersonal problems across eight domains. The global scale has acceptable reliability ( $\alpha = 0.87$ ) and test-retest reliability (r = 0.70; Barkham et al., 1996). Internal consistency for the global scale was high ( $\alpha = 0.93$ ) in the present study.

Eating Disorder Examination Questionnaire, version 6.0 (EDE-Q, Fairburn, 2008). The EDE-Q is a widely used measure of eating pathology. It has four attitudinal subscales: Restraint, Weight Concern, Shape Concern and Eating Concern. It has acceptable psychometric and clinical validity (e.g., Luce, & Crowther, 1999), particularly at the global score level. High internal consistency of the global EDE-Q was observed in the present sample ( $\alpha = 0.94$ ), and their mean score was 2.00 (SD = 1.31).

Brief Fear of Negative Evaluation Scale - Straightforwardly Worded (BFNE-S, Rodebaugh, Woods, Thissen, Heimberg, Chambless, & Rapee, 2004). The BFNE-S is an eight-item self-report measure of fear about being negatively evaluated. It contains straightforwardly worded items from the BFNE scale (Leary, 1983). The eight items are summed to create a total score with higher scores indicating a higher fear of negative evaluation. The BFNE-S has excellent internal consistency ( $\alpha = .92$ ) and strong construct validity in clinical samples (Weeks et al., 2005). The BFNE-S has demonstrated predictive utility for social anxiety symptoms as measured by the Liebowitz Social Anxiety Scale (Weeks et al., 2005). High internal consistency was observed in the present sample ( $\alpha = 0.95$ ), and their mean score was 23.0 (SD = 8.99).

Patient-Reported Outcomes Measurement Information System – Depression and Anxiety (PROMIS-D and PROMIS-A, Pilinkos et al., 2011). The PROMIS-D is an eight-item questionnaire measuring depression. It has high internal consistency ( $\alpha = 0.95$ ) and convergent validity (r = 0.83). The PROMIS-A is a seven-item questionnaire measuring anxiety (Pilinkos et al., 2011), with high internal consistency ( $\alpha = 0.93$ ) and convergent validity (r = 0.80). High internal consistency was observed in the present study for PROMIS-D ( $\alpha = 0.94$ ) and PROMIS-A ( $\alpha$  = .94). Their mean scores were 2.14 (SD = 0.92) and 2.27 (SD = 0.86), respectively.

### **Procedure**

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The IR-ED was developed through several iterations by the research team. An initial pool of 28 items was developed by SJ and GW based upon clinical experience and a prior literature search of interpersonal problems in eating disorders. The pool was shared with research colleagues in Australia (BR, AF, SB, PM), who revised and added items based upon their own clinical experience. The resultant pool consisted of 49 items, which were reviewed iteratively by the research teams, leading to similar items being omitted or merged. This iterative process led to a final pool of 26 items, which the research team as a whole reviewed and agreed on in terms of face validity. Each item is rated on a five-point Likert scale, where higher scores indicate a greater presence of the specific interpersonal issue over the past 28 days. After reading the information sheet and providing consent, participants completed all measures (Time 1). The IR-ED was completed again by a subset of participants two weeks later, to determine test-retest reliability (Time 2). Ethical approval for this study was provided by the relevant Ethical Review Committee.

#### **Data analysis**

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Initially, exploratory factor analysis (SPSS principal analysis factoring) was used to determine whether the IR-ED had a meaningful factor structure. To protect against heterogeneity influencing this analysis, only female participants were used for this analysis, and those with an eating disorder diagnosis were excluded, resulting in N = 364. This number of participants was well above the recommended guideline of five to 10 participants per questionnaire item (Gorsuch, 1983). Specific factors were not hypothesised a priori. Nor was it hypothesised that the resulting factors would be correlated. However, as that was a possible outcome, different rotations were compared (Varimax and Direct Oblimin), and the most appropriate model was used based on the coherence of the factors that emerged (i.e., the items could be conceptually grouped into meaningful scales). Factors were retained if they had an eigenvalue of >1.0 (Dancey & Reidy, 2004) and following visual inspection of scree plots and other characteristics (see below for further detail). Tang et al. (1998) recommend that individual items should be retained only if they load onto a specific factor by at least 0.4. For this study, a more stringent cut-off of 0.5 was used to ensure a more robust measure. Individual items were excluded if substantial cross-loading was detected (i.e., the difference in loadings between factors was less than 0.2), to ensure that the factors were as distinct as possible. Parallel analysis was conducted to exclude the possibility of inclusion of inappropriately weak factors, using the online engine (https://analytics.gonzaga.edu/parallelengine/) developed by Patil, Singh, Mishra, & Donavan (2008). The criteria set were 26 variables, 380 participants, 100 random correlation matrices, 95% percentile of eigenvalues, and 1000 seeds. Cronbach's alpha was used to determine the internal consistency of the emergent factors within the IR-ED.

The test-retest reliability of the IR-ED scores was analysed using intraclass correlations

and paired t-tests, based on the 35 male and 107 female participants who completed the study at Time 1 and Time 2 but who did not report any history of an eating disorder (thus excluding 12 participants for this analysis). Pearson's correlations were also used to determine the association of scores on the IR-ED and the IIP-32, based on the participants who completed both measures and who did not report any eating disorder history (n = 500). For those correlations, Bonferroni's correction was used to correct for multiple tests (n = 24), resulting in an acceptable alpha of .002.

Multiple regression analyses were used to determine the relative utility of the IR-ED and IIP-32 to explain variance in eating pathology (EDE-Q global score). This analysis was repeated to determine which of the IR-ED and IIP-32 scales predicted general psychopathology (anxiety, depression and fear of negative social evaluation). These analyses included all male (n = 134) and female (n = 364) participants who did not report an eating disorder diagnosis. Partial correlations were used to determine whether levels of anxiety, depression and fear of negative social evaluation were uniquely associated with IR-ED subscales, to exclude the possibility that apparent links between IIP-32 and IR-ED scores with eating pathology were indirect, and were actually due to associations of interpersonal problems with mood and anxiety symptoms. Again, Bonferroni's correction was used to correct for multiple correlations (n = 24), resulting in an acceptable alpha of .002.

**Results** 223

#### **Factor structure of the IR-ED**

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Table 1 shows the results of the factor analysis. A Varimax rotation provided the best solution (with strong, psychologically meaningful factors). It revealed three factors, based on 15 of the 26 original IR-ED items. The remaining items were excluded due to loading below 0.5 on all scales. No other items were excluded due to cross-loading, as none had loadings within 0.2 of the strongest factor loading. Three factors were chosen because they met all the

following criteria - eigenvalue > 1.0 (this applied to five factors that came before the scree point (four met this criterion, as eigenvalues for the first four factors were 4.341, 4.237, 3.069, 2.203 and 2.011, levelling off after that); accounted for a substantial additional cumulative amount of variance (three factors met this criterion, as the variance accounted for was 16.7%, 16.3%, 11.8%, 8.47% and 7.61%); and contained items that loaded most strongly on the relevant factor (this applied to three factors, as no items loaded most strongly on the final two of the five strongest factors). Parallel analysis was also conducted on the data set, and suggested that all of the five factors could have been included, but this was not done, given the lack of items loading on any factor after the third. However, this analysis did offer reassurance that the use of three factors was not excessive.

The first of the three factors was labelled Food-Related Isolation. It consisted of items 3, 5, 15, 18 and 22, and accounted for 16.70% of the variance in scores. The second factor was Avoidance of Body Evaluation, which contained items 1, 6, 9, 11, 21 and 24, and accounted for 16.30% of the variance in scores. The third factor was Food-Related Interpersonal Tension, which consisted of items 8, 12, 14 and 16, and accounted for 11.80% of variance in scores. All IR-ED scales had acceptable internal consistency (Table 1). IR-ED scales were all significantly correlated for this non-clinical group, in the moderate to strong range: Food-Related Isolation with Avoidance of Body Evaluation -r = .65, p < .001; Food-Related Isolation with Food-Related Interpersonal Tension -r = .55, p < .001; Avoidance of Body Evaluation with Food-Related Interpersonal Tension – r = .43, p < .001).

Finally, item mean scores on the three IR-ED scales were calculated (sum of the relevant items/number of items), and are reported in Table 1. A global score on the IR-ED was calculated from the mean of the three subscales -M = 1.50; SD = 0.62; range = 1.00-3.83. The final, 15item version of the IR-ED and scoring key are presented in Appendix 1.

#### Test-retest reliability of the IR-ED

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All IR-ED factors demonstrated significant (P < .001, in all cases) and strong intraclass correlations (ICCs) between time 1 and time 2 for males and females, as follows – Food-Related Isolation (males = .89; females = .90); Avoidance of Body Evaluation (males = .83; females = .90); and Food-Related Interpersonal Tension (males = .77; females = .87). There were no significant differences across time on the IR-ED subscales for either gender (Bonferroni corrected p > .05 in all cases), apart from the scores for females on the Avoidance of Body Evaluation scale, where there was a small but statistically significant rise in scores across the two weeks (M = 1.77, SD = 1.00 vs M = 1.88, SD = 1.07; t = 2.72; p = .008). Therefore, the IR-ED scores demonstrated broadly acceptable test-retest reliability.

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# Association between generic and eating-specific measures of interpersonal problems

Table 2 presents individual correlation coefficients between the subscales of the IR-ED and the IIP-32. As stated above, an acceptable alpha value of .002 was used to reduce the risk of Type 1 errors. Most IR-ED subscales were significantly associated with the IIP-32 subscales. However, the correlations were weak to moderate, suggesting that the IR-ED and IIP-32 were measuring relatively distinct constructs.

# Association of interpersonal problems with general and eating psychopathology

Correlations between measures of psychopathology and the IR-ED are presented in Table 3 for females and males separately. For the EDE-Q, partial correlations were conducted controlling for anxiety, depression and fear of negative evaluation (PROMIS-A, PROMIS-D, and BFNE-S scores), in order to understand the specific link between interpersonal problems and eating pathology. As stated above, the alpha value was set at .002 to reduce the risk of Type 1 errors. Most of the IR-ED subscales were significantly associated with the PROMIS-A, PROMIS-D, and BFNE-S. The IR-ED scales were also correlated with Global EDE-Q scores when anxiety, depression and fear of negative evaluation were controlled for. Correlations were mostly weak to moderate in strength.

To determine which aspects of interpersonal problems were associated most strongly with eating pathology, multiple regression analyses were used (for females and males separately). The individual scale scores of the IIP-32 and the IR-ED were entered simultaneously as predictors of the global EDE-Q score (see Table 4). For females, the three IR-ED scales were the strongest predictors of eating pathology, with only the IIP-32 Too Dependent scale contributing significantly from the more generic elements of interpersonal problems. IR-ED Avoidance of Body Evaluation was the strongest individual factor. For males, only Avoidance of Body Evaluation was associated with eating pathology. Thus, one interpersonal factor was the dominant concern relating to eating pathology - avoidance of people and situations due to body concerns.

Similar analyses were conducted for the three broader measures of psychopathology in the whole of this sample (full analyses available on request from the corresponding author). All three regression analyses showed a significant overall effect of the IR-ED and IIP-32 on the dependent variables (F > 15.0; p < .001; Adjusted  $R^2$  > 0.5 in all cases). In the case of depression (PROMIS-D), the IR-ED was the better predictor, with IR-ED Negative Body Evaluation (t = 3.45; beta = .314; p < .001) contributing most strongly to poorer mood. In contrast, for anxiety (PROMIS-A), the IIP-32 was the stronger predictor, with the IIP-32 Dependent scale being the most powerful correlated of anxiety (t = 3.30; beta = .275; p < .001). The same pattern was found for fear of negative evaluation (FNEB), where the IIP-32 Dependent scale was again the most powerful (t = 5.34; beta = .432; p < .001).

Discussion 300

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Using a non-clinical sample, the IR-ED demonstrated an acceptable three-factor solution which consisted of Food-Related Isolation, Avoidance of Body Evaluation and Food-Related Interpersonal Tension. The IR-ED was associated with a generic measure of interpersonal problems (IIP-32) and showed equivalence to the IIP-32 in relation to other areas

of psychopathology (anxiety, depression and social anxiety). Importantly, the IR-ED predicted more variance in eating pathology compared to the generic measure of interpersonal problems. Future research using a larger sample of men is needed to demonstrate whether the same factor structure holds among males as well as females.

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# STUDY 2 – CONFIRMATORY ANALYSIS OF THE IR-ED'S STRUCTURE

The first aim of study two was to cross-validate the IR-ED scores in an independent undergraduate sample using confirmatory factor analyses comparing unitary, uncorrelated three-factor, correlated three-factor, and bifactor models. The three-factor models assume the factors represent theoretically distinct constructs beyond the total scale, and therefore imply that the calculation of subscale scores will result in a more meaningful interpretation. However, it may be premature to interpret subscale scores as representing a meaningful construct distinct from a general interpersonal difficulty factor (Rodriguez, Reise, & Haviland, 2016). Interpersonal problems in response to disordered eating are diverse and, although the datadriven exploratory factor analysis in Study 1 statistically distinguished between three components of interpersonal problems, the majority of variance across these three factors may still be common (i.e., someone experiencing one type of interpersonal problem is likely to experience others). Such a pattern of findings would suggest that the boundaries between the problems assessed by the three factors in the IR-ED provide little independent or incremental utility compared to a total score. For instance, if the shared variance amongst the subscales explains most of the variance in subscale scores, then each subscale mostly reflects a single underlying 'interpersonal problems' construct. In this case, specifying distinct latent variables in models using the IR-ED may result in redundancy and multicollinearity problems. Conceptually, identifying a predominant underlying general interpersonal problems factor would be more parsimonious and may help to simplify case formulation and treatment planning. Alternatively, if each subscale assesses substantive unique group factors that are separate to the

general factor and have unique predictive utility, this may assist the development of more targeted interventions.

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Adopting a bifactor modelling approach is one way to inform researchers and clinicians on the most appropriate psychometric structure of the IR-ED, including whether total and/or subscale scores should be used when interpreting the measure (Reise, Moore, & Haviland 2010; Rodriguez et al., 2016). The bifactor measurement model stipulates that the variance in item responses can be accounted for by a general factor representing shared variance amongst all the items in addition to a set of group factors that explain variance beyond that explained by the general factor (Reise et al., 2010; Rodriguez et al., 2016). Therefore, the general factor reflects the broad construct a scale is attempting to measure, whilst the group factors reflect more specific subdomains. For the present study, bifactor modelling will assist in determining whether interpersonal problems in eating disorders are best conceptualised as unidimensional or multidimensional.

The second aim was to investigate whether the IR-ED can predict eating disorder symptoms, namely dietary restraint, shape concern, weight concern, and eating concern in a non-clinical sample, after controlling for more generic measures of interpersonal functioning. To determine whether the interpersonal problems measured by the IR-ED are specific to eating disorder symptoms, it was also important to control for co-morbid psychopathology, such as depression and anxiety, given that previous research has demonstrated these factors to be associated with interpersonal problems in eating disorders (Arcelus et al., 2013).

The first hypothesis was that a bifactor model would provide the best fit relative to the three-factor uncorrelated (orthogonal) model from study one, a three-factor correlated model (factors were free to correlate), and a unidimensional model. The rationale for this hypothesis was that interpersonal problems resulting from disordered eating are expected to co-occur and interact with each other, such that a substantial proportion of variance across the factors is shared. A bifactor model also enables the separability of specific interpersonal problems to be modelled, and it is plausible that the three factors will also explain a substantive proportion of unique reliable variance in subscale scores. The second hypothesis was that the IR-ED would uniquely predict eating disorder psychopathology (dietary restraint, shape concerns, weight concerns, eating concerns) after controlling for general interpersonal measures, depression, and anxiety in a non-clinical sample.

#### **Participants**

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Participants were recruited through convenience sampling through the School of Psychology and Speech Pathology's participant pool. They were not compensated for participation. MacCallum, Browne, and Sugawara's (1996) recommendations suggest that a minimum sample size of 200 should be adequate to test models with the degrees of freedom in the proposed models, although a larger sample will enable greater precision of parameter estimates. The final sample comprised 396 cases (306 females, 86 males, and 4 identifying as "other") with ages ranging from 17-69 years (M = 21.76, SD = 6.13).

#### **Measures and Procedure**

The same measures as Study 1 were administered in Study 2. This study was granted ethics approval from the relevant University Ethics Committee (RDHS-58-16). The questionnaires were administered online using Qualtrics. Participants were first required to read and indicate their agreement to an information sheet and consent form about the study. After completing the survey (time 1), participants were presented with a debriefing document that detailed the aims of the study, and were provided with the contact details of counselling services. Participants were asked to complete the IR-ED again after two-weeks (time 2). Participants received course credit for their participation.

#### **Data Analysis**

Preliminary data screening to assess normality, univariate and multivariate outliers,

multicollinearity and the normality, linearity, and homoscedasticity of residuals was performed in SPSS 23.0. As for study 1, the test-retest reliability of the IR-ED scores was analysed using intraclass correlations and paired t-tests. Confirmatory factor analysis (CFA) using maximum likelihood estimation was conducted in Mplus 8.0 (Muthén & Muthén, 1998-2017) to compare the relative fit of each competing IR-ED measurement model. The IR-ED bifactor model was compared to unidimensional, three-factor correlated, and three-factor uncorrelated models. A number of fit indices were used to evaluate the competing IR-ED models including the chisquare goodness of fit statistic ( $\chi^2$ ), comparative fit index (CFI), the Tucker-Lewis index (TLI), and the root mean square error of approximation (RMSEA) with 90% confidence intervals (CIs). Values greater than .90 and .95 for the CFI and TLI are indicative of a good and excellent fit, respectively (Hu & Bentler, 1999; Tabachnick & Fidell, 2013). For the RMSEA, values of .06 or less indicate a good-fitting model, with lower values corresponding with a closer fit, and the upper CI limit should not exceed .10 (Tabachnick & Fidell, 2013). Model comparisons were evaluated using chi-square difference tests.

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Several statistical indices were calculated to assess the degree to which the variance in the total and subscale scores could be attributed to variance associated with a single latent variable (Rodriguez et al., 2016). The coefficient omega (ω) represents the proportion of total score variance that is attributable to all common factors (i.e. both the general and group factors). Alternatively, coefficient omega hierarchal ( $\omega_H$ ) represents the percentage of variance in IR-ED total scores that is attributable to a single general factor. Explained common variance (ECV) reflects the percentage of common variance that can be explained by the general factor with higher values (greater than .70 or .80) suggesting the presence of a strong general factor in addition to providing support for the unidimensionality of the scale's items (Rodriguez et al., 2016). Item explained common variance (I-ECV) reflects the percentage of variance in each IR-ED item that is attributable to the general factor (Rodriguez et al., 2016). The percent uncontaminated correlations (PUC) can be used in conjunction with the ECV do determine the dimensionality of the model (Reise et al., 2010). PUC represents the proportion of IR-ED item covariance's that can be accounted for by the variance that is attributable to the general factor and group factors (Rodriguez et al., 2016). Reise and colleagues (2010) suggest that when PUC values are less than .80, ECV values are greater than .60, and ω<sub>H</sub> is greater than .70 then the multidimensionality within the data does not have enough impact to prohibit the interpretation of the model as unidimensional.

A structural equation model was used to assess if the group factors and general interpersonal problem factor accounted for unique variance in eating disorder symptoms (as indicated by the dietary restraint, weight concern, shape concern, and eating concerns subscale scores from the EDE-Q) beyond the variance accounted for by more generic measures of interpersonal problems (BFNE-S and IIP-32). The model was run again controlling for depression and anxiety. Confidence intervals (95%) were calculated around the parameter estimates of the final model using 1000 bootstrapping resamples.

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#### **Preliminary Analyses**

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No data were missing on key outcome variables (IR-ED, EDEQ). Missing data at Time 1 were observed for 12 cases (22 missing values in total) on depression and anxiety measures. Little's Missing Completely At Random (MCAR) test was not significant,  $\gamma^2(125) = 127.63$ , p = .42, so missing data were imputed using expectation-maximization in SPSS. Histograms of model variables were inspected for normality. All measures demonstrated acceptable skewness (< 2) and kurtosis (< 7), thereby satisfying the assumption of normality (Tabachnick & Fidell, 2013). Inspection of box plots and Mahalanobis Distance revealed no problematic univariate or multivariate outliers. The assumptions of linearity and homoscedasticity of residuals were met. In addition to descriptive statistics, Table 5 shows that, in the current sample, scores on all measures demonstrated excellent internal consistency, and the measures were moderately and positively correlated with each other.

# Test-retest reliability of the IR-ED

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A total of 304 participants (242 females, 59 males, 3 no gender given) provided re-test data for the IR-ED at time 2. Participants who did versus did not provide time 2 data did not significantly differ on age, gender, or any measure (all ps > .29). All IR-ED factors demonstrated significant and strong ICCs between time 1 and time 2 for males and females, as follows – Total score (males - r = .80, females - r = .90, Food-Related Isolation (males - r = .67; females - r = .89); Avoidance of Body Evaluation (males - r = .84; females - r = .89); and Food-Related Interpersonal Tension (males - r = .81; females - r = .87). For females, there was a significant but small reduction in Avoidance of Body Evaluation (M = 2.22, SD = 1.00 vs M = 2.13, SD = .99, p = .02), but changes were small and non-significant for the other subscales and total score (mean changes = -.02 to .72, ps = .06 to .63). For males, changes on total and subscale scores were very small (mean change ranged from -.003 to .025) and non-significant (ps = .68-.96). Therefore, scores on the IR-ED demonstrated broadly acceptable test-retest reliability and stability.

# **IR-ED Measurement Models**

The unidimensional,  $\chi^2(90) = 555.25$ , CFI = .782, TLI = .746, RMSEA = .114 (90% CI = .105-.123), and uncorrelated three-factor,  $\chi^2(90) = 476.83$ , CFI = .819, TLI = .789, RMSEA = .104 (90% CI = .095-.113), models provided a poor fit to the data. The correlated three-factor model,  $\chi^2(87) = 175.97$ , CFI = .958, TLI = .950, RMSEA = .051 (90% CI = .040–.061), and the bifactor model,  $\chi^2(75) = 129.06$ , CFI = .975, TLI = .965, RMSEA = .042 (90% CI = .030– .055), provided an excellent fit to the data. A significant chi-square difference test indicated that the bifactor model fit the data significantly better than the correlated three-factor model,  $\Delta \chi^2(12) = 46.91$ , p < .001. The standardized factor loadings for the one-factor, three-factor uncorrelated, three-factor correlated, and bifactor models are presented in Table 6.

A majority of the IR-ED items had stronger loadings on the general factor than on the group factors with the exception of three items from group factor two (1, 6, and 7). Loadings were greater than .50 on the general factor, indicating that the items mostly represent the general interpersonal problems factor and support computation of a total score (Reise et al., 2010). The omega (ω) coefficients for the general IR-ED factor and group factors were high, ranging from .85 to .95. Omega hierarchal (ω<sub>H</sub>) suggested that 82.3% of variance in IR-ED total scores can be accounted for by individual differences on the general factor. Group factor one (Food Related Isolation) explained very little variance (5.6%) in the subscale scores independent from the general IR-ED factor. Evidence of some multidimensionality was found as group factor two (Avoidance of Body Evaluation) and group factor three (Food-related Interpersonal Tension) accounted for a moderate proportion of subscale score variance (48.0% and 32.6%, respectively). The general factor accounted for 68.7% of the common variance whilst 31.3% of the common variance was attributable to the three group factors. The ECV value provided support for a robust general factor, though failed to reach the benchmark (>.70) needed to unambiguously indicate unidimensionality. A majority (66.7%) of the IR-ED items had I-ECV values less than .80, signifying they are poorer indicators of the general IR-ED factor and contribute more to the variance in their respective group factors. The PUC demonstrated that the general factor accounted for 70.5% of the item correlations in the IR-ED. Furthermore, the average relative parameter bias across the IR-ED items was acceptable (11.6%).

## **Structural Equation Models**

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Due to the inability to rule out multidimensionality from the CFAs, the bifactor model was employed in all structural models. An initial measurement model with the IR-ED bifactor model plus the eating disorder symptoms measurement model, but without any freed covariances between latent variables, provided a poor fit to the data,  $\chi^2(137) = 465.745$ , p <

.001, CFI = .908, TLI = .885, RMSEA = .078 (90% CI = .070-.086). The initial structural model with the general factor and three group factors predicting eating disorder symptoms did not converge. Given that group factor 1 explained a very small proportion of unique variance in the bifactor measurement model, it was removed as a predictor of eating disorder symptoms. This time the model converged, providing an excellent fit,  $\chi^2(134) = 275.252$ , p < .001, CFI = .960, TLI = .949, RMSEA = .048 (90% CI = .043-.060). However, factor 3 of the IR-ED was a nonsignificant predictor of eating disorder symptoms so this pathway was removed and the model rerun. This final model also provided an excellent fit to the data,  $\chi^2(135) = 275.660$ , p < .001, CFI = .961, TLI = .950, RMSEA = .048 (90% CI = .043-.060). The model accounted for 55.9%  $(R^2)$  of the variance in eating disorder symptoms.

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To examine the independent contribution of the IR-ED beyond the BFNE-S and IIP-32, these measures were entered in the model as unique predictors of global EDE-Q. First, we tested a measurement model with the IR-ED bifactor model, EDE-Q, BFNE-S, and IIP-32 measurement models without the structural pathways, which provided a poor fit to the data,  $\chi^2(545) = 1932.514$ , p < .001, CFI = .826, TLI = .810, RMSEA = .080 (90% CI = .076-.084). The structural pathways were then freed, and the general factor and second group factor of the IR-ED were freed to correlate with BFNE-S and IIP-32. This model provided an adequate fit to the data,  $\chi^2(536) = 1300.67$ , p < .001, CFI = .904, TLI = .894, RMSEA = .060 (90% CI = .056-.064). The modification indices were observed to identify sources of model strain and the largest modification index (118) was between items 5 and 6 of the BFNE-S. Both of these items begin with the same sentence structure "I'm afraid that..." which is indicative of common method variance. The residual variances of these two items were freed to correlate and the model was rerun. No further modifications were deemed theoretically defensible. Results revealed an acceptable fit,  $\chi^2(535) = 1189.31$ , p < .001, CFI = .918, TLI = .909, RMSEA = .056 (90% CI = .051-.060) and demonstrated the IR-ED general factor (b = .50, 95% CI = .35-.64), IR-ED group factor two (b = .47, 95% CI = .32-.61), BFNE-S (b = .34, 95% CI = .22-.46), and IIP-32 (b = -.21, 95% CI = -.37- -.06) all uniquely predicted global EDE (see Figure 1). The model explained 61% (R<sup>2</sup>) of the variance in eating disorder symptoms. Finally, the model was rerun controlling for depression and anxiety symptoms. The relationships between the IR-ED, IR-ED group factor two, BFNE-S, IIP-32 and eating disorder symptoms remained significant after controlling for depression and anxiety. Depression and anxiety were not significantly associated with eating disorder symptoms in this model, so the penultimate model was preferred.

**Discussion** 

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Confirmatory Factor Analysis using an independent non-clinical sample suggested that the IR-ED contains a strong general interpersonal problems factor, which can be assessed using a total IR-ED score, and that is related to eating pathology. Variance in the Food-Related Interpersonal Tension and Food-Related Isolation subscales is mostly accounted for by the general factor, so these subscales cannot be meaningfully differentiated from the total score and therefore should not be calculated. In contrast, Avoidance of Body Evaluation appears to be a distinct factor and therefore could be considered separately from the total score to inform formulation and subsequent intervention.

#### STUDY 3 – CLINICAL VALIDATION OF THE IR-ED

Using a clinical sample with diagnosed eating disorders, the aims of Study 3 were to: (a) report descriptive statistics for the IR-ED, relative to the scores of non-clinical individuals; (b) report bivariate correlations between the IR-ED and clinical characteristics (fear of negative evaluation, anxiety, depression, generic interpersonal problems, and eating disorder symptoms); and (c) demonstrate unique variance between the IR-ED subscales and eating disorder symptoms. Although findings from Study 2 using an undergraduate sample suggested that a total score should be used, in Study 3 we examined the IR-ED total and subscale scores

within a clinical sample as future psychometric investigations with larger clinical samples might reveal greater differentiation between the subscales than was found in the non-clinical sample.

## **Participants**

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The clinical sample comprised 107 patients (96% female; 71% Anglo-Australian ethnicity; 55% employed) who were consecutively referred to a public mental health service with a dedicated outpatient eating disorders service. All patients had a confirmed DSM-5 eating disorder diagnosis (32% anorexia nervosa, 37% bulimia nervosa, 25% other specified feeding or eating disorder, 6% unspecified feeding or eating disorder). Patients ranged in age from 16 to 63 years (M = 24.3 years, SD = 9.5 years) and illness duration ranged from 4 months to 51 years (M = 5.9 years; SD = 8.1 years). Exclusion criteria included current psychosis, schizophrenia, or schizoaffective disorder, significant alcohol or substance abuse/dependence, medical instability, or BMI below 14 kg/m2. This study received approval from the Institution's Human Research Ethics Committee (Approval Number QI 2014/39) and all patients provided written informed consent for their data to be included. No patients were compensated for taking part.

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# **Procedure**

As part of routine clinical practice, patients attended an assessment at the clinic, which included completion of self-report measures and administration of the Eating Disorder Examination interview (EDE Version 12; Fairburn & Cooper, 1993). The EDE was administered by clinical psychologists trained in its administration and specialising in eating disorder treatment. The EDE scores have good convergent and concurrent validity, have good inter-rater reliability, and discriminate well between groups with and without an eating disorder (Berg et al., 2012; Fairburn & Cooper, 1993).

#### **Measures**

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The measures from Study 2 were administered in Study 3.

#### **Data Analysis**

Mean IR-ED total and subscale scores were compared between the undergraduate sample in Study 2 and the clinical sample in Study 3 using a MANOVA (to correct for any potential intercorrelations of the IR-ED scales). Pearson's bivariate correlation coefficients were also calculated between the IR-ED and BFNE-S, PROMIS anxiety and depression scales, IIP-32, and EDE-Q global scores. An initial multiple regression analysis was then conducted to examine the proportion of variance in EDE-Q Global scores that could be explained by the IR-ED subscales, and the unique predictive utility of each subscale. BFNE, PROMIS anxiety and depression, and IIP-32 total scale scores were then added to the model to investigate whether IR-ED subscales continued to explain unique variance in EDE-Q Global scores.

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> The means (SDs) for IR-ED Total score, Food-Related Isolation, Avoidance of Body Evaluation, and Food-Related Interpersonal Tension scales were 3.12 (0.93), 3.05 (0.98), 3.13 (0.98), and 3.15 (1.15), respectively. The MANOVA used to compare the student sample (from Study 2) with the clinical sample (this study) showed a significant overall effect (F(3,499) =84.5; p < .001; partial eta<sup>2</sup> = 0.337). Correcting for any intercorrelations, the clinical group had significantly higher mean scores on: Food-Related Isolation (F(1,156.6) = 248.6; p < .001; partial eta<sup>2</sup> = 0.337); Avoidance of Body Evaluation (F(1,100.0) = 113.1; p < .001; partial eta<sup>2</sup> = 0.184); and Food-Related Interpersonal Tension (F(1,151.4) = 175.5, p < .001; partial eta<sup>2</sup> = 0.259). All effect sizes were very large. Table 7 shows that the IR-ED subscales were significantly and positively correlated with the BFNE-S, PROMIS anxiety and depression scales, IIP-32 subscales and EDE-Q global.

Multiple regression analysis was used to test if the three IR-ED scales significantly

predicted EDE-Q Global scores. Together, the three predictors explained 46.0% of the variance in EDE-Q Global (adjusted  $R^2 = .46$ ), F(3,102) = 30.77, p < .001. However, the only significant individual predictor variable was IR-ED Avoidance of Body Evaluation subscale score (B = .75, 95% CI = .43-1.07, SE B = .16,  $\beta$  = .57, p < .001). IR-ED Food-related isolation (B = .21, 95% CI = -.19 - .60, SE B = .30,  $\beta$  = .21, p = .31) and Food-related interpersonal tension (B = -.01, 95% CI = -.26 - .24, SE B = .13,  $\beta$  = .01, p = .92) did not contribute significantly to the prediction of EDE-Q Global. When BFNE-S, IIP-32 Total, PROMIS anxiety and depression scores were included in the regression analysis, the predictor variables together explained 59% of the variance in EDE-Q Global (adjusted  $R^2 = .59$ ), F(7.98) = 22.23, p < .001. IR-ED Avoidance of Body Evaluation (B = .71, 95% CI = .41 – 1.00, SE B = .15,  $\beta$  = .53, p < .001), BFNE-S (B = .03, 95% CI = .01 – .06, SE B = .01,  $\beta$  = .22, p = .009), IIP-32 Total (B = -.54, 95% CI = -.97 - -.11, SE B = .22,  $\beta = .27$ , p = .01), and anxiety (B = .03, 95% CI = .003 - .06, SE B = .01,  $\beta$  = .22, p = .029) emerged as significant individual predictors. IR-ED Food-related isolation (B = .08, 95% CI = -.28 – .44, SE B = .18,  $\beta$  = .06, p = .439) and IR-ED Food-related interpersonal tension (B = .02, 95% CI = -.21 - .24, SE B = .11,  $\beta$  = .01, p = .140) did not contribute significantly to the prediction of EDE-Q Global.<sup>1</sup>

<sup>&</sup>lt;sup>1</sup> Exploratory regression analyses were conducted separately in patients with primary AN and primary BN with the IR-ED subscales predicting EDE-Q Global scores. For patients with primary AN, the three IR-ED subscales explained 67% of the variance in EDE-Q Global (adjusted  $R^2 = .67$ ), F(3,31) = 21.61, p < .001. As for the full sample, Avoidance of Body Evaluation subscale score was a significant individual predictor (B = 1.30, 95% CI = .69-1.91, SE B = .30,  $\beta$  = .97, p < .001), but Food-related isolation (B = -.03, 95% CI = -.76 - .71, SE B = .36,  $\beta$  = -.02, p = .94) and Food-related interpersonal tension (B = -.21, 95% CI = -.63 - .21, SE B = .21,  $\beta$  = -.18, p = .31) were not. For patients with primary BN, the three IR-ED subscales explained 37% of the variance in EDE-Q Global (adjusted  $R^2 = .37$ ), F(3,35) = 21.61, p < .001. Avoidance of Body Evaluation subscale (B = .33, 95% CI = -.08 - .75, SE B

#### **Discussion**

A sample of individuals diagnosed with an eating disorder scored significantly higher on all IR-ED subscales compared to the non-clinical sample from study two. The IR-ED was also able to predict a large amount of variance in eating pathology beyond generic interpersonal problems, fear of negative evaluation, anxiety and depression, which was accounted for by the Avoidance of Body Evaluation subscale. This study suggests that the IR-ED has a unique and positive association with eating pathology in a clinical sample.

#### **GENERAL DISCUSSION**

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Various models of eating disorders stress the importance of understanding interpersonal problems in assessing, formulating, and treating eating disorders (e.g., Fairburn et al., 2003; Lampard et al., 2011; Murphy et al., 2012; Rieger et al., 2010; Steiger et al., 1999). However, existing generic measures of interpersonal problems are only weakly associated with eating disorder pathology (e.g., McEvoy et al., 2013; Raykos et al., 2017). This series of studies aimed to develop a measure of interpersonal problems that is specific to eating pathology, and to determine whether this focus resulted in greater utility. The resulting measure – the IR-ED – demonstrates strong psychometric properties (factor structure, internal consistency, test-retest reliability), is moderately associated with a generic measure of interpersonal problems (the IIP-32), and has comparable associations to the IIP-32 with other pathologies (e.g., anxiety). However, the IR-ED scores have much greater clinical validity for use in eating disorders than

<sup>= .20,</sup>  $\beta$  = .30, p = .11), Food-related isolation (B = .45, 95% CI = -.04 - .94, SE B = .24,  $\beta$  = .43, p = .07), and Food-related interpersonal tension (B = -.05, 95% CI = -.40 - .30, SE B = .17,  $\beta$  = -.06, p = .77) were not statistically significant predictors. These post-hoc exploratory analyses within the AN and BN subsamples must be interpreted cautiously due to low power, and they must be replicated in larger samples.

the IIP-32, with stronger associations with eating attitudes and self-reported diagnoses in a nonclinical sample, and with confirmed eating disorder diagnoses in a clinical sample. These findings suggest that basing a measure of interpersonal problems on the specific disorder being considered may be more effective than using a generic measure. However, whether this improvement applies to other disorders is a matter for empirical investigation.

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In Study 1, three correlated subscales emerged from the IR-ED - Food-Related Isolation, Avoidance of Body Evaluation, and Food-Related Interpersonal Tension – reflecting different domains of interpersonal problems relative to eating pathology. Food-Related Isolation captures a theme of not wanting to eat in front of others and the secrecy that can surround eating for people who experience eating pathology. Such concerns relate to the roles of shame and self-criticism in the maintenance cycle of eating disorders (e.g., Danakalis et al., 2016), impacting upon an individual's likelihood to isolate themselves when eating and to disengage from social activities where eating may be likely. Avoidance of Body Evaluation relates to themes of social withdrawal from activities or scenarios where an individual's body may be viewed or evaluated by others, as shown to be relevant to the onset and maintenance of eating pathology (e.g., Fairburn, 2008) via the impact of restriction and efforts to manage body size and weight. Food-Related Interpersonal Tension relates to how an individual's eating behaviours influence the way other people interact with and behave towards them. For example, people with eating disorders can isolate themselves and get involved in interpersonal role disputes (e.g., Murphy et al., 2012).

Study 2 found evidence of a bifactor structure for the IR-ED, consisting of a general interpersonal problem factor and one group factor (Avoidance of Body Evaluation). Avoidance of Body Evaluation uniquely explained 48% of its respective subscale variance beyond that explained by the general IR-ED factor, and contained three items that loaded higher onto the group factor rather than the general factor. This group factor also demonstrated acceptable reliability, suggesting it reflects a well-defined and stable subscale. These findings indicate that Avoidance of Body Evaluation is a distinct construct from general interpersonal problems in eating disorders, and should be considered as such when formulating a case and adapting interventions for clients. In contrast, the Food-related Isolation and Food-related Interpersonal Tension subscales appear to be ill-defined and unstable, suggesting that the items within these subscales should be integrated within the total score but not be scored separately. During case formulation, treatment planning and treatment outcome evaluation, clinicians may choose to use a total score as a broad assessment of eating disorder related interpersonal problems, but they also have the option to use the Avoidance of Body Evaluation subscale score to capture a specific interpersonal problem that uniquely predicts eating disorder symptoms. It is important to emphasise that the items measured by the other two IR-ED subscales still appear to assess important components of interpersonal problems related to eating disorders, as evidenced by their strong loading on the general IR-ED factor, but scoring them separately from the total score is not meaningful.

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Study three found evidence of discriminant validity of the IR-ED. Patients with a confirmed eating disorder scored significantly higher overall and on all three IR-ED scales than individuals in the community sample (Study 2). Consistent with findings from the two community samples, Avoidance of Body Evaluation emerged as the strongest predictor of eating pathology. These findings are consistent with the earlier conclusion that Avoidance of Body Evaluation is a distinct construct from generic interpersonal problems that may be important to consider in the assessment of patients with eating disorders. The specificity of this interpersonal problem is highly amenable to case formulation (e.g., as a specific element within the 'Life' element of Fairburn et al.'s [2003] transdiagnostic model, potentially explaining the over-evaluation of shape) and treatment planning (e.g., exposure-based methods, behavioural experiments, or imagery re-scripting interventions that directly address avoidance of body evaluation may augment current evidence-supported protocols for eating disorders).

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This series of studies has several strengths, including initial validation and crossvalidation in independent samples, and preliminary results within a clinical sample with eating disorders. To summarise, convergent validity of the IR-ED scores is demonstrated by their strong Cronbach's alphas ( $\geq$  .8 in all cases) and the associations of the scores with the other inventories. Discriminant validity of the scores was shown by the size of the correlations between the IR-ED scales (r = .43-.63), the differences between males and females, and the differences between those with and without eating disorders. However, there are several ways in which the research needs to be consolidated and extended. These will include confirmation of the factor structure within a larger clinical sample, and determining the utility of the measure among younger people, larger samples of males, and in different cultures. Confirmation of the factor structure within a clinical sample is particularly critical before findings using the three subscales in Study 3 can be considered reliable. For completeness, and in case the three subscales are found to be more separable in future clinical samples than they were in our undergraduate sample (Study 2), we investigated all three subscales as predictors of eating disorder symptoms. However, if future studies find that the IR-ED is best considered unidimensional in clinical samples then only a total score should be used as a predictor.

There is also a need to investigate whether there are differences between different diagnostic groups. Post-hoc exploratory analyses suggested that there may be differences across individuals with principal AN versus BN diagnoses (footnote 1), although the small sample sizes militate against strong conclusions being drawn from this study. Future research with larger samples is required to examine differences.

It is important that future research evaluates measurement invariance of the IR-ED across groups (e.g., clinical and non-clinical samples, different eating disorders, males and

females), to ensure that items are being interpreted in similar ways and that any identified differences can be meaningfully interpreted. Finally, it should also be considered in future research whether such a measure should be based on clinically-generated items that reflect interpersonal problems found specifically in those with eating disorders, by adding items generated by patients with eating disorders themselves.

The IR-ED has the potential to be a valuable tool across tasks and therapies. Initial assessment of interpersonal problems specific to eating and body concerns might indicate ways of understanding the origins and maintenance of eating disorders. However, the IR-ED might also be used to identify central or supplementary targets for interventions (e.g., interpersonal issues to address in interpersonal psychotherapy or in enhanced CBT), as well as potential moderators (e.g., is there a need to individualise therapy to enhance its impact, as suggested within existing protocols?). If this is the case, then the IR-ED might be used to evaluate progress in treatment and prevention programmes (e.g., as an index of the outcome of stigma-reduction programmes), as well as in initial identification of interpersonal problems.

#### **Conclusions**

This study showed evidence of construct and convergent validity, as well as internal consistency and test-retest reliability, for a measure of eating-specific interpersonal problems – the IR-ED. Compared to more generic measures of interpersonal problems, the IR-ED provides greater insight into eating-specific interpersonal problems. Future research should aim to confirm the structure in a clinical sample and investigate the roles of eating-specific interpersonal problems in the onset, maintenance, and treatment of eating disorders.

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#### **Conflict of Interest**

The authors have no conflict of interest to declare.

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Table 1 Principal Analysis Factoring (Varimax rotation) of the IR-ED measure for females who did not report an eating disorder diagnosis (n = 364), with item mean scores and internal consistency of resulting scales.

	Items	Factor 1 Food-Related Isolation	Factor 2 Avoidance of Body Evaluation	Factor 3 Food-Related Interpersonal Tension
1	I find it hard to spend time with others because I	.408	.730	.109
	worry what they think about my body	.400	.750	.10)
2	I worry what others would think of my if they knew how I eat	.568	.478	.312
3	I avoid social situations where eating is involved	.729	.306	.161
4	My appearance allows me to stand out amongst my peers	.131	.189	.096
5	I avoid getting into conversations with others about food	.649	.266	.196
6	I avoid socialising with people who are likely to comment on my body or appearance	.353	.610	.168
7	Eating the way I do helps me to cope with my anxiety in social situations	.363	.366	.296
8	Other people try to pressure me into eating differently	.221	.174	.647
9	I avoid intimacy because I worry what others will think of my body	.168	.712	.121
10	Others admire my ability to control what I eat	.103	088	.067
11	I avoid certain activities that would mean other people might judge my body	.138	.746	.179
12	My pattern of eating often leads to disagreements or tension with others	.317	.123	.690
13	My appearance helps me feel that I fit in and am more accepted by others	033	053	.068
14	Other people try to pressure me into changing my appearance	011	.295	.551
15	My eating patterns make it hard for me to socialise as much as I would like to	.725	.286	.145
16	Other people worry about what I eat	.275	.049	.736

17	When I experience tension with others, I focus			
	more on controlling my eating / weight	.292	.279	.304
18	I prefer to eat alone to avoid conflict with others			
	about what I eat	.630	.269	.381
19	I worry that I spend too much time taking with			
	other people about my appearance	.126	.397	.235
20	Controlling my weight helps me to feel more			
	confident in social situations	.160	.318	.163
21	Worrying about my weight and appearance makes it			
	difficult to feel really "connected" when I am with	.376	.656	.147
	other people			
22	My eating patterns cause me to withdraw from			
	others	.705	.262	.289
23	Eating the way I do makes it more likely that others			
	will show concern for me	.369	111	.511
24	It is difficult to meet new people as I worry they are		-0-	
	judging me or my appearance	.352	.697	.145
25	My eating patterns cause me to be secretive or	502	245	407
	deceptive with others	.583	.245	.407
26	Eating the way I do helps me to communicate my	222	117	205
	feelings and needs to others	.232	.117	.285
	Eigenvalue	4.341	4.237	3.069
	Variance explained (Rotated)	16.70%	16.30%	11.80%
	Cronbach's alpha	0.838	0.892	0.800
	Item mean	1.38	1.69	1.42
	(SD)	(0.71)	(0.89)	(0.62)

Note. Items where loadings are in bold were retained in that factor in the final version of the IR-ED.

Table 2 Pearson's correlations (r) between the IR-ED and IIP-32, for all participants who did not report an eating disorder diagnosis (n=500)

				IIP-32 scale							
	Hard to be	Hard to be	Too	Too	Too	Hard to be	Hard to be	Too			
IR-ED scale	sociable	assertive	aggressive	open	caring	supportive	involved	dependent			
Food-Related Isolation	.53*	.31*	.22*	21*	.35*	.27*	.43*	.35*			
Avoidance of Body Evaluation	.56*	.36*	.30*	18*	.41*	.33*	.51*	.51*			
Food-Related Interpersonal Tension	.26*	.15*	.23*	.02	.23*	.19*	.18*	.30*			

Note. <sup>a</sup> IR-ED = Interpersonal Relationships in Eating Disorders; IIP-32 = Inventory of Interpersonal Problems. \*p < .001

Table 3 Pearson's correlations (r) between the IR-ED and measures of fear of negative evaluation, depression, anxiety and eating disorder psychopathology

		Female (n = 364) Male				(n = 134)		
	Food-	Avoidance Food		Food-	Avoidance	Food-Related		
	Related	of Body	Related	Related	of Body	Interpersonal		
	Isolation	Evaluation	Interpersonal	Isolation	Evaluation	Tension		
			Tension					
BFNE-S	.45*	.56*	.30*	.31*	.54*	.17		
PROMIS-anxiety	.40*	.50*	.28*	.23	.47*	.14		
PROMIS-depression	.40*	.58*	.26*	.39*	.58*	.26*		
EDE-Q Global	.40*a	.54* a	.34* a	.31* a	.34* a	.23 a		

Note. <sup>a</sup> Partial Pearson's correlations controlling for PROMIS-A, PROMIS-D and BFNE-S. PROMIS-

D = Patient-Reported Outcomes Measurement Information System – Depression; PROMIS-A =

Patient-Reported Outcomes Measurement Information System – Anxiety; BFNE-S = Brief Fear of

Negative Evaluation-Straightforwardly worded scale; EDE-Q = Eating Disorders Examination

809 Questionnaire. \* p < .001

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811 Table 4

Multiple regressions using Interpersonal Relationships in Eating Disorders scores (IR-ED) and 812

Inventory of Interpersonal Problems scores (IIP-32) as predictors of Eating Disorder Examination 813

Questionnaire global score (EDE-Q) 814

Independent variables	t	p	Beta
<b>Females:</b> N = 364, F(df = 11,363) = 39.8, P < .001, Adjusted R $^2$ = .54			
IR-ED Food-Related Isolation	3.04	.003	.160
IR-ED Avoidance of Body Evaluation	9.66	.001	.537
IR-ED Food-Related Interpersonal Tension	2.34	.02	.100
IIP Hard to be Sociable	2.16	.04	122
IIP Hard to be Assertive	0.26	NS	013
IIP Too Aggressive	0.04	NS	002
IIP Too Open	0.79	NS	033
IIP Too Caring	2.41	.02	102
IIP Hard to be Supportive	0.30	NS	014
IIP Hard to be Involved	1.46	NS	074
IIP Too Dependent	3.37	.001	.174
<b>Males:</b> $N = 134$ , $F(df = 11,132) = 5.47$ , $P < .001$ , Adjusted $R^2 = .27$			
IR-ED Food-Related Isolation	1.80	NS	.224
IR-ED Avoidance of Body Evaluation	2.89	.005	.338
IR-ED Food-Related Interpersonal Tension	0.28	NS	.031
IIP Hard to be Sociable	1.42	.04	.162
IIP Hard to be Assertive	1.86	NS	191
IIP Too Aggressive	0.40	NS	.035
IIP Too Open	0.14	NS	.013
IIP Too Caring	1.49	.02	.135
IIP Hard to be Supportive	0.67	NS	078
IIP Hard to be Involved	1.18	NS	156
IIP Too Dependent	1.15	NS	.124

Table 5 816 Descriptive Statistics, Bivariate Correlations, and Internal Consistencies in the 817 undergraduate sample (Study 2) 818

	Mean	SD	1	2	3	4	5	6
1. IR-ED total	1.83	.77	.93					
2. EDE-Q Global	2.00	1.46	.66*	.96				
3. IIP-32	1.18	.65	.66*	.47*	.94			
4. BFNE-S	22.29	8.52	.57*	.59*	.62*	.95		
5. PROMIS-Dep	18.01	7.90	.56*	.51*	.58*	.59*	.95	
6. PROMIS-Anx	19.29	8.00	.54*	.49*	.58*	.57*	.82*	.95

Note. Internal consistencies are on the diagonals. SD = standard deviation; IR-ED =

Interpersonal Relationships – Eating Disorders; EDE-Q Global = Eating Disorder

Examination Global Score; IIP = Inventory of Interpersonal Problems; BFNE-S = Brief Fear

of Negative Evaluation Straightforwardly Worded; PROMIS-Dep = Patient Reported

Outcome Measurement Information System (PROMIS): Depression and Anxiety. \*p < .001.

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Table 6. Standardized Factor Loadings for the Measurement Models of the Interpersonal Relationships – Eating Disorders

			ree-fac correla			ree-fac orrelate			Bifact	tor model	
Item		F1	F2	F3	F1	F2	F3	General	F1	F2	F3
1. I find it hard to spend time with others because I worry what they think about my body	.67		.78			.78		.53		.57	
2. I avoid social situations where eating is involved	.670	.72			.72			.71	.12		
3. I avoid getting into conversations with others about food	.67	.71			.72			.71	.09		
4. I avoid socialising with people who are likely to comment on my body or appearance	.71		.74			.75		.61		.43	
5. Other people try to pressure me into eating differently	.64			.82			.81	.64			.50
6. I avoid intimacy because I worry what others will think of my body	.66		.79			.78		.52		.60	
7. I avoid certain activities that would mean other people might judge my body	.63		.77			.76		.48		.62	
8. My pattern of eating often leads to disagreements or tension with others	.67			.71			.77	.72			.26
9. Other people try to pressure me into changing my appearance	.53			.60			.59	.45			.41
10. My eating patterns make it hard for me to socialise as much as I would like to	.75	.80			.80			.76	.21		
11. Other people worry about what I eat	.61			.82			.79	.60			.58
12. I prefer to eat alone to avoid conflict with others about what I eat	.75	.76			.78			.80	.00		
13. Worrying about my weight and appearance makes it difficult to feel really "connected" when I am with other people	.75		.78			.79		.65		.45	

(Table continues)

			ree-fac correla			ree-fact orrelate			Bifac	tor model	
Item		F1	F2	F3	F1	F2	F3	General	F1	F2	F3
14. My eating patterns cause me to withdraw from others	.80	.88			.86			.81	.51		
15. It is difficult to meet new people as I worry they are judging me or my appearance	.72		.82			.82		.59		.57	
Coefficient omega								$\omega$ =.95	ωs=.90	ωs=.91	$\omega$ s=.85
	ECV							.93	.29	.72	.52
	PUC							.71			

Note. N = 396. F1 = food-related isolation factor; F2 = avoidance of body evaluation factor; F3 = food related interpersonal tension factor,

 $\omega$  = omega;  $\omega$ s= omega subscale;  $\omega_H$  = omegaH; ECV = explained common variance; PUC = percent uncontaminated correlations.

Table 7

Spearman's rho correlations between the IR-ED and the BFNE-S, anxiety, depression, and IIP-32

in the clinical sample (N = 107)

BFNE-S .38*** .40*** PROMIS-anxiety .39*** .40***	body evaluation .41*** .38***	interpersonal tension .20* .30***
	.41*** .38***	.20*
	.38***	
PROMIS-anxiety .39*** .40***		.30***
•	48***	
PROMIS-depression .47*** .47***		.34**
EDE-Q Global .67*** .60***	.68***	.43***
IIP-32		
Hard to be sociable .55*** .56***	.57***	.34***
Hard to be assertive .32** .36***	.42***	.13
Too aggressive .38*** .38***	.28**	.35***
Too open26**26**	29*	22
Too caring .25* .26*	.29**	.13
Hard to be supportive .38*** .36***	.37***	.31**
Hard to be involved .60*** .58***	.62***	.41***
Too dependent .30** .29**	.33**	.19 <sup>a</sup>

Note. BFNE = Brief Fear of Negative Evaluation scale, IR-ED = Interpersonal Relationships in Eating Disorders; IIP-32 = Inventory of Interpersonal Problems; EDE-Q = Eating Disorder Examination Questionnaire; PROMIS = Patient-Reported Outcomes Measurement Information System – Depression and Anxiety scales.  $^ap = .05 *p < .05 **p < .01 ***p < .001$ 

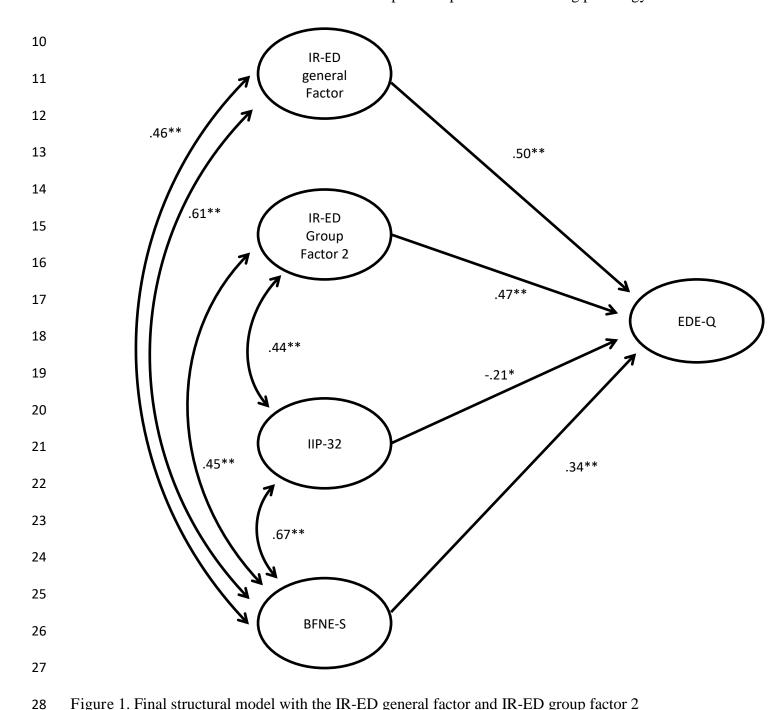


Figure 1. Final structural model with the IR-ED general factor and IR-ED group factor 2 predicting EDE-Q, controlling for general IIP-32 and BFNE-S. IR-ED = Interpersonal 29 Relationships in Eating Disorder (Group factor 2 = Avoidance of Body Evaluation), IIP = 30 Inventory of Interpersonal Problems, BFNE = Brief Fear of Negative Evaluation Scale – 31 straightforwardly worded items version, EDE-Q = Eating Disorders Examination 32 Questionnaire. Parameters are standardised. Only significant pathways are included. 33 \*p < .01, \*\*p < .001. 34

# **Appendix 1: Interpersonal Relationships – Eating Disorders: 15-item version**

Instructions: Thinking about your experiences with others over the past 28 days, how much would you 36

say that the following statements applied to you?

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		Not at all	A little bit	Moderately	Quite a bit	All the time
1.	I find it hard to spend time with others because I worry what they think about my body					
2.	I avoid social situations where eating is involved					
3.	I avoid getting into conversations with others about food					
4.	I avoid socialising with people who are likely to comment on my body or appearance					
5.	Other people try to pressure me into eating differently					
6.	I avoid intimacy because I worry what others will think of my body					
7.	I avoid certain activities that would mean other people might judge my body					
8.	My pattern of eating often leads to disagreements or tension with others					
9.	Other people try to pressure me into changing my eating					
10.	My eating patterns make it hard for me to socialise as much as I would like to					
11.	Other people worry about what I eat					
	I prefer to eat alone to avoid conflict with others about what I eat					
13.	Worrying about my weight and appearance makes it difficult to feel really "connected" when I am with other people					
14.	My eating patterns cause me to withdraw from others					
15.	It is difficult to meet new people as I worry they are judging me or my appearance					

# Scoring key (item means)

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Food-Related Isolation scale: Total items 2, 3, 10, 12 and 14, and divide by 5 41 42

Avoidance of Body Evaluation scale: Total items 1, 4, 6, 7, 13 and 15, and divide by 6

Food-Related Interpersonal Tension scale: Total items 5, 8, 9 and 11, and divide by 4