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**Article:**

Draper, FC, Baraloto, C, Brodrick, PG et al. (13 more authors) (2019) Imaging spectroscopy predicts variable distance decay across contrasting Amazonian tree communities. *Journal of Ecology*, 107 (2). pp. 696-710. ISSN 0022-0477

<https://doi.org/10.1111/1365-2745.13067>

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1 **Imaging spectroscopy predicts variable**  
2 **distance decay across contrasting**  
3 **Amazonian tree communities**

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18

19 **Summary**

20 1. The forests of Amazonia are among the most biodiverse on Earth, yet accurately quantifying  
21 how species composition varies through space (i.e. beta-diversity) remains a significant  
22 challenge. Here we use high-fidelity airborne imaging spectroscopy from the Carnegie Airborne  
23 Observatory to quantify a key component of beta-diversity, the distance decay in species  
24 similarity through space, across three landscapes in Northern Peru. We then compared our  
25 derived distance decay relationships to theoretical expectations obtained from a Poisson Cluster  
26 Process, known to match well with empirical distance decay relationships at local scales.

27 2. We used an unsupervised machine learning approach to estimate spatial turnover in species  
28 composition from the imaging spectroscopy data. We first validated this approach across two  
29 landscapes using an independent dataset of forest composition in 49 forest census plots (0.1-1.5  
30 ha). We then applied our approach to three landscapes, which together represented terra firme  
31 clay forest, seasonally-flooded forest and white-sand forest. We finally used our approach to  
32 quantify landscape-scale distance decay relationships and compared these with theoretical  
33 distance decay relationships derived from a Poisson Cluster Process.

34 3. We found a significant correlation of similarity metrics between spectral data and forest plot  
35 data, suggesting that beta-diversity within and among forest types can be accurately estimated  
36 from airborne spectroscopic data using our unsupervised approach. We also found that estimated  
37 distance decay in species similarity varied among forest types, with seasonally-flooded forests  
38 showing stronger distance decay than white-sand and terra firme forests. Finally, we  
39 demonstrated that distance decay relationships derived from the theoretical Poisson Cluster  
40 Process compare poorly with our empirical relationships.

41 4. Synthesis: Our results demonstrate the efficacy of using high-fidelity imaging spectroscopy to  
42 estimate beta-diversity and continuous distance decay in lowland tropical forests. Furthermore,  
43 our findings suggest that distance decay relationships vary substantially among forest types,  
44 which has important implications for conserving these valuable ecosystems. Finally, we  
45 demonstrate that a theoretical Poisson Cluster Process poorly predicts distance decay in species  
46 similarity as conspecific aggregation occurs across a range of nested scales within larger  
47 landscapes.

48

49

## 50 **Introduction**

51 The forests of Amazonia are highly diverse, supporting as many as 16,000 tree species (ter  
52 Steege et al., 2013). The importance of this diversity, beyond its intrinsic value as a natural  
53 wonder, is increasingly well documented, for example, by underpinning key biogeochemical  
54 cycles and determining the resilience of Amazonian forests to climate change (Sakschewski et al.  
55 2016). Despite this recognition of the importance of diversity, accurately quantifying how  
56 species composition varies through space (i.e., beta-diversity) in Amazonia remains a significant  
57 challenge given the remoteness of the largest tropical forest on Earth. Over recent years large  
58 networks of forest plots (e.g. RAINFOR, ATDN and CTFS) have provided invaluable insight  
59 into the spatial ecology of Amazon forests (Duque et al., 2017; Phillips et al., 2004; ter Steege et  
60 al., 2006). However, even summed together these networks represent only ~2000 ha of forest,  
61 with many plots in localised clusters. Therefore, using plot data alone to assess continuous  
62 spatial phenomena such as turnover in species composition represents a significant current  
63 limitation to understanding tropical biodiversity.

64 An alternative, yet complementary, approach to quantifying biodiversity is through the use of  
65 remotely sensed data integrated with existing plot data. Such an approach enables the acquisition  
66 of contiguous data over vast swaths of forests irrespective of accessibility, potentially  
67 transforming the power of an entirely ground-based approach. Multispectral data from satellite  
68 based remote sensing, in conjunction with plot data, has been used successfully to broadly  
69 classify different forest types (Draper et al., 2014; Salova et al., 2005) and to provide general  
70 assessment of species turnover in Amazonia (Thessler, 2008; Tuomisto et al., 2003). However,  
71 current satellite based multispectral sensors (e.g. Landsat) lack the spatial and spectral resolution  
72 required to sufficiently differentiate the high species-level diversity occurring within tropical  
73 forests (Rocchini, 2007a, 2007b; Rocchini et al., 2016). Recent advances in high-fidelity, laser  
74 guided imaging spectroscopy present a viable solution, and have been used successfully to  
75 estimate beta-diversity in Neotropical forests (Féret & Asner 2014a; b; Somers et al. 2015).

76 A key component of beta-diversity is the variation in species composition as a function of  
77 geographic distance (hereafter referred to as distance decay). Distance decay is a particularly  
78 useful concept as it allows for an understanding of the relative importance of different processes

79 that may determine patterns of beta-diversity, such as environmental filtering and dispersal  
80 limitation (Soininen, McDonald, & Hillebrand 2007; Tuomisto, Ruokolainen, & Yli-Halla 2003).  
81 Understanding variation in distance decay relationships among different landscapes and forest  
82 types also has important implications for designing effective conservation strategies (Socolar et  
83 al. 2016). For example, the gradient of distance decay can help to understand if conserving  
84 species in a given landscape or forest type will be maximised by many small or few large  
85 protected areas (Nekola & White 1999). Furthermore, distance decay relationships can be used to  
86 formally test theoretical predictions of community assembly, for example from neutral theory  
87 and sampling area theory (Chave & Leigh, 2002; Condit et al., 2002; Hubbell, 2001; Morlon et  
88 al., 2008).

89 One particularly significant theoretical model suggests that distance decay relationships are  
90 defined by the spatial aggregation of tree species, which can be characterized by a Poisson  
91 Cluster Process (hereafter PCP) (Plotkin et al. 2000; Morlon et al. 2008). This model is useful  
92 because it correctly recognises that tree species are spatially aggregated (Condit et al. 2000), but  
93 does not attempt to ascribe a particular community assembly mechanism. Furthermore, this  
94 model has accurately characterised species area curves, and distance decay relationships in  
95 number of tropical forests (Plotkin et al. 2000; Morlon et al. 2008). Importantly, while this  
96 model has found relatively good agreement at small scales ( $\leq 50$  ha), it has not been possible to  
97 test this model at larger spatial scales. A key limitation of the PCP approach is that it assumes a  
98 single scale of aggregation, in this paper we test the validity of this assumption at larger spatial  
99 scales ( $>1000$  ha).

100 Within western Amazonia, several plot based studies have examined distance decay relationships  
101 in tree communities, and most of these studies find an initial rapid decay in species similarity  
102 over the first few kilometres followed by a far more gradual decay over greater distances (Condit  
103 et al., 2002; Duque et al., 2009; Tuomisto et al., 2003). However, this relationship varies  
104 substantially with the spatial scale of study (Morlon et al., 2008; Phillips et al., 2003; Tuomisto  
105 et al., 2003), forest type (Draper et al., 2018), underlying geology (Phillips et al., 2003) and  
106 taxonomic group (Kristiansen et al., 2012; Tuomisto et al., 2003). Importantly, all of these plot  
107 based studies have been data limited, either using a relatively small number of plots (typically  $<$   
108 50 ha) to interpolate distance decay over tens to hundreds of kilometres (Condit et al., 2002;

109 Tuomisto et al., 2003), or using spatially continuous data to investigate distance decay over small  
110 spatial scales ( $\leq 50$  ha) (May et al., 2016; Morlon et al., 2008).

111 Here we apply a sequence of unsupervised machine learning techniques (Féret & Asner 2014b)  
112 to continuous high-fidelity spectral datasets to quantify contiguous beta-diversity and associated  
113 distance decay relationships at a landscape scale ( $>1000$  ha) across three lowland landscapes in  
114 Amazonian Peru. At each of these landscapes we apply our method to one of three distinct forest  
115 types: white-sand forest, seasonally-flooded forest and terra firme clay forest. We also use an  
116 extensive network of 49 forest census plots across two landscapes to thoroughly validate our  
117 approach and to answer the following questions.

- 118 1. Does high spatial resolution imaging spectroscopy accurately predict turnover in tree  
119 species composition across different forest types in lowland Amazonia?
- 120 2. How does distance decay in tree species composition vary across different forest types in  
121 lowland Amazonia?
- 122 3. How well does a theoretical PCP predict distance decay in tree species composition across  
123 a range of forest types in lowland Amazonia?

124

125

## 126 **Methods**

### 127 Study landscapes

128 Three distinct landscapes were used in this study: Allpahuayo Mishana, Jenaro Herrera, and  
129 Quebrada Braga. These landscapes are all located in the department of Loreto, Peru (Fig. 1), and  
130 were selected because they harbour at least one of the three most common forest types  
131 encountered across western Amazonia: terra firme clay forest, seasonally-flooded forest, and  
132 white-sand forests (Baraloto et al. 2011). The first landscape, Allpahuayo Mishana, is a national  
133 reserve located close to the city of Iquitos that contain a mosaic of terra firme clay and white-  
134 sand forest (Fine et al., 2010; García Villacorta et al., 2003). These white-sand forests have  
135 exceptionally nutrient poor sandy soils of cratonic origin, and harbour numerous endemic tree  
136 species (Fine et al. 2010). The second landscape, Jenaro Herrera, is a centre of research of the  
137 Instituto de Investigaciones de la Amazonía Peruana (IIAP). Jenaro Herrera is made up primarily  
138 of terra firme forest, although there are some small patches of white-sand forest, seasonally-  
139 flooded forest, and palm swamp forest (Honorio Coronado et al., 2008; Honorio Coronado et al.,  
140 2009). Finally, the Quebrada Braga landscape is located south of Jenaro Herrera, and is  
141 surrounded by the Ucayali river on three sides, these low-lying forests are inundated seasonally  
142 with nutrient-rich white water (Nebel et al. 2001).

### 143 Airborne data

144 We used the Carnegie Airborne Observatory (CAO) Airborne Taxonomic Mapping System  
145 (AToMS) to obtain fused high fidelity imaging spectroscopy and Light Detection and Ranging  
146 (LiDAR) data for all three of our landscapes (Asner et al. 2012). CAO flights took place between  
147 June and September 2012 at an altitude of approximately 2000 above ground level, with an  
148 average flight speed of  $60 \text{ m s}^{-1}$ , and a mapping swath of  $\sim 1.2 \text{ km}$ . Spectral radiance data were  
149 collected between 380 and 2510 nm at 5 nm increments (Asner et al. 2012). These measurements  
150 were subsequently resampled to 10-nm resolution, resulting in 214 contiguous spectral bands at a  
151 ground-level resolution (pixel size) of 2 m. LiDAR data were obtained from a dual laser  
152 waveform scanner that was operated at 200 kHz, with a  $17^\circ$  scan half-angle from nadir, yielding  
153 a point density of 4 laser shots  $\text{m}^{-2}$  (up to 16 returns  $\text{m}^{-2}$ ). Lidar data were used to produce maps  
154 of tree canopy height and ground surface at 1-m spatial resolution. Spectral and LiDAR data

155 were precisely geo-located using an embedded high resolution Global Positioning System-  
156 Inertial Measurement Unit (GPS-IMU).

157 The spectral radiance data were atmospherically corrected to apparent surface reflectance with  
158 the ACORN-5 model (ImSpec LLC, Glendale, CA USA). Images were then processed to exclude  
159 pixels that were not fully sunlit (i.e. shaded by another tree), covered by cloud, or represented a  
160 non-forested land surface. Shade masks were built using LiDAR-derived ray tracing models  
161 (Asner et al. 2007), clouds were masked manually, and non-forested land surfaces were  
162 identified using a LiDAR derived map of tree canopy height where pixels with a canopy < 3 m  
163 were considered non-forested. In addition, spectral bands that contained sampling noise  
164 (wavelengths < 400 nm and > 2500 nm) or that were dominated by atmospheric water vapour  
165 (wavelengths 1350-1480 nm and 1780-2032 nm), were not used in this analysis.

166 Estimating beta-diversity from spectral data

167 To estimate beta-diversity from spectral data, we used the ‘spectral species distribution’ (SSD)  
168 approach, building on the previous work of Féret & Asner (2014 a, b) and more generally on the  
169 foundations of the spectral variation hypothesis (Palmer et al., 2002). Our approach assumes that  
170 the spectral properties of a landscape vary with species composition, and therefore we are able to  
171 use variation in spectral composition as a proxy for variation in species composition. At each of  
172 the three sites, we independently applied a seven-step analysis procedure to generate our mapped  
173 estimates of tree species compositional change as follows.

174 (1) We performed a principal component analysis (PCA) on our processed spectral image in  
175 order to reduce the high dimensionality of the spectral data and to isolate and remove  
176 sampling artefacts such as cross-track brightness gradients.

177 (2) We manually selected components associated with biological gradients by visually  
178 examining the first 35 components, and removing any that showed obvious artefacts,  
179 such as clear striping. This left 4-8 useful components that were used in steps 3-7. At all  
180 landscapes the first three components were always selected and the together the  
181 components represented >60% of the variance.

182 (3) We applied k-means clustering to the selected components, clustering each pixel into one  
183 of 50 possible ‘spectral species’. Spectral species being simply clusters of pixels that  
184 have similar reflectance values, which may, but equally may not, trace onto actual

185 species. This process reduces the multi-layer image of PCs into a single layer image  
186 containing the spatial distribution of spectral species. Due to the large size of the dataset,  
187 k-means was applied using the ‘mini-batch k-means’ function in the Python package  
188 scikit learn, which provides near-equivalent performance at rapid computational speed  
189 (Pedregosa et al. 2012). Mini-batches of 10,000 pixels were used, each with 20 random  
190 starts.

191 (4) We then divided the resulting spectral species distribution image into 1 ha mapping  
192 kernels. Kernels in which > 66 % of pixels corresponded to either shade, non-vegetated  
193 ground, or were clouded were excluded from all further analysis. This led to a ~20 % loss  
194 of area from each landscape (Table 1).

195 (5) We then converted the image into a spectral species abundance matrix where each row  
196 corresponded to an individual kernel and each column to a spectral species, from which  
197 we calculated a Bray-Curtis distance matrix.

198 (6) We then applied Non-metric multi-dimensional scaling (NMDS) to the distance matrix in  
199 order to extract the most important compositional gradients in the spectral species data.  
200 The NMDS was optimized for three axes and run for 30 iterations.

201 (7) Finally, we re-projected the three NMDS axis scores into a raster format so that spatial  
202 variation in spectral species composition could be visualized.

203 The PCA and k-means analysis were undertaken using the Python package Sci-kit learn  
204 (Pedregosa et al. 2012). All beta-diversity analyses (steps 5 and 6) were performed in the R  
205 statistical environment using the Vegan package (Oksanen et al. 2013).

206 Plot inventory beta-diversity estimates

207 To validate our approach, we compared our estimates of beta-diversity derived from spectral data  
208 to measured beta-diversity obtained from inventory plot data at Allpahuayo Mishana and Jenaro  
209 Herrera. Our plot dataset consisted of 37 existing forest inventory plots distributed across  
210 Allpahuayo Mishana in white-sand and terra firme forest types, and 12 forest plots distributed  
211 across Jenaro Herrera in terra firme, white-sand and palm swamp forest types (Fig. 2). Plots  
212 varied in size from 0.1-1.5 ha, and five different sampling protocols were used as described  
213 below.

214 We used 12 large rectangular permanent sampling plots (0.5 to 1.5 ha), in which all tree stems  
215 with a diameter greater than 10 cm have been tagged and identified. Seven of these rectangular  
216 plots were one ha in size and belong to the RAINFOR Network, two of these plots were 1.5 ha in  
217 size (Peacock et al., 2007; Martinez & Phillips 2000). We also used three rectangular 0.5 ha plots  
218 in which all stems greater than 5 cm have been identified (Honorio Coronado et al., 2008). We  
219 further used 16 small 0.1 ha plots, in which all stems greater than 2.5 cm in diameter were  
220 identified. Six of these 0.1 ha plots were ‘Gentry’ plots consisting of ten 2 x 50 m intersecting  
221 transects (Gentry, 1982; Phillips et al., 2003). These six Gentry plots, alongside the seven 1 ha  
222 RAINFOR plots were downloaded from the ForestPlots.net online repository (Lopez-Gonzalez  
223 et al., 2009; Lopez-Gonzalez et al., 2011). The ten remaining 0.1 ha plots were rectangular 20 x  
224 50 m plots (Zárate et al., 2006). We used four 0.5 ha modified Gentry plots, within which all  
225 stems greater than 2.5 cm in diameter were identified (Baraloto et al. 2011). The remaining 14  
226 plots were circular plots in which all species greater than 10 cm dbh were identified (Baldeck et  
227 al., 2016); two of these circular plots were 0.25 ha and 12 were 0.14 ha. Summary details of the  
228 inventory plot dataset are given in Table 2, and full details of all plots are given in table S.1.

229 GPS coordinates were taken in the centre of each plot to determine its position within the  
230 landscape. There are significant uncertainties associated with using a GPS underneath a forest  
231 canopy, particularly for smaller inventory plots. Our approach partially mitigates these  
232 uncertainties as our aim is to align these plots with spectral species composition estimates at a 1  
233 ha scale, and therefore, GPS locations need only be located in the correct 1 ha kernel. Ultimately,  
234 we removed five plots from this aggregate dataset in Allpahuayo Mishana (four 0.1 ha and one  
235 0.5 ha), that were located < 10 m from a kernel boundary between white-sand forest and terra  
236 firme forest according to our spectrally derived map of estimated beta-diversity. As these plots  
237 were larger than 10 m in any dimension, there is a high likelihood that much of the area of these  
238 plots was situated in an incorrect kernel. These five boundary plots introduced additional  
239 variation in the relationship, as shown in Figure 3.

240 Because morpho-species were not standardised across datasets, it was necessary to exclude all  
241 individuals not identified to species level from the dataset before calculating beta-diversity.  
242 These exclusions led to a loss of 5-20% of individuals, which is likely to slightly increase the  
243 similarity among plots. However, patterns of beta-diversity among Amazonian tree census plots

244 have been shown to be generally robust to the exclusion of similar proportions of morpho-  
245 species (Pos et al. 2014).

246 Given that estimates of beta-diversity are sensitive to the number of individuals per plot, and that  
247 our dataset was made up of plots of different sizes (and different numbers of individuals), it was  
248 necessary to standardise our plot dataset by stem number before calculating beta-diversity. We  
249 did this by using a bootstrap resampling process. This process consisted of first establishing the  
250 minimum number of individuals in any plot, in this case 65, and then sampling (without  
251 replacement) 65 individuals from each plot at random. A Bray-Curtis distance matrix was then  
252 constructed using this subsample of 65 individuals per plot. Using this distance matrix, NMDS  
253 ordinations were performed. NMDS axis scores were then extracted for each plot. This process  
254 was then repeated 1000 times with a different set of 65 individuals per plot in order to develop  
255 confidence intervals for NMDS axis scores. Finally, we were able to compare NMDS axis scores  
256 derived from this plot inventory data with the corresponding NMDS axis scores derived from the  
257 spectral data.

#### 258 Estimating spectral distance decay

259 To estimate the distance decay in species composition from spectral data within forest types, it  
260 was first necessary to isolate pixels that correspond to the forest type of interest. At Allpahuayo  
261 Mishana, the target forest type was white-sand forest. Using our validation data, we  
262 demonstrated that at this site white-sand forests can be readily separated from terra firme forests  
263 based on spectral composition (Fig. 2 and 3). Therefore, pixels with a value of greater than 0.3  
264 on the first NMDS axis were classified as white-sand forest.

265 At Jenaro Herrera the target forest type was terra firme forest. We first used our spectral data to  
266 exclude small patches of white-sand forest from our analysis; to do this, we excluded all pixels  
267 with a value of greater than 0.2 on the second NMDS axis as this was shown to represent white  
268 sand forests in the validation data (Fig. 3). We then used the LiDAR derived DEM to separate  
269 pixels of seasonally-flooded forest from terra firme forest. Kernels with a mean elevation greater  
270 than 118 m were considered to be terra firme forest.

271 At Quebrada Braga the target forest type was seasonally-flooded forest. We used our LiDAR  
272 derived DEM to isolate those forests that are seasonally-flooded from those that are not. We

273 were able to use existing plot data to identify the elevation of seasonally-flooded forests (Kvist &  
274 Nebel 2001; Nebel et al. 2001). All kernels that had a mean elevation of 113-117 m a.s.l. were  
275 deemed to be seasonally-flooded. As this landscape is surrounded on three sides by a white-  
276 water river, we assume that seasonal flooding provides uniformly high nutrient deposition and  
277 that there are no further edaphic gradients.

278 To visualise the distance decay across each landscape we calculated the mean similarity (inverse  
279 Bray Curtis) for all paired plots within bins of 100 m, (i.e. the mean similarity between plots  
280 located 0-100 m apart, 100-200 m apart etc.). We have presented the ensemble mean and  
281 standard deviation with each distance bin and do not assume independence among these pairwise  
282 distances. Additionally, we calculated the first order derivative of similarity every 100 m across  
283 each landscape. We used a LOESS smoothing function (span = 0.35), to demonstrate how the  
284 derivative varies with distance across each landscape.

285 Theoretical distance decay

286 To assess the extent to which our empirical spectral distance decay relationships could be  
287 reproduced by a PCP, we applied the theoretical framework outlined by Morlon et al. (2008).  
288 Because we applied this approach to 50 spectral species rather than hundreds or thousands of  
289 species, it was essential that our measure of similarity was calculated using abundance rather  
290 than occurrence data. Therefore, we did not fit the general formula supplied by Morlon et al.  
291 (2008) which had been developed to using the Sorensen index. Instead, we simulated maps of  
292 spectral species distributions with a PCP, which we parameterised using fits of Ripley's K curves  
293 to our spectral species maps. Subsequently, we were able to derive abundance-based distance  
294 decay relationships from these theoretically derived maps of spectral species distributions.

295 The PCP is a stochastic mathematical process of assigning clusters of objects (here spectral  
296 species) in space according to the following: 1. Cluster centres for each object are randomly  
297 distributed across a landscape assuming a constant cluster density. The number of individuals in  
298 each cluster is drawn from a Poisson distribution. 3. Individuals within each cluster are then  
299 distributed based on a radially symmetrical Gaussian distribution .

300 In this study, a PCP was produced for each of the 50 spectral species across each of the three  
301 landscapes according to the following process:

- 302 (1) Empirical Ripley's K curves were derived for each spectral species in each landscape  
303 using the R package Spatstat (Baddeley & Turner 2005).
- 304 (2) When a Ripley's K curve is calculated for a PCP, it can be shown to have the functional  
305 form presented in Equation 1 (Plotkin et al. 2000). Consequently, we use an inverse  
306 modelling framework to match each empirically derived Ripley's K curve with Equation  
307 1 by adjusting  $\rho$  (the density of clusters across the landscape), and  $\mu$  (the intensity of  
308 individuals within each cluster).

309 Equation 1:  $K(d)^{PCP} = \pi d^2 + \rho^{-1} \left( 1 - \exp \left( \frac{-d^2}{4\mu^2} \right) \right)$

- 310 (3) Species likelihood probabilities were then determined for each spectral species using the  
311  $\rho$  and  $\mu$  values in a PCP in concert with the radial Gaussian probability function defined  
312 in Equation 2. Probabilities from each clump were overlaid on top of one another and the  
313 maximum likelihood was used.

314 Equation 2:  $h(x, y) = (2\pi\mu^2)^{-1} \exp \left( \frac{-(x^2+y^2)}{2\mu^2} \right)$

- 315 (4) The 50 species likelihood maps (one per spectral species) were then normalized based on  
316 the abundance of each spectral species in the empirical maps. These likelihoods were  
317 then used to weight a random draw that was used to condense the likelihoods into a  
318 single, theoretically-based spectral species map.
- 319 (5) A one ha grid was then fit over the simulated spectral species distribution map and the  
320 Bray Curtis distance among one ha kernels was calculated in exactly the same way as  
321 was done with the empirical data. From this grid, theoretical distance decay relationships  
322 were calculated in exactly the same manner as was done with the empirical spectral data  
323 (i.e. by calculating the mean similarity (inverse Bray Curtis) for all paired plots within  
324 bins of 100 m.
- 325 (6) Steps 2 to 5 were then repeated 20 times, to generate 20 distinct theoretical spectral  
326 species maps and associated distance decay curves. The final curves presented were the  
327 mean of means within each 100 m bin and the standard deviations of the means.

328

## 329 **Results**

### 330 Validation with forest plot data

331 At Allpahuayo Mishana, our estimates of species compositional turnover derived from spectral  
332 data were strongly correlated with field plot-based measures of beta-diversity ( $R^2 = 0.85$ ;  $P <$   
333  $0.001$ ; Fig. 3). However, the residual variance was higher among only terra firme forest plots ( $R^2$   
334  $= 0.29$ ;  $P = 0.05$ ) than among only white-sand forest plots ( $R^2 = 0.76$ ;  $P < 0.001$ ). At Jenaro  
335 Herrera, there was also a highly significant relationship between beta-diversity estimated with  
336 our spectral approach and field-measured beta-diversity ( $P = <0.001$ ), although there was more  
337 residual variance at this site than at Allpahuayo Mishana ( $R^2=0.68$ ). Most of the variation in the  
338 relationship between spectral and plot data came from palm swamp forests, which were poorly  
339 distinguished in the second NMDS axis; instead, the third NMDS axis was more useful at  
340 identifying areas of palm swamp (figure S.2). The relationship between spectral composition and  
341 species composition was consistent across two landscapes, and among different field plot  
342 datasets that were established using different sampling protocols with different stem diameter  
343 size limits.

### 344 Mapping beta-diversity

345 Our spectrally-derived maps of estimated tree species composition demonstrate clear gradients  
346 across the three study landscapes (Fig. 4). However, the underlying determinants of these  
347 floristic gradients appear to be different among the three sites. At Allpahuayo Mishana, the three  
348 NMDS axes show similar spatial patterns (Figs 4 and S.1), with NMDS axes 2 and 3 additionally  
349 containing a substantial element of sampling artefact (i.e. clear striping). This relative uniformity  
350 across NMDS axes suggests there is a single predominant floristic gradient at this site, because,  
351 if multiple important floristic gradients were present, we would expect them to be reflected in  
352 different NMDS axes. Combined with field validation data, our spectrally-derived maps indicate  
353 that the primary floristic gradient at this site reflects an underlying edaphic gradient from nutrient  
354 rich terra firme clay soils, to nutrient poor white-sand soils. These white-sand forests were  
355 always found at higher elevations ( $>145$  m a.s.l.) at Allpahuayo Mishana.

356 Our estimates of tree species composition also suggest that there is a strong spatial gradient in  
357 floristic composition at Quebrada Braga. Similar to Allpahuayo Mishana, consistency among  
358 NMDS axes suggests there is a single primary floristic gradient at Quebrada Braga (Fig. 4 and  
359 S.2). Somewhat surprisingly, this floristic gradient did not correspond strongly with elevation.  
360 The Quebrada Braga landscape is seasonally flooded by the large and nutrient-rich Ucayali  
361 River, which surrounds this landscape on three sides. Therefore, elevation will primarily  
362 determine the intensity and duration of this seasonal flooding.

363 Jenaro Herrera appears to be a more complex landscape than the other two, as it contains three  
364 distinct floristic gradients, demonstrated by three distinctive NMDS axes (Fig. 4 and S.3). This  
365 landscape appears to contain two forms of flooded forest, one flooded by nutrient-rich white  
366 water from the large Ucayali River and another flooded by nutrient-poor black water. In addition,  
367 there are patches of white-sand forest as well as forests that have been significantly impacted by  
368 anthropogenic activities.

#### 369 Empirical spectral distance decay

370 We observed a consistent pattern of a rapid decline in floristic similarity over distances of 500 m  
371 or less across all three forest types. Beyond this initial steep decay in similarity, three patterns  
372 distinguish these landscapes. In white-sand forests at Allpahuayo Mishana, after a rapid decay in  
373 similarity over the initial 800 m there was almost no discernible decrease in similarity with  
374 increasing distance (Fig. 5 panels A and D).

375 In seasonally-flooded forests at Quebrada Braga, we found a constant decay in floristic similarity  
376 with increasing distance. As with the other two landscapes, this decline was steepest over the  
377 initial 700 metres. However, the decline in compositional similarity persisted over the entirety of  
378 this landscape, as demonstrated by the consistently negative differential values (Fig. 5 panels B  
379 and D).

380 Finally, in terra firme forests at Jenaro Herrera we found a steep decay in compositional  
381 similarity over 500 m, followed by a more gradual decline up to distances of 3 km (Fig. 5 panels  
382 C and D). Beyond 3 km there was no discernible decrease in similarity with increasing distance  
383 up to 10 km. Additionally, at Jenaro Herrera there was greater overall variation in compositional  
384 similarity across all distances compared with the other two sites, as shown by the wider error

385 bars. We attribute this variation to the greater environmental variation at this site, as well as  
386 greater overall species diversity in terra firme forests as opposed to both white-sand forests and  
387 seasonally-flooded forests.

#### 388 Theoretical distance decay model

389 Overall the theoretical models derived from our PCP approach poorly represented the three  
390 empirical (spectrally derived) distance decay relationships (Fig. 5). At Allpahuayo Mishana,  
391 although the form of the theoretical distance decay relationship was very similar to that derived  
392 from the empirical data, the theoretically derived distance decay generally overestimates  
393 similarity relative to the empirical data (Fig. 5 panel A). Similarly, Fig. 5 panel B shows that at  
394 Jenaro Herrera, the general pattern of the distance decay relationship was reasonably  
395 characterized relative to the empirical relationship, but the overall distance magnitude was not.  
396 At Quebrada Braga, we found a very different pattern, with the PCP models predicting a  
397 sustained sharp decrease in similarity over the first kilometre, which was not reflected in the  
398 empirical data (Fig. 5 panel C). However, the shallow but continuous decline in similarity  
399 beyond the first kilometre demonstrated by the PCP at Quebrada Braga showed reasonable  
400 agreement with the empirically-based relationship (Fig. 5 panel C).

401

402

403 **Discussion**

404 Our results demonstrate that distance decay relationships vary among forest types in lowland  
405 Amazonia at a landscape scale. This is significant, because in contrast with previous plot-based  
406 studies, we are able to investigate this distance decay relationship continuously across landscapes  
407 while simultaneously maintaining high resolution. Within terra firme forests, our estimated  
408 distance decay curves are broadly consistent with a number of previous studies in this region  
409 (Condit et al., 2002; Duque et al., 2009), showing both rapid decay in similarity over short  
410 distances, followed by almost no decay at distances greater than 4 km. The two other forest types  
411 that we investigated also demonstrate this initial rapid decline in similarity over the first  
412 kilometre, supporting the idea that canopy tree species across forest types are spatially  
413 aggregated over scales less than one kilometre (Condit et al. 2000). However, beyond this first  
414 kilometre, patterns of distance decay sharply differ among different forest types.

415 The variation in distance decay among forest types is particularly apparent in seasonally-flooded  
416 forest, which shows a strong and relatively continuous decline in similarity with increasing  
417 distance. There are few plot based estimates of distance decay relationships in seasonally-  
418 flooded forests with which to compare our data (but see Wittmann et al., 2006; Draper et al.,  
419 2018). Nevertheless, our broad pattern of continuous decline in similarity appears to be  
420 consistent with these plot-based analyses. Much of the variation in spectral species composition  
421 across the Quebrada Braga landscape appears to be broadly independent of elevation. As  
422 elevation here should be a reasonable proxy for flooding duration and intensity, our data suggest  
423 that flooding duration and intensity are not the most important determinant of species  
424 composition in this landscape. This contrasts with a number of previous studies that have found  
425 flooding depth and duration to be the most important determinants of species composition (Assis  
426 et al., 2015; Junk et al., 2011; Wittmann et al., 2004; Wittmann et al., 2006).

427 Instead, our results appear to emphasize the importance of disturbance in determining species  
428 composition at this site. Disturbance has been recognised as an important driver of beta-diversity  
429 in West Amazonian floodplain forests (Puhakka et al., 1992; Salo et al., 1986). This may be  
430 especially true in Quebrada Braga as it is surrounded by the large and dynamic Ucayali River,  
431 which migrates laterally over decadal timescales (Salo et al., 1986; Schwenk et al., 2017).

432 Therefore, while some areas might have experienced large-scale disturbance relatively recently,  
433 other areas may not have been disturbed for many decades or centuries. Such disturbance  
434 patterns would also be spatially auto-correlated, and therefore consistent with the distance decay  
435 patterns we observe. The discrepancy between our study and previous plot-based studies (e.g.  
436 Assis et al., 2015; Junk et al., 2011; Wittmann et al., 2004; Wittmann et al., 2006) may arise  
437 from plot-based studies sampling predominantly mature seasonally-flooded forests over  
438 disturbed forests, whilst our study samples the whole landscape without this apparent bias.  
439 Fluvial disturbance is not the only form of large-scale spatially auto-correlated disturbance that  
440 may be driving beta-diversity patterns in Amazonian forests; for example, in central Amazonia  
441 large blow-down events have an important role in driving turnover in species composition  
442 (Marra et al. 2014).

443 In white-sand forests, the initial rapid decline in similarity with increasing distance is even more  
444 pronounced than in the other forest types and does not persist beyond the initial 800 m. This  
445 initial rapid decay may reflect the patchiness of white-sand forests at Allpahuayo Mishana.  
446 Patches of white-sand forests at this site are frequently smaller than 800 metres across, and  
447 ecological similarity is likely to be higher within a patch than between patches. In this way,  
448 white-sand forest tree communities may be functioning as meta-communities, separated by terra  
449 firme forests (Adeney et al., 2016; Palacios et al., 2016). The lack of declining similarity with  
450 increasing distance beyond 800 m is consistent with some published distance decay curves for  
451 white-sand forests in this region (Draper et al., 2018), whilst others that have been developed for  
452 much broader spatial scales appear to show a more constant decay (García-Villacorta et al.,  
453 2016; Guevara et al. 2016), presumably because they include several compositionally distinct  
454 floras.

455 Jenaro Herrera presents a different, and perhaps more complex pattern than in the other  
456 landscapes, indicated by the three NMDS axes showing distinct spatial patterns that reflect  
457 different underlying gradients. For example, patches of white-sand forests and terra firme forests  
458 are clearly distinct in NMDS axis 2, whilst palm swamp forests appear more strongly in the third  
459 NMDS axis. Furthermore, unlike the other two landscapes, Jenaro Herrera appears to show a  
460 strong anthropogenic disturbance gradient, which can be seen in high values in NMDS axis 1  
461 that cluster near the town (fig. S.1). This apparently high level of anthropogenic disturbance is in

462 some ways unsurprising as Jenaro Herrera supports a larger population than the other two sites  
463 and is surrounded by forests that are accessible and without formal legal protection. This  
464 contrasts with the other two landscapes, with Allpahuayo Mishana being accessible but protected  
465 and Quebrada Braga being unprotected but further from human development and due to seasonal  
466 flooding, relatively inaccessible.

467 A clear feature revealed by our LiDAR-derived DEM at Jenaro Herrera is the sharp increase in  
468 elevation that bisects the landscape from West to East (Fig. 4). This geological feature appears to  
469 be a boundary between the upland Tertiary Iquitos geanticline and Pleistocene alluvial terraces  
470 (Dumont et al., 1990; Dumont et al., 1991; Rasanen et al., 1992). Interestingly, this boundary  
471 appears to have little impact on floristic composition unlike other geological features in this  
472 region (Higgins et al., 2011, 2012). While field data will be required to confirm that there is little  
473 floristic turnover across this boundary, the boundary does not appear in local floristic  
474 classifications nor in maps of forest types (López Parodi & Freitas 1990; Honorio et al. 2008).

475 We were able to validate our approach by comparing our spectrally-derived estimates of beta-  
476 diversity with an extensive network of 53 forest plots distributed across two sites. Overall, this  
477 comparison provides compelling evidence that high fidelity imaging spectroscopy can be used to  
478 understand the spatial organisation of biodiversity in hyper-diverse tropical forests. Our results  
479 show highly significant linear relationship between spectrally-derived and plot-based estimates  
480 of beta-diversity consistent with previous studies that have used similar unsupervised approaches  
481 (Baldeck & Asner, 2013; Féret & Asner, 2014a, 2014b; Somers et al., 2015). Importantly, this  
482 strong relationship is preserved across plots using both 2 cm and 10 cm diameter cut-offs. As the  
483 spectral signal is derived entirely from the uppermost canopy layer, our results suggest that  
484 canopy level species composition may an excellent proxy for species composition in understory  
485 strata in these landscapes. The weaker relationship between spectral similarity and floristic  
486 similarity in terra firme forests may reflect the fact that fewer canopy species were recorded in  
487 this forest type. This is because the majority of stems recorded in the 0.1 ha plots are < 10 cm  
488 dbh, which will not reach the forest canopy in these tall forests. In the shorter stature white-sand  
489 forests, a larger proportion of small-stemmed trees will reach the canopy and therefore will be  
490 included in the spectral data.

491 Across all forest types, the distance decay relationships derived from the theoretical PCP  
492 compared poorly with the comparable empirical data. This mismatch suggests that the decay in  
493 community composition cannot be easily predicted by the clustering of conspecific individuals  
494 following a PCP. Major limitations of the PCP approach include the assumption that conspecific  
495 individuals are aggregated at a single scale, and the assumption that each clump of individuals  
496 throughout the landscape has the same Gaussian dispersal pattern (Morlon et al. 2008). The  
497 single scale of aggregation assumption may be largely correct at small spatial scales ( $\leq 50$  ha) in  
498 relatively homogenous environments (Morlon et al. 2008), where trees are aggregated mainly at  
499 small scales  $< 50$  m (Condit et al. 2000). However, at larger spatial scales ( $> 500$  ha), conspecific  
500 individuals aggregate at a range of different scales due to dispersal limitation, environmental  
501 specificity, Janzen-Connell effects, and competition among individuals (Levin, 1992; Wiegand et  
502 al., 2007). Similarly, a species is unlikely to have constant density across a 50 ha plot, however  
503 across a landscape  $> 1000$  ha, assuming a constant density becomes an even less plausible  
504 assumption. Our results demonstrate that these assumptions would need to be relaxed in order to  
505 reasonably predict distance decay relationships a landscape scales from theoretical spatial point  
506 process models such as the PCP.

507 Furthermore, our theoretical approach calculates PCP distributions for each spectral species  
508 independently; these distributions are then combined into a single map using random draws  
509 weighted by the landscape abundance of each spectral species in the empirical spectral species  
510 map. Our approach does not include interactions among species and between species and the  
511 environment, instead assuming the landscape is a homogeneous plane. Incorporating these biotic  
512 and abiotic interactions in future models could provide a way to further explore the relative  
513 influence of neutral and niche processes at landscape scales. Finally, our PCP was parameterized  
514 with by spectral species distributions. It is possible that parameterization based on actual species  
515 distribution data, which would be extremely difficult to collect at such large scales, may lead to  
516 different results. An approach integrating field and spectral species distributions could provide  
517 further insight.

518 A more general limitation of our approach is that we cluster the spectral signal of the entire  
519 landscape into just 50 spectral species and assume they are representative of hyper-diverse  
520 tropical forest landscapes that will contain hundreds (if not thousands) of tree species. While this

521 approach is well supported by both our comparisons with field data, and previous work that has  
522 shown 40 spectral species to be optimal (Feret & Asner 2014a), there are limitations. In general,  
523 it is likely that common canopy species will dominate the spectral signal as they make up a far  
524 greater proportion of the sunlit canopy, whilst rare and or understory species will be under-  
525 represented. Rare species are thought to have more localised and environmentally specific  
526 distributions (Hubbell, 2013), and therefore, the extent to which common species can be used to  
527 investigate spatial patterns of beta-diversity merits further investigation.

528 Additionally, using 50 spectral species elevates the similarity among plots within each forest  
529 type. This is especially evident in white-sand forests, where overall similarity is far higher in our  
530 spectral based analysis than has been found previously in plot based studies (Fine et al., 2010;  
531 García-Villacorta et al., 2016; Guevara et al., 2016, Draper et al., 2018). Many white-sand  
532 specialist tree species share functional characteristics that are likely to make them spectrally  
533 similar, such as increased leaf thickness and toughness, as well as lower concentrations of foliar  
534 N and P (Asner et al., 2016; Fortunel et al., 2014; Fyllas et al., 2009). Therefore, the diversity  
535 within white-sand forests may be poorly represented by our approach, resulting in an artificial  
536 increase in similarity between plots. However, the tight correlation between spectral and plot-  
537 based estimates of species composition in white-sand forests suggests that despite the overall  
538 increase in similarity among plots, our approach is still able to capture the main correlates of plot  
539 diversity.

540 The strength of our approach is that we can apply this method continuously to much larger areas  
541 than would be impossible using field data alone. Therefore, there is great potential for using our  
542 method to quantify beta-diversity and distance decay relationships continuously over far greater  
543 spatial extents. Furthermore, our approach is not only able to quantify beta-diversity, but also to  
544 precisely geo-locate where turnover occurs and therefore to suggest which environmental  
545 features may be important. We suggest that unsupervised spectral-based approaches, such as  
546 ours, can be used to actively guide field efforts to areas containing floristic assemblages that are  
547 poorly represented by current plot networks. We advocate for closer collaboration among  
548 ecologists using field-based data and those using imaging spectroscopy data.

549 In summary, this study demonstrates that distance decay relationships vary substantially among  
550 landscapes and forest types in lowland Amazonia, consistent with much of what has been found

551 previously using field plot-based data. Nevertheless, we also present findings that challenge  
552 previous hypotheses regarding the environmental drivers of tree species composition. In  
553 particular, we suggest that edaphic properties and topography may not always be the most  
554 important determinants of floristic composition, and in dynamic floodplain landscapes,  
555 disturbance may be a more important driver of tree species composition. Comparing estimates  
556 derived from our spectral data with a large dataset of forest plots, we provide compelling  
557 evidence for the validity of our approach, not only in classifying broad forest types, but also in  
558 describing subtle changes in floristic composition. Finally, our results demonstrate that distance  
559 decay relationships are driven by conspecific individuals aggregating at a range of nested scales  
560 across landscapes. Reproducing these patterns from theory will require the assumptions of PCP  
561 models to be relaxed.

## 562 **Acknowledgements**

563 This study was supported through a joint project between the Carnegie Institution for Science  
564 and the International Center for Tropical Botany at Florida International University. CAO data  
565 collection was supported by the John D. and Catherine T. MacArthur Foundation and the Avatar  
566 Alliance Foundation. We thank N Vaughn and D Knapp for processing the CAO data used in  
567 this analysis. Plot installation, fieldwork and botanical identification by the authors and  
568 colleagues has been supported by several grants including a Gordon and Betty Moore  
569 Foundation grant to RAINFOR, the EU's Seventh Framework Programme (283080,  
570 'GEOCARBON') and NERC Grants to OLP (Grants NER/A/S/2000/0053, NE/B503384/1,  
571 NE/F005806/1, and a NERC Postdoctoral Fellowship), and a National Geographic Society for  
572 supporting forest dynamics research in Amazonian Peru (grant #5472-95). OLP is supported by  
573 an ERC Advanced Grant and is a Royal Society-Wolfson Research Merit Award holder. The  
574 Carnegie Airborne Observatory is made possible by grants and donations to GP Asner from the  
575 Avatar Alliance Foundation, Grantham Foundation for the Protection of the Environment,  
576 Gordon and Betty Moore Foundation, the John D. and Catherine T. MacArthur Foundation, W.  
577 M. Keck Foundation, the Margaret A. Cargill Foundation, Mary Anne Nyburg Baker and G.  
578 Leonard Baker, Jr., and William R. Hearst III.

579 **Authors contributions**

580 FCD, GPA and CB conceived the ideas and designed the methodology. GPA collected the  
581 imaging spectroscopy data. CB, OLP, RVM, RZG, CAAG, MF, ENHC, TRB, RGV, PVAF, LF,  
582 AMM, and RJWB collected the field validation data. FCD, PB and GPA analysed the data. FCD  
583 wrote the manuscript with input from GPA, CB and PB. All authors contributed to drafts and  
584 gave final approval for publication.

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792 bosques de arena blanca y de suelo arcilloso en la Amazonía Peruana. *Revista Peruana de  
793 Biología*, **13**.

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797 **Table 1.** Summary of the spectral data used to estimate species composition for the three study  
 798 landscapes

	<b>Allpahuayo Mishana</b>	<b>Jenaro Herrera</b>	<b>Quebrada Braga</b>
Forest type	White-sand forest	Terra firme forest	Seasonally-flooded forest
Total landscape area (ha)	4540	4910	3107
Area of forest type (ha)	794	2309	2522
No. pairwise comparisons	315,218	2,665,740	2,412,585

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802 **Table 2** Summary of field plot inventory data used to calibrate spectral data at Allpahuayo  
 803 Mishana (AM) and Jenaro Herrera (JH)

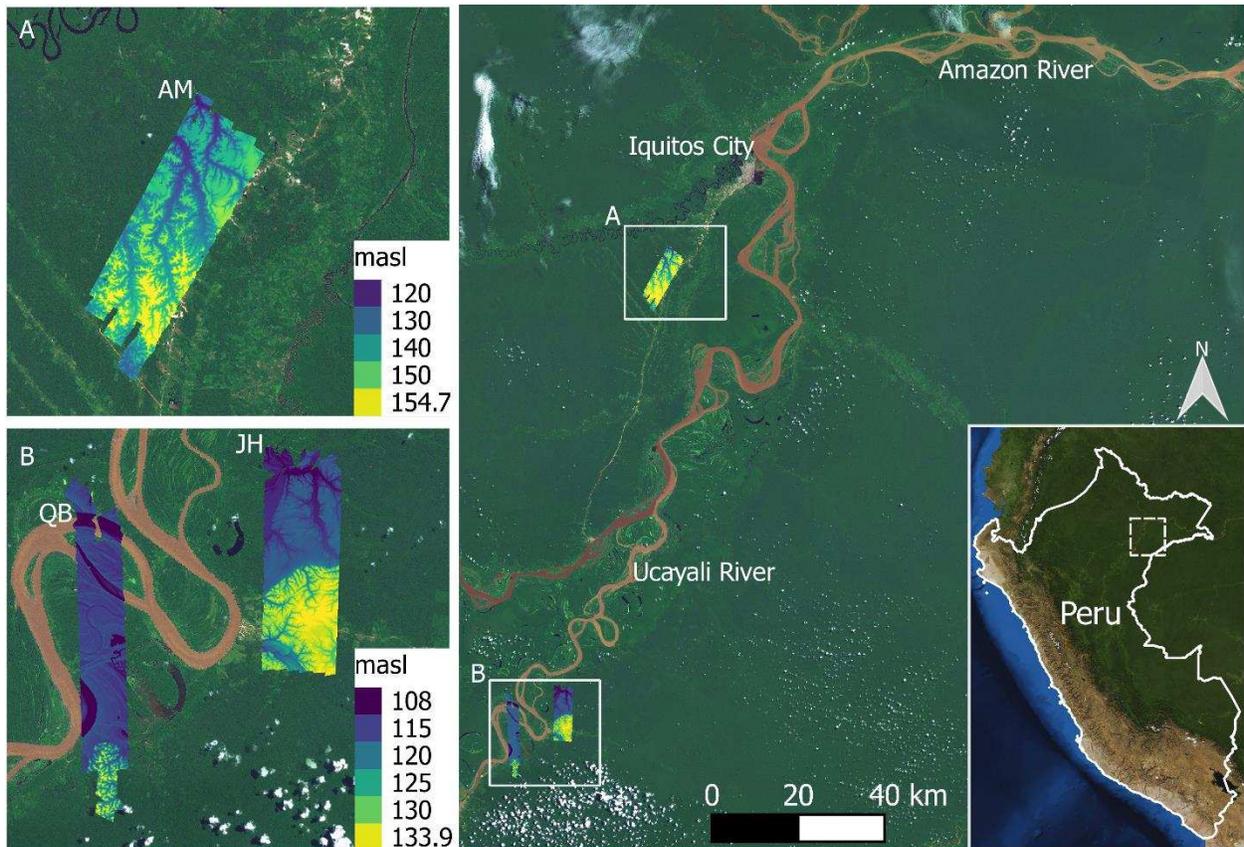
Plot Type	Large rectangular	Small rectangular	Small circular	Large 'Gentry'	Small 'Gentry'
Reference	Vasquez & Phillips 2000; Honorio Coronado et al. 2008	Zarate et al. 2006	Baldeck et al. 2016	Baraloto et al. 2011	(Phillips et al. 2003a)
Site	AM & JH	AM	AM	AM & JH	AM
No. plots	15	10	14	4	6
Plot area (ha)	0.5-1.5	0.1	0.1-0.25	0.5	0.1
Min. dap (cm)	5/10	2.5	10	2.5	2.5
Mean individuals (per plot)	663	358	88	242	260
Mean identified species (per plot)	139	89	33	78	79

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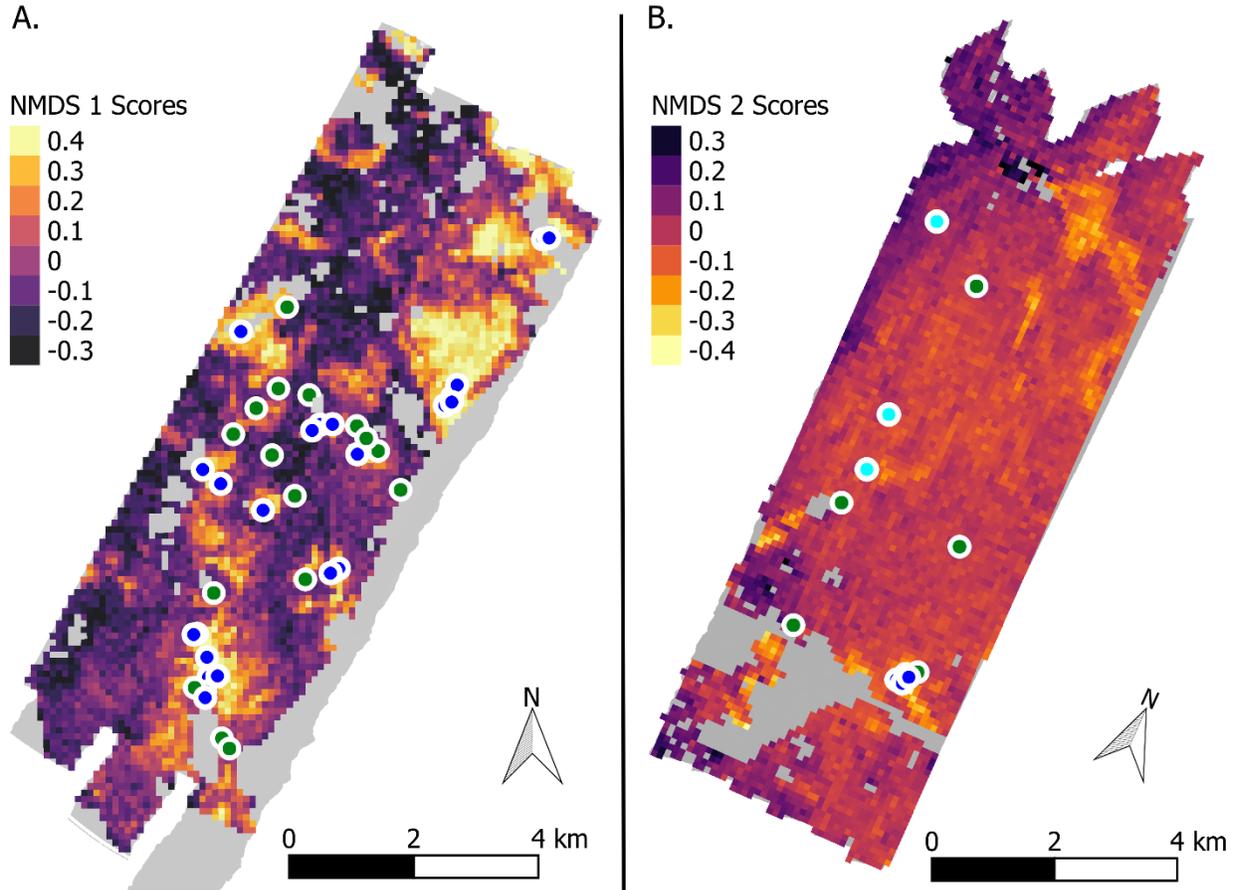
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810 **Figure 1.** Maps of the three study landscapes: Allpahuayo Mishana (AM), Quebrada Braga (QB)  
 811 and Jenaro Herrera (JH). Inset maps A and B show the immediate surroundings of the study  
 812 landscapes as well as the CAO LiDAR-derived digital terrain models for each landscape. The  
 813 third inset map shows the wider study region (dashed white box) in the context of Peru.

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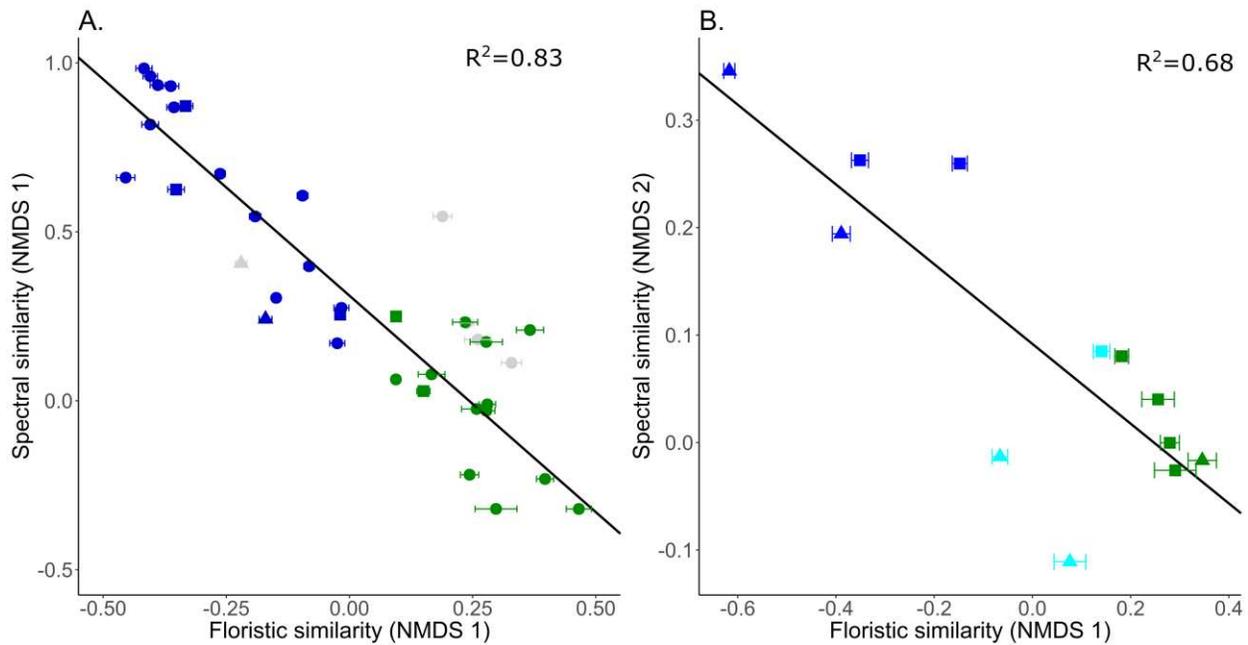


816  
 817 **Figure 2** Distribution of field plots across the Allpahuayo Mishana landscape (panel A) and the  
 818 Jenaro Herrera landscape (panel B). Blue circles represent plots in white-sand forest, green  
 819 circles represent terra firme forest plots and cyan represent palm swamp forest plots. The  
 820 backdrop of the map shows the first NMDS axis of the estimated species composition of  
 821 Allpahuayo Mishana and the second NMDS axis of the estimated species composition of Jenaro  
 822 Herrera, derived from airborne imaging spectroscopy.

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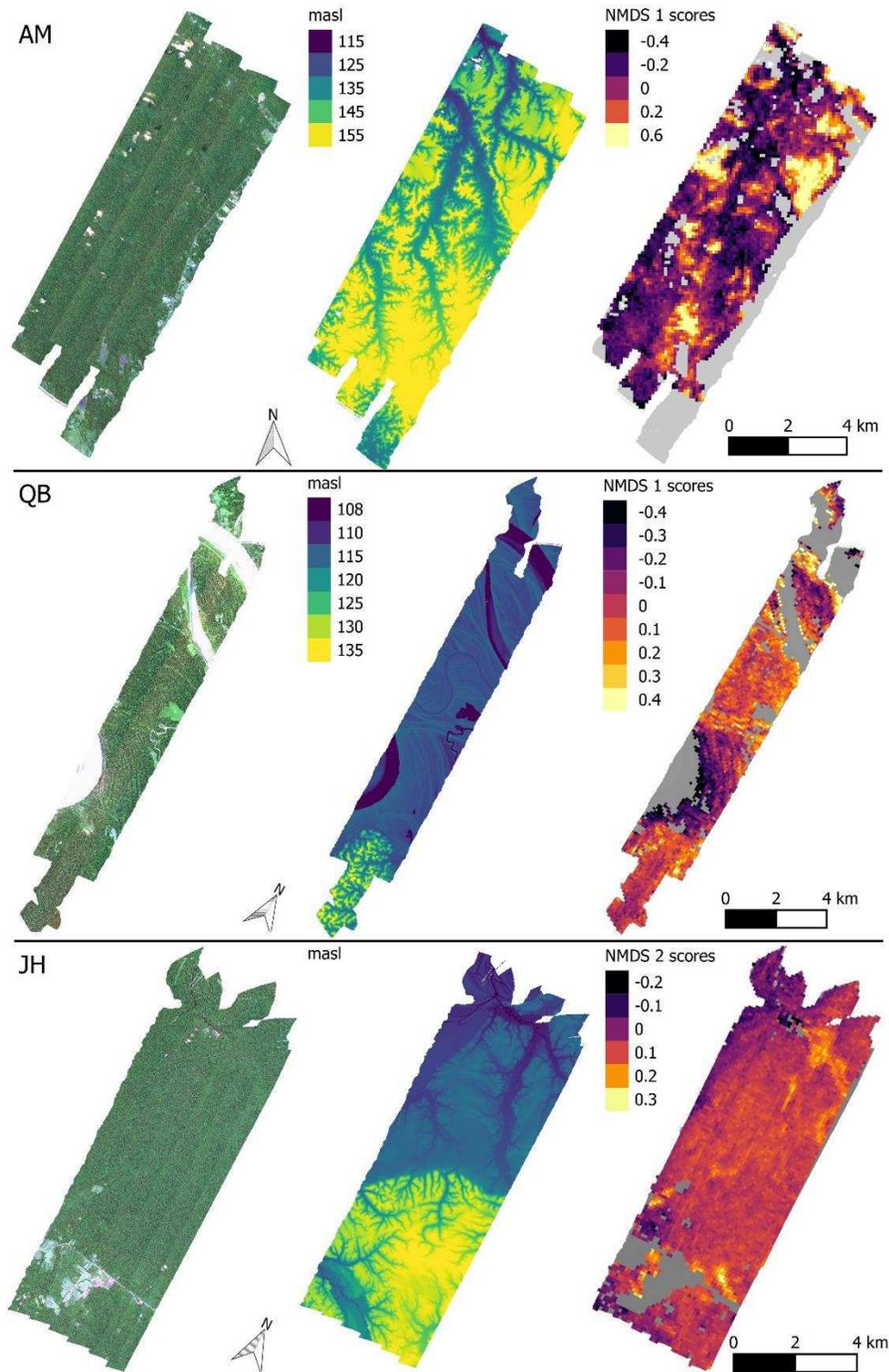
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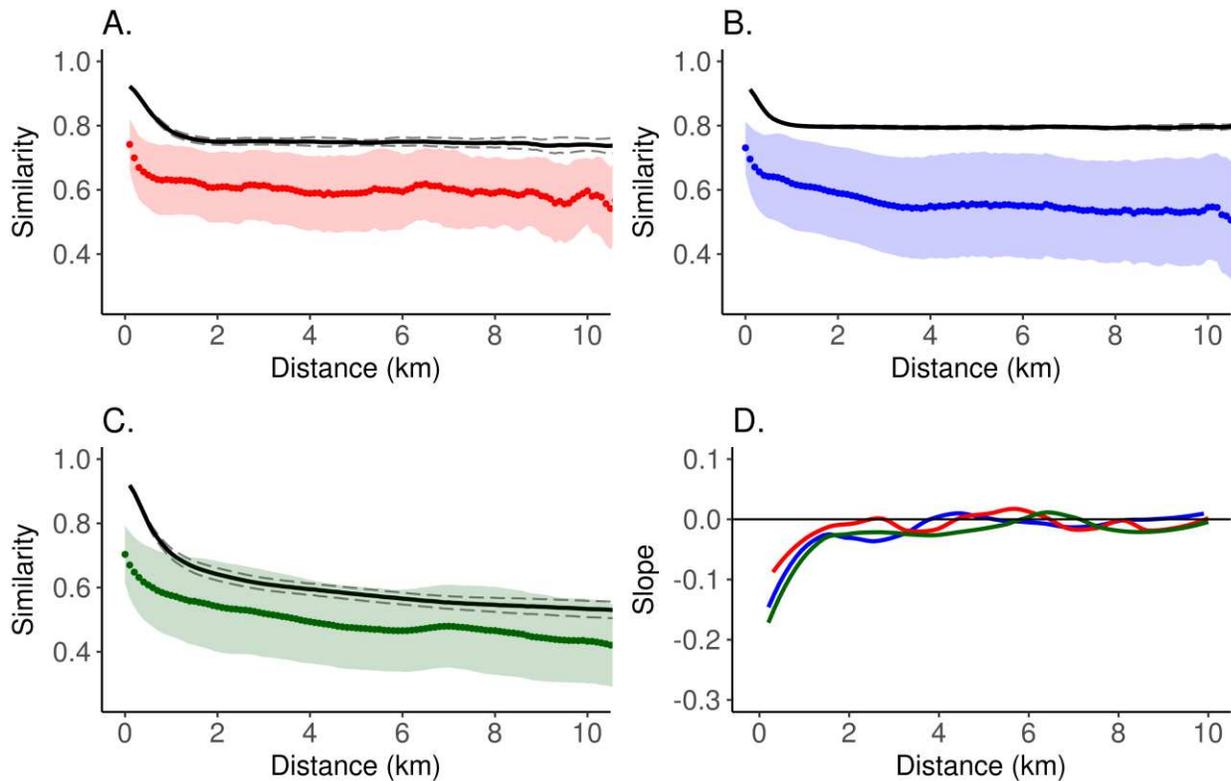
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827 **Figure 3** The relationship between spectrally derived estimates of tree species compositional  
828 turnover (represented by the first axis of the NMDS ordination of spectral species) and measured  
829 tree species compositional turnover (represented by the first axis of the NMDS ordination of tree  
830 species) at Allpahuayo Mishana (panel A.) and Jenaro Herrera (panel B.). Colours represent  
831 different forest types: Dark blue (white-sand forests); green (terra firme clay forests); cyan (palm  
832 swamp forests, grey symbols were those excluded from the analysis as they were < 10 m from a  
833 border between forest types. Error bars signify 95 % confidence intervals around floristic NMDS  
834 axis scores. Symbol shape corresponds to size of forest census plots, square (1 – 1.5 ha), triangle  
835 (0.5 ha), and circle (0.1 – 0.25 ha). Black lines represent linear regressions, both regressions  
836 were highly significant (P ≤ 0.001).

837



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 839 **Figure 4** Maps of the three study landscapes, Allpahuayo Mishana (AM), Jenaro Herrera (JH)  
 840 and Quebrada Braga (QB), The maps show RGB true colour (column 1), LiDAR-derived  
 841 elevation (column 2) and spectrally-derived estimates of tree species composition, summarised  
 842 by a single NMDS axis (column 3).



844

845 **Figure 5.** Distance decay relationships in three examples of forest types in the three different  
 846 landscapes: white-sand forests at Allpahuayo Mishana (panel A); terra firme forest at Jenaro  
 847 Herrera (panel B); seasonally-flooded forest at Quebrada Braga (Panel C). Points indicate mean  
 848 Bray-Curtis indices of similarity every 100 meters, and shaded areas are the standard deviations  
 849 surrounding each 100m point. Panel D shows the loess smoothed line (span=0.35) through the  
 850 first order derivative, calculated every 100m at each site. Colours correspond to different  
 851 landscapes/forest types: red = Allpahuayo Mishana white-sand, blue = Jenaro Herrera terra firme  
 852 clay, green = Quebrada Braga seasonally-flooded. Solid black lines indicate the mean PCP  
 853 theoretical predicted distance decays, and dashed black lines the standard deviations surrounding  
 854 these means.

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