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TITLE PAGE

Title: Efficacy of Pharmacological Therapies for the Treatment of Opioid-induced Constipation: Systematic Review and Network Meta-analysis.

Short running head: Treatment of Opioid-induced Constipation: Network Meta-analysis.

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Abbreviations:	b.i.d	twice-daily
	BM	bowel movement
	CI	confidence interval
	GI	gastrointestinal
	MeSH	medical subject heading
	o.d.	once-daily
	NNT	number needed to treat
	OIC	opioid-induced constipation

	PR	prolonged release
	QALY	quality-adjusted life year
	RCT	randomised controlled trial
	RR	relative risk
	SUCRA	surface under the cumulative ranking curve
	t.i.d	three times daily
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ABSTRACT

Objective: Opioids are increasingly prescribed in the West, and have deleterious gastrointestinal consequences. Pharmacological therapies to treat opioid-induced constipation (OIC) are available, but their relative efficacy is unclear. We performed a systematic review and network meta-analysis to address this deficit in current knowledge. **Design:** We searched MEDLINE, EMBASE, EMBASE Classic, and the Cochrane central register of controlled trials through to December 2017 to identify randomised controlled trials (RCTs) of pharmacological therapies in the treatment of adults with OIC. Trials had to report a dichotomous assessment of overall response to therapy, and data were pooled using a random effects model. Efficacy and safety of pharmacological therapies was reported as a pooled relative risk (RR) with 95% confidence intervals (CIs) to summarise the effect of each comparison tested, and ranked treatments according to their P-score. Results: Twenty-seven eligible RCTs of pharmacological therapies, containing 9149 patients, were identified. In our primary analysis, using failure to achieve an average of ≥ 3 bowel movements (BMs) per week with an increase of ≥ 1 BM per week over baseline, or an average of ≥ 3 BMs per week, to define non-response the network meta-analysis ranked naloxone first in terms of efficacy (RR = 0.65; 95% CI 0.52-0.80, P-score 0.84), and it was also the safest drug. When non-response to therapy was defined using failure to achieve an average of \geq 3 bowel movements (BMs) per week, with an increase of \geq 1 BM per week over baseline, naldemidine was ranked first (RR = 0.66; 95% CI 0.56-0.77, P-score 0.91), and alvimopan second (RR = 0.74; 95% CI 0.57-0.94, P-score 0.71).

Conclusion: In network meta-analysis, naloxone and naldemidine appear to be the most efficacious treatments for OIC. Naloxone was the safest of these agents.

What is already known about this subject?

There is an epidemic of opioid prescribing in the West. These drugs have deleterious effects on the gastrointestinal tract. Effective treatments, including μ-opioid receptor antagonists, prokinetics, and secretagogues are available, but their relative effectiveness is uncertain.

What are the new findings?

We identified 27 RCTs of pharmacological therapies in opioid-induced constipation (OIC), containing 9149 patients.

The μ -opioid receptor antagonists naloxone, naldemidine, alvimopan, and subcutaneous methylnaltrexone, as well as the prokinetic prucalopride, were all more effective than placebo for the treatment of OIC.

In our primary analysis, naloxone was ranked as the best drug, and was also the safest. Naldemidine was the most effective drug when an average of ≥ 3 bowel movements (BMs) per week with an increase of ≥ 1 BM per week over baseline was used to define response.

How might it impact on clinical practice in the near future?

Clinicians should consider the use of naloxone and naldemidine in OIC as a first choice when laxatives fail.

Guidelines for the management of OIC should be updated to include this important information.

INTRODUCTION

Chronic constipation affects up to 20% of individuals in the community. [1] Prescribed medications are implicated as a contributing factor in the aetiology of chronic constipation in a substantial proportion of individuals. [2] In the West, there is an epidemic of opioid prescribing, [3, 4] which in the US has led the surgeon general to discourage strongly the prescription of these drugs for non-malignant pain in adults. [5] As well as the risk of long-term addiction and higher rates of death among opioid users, [6] these drugs have undesired actions on the gastrointestinal (GI) tract, due to the location of μ -opioid receptors. Opioids lead to delayed GI transit and hard, infrequent stools, [7] with up to 50% of individuals taking these drugs reporting constipation, which they attribute to opioid user. [8, 9]

Surveys of patients receiving long-term opioid therapy reveal that opioid-induced constipation (OIC) is associated with significant increases in physician visits and sickness-related absence from work, as well as a significantly lower quality of life, compared with opioid users who do not experience constipation. [10, 11] In addition, up to one-third of patients reduce their opioid dosage, or discontinue the drugs altogether, in order to improve their bowel symptoms, which may negatively impact on their pain control. [9] As a result, a substantial proportion of patients with OIC use over the counter or prescription laxatives, in an attempt to alleviate their symptoms. [8] However, a previous Cochrane review revealed that there is little evidence to support a benefit of laxatives in this patient group, [12] which is mirrored by surveys of patients, where less than 50% report a satisfactory effect. [13]

As a result, pharmacological therapies have been developed in an attempt to provide effective therapy for a disorder that can be difficult to treat. The majority of these are drugs that act on μ -opioid receptors, in order to selectively antagonise the GI effects of

opioids. [14] These include methylnaltrexone, naloxone, alvimopan, bevenopran, naldemedine, and naloxegol. However, other agents that have already demonstrated their efficacy in the treatment of chronic idiopathic constipation and irritable bowel syndrome with constipation, such as prucalopride and lubiprostone, [15, 16] have also been tested in OIC. Prucalopride is a prokinetic, which is a highly selective 5-hydroxytryptamine-4 receptor agonist, and lubiprostone is a secretagogue, which acts on CIC-2 chloride channels in the intestine. Our previous meta-analysis demonstrated that some of these drugs were efficacious and safe in the treatment of OIC. [14] However, there have been new randomised controlled trials (RCTs) conducted for some of these agents, and several new drugs developed in the intervening 5 years. In addition, there have been no head-tohead studies conducted to enable healthcare providers to judge which of these drugs is likely to be the most effective in treating OIC. We have therefore updated our previous systematic review, [14] but also conducted a network meta-analysis, in order to examine these issues.

METHODS

Search Strategy and Study Selection

We updated a previous meta-analysis studying this issue. [14] A search of the medical literature was conducted using MEDLINE (2012 to December 2017), EMBASE and EMBASE Classic (2012 to December 2017), PUBMED (2012 to December 2017) and the Cochrane central register of controlled trials. We also searched clinicaltrials.gov for unpublished trials, or supplementary data for potentially eligible studies. RCTs examining the effect of pharmacological therapies (methylnaltrexone, naloxone, alvimopan, naldemedine, naloxegol, bevenopran, lubiprostone, prucalopride, naronapride, velusetrag, linaclotide, or plecanatide) in adult patients (>90% of participants over the age of 16 years) with OIC were eligible for inclusion (Box 1). The first period of cross-over RCTs were also eligible for inclusion.

A diagnosis of OIC was based on a history of constipation associated with the onset of opioid analgesic use. Studies recruiting patients with organic or chronic idiopathic constipation were ineligible. Trials using any dose of pharmacological therapy were considered eligible, and agents could be compared with each other, or with placebo. Only trials that used a minimum duration of 2 weeks of treatment were considered, in order not to overestimate the efficacy of one pharmacological therapy relative to another, meaning that we excluded three RCTs deemed eligible for the previous version of this metaanalysis. [17, 18, 19] Studies had to report a dichotomous assessment of overall response to therapy. First and senior authors of studies were contacted to provide additional information on trials, where required. The search strategy is provided in the supplementary materials. There were no language restrictions, and abstracts identified by the initial search were evaluated independently by two investigators for eligibility. All potentially relevant papers were obtained and evaluated in detail. Foreign language papers were translated, where required. Abstract books of conference proceedings between 2012 and 2017 were hand-searched to identify potentially eligible studies published only in abstract form. Bibliographies of all identified relevant studies were used to perform a recursive search of the literature. Articles were assessed independently by two investigators using predesigned eligibility forms, according to the pre-defined eligibility criteria. Disagreements

between investigators were resolved by discussion.

Outcome Assessment

The primary outcome assessed was the efficacy of pharmacological therapies, compared with each other or with placebo, in OIC in terms of failure to respond to therapy, with the endpoints of interest used to define response reported below. Secondary outcomes included adverse events occurring as a result of therapy (overall numbers, as well as individual adverse events including diarrhoea, abdominal pain, nausea, or reversal of analgesia).

Data Extraction

All data were extracted independently by two investigators on to a Microsoft Excel spreadsheet (XP professional edition; Microsoft Corp, Redmond, WA, USA) as dichotomous outcomes (response or no response to therapy). Many of the included eligible RCTs used different primary endpoints. However, due to the multitude of endpoints reported within the individual trials, we were able to assess the efficacy of therapies according to the following four dichotomous endpoints to define response to treatment: reporting an average of \geq 3 bowel movements (BMs) per week with an increase of \geq 1 BM per week over baseline; reporting an average of \geq 3 BMs per week; reporting any clinical improvement in symptoms; or reporting the need for use of rescue laxatives. For all included studies the following data were also extracted for each trial, where available: country of origin, setting (primary, secondary, or tertiary care), number of centres, criteria used to define OIC, proportion of female patients, and dose and duration of therapy. Data were extracted as intention-to-treat analyses, with drop-outs assumed to be treatment failures (i.e. no response to therapy), wherever trial reporting allowed. If this was not clear from the original article we performed an analysis on all patients with reported evaluable data.

Quality Assessment and Risk of Bias

Two investigators performed this independently at the study level. Disagreements were resolved by discussion. The Cochrane handbook was used to assess risk of bias, [20] by recording the method used to generate the randomisation schedule and conceal treatment allocation, whether blinding was implemented for participants, personnel, and outcomes assessment, whether there was evidence of incomplete outcomes data, and whether there was evidence of selective reporting of outcomes.

Data Synthesis and Statistical Analysis

A network meta-analysis was performed using the frequentist model, using the statistical package "netmeta" (version 0.9-0, https://cran.r-

project.org/web/packages/netmeta/index.html) in R (version 3.4.2), and reported according to the PRISMA extension statement for network meta-analyses, [21] in order to explore indirect treatment comparisons of the efficacy and safety of each medication. Network

meta-analysis results usually give a more precise estimate, compared with results from standard, pairwise analyses, [22, 23] and can also rank treatments to inform clinical decisions. [24]

We examined the symmetry and geometry of the evidence by producing a network plot with node and connection size corresponding to the number of study subjects and studies respectively. We produced a comparison adjusted funnel plot to explore publication bias of all available comparisons, versus placebo, using Stata version 14 (Stata Corp., College Station, TX, USA). Symmetry around the effect estimate line indicates the absence of publication bias, or small study effects. [25] We produced a pooled RR with 95% CIs to summarise the effect of each comparison tested, using a random effects model as a conservative estimate. There were no indirect comparisons between the active treatment groups, so we were unable to perform consistency modelling to check correlation between direct and indirect evidence. [26] Global statistical heterogeneity across all comparisons was assessed using the I^2 measure from the "netmeta" statistical package. The I^2 measure ranges between 0% and 100%, and is typically considered low, moderate, and high for values of 25% to 49%, 50% to 74%, and \geq 75% respectively. [27] We ranked the treatments according to their P-score. The P-score is a value between 0 and 1, with a higher score indicating a greater probability of the treatment being ranked as best. [28] However, the magnitude of the P-score should be considered, as well as the treatment rank. The mean value of the P-score is always 0.5, so if treatments cluster around this value they are likely to be of similar efficacy. In our primary analysis we combined an average of \geq 3 BMS per week with an increase of ≥ 1 BM per week over baseline and an average of ≥ 3 bowel movements (BMs) per week, but also analysed these separately. The number needed to treat (NNT), with a 95% CI, was calculated for each drug compared with placebo using the formula NNT = 1 / (control event rate x (1 - RR)).

We also performed analyses to assess the overall safety of each medication, including overall numbers of adverse events, as well as occurrence of diarrhoea, abdominal pain, nausea, or reversal of analgesia. We then performed a series of a priori subgroup analyses to test the robustness of our primary results. Firstly, we included only those studies with a low risk of bias. Secondly, we included only those studies with treatment duration of \geq 6 weeks, to account for a more prolonged response that may better reflect "real-world" use. Finally, we repeated our primary analysis using a Bayesian model to assess the robustness of our findings. We compared the relative efficacy of therapies for our primary outcome using the "mvmeta" commands in Stata, using a random effects model. We ranked the treatments according to their surface under the cumulative ranking curve (SUCRA) value. The SUCRA value is the equivalent to the P-score used in the frequentist model of our primary analyses. [28]

RESULTS

The search strategy generated 2523 citations, 48 of which appeared to be relevant to the systematic review and were retrieved for further assessment (Figure 1). Of these, 21 were excluded for various reasons, leaving a total of 27 eligible articles, [29, 30, 31, 32, 33, 34, 35, 36, 37, 38, 39, 40, 41, 42, 43, 44, 45, 46, 47, 48, 49, 50, 51, 52, 53, 54, 55] which reported 27 separate placebo-controlled trials, containing a total of 9149 patients. Twentytwo of these RCTs, reported in 22 articles, [29, 30, 31, 32, 33, 34, 35, 36, 37, 38, 39, 40, 41, 42, 43, 44, 45, 46, 47, 48, 49, 50] studied the effect of μ -opioid-receptor antagonists in OIC, three assessed lubiprostone, [51, 52, 53] and two prucalopride. [54, 55] We did not identify any trials of naronapride, velusetrag, linaclotide, or plecanatide in OIC. Five of the 22 trials of µ-opioid-receptor antagonists used naloxone, [29, 30, 31, 32, 33] four used methylnaltrexone, with endpoints of interest reported in six separate articles, [34, 35, 36, 37, 38, 39] five used naldemidine, reported in four articles, [40, 41, 42, 43] four used alvimopan, [44, 45, 46, 47] two naloxegol, reported in one article, [48] and two bevenopran. [49, 50] All studies were published in English, and agreement between investigators for trial eligibility for the 48 articles retrieved was excellent (Kappa statistic = 0.87). Detailed characteristics of individual RCTs are provided in Table 1. Risk of bias for all included trials is reported in Supplementary Table 1. Only 11 trials, reported in 11 articles, [30, 31, 33, 35, 37, 39, 40, 43, 45, 51, 52] were at low risk of bias. We obtained data for four of the RCTs from clinicaltrials.gov. [49, 50, 53, 55] We did not identify any trials making head-to-head comparisons of one drug versus another, meaning that direct evidence was only available in comparison with placebo. Active medications could therefore only be compared with each other using an indirect evidence meta-analysis. Data concerning failure to achieve clinical improvement, need for rescue laxatives, and safety are provided for the reader in the supplementary materials.

Table 1. Characteristics of Randomised Controlled Trials of Pharmacological Therapies Versus Placebo in Opioid-induced

Constipation.

Study	Country,	Patient group	Criteria used to define opioid-induced	Number of	Number of patients	Outcomes
	number of	studied	constipation	patients	assigned to active drug,	reported
	centres, and			(% female)	dosage, schedule, and	
	setting				duration of therapy	
Thomas 2008 [34]	USA and	Advanced illness	Stable opioid regimen for ≥ 2 weeks, ≤ 3	134 (56.7)	63 patients received	≥3 BMs per week
and Chamberlain	Canada, 27 sites,	(life expectancy	BMs* in previous week, and no		methylnaltrexone 0.15mg/kg	Clinical
2009 [38]	primary and	≥ 1 month),	clinically meaningful BM within 48		subcutaneously o.d.† on	improvement
	secondary care	laxative refractory	hours before first study dose		alternate days for 2 weeks	Need for rescue
						laxatives

Michna 2011 [35]	USA, multiple	Chronic non-	Opioid dose equivalent to >50mg/day of	460 (60.2)	298 patients received	\geq 3 BMs per week
and Iyer 2011	sites, setting not	malignant pain,	morphine for ≥ 2 weeks, <3 rescue-free		methylnaltrexone 12mg	≥3 BMs per week
[39]	reported	not laxative	BMs per week associated with ≥ 1 of:		subcutaneously o.d. or on	with an increase of
		refractory	hard or lumpy stools, straining, or		alternate days for 4 weeks	≥ 1 BM per week
			incomplete evacuation			from baseline
						Clinical
						improvement
						Need for rescue
						laxatives
Bull 2015 [36]	Multi-national,	Advanced illness	<3 BMs in previous week, and no	230 (48.7)	116 patients received	Need for rescue
	60 sites,	(life expectancy	clinically meaningful BM within 24		methylnaltrexone 8 or 12mg	laxatives
	secondary and	≥ 1 month), not	hours before first study dose		subcutaneously on alternate	
	tertiary care	laxative refractory			days for 2 weeks	
Rauck 2017 [37]	USA, 117 sites,	Chronic non-	Opioid dose equivalent to ≥50mg/day of	804 (62.9)	603 patients received	\geq 3 BMs per week
	setting not	malignant pain,	morphine for ≥ 2 weeks, <3 rescue-free		methylnaltrexone 150, 300,	with an increase of
	reported	laxative refractory	BMs per week associated with ≥ 1 of:		or 450mg orally o.d. for 4	≥ 1 BM per week
			Bristol stool form 1 or 2, straining, or		weeks, then as required for 8	from baseline
			incomplete evacuation on \geq 25% of BMs		weeks	

Liu 2002 [29]	USA, 1 site,	Chronic non-	Onset of constipation corresponding to	9 (55.6)	6 patients received naloxone	Clinical
	secondary care	malignant and	use of opioids and on a stable dose of		2 or 4mg orally t.i.d. \pm for 3	improvement
		malignant pain,	opioids		weeks	
		laxative status not				
		reported				
Simpson 2008	4 European	Chronic non-	Opioid dose equivalent to ≥20mg/day to	322 (60.9)	162 patients received	\geq 3 BMs per week at
[30]	countries,	malignant pain,	≤50mg/day of oxycodone, and		oxycodone PR [‡] /naloxone	4 weeks
	multiple sites,	not laxative	constipation caused or aggravated by an		PR orally in a 2:1 fixed dose	Need for rescue
	primary and	refractory	opioid		ratio for 12 weeks	laxatives at 4 weeks
	secondary care					
Meissner 2009	Germany, 28	Chronic non-	Stable oxycodone dose (40, 60, or	202 (62.9)	152 patients received	Need for rescue
[33]	sites, secondary	malignant pain,	80mg/day) with concomitant		naloxone 10, 20, or 40mg	laxatives
	and tertiary care	not laxative	constipation		orally o.d. for 4 weeks	
		refractory				
Lowenstein 2009	Multi-national,	Chronic non-	Stable oxycodone dose (60 to	265 (68.3)	130 patients received	\geq 3 BMs per week
[32]	multiple sites,	malignant pain,	80mg/day) with <3 BMs per week		oxycodone PR/naloxone PR	Clinical
	secondary care	not laxative	caused or aggravated by opioids		orally in a 2:1 fixed dose	improvement
		refractory			ratio for 12 weeks	Need for rescue
						laxatives

Sanders 2015 [31]	UK and	Chronic non-	Stable opioid dose for ≥ 1 month, and <3	40 (52.5)	32 patients received	Clinical
	Germany, 10	malignant pain,	BMs per week associated with ≥ 1 of:		naloxone 2.5, 5, 10, or 20mg	improvement
	sites, setting not	not laxative	hard or small stools, straining, or		orally o.d. for 3 weeks then	
	reported	refractory	incomplete evacuation on \geq 25% of BMs		b.i.d.§ for 3 weeks	
Paulson 2005 [44]	USA, 22 sites,	Chronic non-	Stable opioid dose for ≥ 1 week of	168 (58.3)	114 patients received	Clinical
	secondary and	malignant pain or	\geq 10mg morphine or equivalent, and <3		alvimopan 0.5 or 1mg orally	improvement
	tertiary care	opioid-dependent,	BMs per week associated with ≥ 1 of:		o.d. for 3 weeks	
		not laxative	hard or lumpy stools, straining,			
		refractory	sensation of obstruction, or incomplete			
			evacuation			
Webster 2008 [45]	Multi-national,	Chronic non-	Stable opioid dose for ≥ 1 month of	522 (63.8)	393 patients received	≥3 BMs per week
	113 sites,	malignant pain,	≥30mg morphine or equivalent and		alvimopan 0.5mg b.i.d., 1mg	Clinical
	secondary and	not laxative	history of decreased BMs since starting		o.d., or 1mg b.i.d. orally for	improvement
	tertiary care	refractory	opioids, with ≥ 1 of: hard stools,		6 weeks	
			straining, or incomplete evacuation on			
			25% of BMs			

Irving 2011 [46]	Multi-national,	Chronic non-	Stable opioid dose for ≥ 1 month of	485 (64.0)	321 patients received	\geq 3 BMs per week
	153 sites,	malignant pain,	≥30mg morphine or equivalent and		alvimopan 0.5mg o.d. or	with an increase of
	secondary and	laxative status not	history of decreased BMs since starting		0.5mg b.i.d. orally for 12	≥ 1 BM per week
	tertiary care	reported	opioids, with ≥ 1 of: hard stools or		weeks	from baseline
			straining			Clinical
						improvement
						Need for rescue
						laxatives
Jansen 2011 [47]	Multi-national,	Chronic non-	Stable opioid dose for ≥ 1 month of	518 (63.0)	346 patients received	\geq 3 BMs per week
	148 sites,	malignant pain,	≥30mg morphine or equivalent and		alvimopan 0.5mg o.d. or	with an increase of
	secondary and	laxative status not	history of decreased BMs since starting		0.5mg b.i.d. orally for 12	≥ 1 BM per week
	tertiary care	reported	opioids, with ≥ 1 of: hard stools or		weeks	from baseline
			straining			Need for rescue
						laxatives

Webster 2017 [41]	USA, 49 sites,	Chronic non-	Stable opioid dose for ≥ 1 month of	244 (68.4)	183 patients received	\geq 3 BMs per week
	setting not	malignant pain,	\geq 30mg morphine or equivalent and $<$ 3		naldemedine 0.1, 0.2, or	with an increase of
	reported	laxative refractory	BMs per week associated with ≥ 1 of:		0.4mg orally o.d. for 4	≥1 BM per week
			hard or small stools, straining, sensation		weeks	from baseline
			of obstruction, or incomplete evacuation			Clinical
			on \geq 25% of BMs			improvement
						Need for rescue
						laxatives
Katakami 2017a	Japan and South	Chronic malignant	Stable opioid dose and ≤ 5 BMs during a	227 (40.1)	170 patients received	≥3 BMs per week
[42]	Korea, 102 sites,	pain, laxative	14 day run-in period associated with ≥ 1		naldemedine 0.1, 0.2, or	with an increase of
	setting not	refractory	of: hard stools, straining, sensation of		0.4mg orally o.d. for 2	≥ 1 BM per week
	reported		obstruction, or incomplete evacuation		weeks	from baseline
			on 25% of BMs			Clinical
						improvement
						Need for rescue
						laxatives

Katakami 2017b	Japan, 70 sites,	Chronic malignant	Stable opioid dose and ≤ 5 BMs during a	193 (38.3)	97 patients received	\geq 3 BMs per week
COMPOSE-4	setting not	pain, not laxative	14 day run-in period associated with ≥ 1		naldemedine 0.2mg orally	with an increase of
[43]	reported	refractory	of: hard stools, straining, sensation of		o.d. for 2 weeks	≥ 1 BM per week
			obstruction, or incomplete evacuation			from baseline
			on 25% of BMs			
Hale 2017	Multi-national,	Chronic non-	Stable opioid dose for ≥ 1 month of	547 (60.4)	274 patients received	\geq 3 BMs per week
COMPOSE-1	68 sites, setting	malignant pain,	\geq 30mg morphine or equivalent and \leq 3		naldemedine 0.2mg orally	with an increase of
[40]	not reported	not laxative	BMs per week associated with ≥ 1 of:		o.d. for 12 weeks	$\geq 1 \text{ BM per week}$
		refractory	hard or lumpy stools, straining,			from baseline
			sensation of obstruction or blockage, or			
			incomplete evacuation on \geq 25% of BMs			
Hale 2017	Multi-national,	Chronic non-	Stable opioid dose for ≥ 1 month of	553 (60.5)	277 patients received	≥3 BMs per week
COMPOSE-2	69 sites, setting	malignant pain,	\geq 30mg morphine or equivalent and \leq 3		naldemedine 0.2mg orally	with an increase of
[40]	not reported	not laxative	BMs per week associated with ≥ 1 of:		o.d. for 12 weeks	≥1 BM per week
		refractory	hard or lumpy stools, straining,			from baseline
			sensation of obstruction or blockage, or			
			incomplete evacuation on \geq 25% of BMs			

Chey 2014	USA and	Chronic non-	Stable opioid dose for ≥ 1 month of ≥ 30	641 (61.3)	427 patients received	\geq 3 BMs per week
KODIAC-04 [48]	Europe, 115	malignant pain,	to 1000mg morphine or equivalent and		naloxegol 12.5 or 25mg	with an increase of
	sites, setting not	54.6% of patients	<3 BMs per week associated with ≥ 1 of:		orally o.d. for 12 weeks	≥ 1 BM per week
	reported	laxative refractory	hard or lumpy stools, straining,			from baseline
			sensation of obstruction, or incomplete			Need for rescue
			evacuation on \geq 25% of BMs			laxatives
Chey 2014	USA and	Chronic non-	Stable opioid dose for ≥ 1 month of ≥ 30	696 (63.4)	464 patients received	≥3 BMs per week
KODIAC-05 [48]	Europe, 142	malignant pain,	to 1000mg morphine or equivalent and		naloxegol 12.5 or 25mg	with an increase of
	sites, setting not	53.2% of patients	<3 BMs per week associated with ≥ 1 of:		orally o.d. for 12 weeks	≥ 1 BM per week
	reported	laxative refractory	hard or lumpy stools, straining,			from baseline
			sensation of obstruction, or incomplete			Need for rescue
			evacuation on \geq 25% of BMs			laxatives
Singla 2012 [50]	USA, number of	Chronic non-	<3 BMs per week associated with ≥ 1 of:	131 (48.1)	88 patients received	≥3 BMs per week
	sites and setting	malignant pain,	hard or lumpy stools, straining,		bevenopran 0.1 or 0.25mg	with an increase of
	not reported	not laxative	sensation of obstruction, or incomplete		orally o.d. for 4 weeks	≥ 1 BM per week
		refractory	evacuation on ≥25% of BMs			from baseline

Techner 2012 [49]	USA, number of	Chronic non-	\leq 3 BMs per week associated with \geq 1 of:	81 (69.1)	40 patients received	\geq 3 BMs per week
	sites and setting	malignant pain,	hard or lumpy stools, straining,		bevenopran 0.25mg orally	with an increase of
	not reported	not laxative	sensation of obstruction, or incomplete		o.d. for 4 weeks	$\geq 1 \text{ BM per week}$
		refractory	evacuation on ≥25% of BMs			from baseline
Cryer 2014 [51]	USA and	Chronic non-	Stable opioid dose for ≥ 1 month and <3	439 (64.4)	221 patients received	\geq 3 BMs per week
	Canada, 79 sites,	malignant pain,	BMs per week associated with ≥ 1 of:		lubiprostone 24mcg orally	Need for rescue
	primary,	not laxative	hard or very hard stools, straining, or		b.i.d for 12 weeks	laxatives
	secondary, and	refractory	incomplete evacuation on \geq 25% of BMs			
	tertiary care					
Jamal 2012 [52]	USA and	Chronic non-	Stable opioid dose for ≥ 1 month and <3	431 (63.1)	214 patients received	\geq 3 BMs per week
	Europe, 103	malignant pain,	BMs per week associated with ≥ 1 of:		lubiprostone 24mcg orally	with an increase of
	sites, primary,	not laxative	hard or very hard stools, straining, or		b.i.d for 12 weeks	≥ 1 BM per week
	secondary, and	refractory	incomplete evacuation on \geq 25% of BMs			from baseline
	tertiary care					Need for rescue
						laxatives

NCT00597428	USA and	Chronic non-	Stable opioid dose for ≥ 1 month and <3	437 (61.1)	223 patients received	\geq 3 BMs per week
(unpublished)	Canada, 114	malignant pain,	BMs per week associated with ≥ 1 of:		lubiprostone 24mcg orally	
[53]	sites, primary,	not laxative	hard or very hard stools, straining, or		b.i.d for 12 weeks	
	secondary, and	refractory	incomplete evacuation on \geq 25% of BMs			
	tertiary care					
Sloots 2010 [54]	Multi-national,	Chronic non-	Constipation secondary to chronic daily	196 (61.2)	130 patients received	≥3 BMs per week
	60 sites,	malignant pain,	opioid use		prucalopride 2 or 4mg orally	Clinical
	secondary and	not laxative			o.d. for 4 weeks	improvement
	tertiary care	refractory				
NCT01117051	Belgium, number	Chronic non-	Constipation secondary to chronic daily	174 (72.8)	88 patients received	\geq 3 BMs per week
(unpublished)	of sites and	malignant pain,	opioid use		prucalopride 1 or 2mg orally	
[55]	setting not	not laxative			o.d. for up to 12 weeks	
	reported	refractory				

*BM; bowel movement

†o.d.; once-daily

±t.i.d.; three times daily

‡PR; prolonged release

§b.i.d.; twice-daily

Efficacy

Failure to Achieve an Average of \geq 3 BMs per Week with an Increase of \geq 1 BM per Week Over Baseline, or an Average of \geq 3 BMs per Week

There were 22 RCTs, reported in 20 separate articles, [30, 32, 34, 35, 37, 40, 41, 42, 43, 45, 46, 47, 48, 49, 50, 51, 52, 53, 54, 55] randomising 5212 patients to active treatment and 3288 to placebo, included in this analysis. The network plot is provided in Supplementary Figure 1. There were moderate levels of global statistical heterogeneity (I^2 = 58.8%). The comparison adjusted funnel plot for publication bias, or other small study effects, showed symmetry around the zero line (Supplementary Figure 2). Naloxone was ranked as the most effective treatment (P-score 0.84), and was significantly more effective than placebo in two RCTs (RR 0.65; 95% CI 0.52 to 0.80, NNT = 4; 95% CI 3 to 8) (Figure 2). Naldemedine (five RCTs, RR 0.67; 95% CI 0.59 to 0.77, NNT = 5; 95% CI 4 to 7), alvimopan (three RCTs, RR 0.67; 95% CI 0.57 to 0.80, NNT = 5; 95% CI 4 to 8), subcutaneous methylnaltrexone (two RCTs, RR 0.74; 95% CI 0.58 to 0.94, NNT = 6; 95% CI 4 to 26), and prucalopride (two RCTs, RR 0.74; 95% CI 0.58 to 0.96, NNT = 6; 95% CI 4 to 39) were also significantly more effective than placebo (Figure 2). On indirect comparison of active treatments, significant differences were seen with naloxone compared with oral methylnaltrexone or lubiprostone; naldemedine compared with naloxegol, oral methylnaltrexone, or lubiprostone; and alvimopan compared with lubiprostone (Table 2).

We performed a series of pre-specified subgroup analyses. When considering only studies with a low risk of bias, there were nine RCTs included, reported in eight articles, [30, 35, 37, 40, 43, 45, 51, 52] randomising 2539 patients to active treatment and 1732 to placebo. Alvimopan was ranked as the best treatment (one RCT, P-score 0.89; RR

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Table 2. League Table of Results for Failure to Achieve an Average of ≥3 BMs per Week with an Increase of ≥1 BM per Week Over

Baseline or an Average of \geq 3 BMs per Week.

Naloxone									
0.97 (0.75; 1.25)	Naldemedine								
0.96 (0.73; 1.27)	0.99 (0.80; 1.24)	Alvimopan							
0.88 (0.64; 1.21)	0.91 (0.69; 1.19)	0.91 (0.68; 1.23)	Methylnaltrexone SC						
0.87 (0.62; 1.22)	0.90 (0.68; 1.20)	0.91 (0.66; 1.23)	0.99 (0.70; 1.41)	Prucalopride					
0.83 (0.60; 1.16)	0.86 (0.64; 1.15)	0.86 (0.63; 1.17)	0.95 (0.67; 1.34)	0.95 (0.66; 1.37)	Bevenopran				
0.76 (0.58; 1.01)	0.79 (0.63; 0.99)	0.79 (0.62; 1.02)	0.87 (0.65; 1.17)	0.88 (0.64; 1.19)	0.92 (0.68; 1.25)	Naloxegol			
0.71 (0.51; 0.99)	0.74 (0.56; 0.97)	0.74 (0.55; 1.00)	0.81 (0.58; 1.14)	0.82 (0.57; 1.16)	0.86 (0.60; 1.22)	0.93 (0.69; 1.26)	Methylnaltrexone		
0.71 (0.55; 0.92)	0.73 (0.60; 0.90)	0.74 (0.58; 0.93)	0.81 (0.61; 1.07)	0.81 (0.60; 1.10)	0.85 (0.63; 1.15)	0.93 (0.74; 1.17)	1.00 (0.75; 1.33)	Lubiprostone]
0.65 (0.52; 0.80)	0.67 (0.59; 0.77)	0.67 (0.57; 0.80)	0.74 (0.58; 0.94)	0.74 (0.58; 0.96)	0.78 (0.61; 1.01)	0.85 (0.71; 1.01)	0.91 (0.71; 1.17)	0.92 (0.79; 1.07)	Placebo

Relative risk with 95% confidence intervals in parentheses. Comparisons, column versus row, should be read from left to right, and are ordered relative to their overall efficacy. The treatment in the top left position is ranked as best after the network meta-analysis of indirect effects.

compared with placebo = 0.58; 95% CI 0.42 to 0.79) (Supplementary Figure 3). There was moderate global statistical heterogeneity ($I^2 = 70.2\%$). When restricting to studies with treatment duration of ≥ 6 weeks 13 RCTs were included, reported in 11 articles, [32, 37, 40, 45, 46, 47, 48, 51, 52, 53, 55] randomising 3981 patients to active treatment and 2531 to placebo. Naloxone was the highest ranked treatment (one RCT, P-score 0.89; RR compared with placebo = 0.66; 95% CI 0.53 to 0.84), with low global statistical heterogeneity ($I^2 = 36.2\%$) (Supplementary Figure 4). Results for our primary analysis were almost identical when a Bayesian model was used, although bevenopran moved from being ranked sixth to being ranked third.

Failure to Achieve an Average of \geq 3 BMs per Week with an Increase of \geq 1 BM per Week Over Baseline

There were 14 separate RCTs, reported in 12 articles, [35, 37, 40, 41, 42, 43, 46, 47, 48, 49, 50, 52] which randomised 3802 patients to active treatment and 2209 to placebo, providing data for this analysis. The network plot is provided in Supplementary Figure 5. There were moderate levels of global statistical heterogeneity ($I^2 = 70.6\%$). The comparison adjusted funnel plot for publication bias, or other small study effects, showed symmetry around the zero line. From the network meta-analysis naldemedine was ranked as the most effective treatment in terms of this endpoint (P-score 0.91), and was significantly more effective than placebo (five RCTs, RR 0.66; 95% CI 0.56 to 0.77) (Figure 3). Alvimopan (two RCTs, RR 0.74; 95% CI 0.57 to 0.94) was also significantly more effective than placebo. On indirect comparison of active treatments, there were no significant differences seen (Supplementary Table 2).

Failure to Achieve an Average of \geq 3 BMs per Week

There were nine separate RCTs, [30, 32, 34, 35, 45, 51, 53, 54, 55] which randomised 1708 patients to active treatment and 1241 to placebo, providing data for this analysis. The network plot is provided in Supplementary Figure 6. There was no statistical heterogeneity ($I^2 = 0\%$). With only nine studies, there were too few to assess for risk of publication bias, or other small study effects. Alvimopan was ranked as the most effective treatment in terms of this endpoint (P-score 0.96), and was significantly more effective than placebo (one RCT, RR 0.58; 95% CI 0.48 to 0.70) (Figure 4). Naloxone (two RCTs, RR 0.65; 95% CI 0.56 to 0.75), subcutaneous methylnaltrexone (two RCTs, RR 0.75; 95% CI 0.64 to 0.88), and prucalopride (two RCTs, RR 0.75; 95% CI 0.62 to 0.92) were also significantly more effective than placebo (Figure 4). On indirect comparison of active treatments, significant differences were seen with alvimopan compared with subcutaneous methylnaltrexone or lubiprostone; and with both naloxone and subcutaneous methylnaltrexone compared with lubiprostone (Supplementary Table 3).

DISCUSSION

In terms of our primary endpoint, this systematic review and network meta-analysis demonstrated that some μ -opioid-receptor antagonists, including naloxone, naldemidine, alvimopan, and subcutaneous methylnaltrexone, as well as the prokinetic prucalopride, were all more effective than placebo for the treatment of OIC. Of these drugs, naloxone was ranked as the most likely to be superior to placebo, and was significantly better than subcutaneous methylnaltrexone and lubiprostone. Naldemidine was the next best drug, and was superior to naloxegol and methylnaltrexone. Naloxone remained the best drug in some of our subgroup analyses, including when only RCTs with a treatment duration of ≥ 6 weeks were considered in the analysis, and in terms of reducing the need for rescue laxatives. However, when failure to achieve an average of ≥ 3 BMs per week with an increase of ≥ 1 BM over baseline was used to define non-response to therapy, which is a more rigorous endpoint, naldemidine was the drug ranked first. In terms of safety, naloxone was the drug ranked first in terms of safety, and lubiprostone last.

We performed a contemporaneous and exhaustive literature search, which included searching the "grey" literature and clinicaltrials.gov, allowing us to analyse data from 27 RCTs of μ -opioid-receptor antagonists, lubiprostone, or prucalopride versus placebo, containing 9149 patients with OIC. In addition, the literature search, eligibility assessment, and data extraction were undertaken independently by two reviewers. We used an intention-to-treat analysis, wherever trial reporting allowed, and pooled data with a random effects model, to provide a more conservative estimate of the efficacy and safety of individual pharmacological therapies in OIC. Finally, we attempted to contact authors of individual studies and accessed clinicaltrials.gov in order to obtain extra information, where required.

Limitations include the fact that only 11 of the RCTS were at low risk of bias, and original authors did not respond to all our queries concerning the methodology used in individual studies. This may mean the efficacy of pharmacological therapies in OIC has been overestimated. [56] The vast majority of trials recruited individuals in secondary or tertiary care, or did not report the study setting, so involved individuals may not be generalisable to OIC patients consulting in primary care. There were moderate levels of global statistical heterogeneity in some of our analyses, although the comparison adjusted funnel plot for the primary outcome from the network was symmetrical, and not suggestive of publication bias, although two of the trials we identified had not been published as either full papers or conference abstracts, [53, 55] and were only identified during our search of clinicaltrials.gov. Finally, there were limited data for naloxegol, bevenopran, and prucalopride. In the case of naloxegol, it should be pointed out that the two phase III RCTs included a total of 1337 patients, [48] and we identified a phase II trial of the drug, but dichotomous data were not reported, and we could not obtain these either from the authors or from clinicaltrials.gov. [57] Phase III RCTs using bevenopran failed to recruit, and were therefore abandoned, meaning that development of the drug was discontinued, and a phase III RCT of prucalopride, which our search identified, [55] was terminated prematurely based on a business priority decision, meaning firstly that there are unlikely to be any further studies of these two drugs, and secondly that the efficacy of both may have been overestimated in the network meta-analysis.

Our network meta-analysis could also be criticised due to the absence of trials making direct comparisons between pharmacological therapies, meaning that all our conclusions were derived from data making indirect treatment comparisons. However, we believe it is unlikely that pharmaceutical companies would ever conduct head-to-head RCTs of these agents. Even if such a study were to be conducted, it is more than likely it would be designed as a non-inferiority trial, rather than a superiority trial. This is a similar situation to other functional disorders, such as chronic idiopathic constipation, where the only trial to date that has compared two active drugs was designed as a non-inferiority trial. [58] This is the advantage of network meta-analyses, which can circumvent problems such as these, in order to provide a credible ranking system of the likely efficacy and tolerability of all available treatment options for OIC. The results of our study are therefore likely to be important for both patients and policy makers, in order to help inform treatment decisions.

At the time of our previous meta-analysis there had been no guidelines from national or international organisations to aid Gastroenterologists in the treatment of OIC, [14] despite the fact that patients with constipation, who may have opioids as a precipitating cause, are often seen in Gastroenterology outpatient departments. However, a group of experts in the field have made recent recommendations based on the current available evidence. [59] The authors noted that there was no evidence for any benefit of lifestyle modifications or dietary changes in patients with OIC. In terms of laxatives, it was felt that indirect evidence favoured the use of bisacodyl, sodium picosulfate, polyethylene glycol, or senna first-line, but that there was insufficient evidence for the use of either lubiprostone or prucalopride. A treatment algorithm was proposed, which included the use of laxatives first-line, with co-prescription of a μ -opioid receptor antagonist in those in whom there is no improvement in the symptoms of OIC.

Less than 50% of patients with OIC report a satisfactory therapeutic effect of laxatives, [13] and this is supported by evidence from an updated Cochrane review, which failed to demonstrate any evidence of a beneficial effect of laxatives in OIC. [12] As a result, there is a clear need for cost-effective treatments, particularly given the continued increase in opioid prescribing worldwide. [3, 4] This is especially relevant to Gastroenterologists, as up to 20% of patients with chronic abdominal pain disorders will be given opioids. [60] In support of our observation from the primary analysis that naloxone was the drug with the best efficacy and safety profile, there have been two cost-utility analyses of combination oxycodone / naloxone, conducted from a UK and Canadian perspective, in patients with moderate to severe pain and OIC. [61, 62] In both of these studies, combination oxycodone / naloxone was associated with a gain in quality-adjusted life years (QALYs), compared with oxycodone alone. The cost per QALY in the UK study was £5841, with a probability of over 96% that oxycodone / naloxone would be cost-effective at a cost per QALY of £20,000. In the Canadian study the cost per QALY ranged from \$2178 to \$7732. Similarly, naloxegol was cost-effective in a UK population with OIC, with a cost per QALY of £10,849, and a 91% probability of being cost-effective at a cost per QALY of £20,000, [63] but the drug was only effective in our analyses in terms of reducing the need for rescue laxatives. Although there have been cost-effectiveness analyses of lubiprostone and prucalopride, [64, 65] which have demonstrated both are likely to be cost-effective, these have been conducted in patients with chronic idiopathic constipation, rather than OIC.

In summary, our network meta-analysis suggested that, in terms of our primary analysis, naloxone was the drug with the best efficacy profile, and was also the least likely to cause adverse events. In our secondary analyses, naloxone remained most effective in terms of reducing the need for rescue laxatives, and was the second most effective drug when failure to achieve an average of \geq 3 BMs per week was used to define non-response. However, naldemidine was the most efficacious drug when failure to achieve an average of \geq 1 BM over baseline was used to define non-response, which is probably a more rigorous endpoint, and in terms of rates of clinical improvement. All of this, together with evidence from the pharmacoeconomic literature demonstrating that naloxone is highly likely to be cost-effective, lends weight to either

naloxone or naldemidine being the clinician's first choice of pharmacological therapy for OIC when laxatives fail.

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Guarantor of the article: ACF is guarantor.

Specific author contributions: PL, NEB, DMB and ACF conceived and drafted the study. ACF and PL collected all data. ACF and NEB analysed and interpreted the data. ACF, PL, and NEB drafted the manuscript. All authors commented on drafts of the paper. All authors have approved the final draft of the manuscript.

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Box 1.

Randomised controlled trials.

Adults (>90% of participants aged >16 years) receiving opioid or opiate drugs.

Diagnosis of opioid-induced constipation based on either clinical symptoms, a physician's

opinion, or meeting specific diagnostic criteria specified by study investigators,

supplemented by negative investigations where trials deemed this necessary.

Compared pharmacological therapies (methylnaltrexone, naloxone, alvimopan,

naldemedine, naloxegol, bevenopran, lubiprostone, prucalopride, naronapride, velusetrag,

linaclotide, or plecanatide,) with each other, or placebo.

Minimum duration of therapy of 2 weeks.

Dichotomous assessment of overall response to therapy.

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FIGURE LEGENDS

Figure 1. Flow Diagram of Assessment of Studies Identified in the Systematic Review. Figure 2. Forest Plot of the Indirect Evidence for Failure to Achieve an Average of ≥3 BMs per Week with an Increase of ≥1 BM per Week Over Baseline, or an Average of ≥3 BMs per Week.

 I^2 for global statistical heterogeneity = 58.8%.

Note: The P-score is the probability of each treatment being ranked as best in the network analysis. A higher score equates to a greater probability of being ranked first.

Direct comp. is the number of direct comparisons of the indicated medication versus placebo.

Figure 3. Forest Plot of the Indirect Evidence for Failure to Achieve an Average of \geq 3 BMs per Week with an Increase of \geq 1 BM per Week Over Baseline.

 I^2 for global statistical heterogeneity = 70.6%.

Note: The P-score is the probability of each treatment being ranked as best in the network analysis. A higher score equates to a greater probability of being ranked first.

Direct comp. is the number of direct comparisons of the indicated medication versus placebo.

Figure 4. Forest Plot of the Indirect Evidence for Failure to Achieve an Average of ≥ 3 BMs per Week.

 I^2 for global statistical heterogeneity = 0%.

Note: The P-score is the probability of each treatment being ranked as best in the network analysis. A higher score equates to a greater probability of being ranked first. Direct comp. is the number of direct comparisons of the indicated medication versus placebo.